

## Assignment 3

### Solutions

**Exercise 1** A bin contains  $n$  balls, labeled with the numbers  $1, 2, \dots, n$ . Exactly  $m$  balls are drawn uniformly at random from the bin. Let  $M$  be the maximum number of a ball that was drawn.

1. Calculate the distribution of  $M$ , when the samples are being made without replacement.
2. Calculate the distribution of  $M$ , when the samples are being made independently with replacement.

#### Solution

We will calculate  $\mathbb{P}(M = k)$  for every  $1 \leq k \leq n$ .

1. In this case, it is evident that  $\mathbb{P}(M = k) = 0$  for every  $1 \leq k \leq m - 1$ . Moreover, for every  $m \leq k \leq n$ , it holds that  $M = k$  if and only if the ball labeled  $k$  is drawn and the remaining  $m - 1$  balls that are drawn have their labels in  $\{1, 2, \dots, k - 1\}$ . Hence  $|\{M = k\}| = \binom{k-1}{m-1}$ . Since the probability space is uniform, we conclude that

$$\mathbb{P}(M = k) = \frac{|\{M = k\}|}{\binom{n}{m}} = \frac{\binom{k-1}{m-1}}{\binom{n}{m}}.$$

2. For every  $1 \leq i \leq m$ , let  $X_i$  be the label of the  $i$ th ball that was drawn. Then

$$\mathbb{P}(X_i \leq k) = \frac{k}{n}$$

holds for every  $0 \leq k \leq n$  and every  $1 \leq i \leq m$ . Since the samples are independent of one another, it follows that

$$\mathbb{P}(M \leq k) = \mathbb{P}(X_i \leq k \text{ for every } 1 \leq i \leq m) = \prod_{i=1}^m \mathbb{P}(X_i \leq k) = \left(\frac{k}{n}\right)^m.$$

We conclude that

$$\mathbb{P}(M = k) = \mathbb{P}(M \leq k) - \mathbb{P}(M \leq k - 1) = \left(\frac{k}{n}\right)^m - \left(\frac{k-1}{n}\right)^m$$

holds for every  $1 \leq k \leq n$ .

**Exercise 2** A machine  $M$  is capable of sampling from  $\{0, 1\}$  such that  $\mathbb{P}(M = 1) = p$  and  $\mathbb{P}(M = 0) = 1 - p$  for some **unknown**  $p \in (0, 1)$ . For every positive integer  $n$ , let  $(L_n, R_n) \leftarrow M^2$  (i.e., we sample pairs of bits), be sampled independently of one another, and independently of all other samples. Define the algorithm  $A$  as follows:  $A$  will sample  $(L_n, R_n)$  until the first time  $L_n \neq R_n$ , and will then output the left element. Prove that  $A$  will output 1 with probability  $1/2$ .

### Solution

We abuse notation and let  $A$  denote both the algorithm and its output. For every positive integer  $n$  it holds that

$$\mathbb{P}(L_n = R_n) = \mathbb{P}(L_n = 0, R_n = 0) + \mathbb{P}(L_n = 1, R_n = 1) = (1 - p)^2 + p^2.$$

Hence

$$\mathbb{P}(A = 1) = \sum_{n=1}^{\infty} \left( \prod_{i=1}^{n-1} \mathbb{P}(L_i = R_i) \cdot \mathbb{P}(L_n = 1, R_n = 0) \right) = \sum_{n=1}^{\infty} (p^2 + (1 - p)^2)^{n-1} \cdot p(1 - p).$$

Similarly

$$\mathbb{P}(A = 0) = \sum_{n=1}^{\infty} \left( \prod_{i=1}^{n-1} \mathbb{P}(L_i = R_i) \cdot \mathbb{P}(L_n = 0, R_n = 1) \right) = \sum_{n=1}^{\infty} (p^2 + (1 - p)^2)^{n-1} \cdot p(1 - p).$$

In particular  $\mathbb{P}(A = 1) = \mathbb{P}(A = 0)$ . Since, clearly,  $\{A = 1\}$  and  $\{A = 0\}$  are disjoint events and  $\mathbb{P}(A = 1 \vee A = 0) = 1$ , we conclude that  $\mathbb{P}(A = 1) = 1/2$ .

**Exercise 3** Let  $(\Omega, \mathbb{P})$  be a probability space and let  $X, Y : \Omega \rightarrow \mathbb{R}$  be random variables. Prove that for every  $m \in \mathbb{R}$  it holds that

$$|\mathbb{P}(X = m) - \mathbb{P}(Y = m)| \leq \mathbb{P}(X \neq Y).$$

### Solution

Let  $m \in \mathbb{R}$  be arbitrary. Then

$$\begin{aligned} |\mathbb{P}(X = m) - \mathbb{P}(Y = m)| &= |\mathbb{P}(X = m, Y = m) + \mathbb{P}(X = m, Y \neq m) - \mathbb{P}(X = m, Y = m) - \mathbb{P}(X \neq m, Y = m)| \\ &\leq \mathbb{P}(X = m, Y \neq m) + \mathbb{P}(X \neq m, Y = m) \\ &= \sum_{k \in \mathbb{R} \setminus \{m\}} (\mathbb{P}(X = m \wedge Y = k) + \mathbb{P}(Y = m \wedge X = k)) \\ &\leq \sum_{k \in \mathbb{R}} \mathbb{P}(X = k \wedge Y \neq k) \\ &= \mathbb{P}(X \neq Y), \end{aligned}$$

where the first inequality holds by the triangle inequality and the last inequality holds since we added more non-negative terms.

**Exercise 4** A library has a total of  $N$  books.  $N_1$  of the books are in English and  $N_2$  of the books are in Hebrew ( $N$  could be larger than  $N_1 + N_2$ ). Alice chooses  $n$  different books from the library uniformly at random. Let  $X_1$  be the number of books in English that Alice chose and let  $X_2$  be the number of books in Hebrew that Alice chose.

1. Calculate the distribution of  $X_1 + X_2$ .
2. After Alice returned all the books she borrowed, Bob came to the library and chose books to borrow in the following way: For every book in the library, he tossed a coin whose outcome is heads with some probability  $p \in (0, 1)$ , all coin tosses being mutually independent. He borrowed each book if and only if the outcome of the corresponding coin toss was heads. Let  $Y_1$  be the number of books in English that Bob chose and let  $Y_2$  be the number of books in Hebrew that Bob chose. Prove that the distribution of  $Y_1 + Y_2$ , conditioned on the event that Bob took exactly  $n$  books, is equal to the distribution of  $X_1 + X_2$ .

### Solution

1. Let  $X = X_1 + X_2$ . Then  $\{X = k\}$  is the event that exactly  $k$  of the books that Alice chose are either in English or in Hebrew. There are  $\binom{N_1+N_2}{k} \binom{N-N_1-N_2}{n-k}$  such choices and  $\binom{N}{n}$  ways to choose  $n$  books from the library. Since the probability space is uniform, we conclude that

$$\mathbb{P}(X = k) = \frac{\binom{N_1+N_2}{k} \binom{N-N_1-N_2}{n-k}}{\binom{N}{n}}.$$

2. Let  $Z$  be the number of books that Bob borrowed which are not in English or Hebrew, and let  $Y = Y_1 + Y_2$ . It follows by Bayes' rule that

$$\mathbb{P}(Y = k | Y + Z = n) = \frac{\mathbb{P}(Y + Z = n | Y = k) \mathbb{P}(Y = k)}{\mathbb{P}(Y + Z = n)}.$$

Note that  $Y_1 \sim \text{Bin}(N_1, p)$ , since exactly  $k$  of the books Bob borrowed will be in English if and only if in exactly  $k$  of the  $N_1$  coin tosses corresponding to the English books in the library, the outcome is heads, regardless of the outcome in the remaining  $N - N_1$  coin tosses. Similarly  $Y_2 \sim \text{Bin}(N_2, p)$ . As was proved in Lecture 7,  $Y_1 = W_1 + W_2 + \dots + W_{N_1}$  and  $Y_2 = W'_1 + W'_2 + \dots + W'_{N_2}$ , where  $W_i \sim \text{Ber}(p)$  for every  $1 \leq i \leq N_1$  and  $W'_i \sim \text{Ber}(p)$  for every  $1 \leq i \leq N_2$ . Moreover,  $W_1, \dots, W_{N_1}, W'_1, \dots, W'_{N_2}$  are mutually independent. Therefore  $Y = Y_1 + Y_2 = W_1 + W_2 + \dots + W_{N_1} + W'_1 + W'_2 + \dots + W'_{N_2} \sim \text{Bin}(N_1 + N_2, p)$ . Similarly  $Y + Z \sim \text{Bin}(N, p)$  and  $Z \sim \text{Bin}(N - N_1 - N_2, p)$ . Therefore

$$\mathbb{P}(Y + Z = n) = \binom{N}{n} p^n (1-p)^{N-n},$$

$$\begin{aligned} \mathbb{P}(Y + Z = n | Y = k) &= \frac{\mathbb{P}(Y + Z = n \wedge Y = k)}{\mathbb{P}(Y = k)} = \frac{\mathbb{P}(Z = n - k \wedge Y = k)}{\mathbb{P}(Y = k)} \\ &= \frac{\mathbb{P}(Z = n - k) \cdot \mathbb{P}(Y = k)}{\mathbb{P}(Y = k)} = \mathbb{P}(Z = n - k) \\ &= \binom{N - N_1 - N_2}{n - k} p^{n-k} (1-p)^{N-N_1-N_2-n+k}, \end{aligned}$$

$$\mathbb{P}(Y = k) = \binom{N_1+N_2}{k} p^k (1-p)^{N_1+N_2-k}.$$

We conclude that

$$\begin{aligned} \mathbb{P}(Y = k \mid Y + Z = n) &= \frac{\binom{N-N_1-N_2}{n-k} p^{n-k} (1-p)^{N-N_1-N_2-n+k} \cdot \binom{N_1+N_2}{k} p^k (1-p)^{N_1+N_2-k}}{\binom{N}{n} p^n (1-p)^{N-n}} \\ &= \frac{\binom{N_1+N_2}{k} \binom{N-N_1-N_2}{n-k}}{\binom{N}{n}} \\ &= \mathbb{P}(X = k). \end{aligned}$$

**Exercise 5** Let  $X \sim \text{Geom}(\lambda n^{-1})$ , for some real number  $\lambda \geq 0$ .

1. Calculate  $\mathbb{P}(X > k)$  for every non-negative integer  $k$ .
2. Prove that

$$\mathbb{P}(n^{-1}X > t) = \left(1 - \frac{\lambda}{n}\right)^{\lfloor tn \rfloor},$$

for all  $t \geq 0$ .

3. Conclude that

$$\lim_{n \rightarrow \infty} \mathbb{P}(n^{-1}X > t) = e^{-\lambda t},$$

for all  $t \geq 0$ .

**Solution**

1. Let  $p = \lambda n^{-1}$ . Since the events  $\{X = i\}$  and  $\{X = j\}$  are disjoint for all  $i \neq j$ , it holds that

$$\mathbb{P}(X > k) = \sum_{i=k+1}^{\infty} \mathbb{P}(X = i) = \sum_{i=k+1}^{\infty} p \cdot (1-p)^{i-1} = p \cdot \frac{(1-p)^k}{1-(1-p)} = (1-p)^k = \left(1 - \frac{\lambda}{n}\right)^k.$$

2. For every  $t \geq 0$  we have

$$\begin{aligned} \mathbb{P}(n^{-1}X > t) &= \mathbb{P}(X > tn) \\ &= \mathbb{P}(X > \lfloor tn \rfloor) \\ &= \left(1 - \frac{\lambda}{n}\right)^{\lfloor tn \rfloor}, \end{aligned}$$

where the second equality holds since the support of  $X$  consists of non-negative integers and the last equality holds by the previous part of this exercise.

3. By the previous part of this exercise

$$\mathbb{P}(n^{-1}X > t) = \left(1 - \frac{\lambda}{n}\right)^{\lfloor tn \rfloor} = \left(1 - \frac{\lambda}{n}\right)^{n(\lfloor tn \rfloor/n)}$$

holds for every  $t \geq 0$ . Therefore

$$\log \left( \mathbb{P} \left( n^{-1} X > t \right) \right) = \frac{\lfloor tn \rfloor}{n} \cdot \log \left( \left( 1 - \frac{\lambda}{n} \right)^n \right).$$

Since  $\lim_{n \rightarrow \infty} \lfloor tn \rfloor / n = t$ , it follows that

$$\lim_{n \rightarrow \infty} \log \left( \mathbb{P} \left( n^{-1} X > t \right) \right) = -\lambda t,$$

which is equivalent to

$$\lim_{n \rightarrow \infty} \mathbb{P} \left( n^{-1} X > t \right) = e^{-\lambda t}.$$