Ourestion (2) (a) We have  $V_{\hat{e}} = F^{-1}(u_{\hat{e}})$  for some  $F^{-1}$ V = F<sup>-1</sup>(U), Vis the random variable generating values {vi}i=s F-tenisto hence Fis bijective function—oneto one onto. Brokerty of Uniform (0,1) U, for 0 < p < 1, p = p ( U < p) Hence consider  $F(y) = P(U \le F(y)), F(y)$  E[0,1]as F(y) represents probability Since F-+ is byjective, U = F(y) € F(y)) = F+(F(y)) Ex Ex(n) EX EX  $F(y) = P(V \in y)$  $= F_{V}(y)$ >> Fv = F or distribution of V is same as that of F.

(b) Before Poroceedingly, let's mole that

O= F=1 (covers are values between othern) as  $F(y) = P((Y) \leq y)$ and F is an increasing function (can be constant at hards also) as  $[-\infty] \int_{-\infty}^{4} f_{y}(y) dy$  and  $f_{y}(y) > 0$ Noue for an increasing function f 0 = b => f(a) = f(b) & hence 1(a = b)  $= 1(f(a) \in f(b))$ as  $a \le b \ 8 \ f(a) \le f(b)$  will be both true or both false at the same time. and hence  $P(Y \in y) = P(F(Y) \in F(y))$ Name F(y) = Z., Z (TO, 1) when  $P(F(Y) \leq Z) = Z \forall Z \in [0,1].$ This is same for a uniform (0,1) v.v.  $P(U \leq z) = z + z \in [0,1]$ hence F(Y) and The Unif (0,1) have same CDFs (distribution) and are equivalent. We will use this fact later in steps.

 $P(D > d) = P\left(\max_{x} \left| \sum_{i=1}^{n} 1(Y_i \leq x) - F(x) \right| > d \right)$ as discussed earlier,  $1(Y_i \leq x) = 1(F(Y_i) \leq F(x))$  $= \rho \left( \max_{x} \left| \frac{\sum_{i=1}^{n} 1(F(Y_i) \in F(x))}{\sum_{i=1}^{n} 1(F(Y_i) \in F(x))} - F(x) \right| \right)$ Substitute F(x) = y, as  $0 \le F(x) \le 1,0 \le y \le 1$  $= P\left(\max_{0 \in \mathcal{Y} \subseteq I} \left| \frac{\sum_{i=1}^{n} 1(F(\mathcal{Y}_{i}) \in \mathcal{Y})}{\sum_{i=1}^{n} 1(F(\mathcal{Y}_{i}) \in \mathcal{Y})} - \mathcal{Y}\right| \mathcal{Y}_{i}d\right)$ From my claim earlier, F (Yi) will be acting as a new uniform random variable Onif (0,1) hence F(Yi) = Unif (0,1) = Ui(say) = P(maxosys1 = 1(Uisy) = P(E>,d) So the fremions step is justified only cherause if I take n- undependent uniform variables samples from [0,1] & those from F(Yi)s, they will have same probability for same event (say event being zd), as F(Yi)s are acting as a new uniform variable, completely randomised now & have same CDFs as a separate uniform vandom variable

and hence the chances of both quantities E & D (> d is same. Taking F(Yi) sample is like taking frist lot of n- independent uniform samples and taking Vi sample is like taking another lot of nundependent samples, & prob. for any event for these 2 lots will be same (due to randomness, can't differentiate which is which). Hence P(D>d) = P(E>d) proved. Importance of this equalion > P(D>d) is same for whatever distribution Fis chosen. It means it is undefendent cof that. F. distribution. This can be useful in checking if the data guien for a distribution F indeed is of Fior mot? If the data belonged to F, then P(D>d) must not vary too much from P(E)d) computed for a uniform (0,1) variable. If it values too much, then we can

know that the data does not fit F very well. Hence it forms a Goodness-to-Fit Itst.