

Chapter 3 Part 2 Group Work

SOLUTIONS

Problem 1

Some biology students were checking eye color in a large number of fruit flies. For the individual fly, suppose that the probability of white eyes is $1/4$ and the probability of red eyes is $3/4$, and that we may treat these observations as independent Bernoulli trials. What is the probability that at least four flies have to be checked for eye color to observe a white-eyed fly?

Solution

$Y \sim \text{geom}(p = 0.25)$, where Y is the number of the trial where the first success occurs.

$$P(Y > 3) = (1 - p)^3 = (.75)^3 = 0.421875$$

$$\text{Alternatively, } P(Y \geq 4) = 1 - P(Y < 4) = 1 - p(1) - p(2) - p(3) = 1 - .25 - .75 * .25 - .75^2 * .25 = 0.421875$$

Problem 2

Suppose that Y is a random variable with a geometric distribution. Show that

- $\sum_{y=1}^{\infty} p(y) = \sum_{y=1}^{\infty} q^{y-1}p = 1$
- $\frac{p(y)}{p(y-1)} = q$, for $y = 2, 3, \dots$. This ratio is less than 1, implying that the geometric probabilities are monotonically decreasing as a function of y . If Y has a geometric distribution, what value of Y is the most likely (has the highest probability)?

Solution

Part a

$$\sum_{y=1}^{\infty} pq^{y-1} = p \sum_{y-1=0}^{\infty} q^{y-1} = p \frac{1}{1-q} = \frac{p}{1-(1-p)} = \frac{p}{p} = 1$$

Part b Since $p(y)$ is monotonically decreasing, and the support of Y is $1, 2, 3, \dots$, then the largest probability is when $Y = 1$.

Problem 3

About 7 months into Donald Trump's 2nd term as president (August 2025), a Gallup poll found that a record low of 39% of adults approved of how the Supreme Court is handling its job.

- Find the probability distribution for Y , the number of calls until the first person is found who *does* express approval of the U.S. Supreme Court.

- b. On average, how many calls are needed until the 1st approval is found?
- c. Find the probability distribution for Z , the number of calls until the 50th person is found who approves of the U.S. Supreme Court.
- d. On average, how many calls are needed until the 50th approval is found?

Solution

Part a $Y \sim \text{geom}(p = 0.39) \implies p(y) = (.61)^{y-1}(.39)$

Part b $E(Y) = \frac{1}{p} = \frac{1}{.39} = 2.56$

Part c $Z \sim \text{nbinom}(0.39, 50) \implies p(z) = \binom{z-1}{49} (.39)^{50} (.61)^{z-50}$

Part d $E(Z) = \frac{r}{p} = \frac{50}{.39} = 128.2$

Problem 4

The employees of a firm that manufactures insulation are being tested for indications of asbestos in their lungs. The firm is requested to send three employees who have positive indications of asbestos on to a medical center for further testing. If 40% of the employees have positive indications of asbestos in their lungs, find the probability that 10 employees must be tested in order to find three positives.

Solution

$Y \sim \text{nbinom}(r = 3, p = 0.4)$

$P(Y = 10) = \binom{9}{2} (0.4)^3 (0.6)^7 = 0.06449$

Problem 5

A jury of 6 persons was selected from a group of 20 potential jurors, of whom 8 were Black and 12 were White. The jury was supposedly randomly selected, but it contained only 1 Black member. Do you have any reason to doubt the randomness of the selection?

Solution

$$\frac{\binom{8}{1} \binom{12}{5}}{\binom{20}{6}} = \frac{8 * 792}{38760} = 0.163$$

There is about a 16% chance of this happening, which is not particularly rare. There is not strong evidence to doubt the randomness of selection in this case.

Problem 6

The number of typing errors made by a typist has a Poisson distribution with an average of four errors per page. If more than four errors appear on a given page, the typist must retype the whole page. What is the probability that a randomly selected page does not need to be retyped?

Solution

Unit of interest = 1 page. $Y \sim \text{Poisson}(\lambda = 4)$.

$$P(Y \leq 4) = \sum_{y=0}^4 \frac{4^y e^{-4}}{y!} = e^{-4} [4^0/0! + 4/1! + 4^2/2! + 4^3/3! + 4^4/4!] = 0.6288$$

An example of streamlining this calculation in R:

```
y <- 0:4
lambda <- 4
p_y <- exp(-4)*4^y/factorial(y)
p_y
```

```
## [1] 0.01831564 0.07326256 0.14652511 0.19536681 0.19536681
```

```
sum(p_y)
```

```
## [1] 0.6288369
```

Problem 7

Let Y be a random variable with probability distribution $p(y) = \frac{y}{6}$, $y = 1, 2, 3$

- Find an expression for the moment generating function of Y . That is, write $E(e^{tY})$ as a sum.
- Use the mgf to show that $E(Y) = 7/3$
- Use the mgf to show that $E(Y^2) = 6$
- Find $V(Y)$

Solution

part a

$$\begin{aligned} E(e^{tY}) &= \sum e^{ty} f(y) \\ &= \sum e^{ty} \left(\frac{y}{6} \right) \end{aligned}$$

part b Recall that $E(Y)$ is the “first moment.” So we need to find the first derivative of the mgf and evaluate it at $t = 0$

$$\begin{aligned}
m(t) &= \sum_{y=1}^3 e^{ty} \left(\frac{y}{6} \right) \\
m'(t) &= \sum_{y=1}^3 e^{ty}(y) \left(\frac{y}{6} \right) \\
E(y) = m'(0) &= \sum_{y=1}^3 e^{0y}(y) \left(\frac{y}{6} \right) \\
&= \sum_{y=1}^3 \frac{y^2}{6} \\
&= \frac{1}{6} \sum_{y=1}^3 y^2 \\
&= \frac{1}{6} [1^2 + 2^2 + 3^2] \\
&= \frac{14}{6} \\
&= \frac{7}{3}
\end{aligned}$$

part c Recall that $E(y^2)$ is the “second moment” so we need to take the second derivative of $m(t)$ and evaluate it at 0.

$$\begin{aligned}
m'(t) &= \sum_{y=1}^3 e^{ty} \left(\frac{y^2}{6} \right) \\
m''(t) &= \sum_{y=1}^3 e^{ty}(y) \left(\frac{y^2}{6} \right) \\
m''(0) &= \sum_{y=1}^3 e^{0y}(y) \left(\frac{y^2}{6} \right) \\
E(y^2) = m''(0) &= \sum_{y=1}^3 \frac{y^3}{6} \\
&= \frac{1^3}{6} + \frac{2^3}{6} + \frac{3^3}{6} \\
&= 6
\end{aligned}$$

part d

$$V(y) = E(y^2) - [E(y)]^2 = 6 - \left(\frac{7}{3} \right)^2 = \frac{5}{9}$$

Problem 8

Obtain an expression for the mgf of the Poisson distribution. Use the mgf to show that $E(Y) = V(Y) = \lambda$ for a Poisson random variable.

Solution

Let Y be Poisson with parameter $\lambda > 0$. The probability distribution is

$$p(y) = \frac{e^{-\lambda} \lambda^y}{y!}, \quad y = 0, 1, 2, \dots$$

The moment generating function is defined as

$$\begin{aligned} m(t) &= E(e^{tY}) = \sum_{y=0}^{\infty} e^{ty} p(y) \\ &= \sum_{y=0}^{\infty} e^{ty} \frac{e^{-\lambda} \lambda^y}{y!} \\ &= e^{-\lambda} \sum_{y=0}^{\infty} \frac{(\lambda e^t)^y}{y!} \\ &= e^{-\lambda} \exp(\lambda e^t) \\ &= \exp(\lambda(e^t - 1)). \end{aligned}$$

Take the first derivative to find the mean

$$M'_Y(t) = \lambda e^t \exp(\lambda(e^t - 1)).$$

Evaluate at $t = 0$:

$$E(Y) = M'_Y(0) = \lambda.$$

Take the second derivative to find $E(Y^2)$:

$$m''(t) = \lambda e^t \exp(\lambda(e^t - 1)) (1 + \lambda e^t).$$

Evaluate at $t = 0$:

$$M''_Y(0) = \lambda(1 + \lambda).$$

Thus,

$$\text{Var}(Y) = m''(0) - (m'(0))^2 = \lambda(1 + \lambda) - \lambda^2 = \lambda.$$