Lecture 5

0.1 Independence of More Than Two Events

Definition 0.1. Three events A, B, and C in some probability space (Ω, \mathbb{P}) are said to be independent if the following four equations hold:

a.
$$\mathbb{P}(A \cap B) = \mathbb{P}(A) \cdot \mathbb{P}(B)$$

b. $\mathbb{P}(A \cap C) = \mathbb{P}(A) \cdot \mathbb{P}(C)$
c. $\mathbb{P}(B \cap C) = \mathbb{P}(B) \cdot \mathbb{P}(C)$
d. $\mathbb{P}(A \cap B \cap C) = \mathbb{P}(A) \cdot \mathbb{P}(B) \cdot \mathbb{P}(C)$

Remark 0.2. For each of the four equations in Definition 0.1, one can construct a probability space and define three events such that this equation does not hold but the other three do. This is left as an exercise.

Below is another definition of the independence of three events; we will later show that it is equivalent to Definition 0.1.

Definition 0.3. Three events A, B, and C in some probability space (Ω, \mathbb{P}) are said to be independent if the following eight equations hold:

1.
$$\mathbb{P}(A \cap B \cap C) = \mathbb{P}(A) \cdot \mathbb{P}(B) \cdot \mathbb{P}(C)$$
2.
$$\mathbb{P}(A \cap B \cap C^c) = \mathbb{P}(A) \cdot \mathbb{P}(B) \cdot \mathbb{P}(C^c)$$
3.
$$\mathbb{P}(A \cap B^c \cap C) = \mathbb{P}(A) \cdot \mathbb{P}(B^c) \cdot \mathbb{P}(C)$$
4.
$$\mathbb{P}(A^c \cap B \cap C) = \mathbb{P}(A^c) \cdot \mathbb{P}(B) \cdot \mathbb{P}(C)$$
5.
$$\mathbb{P}(A \cap B^c \cap C^c) = \mathbb{P}(A) \cdot \mathbb{P}(B^c) \cdot \mathbb{P}(C^c)$$
6.
$$\mathbb{P}(A^c \cap B \cap C^c) = \mathbb{P}(A^c) \cdot \mathbb{P}(B) \cdot \mathbb{P}(C^c)$$
7.
$$\mathbb{P}(A^c \cap B^c \cap C) = \mathbb{P}(A^c) \cdot \mathbb{P}(B^c) \cdot \mathbb{P}(C)$$
8.
$$\mathbb{P}(A^c \cap B^c \cap C^c) = \mathbb{P}(A^c) \cdot \mathbb{P}(B^c) \cdot \mathbb{P}(C^c)$$

Before proving the equivalence of these two definitions, we extend both of them to n events.

Definition 0.4. [Independent Events – First definition] Events A_1, \ldots, A_n are said to be independent if for every $\mathcal{I} \subseteq \{1, \ldots, n\}$ of size at least 2 it holds that

$$\mathbb{P}\left(\bigcap_{i\in\mathcal{I}}A_i\right)=\prod_{i\in\mathcal{I}}\mathbb{P}\left(A_i\right)$$

Definition 0.5. [Independent Events – Second definition] Events A_1, \ldots, A_n are said to be independent if

$$\mathbb{P}\left(\bigcap_{i=1}^{n} B_{i}\right) = \prod_{i=1}^{n} \mathbb{P}\left(B_{i}\right),$$

for all choices of $B_i \in \{A_i, A_i^c\}$.

Theorem 0.6. Definition 0.4 and Definition 0.5 are equivalent.

Proof. We will prove the assertion of the theorem for n=3, while the general case follows similar lines. We first show that Definition 0.5 implies Definition 0.4. Equation d in the first definition is the same as Equation 1 in the second definition, hence there is no need to prove it. We next prove Equation a. Observe that

$$A \cap B = (A \cap B \cap C) \cup (A \cap B \cap C^c),$$

and this union is of disjoint events. Hence

$$\mathbb{P}(A \cap B) = \mathbb{P}(A \cap B \cap C) + \mathbb{P}(A \cap B \cap C^{c})$$

$$= \mathbb{P}(A) \cdot \mathbb{P}(B) \cdot \mathbb{P}(C) + \mathbb{P}(A) \cdot \mathbb{P}(B) \cdot \mathbb{P}(C^{c})$$

$$= \mathbb{P}(A) \cdot \mathbb{P}(B) (\mathbb{P}(C) + \mathbb{P}(C^{c}))$$

$$= \mathbb{P}(A) \cdot \mathbb{P}(B).$$

By symmetry we get that Equations b and c hold as well.

We now prove that Definition 0.4 implies Definition 0.5. As in the proof of the opposite implication, Equation d in the first definition is the same as Equation 1 in the second definition, and so there is no need to prove Equation 1. We next prove Equation 2. Using Equations a and d we obtain

$$\mathbb{P}(A \cap B \cap C^c) + \mathbb{P}(A) \cdot \mathbb{P}(B) \cdot \mathbb{P}(C) = \mathbb{P}(A \cap B \cap C^c) + \mathbb{P}(A \cap B \cap C)$$
$$= \mathbb{P}(A \cap B) = \mathbb{P}(A) \cdot \mathbb{P}(B).$$

Therefore

$$\mathbb{P}\left(A \cap B \cap C^{c}\right) = \mathbb{P}\left(A\right) \cdot \mathbb{P}\left(B\right) - \mathbb{P}\left(A\right) \cdot \mathbb{P}\left(B\right) \cdot \mathbb{P}\left(C\right)$$
$$= \mathbb{P}\left(A\right) \cdot \mathbb{P}\left(B\right) \cdot \left(1 - \mathbb{P}\left(C\right)\right) = \mathbb{P}\left(A\right) \cdot \mathbb{P}\left(B\right) \cdot \mathbb{P}\left(C^{c}\right).$$

Equations 3 and 4 can be derived similarly. We next prove Equation 5. By Equation a the events A and B are independent according to the definition for two events. Hence, by a homework exercise, the events A and B^c are also independent. Combining this with Equation 3, which we already proved, and we get that

$$\mathbb{P}(A \cap B^c \cap C^c) + \mathbb{P}(A) \cdot \mathbb{P}(B^c) \cdot \mathbb{P}(C) = \mathbb{P}(A \cap B^c \cap C^c) + \mathbb{P}(A \cap B^c \cap C)$$
$$= \mathbb{P}(A \cap B^c) = \mathbb{P}(A) \cdot \mathbb{P}(B^c).$$

Therefore

$$\mathbb{P}(A \cap B^{c} \cap C^{c}) = \mathbb{P}(A) \cdot \mathbb{P}(B^{c}) - \mathbb{P}(A) \cdot \mathbb{P}(B^{c}) \cdot \mathbb{P}(C)$$
$$= \mathbb{P}(A) \cdot \mathbb{P}(B^{c}) \cdot (1 - \mathbb{P}(C)) = \mathbb{P}(A) \cdot \mathbb{P}(B^{c}) \cdot \mathbb{P}(C^{c}).$$

Equations 6, 7 and 8 can be derived similarly.

0.2 Conditional Independence

Let A, B, and C be events in some probability space (Ω, \mathbb{P}) such that $\mathbb{P}(C) > 0$. One could ask whether A and B are independent in the conditional space $(\Omega, \mathbb{P}(\cdot \mid C))$, namely, if

$$\mathbb{P}(A \cap B \mid C) = \mathbb{P}(A \mid C) \cdot \mathbb{P}(B \mid C) \tag{1}$$

holds. We will present two examples, demonstrating that independence in the conditional probability space does not necessarily imply, nor is it implied by independence in the original probability space.

Example 1:

1. Toss a fair coin twice, the two tosses being independent. Let A be the event that the outcome of the first toss is heads, let B be the event that the outcome of the second toss is heads, and let C be the event that the outcome of exactly one of the two tosses is heads. Then

$$\mathbb{P}(A \cap B) = \frac{1}{4} = \frac{1}{2} \cdot \frac{1}{2} = \mathbb{P}(A) \cdot \mathbb{P}(B),$$

that is, A and B are independent in (Ω, \mathbb{P}) .

On the other hand, denoting heads by 1 we have

$$\mathbb{P}(A \mid C) = \frac{\mathbb{P}(A \cap C)}{\mathbb{P}(C)} = \frac{\mathbb{P}(\{(1,0)\})}{\mathbb{P}(\{(1,0),(0,1)\})} = \frac{1/4}{1/2} = \frac{1}{2}$$

and, similarly,

$$\mathbb{P}\left(B\mid C\right) = \frac{1}{2}.$$

On the other hand

$$\mathbb{P}\left(A\cap B\mid C\right)=0.$$

We conclude that A and B are not independent in $(\Omega, \mathbb{P}(\cdot \mid C))$.

2. Toss a fair coin three times, all coin tosses being mutually independent. Let A be the event that the outcome of the first two tosses is heads, let B be the event that the outcome of the last two tosses is heads, and let C be the event that the outcome of the second toss is heads. Then

$$\mathbb{P}(A \cap B) = \frac{1}{8} \neq \frac{1}{4} \cdot \frac{1}{4} = \mathbb{P}(A) \cdot \mathbb{P}(B),$$

that is, A and B are not independent in (Ω, \mathbb{P}) .

On the other hand,

$$\mathbb{P}(A \mid C) = \frac{\mathbb{P}(A \cap C)}{\mathbb{P}(C)} = \frac{\mathbb{P}(A)}{\mathbb{P}(C)} = \frac{1/4}{1/2} = \frac{1}{2}$$

and, similarly,

$$\mathbb{P}\left(B\mid C\right) = \frac{1}{2}.$$

Moreover

$$\mathbb{P}(A \cap B \mid C) = \frac{\mathbb{P}(A \cap B \cap C)}{\mathbb{P}(C)} = \frac{1/8}{1/2} = \mathbb{P}(A \mid C) \cdot \mathbb{P}(B \mid C).$$

We conclude that A and B are independent in $(\Omega, \mathbb{P}(\cdot \mid C))$.

1 Random Variables

Definition 1.1 (Random Variable). Let (Ω, \mathbb{P}) be a probability space and let S be a set. A random variable over the set S is a function $X : \Omega \to S$. Whenever S is not explicitly specified, we assume that $S = \mathbb{R}$.

Example 2:

1. A fair die is rolled twice, the two rolls being independent. Let X be the sum of the two results. That is

$$\Omega = \{1, 2, 3, 4, 5, 6\}^2 = \{(i, j) : 1 \le i, j \le 6\},$$

$$\mathbb{P}((i, j)) = \frac{1}{36} \text{ for every } (i, j) \in \Omega,$$

$$X : \Omega \to \mathbb{R},$$

$$X((i, j)) = i + j.$$

2. A coin with probability 1/3 for heads is tossed 4 times, all coin tosses being mutually independent. Let Y_1 be the number of heads in the first two tosses, let Y_2 be the number of heads in the second and third tosses, and let Y_3 be the number of heads in the last two tosses. In the following we denote heads by 1 and tails by 0. Then

$$\Omega = \{0, 1\}^4,
\mathbb{P}\left((\omega_1, \omega_2, \omega_3, \omega_4)\right) = \left(\frac{1}{3}\right)^{\omega_1 + \omega_2 + \omega_3 + \omega_4} \cdot \left(\frac{2}{3}\right)^{4 - (\omega_1 + \omega_2 + \omega_3 + \omega_4)} \text{ for every } (\omega_1, \omega_2, \omega_3, \omega_4) \in \Omega,
Y_1, Y_2, Y_3 : \Omega \to \mathbb{R},
Y_1((\omega_1, \omega_2, \omega_3, \omega_4)) = \omega_1 + \omega_2,
Y_2((\omega_1, \omega_2, \omega_3, \omega_4)) = \omega_2 + \omega_3,
Y_3((\omega_1, \omega_2, \omega_3, \omega_4)) = \omega_3 + \omega_4.$$

3. Let (Ω, \mathbb{P}) be a probability space, and let $A \subseteq \Omega$ be an event. Then the *indicator of* A, denoted by 1_A , is defined as follows

$$1_{A}(\omega) = \begin{cases} 1 & \omega \in A \\ 0 & \omega \notin A \end{cases}$$