# **HW 02 SOLUTIONS**

#### **Practice Problems**

#### 2.37

- a. There are 6! = 720 possible itineraries.
- b. In the 720 orderings, exactly 360 have Denver before San Francisco and 360 have San Francisco before Denver. So, the probability is 0.5.

#### 2.41

If the first digit cannot be zero, there are 9 possible values. For the remaining six digits, there are 10 possible values each. Thus, the total number is  $9 \cdot 10^6$ .

#### 2.43

The number of ways to choose 3 objects from 9, 6, 5, and 1 is  $\binom{9}{3}\binom{6}{5}\binom{1}{1} = 504$  ways.

#### 2.51

There are  $\binom{50}{3} = 19,600$  ways to choose 3 winners. Each of these is equally likely.

- a. There are  $\binom{4}{3}=4$  ways for the organizers to win all of the prizes. The probability is  $\frac{4}{19600}$ .
- b. There are  $\binom{4}{2}\binom{46}{1}=276$  ways the organizers can win two prizes and one of the other 46 people to win the third prize. So, the probability is  $\frac{276}{19600}$ .

- c. There are  $\binom{4}{1}\binom{46}{2}=4140$  ways. The probability is  $\frac{4140}{19600}$
- d. There are  $\binom{46}{3}=15,180$  ways. The probability is  $\frac{15180}{19600}$ .

## 2.57

There are  $\binom{52}{2}=1326$  ways to draw two cards from the deck. The probability is  $\frac{4\cdot12}{1326}=0.0362$ .

## 2.59

There are  $\binom{52}{5} = 2,598,960$  ways to draw five cards from the deck.

- a.  $\binom{4}{1}^5 = 1024$  ways. So, the probability is  $\frac{1024}{2,598,960} = 0.000394$ .
- b. There are 9 different types of "straight" hands. So, the probability is  $\frac{9.45}{2,598,960} = 0.00355$ . Note that this includes "straight flush" and "royal straight flush" hands.

## 2.61

a. 
$$\frac{364(364)(364)\cdots(364)}{365^n} = \frac{364^n}{365^n}$$

b. With 
$$n=253,$$
 
$$1-\left(\frac{364}{365}\right)^{253}\approx 0.5005.$$

## 2.71

a. 
$$P(A|B) = \frac{0.1}{0.3} = \frac{1}{3}$$
.

b. 
$$P(B|A) = \frac{0.1}{0.5} = \frac{1}{5}$$
.

c. 
$$P(A|A \cup B) = \frac{0.5}{0.5 + 0.3 - 0.1} = \frac{5}{7}$$
.

d.  $P(A|A \cap B) = 1$ , since A has occurred.

e. 
$$P(A \cap B|B) = \frac{0.1}{0.5 + 0.3 - 0.1} = \frac{1}{7}$$
.

### 2.75

- a. Given the first two cards drawn are spades, there are 11 spades left in the deck. The probability is  $\frac{\binom{11}{3}}{\binom{50}{3}} = 0.0084$ . Note: this is also equal to  $P(S_3S_4S_5|S_1S_2)$ .
- b. Given the first three cards drawn are spades, there are 10 spades left. The probability is  $\frac{\binom{10}{2}}{\binom{49}{2}}=0.0383$ . Note: this is also equal to  $P(S_4S_5|S_1S_2S_3)$ .
- c. Given the first four cards drawn are spades, there are 9 spades left. The probability is  $\frac{\binom{9}{1}}{\binom{48}{1}} = 0.1875. \text{ Note: this is also equal to } P(S_5|S_1S_2S_3S_4).$

#### 2.79

If A and B are mutually exclusive,  $P(A \cap B) = 0$ . But, P(A)P(B) > 0. So they are not independent.

#### 2.85

$$P(A\mid \overline{B}) = P(A\cap \overline{B})/P(\overline{B}) = \frac{P(\overline{B}\mid A)P(A)}{P(\overline{B})} = \frac{[1-P(B\mid A)]P(A)}{P(\overline{B})} = \frac{[1-P(B)]P(A)}{P(\overline{B})} = \frac{P(\overline{B})P(A)}{P(\overline{B})} = \frac{P(\overline{B})P(A)}{P(\overline{B})} = \frac{P(B)P(A)}{P(B)} = \frac{P(B)P(B)}{P(B)} = \frac{P(B)P(B)}{$$

$$P(\overline{B} \mid \overline{A}) = P(\overline{B} \cap \overline{A})/P(\overline{A}) = \frac{P(\overline{A} \mid \overline{B})P(\overline{B})}{P(\overline{A})} = \frac{[1 - P(A \mid \overline{B})]P(\overline{B})}{P(\overline{A})}.$$
 From the above,  $A$  and  $\overline{B}$  are independent. So  $P(\overline{B} \mid \overline{A}) = \frac{[1 - P(A)]P(\overline{B})}{P(\overline{A})} = \frac{P(\overline{A})P(\overline{B})}{P(\overline{A})} = P(\overline{B}).$  So,  $\overline{A}$  and  $\overline{B}$  are independent.

#### 2.91

If A and B are mutually exclusive,  $P(A \cup B) = P(A) + P(B)$ . This value is greater than 1 if P(A) = 0.4 and P(B) = 0.7. So they cannot be mutually exclusive. It is possible if P(A) = 0.4 and P(B) = 0.3.

# **Submitted Problems**

## 2.38

By the mn rule,  $4 \cdot 3 \cdot 4 \cdot 5 = 240$ .

#### 2.42

There are three different positions to fill using ten engineers. Then, there are  $_{10}P_3=\frac{10!}{(10-3)!}=\frac{10!}{7!}=720$  different ways to fill the positions.

## 2.46

There are  $\binom{10}{2}$  ways to choose two teams for the first game,  $\binom{8}{2}$  for the second, etc. So,  $\binom{10}{2}\binom{8}{2}\binom{6}{2}\binom{4}{2}\binom{2}{2}=\frac{10!}{(2!)^5}=113,400$  ways to assign the ten teams to five games.

## 2.48

Same answer:  $\binom{8}{5} = \binom{8}{3} = 56$ .

#### 2.50

Two numbers, 4 and 6, are possible for each of the three digits. So, there are  $2 \cdot 2 \cdot 2 = 8$  potential winning three-digit numbers.

#### 2.58

There are  $\binom{52}{5} = 2{,}598{,}960$  ways to draw five cards from the deck.

- a. To draw three Aces and two Kings:  $\binom{4}{3}\binom{4}{2}=24$  So the probability is  $\frac{24}{2,598,960}$ .
- b. There are  $13 \cdot 12 = 156$  types of full house hands. From part (a), each type can be made in 24 different ways. So the probability is  $\frac{156 \cdot 24}{2,598,960} \approx 0.00144$

4

,---,---

#### 2.64

$$6! \left(\frac{1}{6}\right)^5 = \frac{5}{324}$$

## 2.68a-c

a.  $\binom{n}{n} = \frac{n!}{n!(n-n)!} = 1$ . There is only one way to choose all of the items.

b.  $\binom{n}{0} = \frac{n!}{0!(n-0)!} = 1$ . There is only one way to choose none of the items.

c.  $\binom{n}{r} = \frac{n!}{r!(n-r)!} = \frac{n!}{(n-r)!(n-(n-r)!)} = \binom{n}{n-r}$ . There are the same number of ways to choose r out of n objects as there are to choose n-r out of n objects.

#### 2.72

Note that P(A)=0.6 and  $P(A\mid M)=\frac{0.24}{0.4}=0.6$ . So, A and M are independent. Similarly,  $P(A'\mid F)=\frac{0.24}{0.6}=0.4=P(A')$ , so A' and F are independent.

#### 2.74

a. P(A) = 0.61, P(D) = 0.30,  $P(A \cap D) = 0.20$ . Dependent.

b.  $P(B) = 0.30, P(D) = 0.30, P(B \cap D) = 0.09$ . Independent.

c.  $P(C) = 0.09, P(D) = 0.30, P(C \cap D) = 0.01$ . Dependent.

## 2.80

If  $B \subset A$ , then  $P(A \cap B) = P(A) \neq P(A)P(B)$ , unless B = S (in which case P(B) = 1).

#### 2.86

a. No. It follows from  $P(A \cup B) = P(A) + P(B) - P(A \cap B) \leq 1.$ 

b.  $P(A \cap B) \ge 0.5$ 

c. No.

d.  $P(A \cap B) \le 0.70$ 

## 2.90

a. (1/50)(1/50) = 0.0004

b.  $P(\text{at least one injury})=1-P(\text{no injuries in 50 jumps})=1-\left(\frac{49}{50}\right)^{50}\approx 0.636.$  Your friend is not correct.

## 2.94

Define the events A: device A detects smoke, B: device B detects smoke.

a. 
$$P(A \cup B) = P(A) + P(B) - P(A \cap B) = 0.95 + 0.90 - 0.88 = 0.97$$

b.\$ P(smoke undetected) = 1 - P(A B) = 1 - 0.97 = 0.03\$

## 2.110

Define the events:

• *I*: item is from line I

• II: item is from line II

• N: item is not defective

Then,  $P(N) = P(N \cap (I \cup II)) = P(N \cap I) + P(N \cap II) = 0.92(0.4) + 0.90(0.6) = 0.908$