Solution for Q2

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Understanding the Gaussian Mixture Model (GMM):

So interesting thing to note about GMM is that it is **not** a mere linear combination of random variables (r.v.). In latter, all the component r.v. X_i s return a value whose linear combination is taken.

But in GMM, as clearly stated in the question, actually only one of the X_i s are chosen with probability p_i and from that value is taken. Hence at a particular measurement, **only and exactly** one r.v. comes into action.

Now, for GMM, we will make use of Law of Total Expectation/Law of Iterated Expectation/Adam's Rule which states:-

$$X \sim \sum_{i} p_{i} X_{i} \Rightarrow E[g(X)] = \sum_{i} E[g(X)|(X = X_{i})] P(X = X_{i}) = \sum_{i} E[g(X_{i})] P(X = X_{i}) = \sum_{i} p_{i} E[g(X_{i})] P$$

Where g is a function on a r.v.

Solving the GMM:-

Before moving ahead, let's see what mean and variance of a Gaussian Variable $\mathcal{N}(\mu_i, \sigma_i^2)$

$$E[X_i] = E[\mathcal{N}(\mu_i, \sigma_i^2)] = \mu_i \tag{1}$$

$$E[(X_i - \mu_i)^2] = \sigma_i^2 \tag{2}$$

As $Var(Y) = E[Y^2] - E[Y]^2$ for any r.v. Y,

$$E[X_i^2] = Var(X_i) + E[X_i]^2 = \sigma_i^2 + \mu_i^2$$
(3)

Mean E(X)

Using the above Law and equation (1)...

$$\mu = E[X] = \sum_{i} p_i E[X_i] = \sum_{i=1}^{K} p_i \mu_i$$
 (4)

Variance Var(X)

$$Var(X) = E[(X - \mu)^2]$$

From the above Law,

$$E[(X - \mu)^2] = \sum_{i} p_i E[(X_i - \mu)^2]$$

Opening by the linearity of the expectation operator...

$$E[(X_i - \mu)^2] = E[X_i^2 - 2\mu X_i + \mu^2] = E[X_i^2] - 2\mu E[X_i] + E[\mu^2]$$

From equations (1), (3) and (4)...

$$E[X_i^2] - 2\mu E[X_i] + E[\mu^2] = \sigma_i^2 + \mu_i^2 - 2\mu\mu_i + \mu^2$$

$$\Rightarrow \sum_{i} p_{i} E[(X_{i} - \mu)^{2}] = \sum_{i} p_{i} (\sigma_{i}^{2} + \mu_{i}^{2} - 2\mu\mu_{i} + \mu^{2}) = \sum_{i} [p_{i} (\sigma_{i}^{2} + \mu_{i}^{2}) - 2\mu(\sum_{i} p_{i}\mu_{i}) + \mu^{2}(\sum_{i} p_{i})]$$

Given $\sum_{i} p_i = 1$, so...

$$\Rightarrow \sum_{i} p_{i} E[(X_{i} - \mu)^{2}] = \sum_{i} [p_{i}(\sigma_{i}^{2} + \mu_{i}^{2})] - 2\mu(\mu) + \mu^{2}(1) = \sum_{i} [p_{i}(\sigma_{i}^{2} + \mu_{i}^{2})] - \mu^{2}$$

Hence...

$$Var(X) = \sum_{i=1}^{K} [p_i(\sigma_i^2 + \mu_i^2)] - \mu^2$$

$MGF(X) \Phi_X(t)$

For a Gaussian r.v. $X_i = \mathcal{N}(\mu_i, \sigma_i^2)$, its MGF is...

$$\Phi_{X_i}(t) = E[e^{tX_i}] = e^{\mu_i t + \frac{1}{2}\sigma_i^2 t^2}$$

So using the above Law...

$$\Phi_X(t) = E[e^{tX}] = \sum_i p_i E[e^{tX_i}] = \sum_i p_i \Phi_{X_i}(t) = \sum_{i=1}^K p_i e^{\mu_i t + \frac{1}{2}\sigma_i^2 t^2}$$

Solving the Linear Combination System:

Now we have a r.v. Z such that $Z = \sum_i p_i X_i$ and as said in the very beginning, all the component r.v. X_i s return a value and their linear combination is taken. We shall make extensive use of the **linearity property** of the expectation operator...

Mean E(Z)

$$E[Z] = E[\sum_{i} p_i X_i] = \sum_{i} E[p_i X_i] = \sum_{i} p_i E[X_i] = \sum_{i=1}^{K} p_i \mu_i = \mu$$

Peculiarly it's same as the GMM case..

Variance Var(Z)

$$Var(Z) = E[(Z - \mu)^2] = E[(\sum_{i} p_i X_i - \mu)^2]$$

$$\Rightarrow E[(\sum_i p_i X_i)^2 - 2\mu Z + \mu^2] = E[(\sum_i p_i X_i)^2] - 2\mu E[Z] + \mu^2 = E[(\sum_i p_i X_i)^2] - 2\mu (\mu) + \mu^2 = E[(\sum_i p_i X_i)^2] - \mu E[($$

Now, using this equation : $(\sum_i p_i X_i)^2 = \sum_{1 \le i \le K} (p_i X_i)^2 + 2 \sum_{1 \le i \le j \le K} p_i p_j X_i X_j \dots$

$$E[(\sum_{i} p_{i} X_{i})^{2}] = E[\sum_{1 \leq i \leq K} (p_{i} X_{i})^{2} + 2 \sum_{1 \leq i \leq j \leq K} p_{i} p_{j} X_{i} X_{j}]$$

$$\Rightarrow \sum_{1 \leq i \leq K} p_i^2 E[X_i^2] + 2 \sum_{1 \leq i < j \leq K} p_i p_j E[X_i X_j] = \sum_{1 \leq i \leq K} p_i^2 E[X_i^2] + 2 \sum_{1 \leq i < j \leq K} p_i p_j E[X_i] E[X_j]$$

 $E[X_iX_j] = E[X_i]E[X_j]$ is possible because for $i \neq j, X_i, X_j$ are **independent** random variables.

Using equation (1) and (3)...

$$\Rightarrow E[(\sum_{i} p_{i} X_{i})^{2}] = \sum_{1 \leq i \leq K} p_{i}^{2} (\sigma_{i}^{2} + \mu_{i}^{2}) + 2 \sum_{1 \leq i < j \leq K} p_{i} p_{j} \mu_{i} \mu_{j} = \sum_{1 \leq i \leq K} [p_{i}^{2} \sigma_{i}^{2}] + (\sum_{1 \leq i \leq K} p_{i}^{2} \mu_{i}^{2} s + 2 \sum_{1 \leq i < j \leq K} p_{i} p_{j} \mu_{i} \mu_{j})$$

$$\Rightarrow E[(\sum_{i} p_{i} X_{i})^{2}] = \sum_{1 \leq i \leq K} [p_{i}^{2} \sigma_{i}^{2}] + (\sum_{i} p_{i} \mu_{i})^{2} = \sum_{1 \leq i \leq K} [p_{i}^{2} \sigma_{i}^{2}] + \mu^{2}$$

Finally...

$$Var(Z) = E[(\sum_{i} p_{i}X_{i})^{2}] - \mu^{2} = \sum_{1 \leq i \leq K} [p_{i}^{2}\sigma_{i}^{2}] + \mu^{2} - \mu^{2} = \sum_{1 \leq i \leq K} p_{i}^{2}\sigma_{i}^{2}$$

$\mathbf{MGF}(\mathbf{Z}) \ \Phi_Z(t)$:

Here we will make use of the fact that for r.v. $A = \sum_i A_i$ where A_i s are independent r.v., then...

$$\Phi_A(t) = E[e^{tA}] = E[e^{t\sum_i A_i}] = E[\prod_i e^{tA_i}] = \prod_i E[e^{tA_i}] = \prod_i \Phi_{A_i}(t)$$

As A_i s are independent r.v.s, then e^{tA_i} s are also independent... that's why they're expectation can be split into individual expectations.

Also, as A_i s are independent, p_iA_i s are also independent... So for A=Z take $A_i=p_iX_i$, hence...

$$\Phi_Z(t) = \prod_i \Phi_{p_i X_i}(t)$$

Also, $MGF(p_iX_i)$ is :

$$\Phi_{p_i X_i}(t) = E[e^{(p_i t) X_i}] = e^{\mu_i p_i t + \frac{1}{2} \sigma_i^2(p_i t)^2}$$

$$\Rightarrow \Phi_Z(t) = \prod_i e^{\mu_i p_i t + \frac{1}{2} \sigma_i^2 (p_i t)^2} = e^{\sum_i \mu_i p_i t + \frac{1}{2} \sigma_i^2 (p_i t)^2} = e^{\sum_i (\mu_i p_i) t + \frac{1}{2} \sum_i (p_i^2 \sigma_i^2) t^2}$$

And from results of mean and variance we calculated above...

$$\Rightarrow \Phi_Z(t) = e^{\mu t + \frac{1}{2}\sigma^2 t^2}$$

This form of MGF is exactly same as of a Gaussian Variable with mean $= \mu$ and variance $= \sigma^2 = \sum_i p_i^2 \sigma_i^2$, in other words, Z is also a Gaussian Variable : $Z = \mathcal{N}(\mu, \sigma^2)$.

PDF(Z):

Now that we know that Z is also a Gaussian Random Variable, we can calculate PDF for it easily.. it comes from the definition of the Gaussian itself...

$$PDF(Z) = f_Z(z) = \frac{1}{\sigma\sqrt{2\pi}}e^{-\frac{(z-\mu)^2}{2\sigma^2}}$$

Where $\mu = \sum_i p_i \mu_i$ and $\sigma^2 = \sum_i p_i^2 \sigma_i^2$ as we had calculated above.