Let  $(\Omega, \mathcal{F}, P)$  be a probability space. As  $\mathcal{G}_u \subseteq \mathcal{G}_t$ , we have that

$$E[E[X_t|\mathcal{G}_t]|\mathcal{G}_u] = \int_{G_u} E[X_t|\mathcal{G}_t] dP(\mathcal{G}_u)$$

$$= \int_{G_u} \int_{G_t} X dP(\mathcal{G}_t) dP(\mathcal{G}_u)$$

$$= \int_{G_t} \int_{G_u} X dP(\mathcal{G}_u) dP(\mathcal{G}_t)$$

$$= E[X_t|\mathcal{G}_u]$$

Similarly,

$$E\left[\int_{0}^{t} E\left[Y_{s}|\mathcal{G}_{s}\right] ds \middle| \mathcal{G}_{u}\right] = E\left[\int_{0}^{t} \int_{G_{s}} Y dP(G_{s}) ds \middle| \mathcal{G}_{u}\right]$$

$$= \int_{G_{u}} \int_{0}^{t} \int_{G_{s}} Y dP(G_{s}) ds dP(G_{u})$$

$$= \int_{0}^{t} \int_{G_{u}} \int_{G_{s}} Y dP(G_{s}) dP(G_{u}) ds$$

$$= \int_{0}^{t} E\left[Y_{s}|\mathcal{G}_{u}\right] ds$$

We also know that

$$\begin{split} \max\left(\int_{G_t} X_+ dP(G_t), \int_{G_t} X_- dP(G_t)\right) &\leq \max\left(\int_{\Omega} X_+ dP, \int_{\Omega} X_- dP\right) \\ &= \max(E[X_+], E[X_-]) < \infty \end{split}$$

Therefore,  $E[|X_t||G_t] < \infty$ . Similarly, using Fubini and the steps above, we can also show that

$$\int_0^t E[|Y_s||\mathcal{G}_s]ds < \infty$$

Therefore,

$$E[E[X_t|\mathcal{G}_t]|\mathcal{G}_u] - E\left[\int_0^t E[Y_s|\mathcal{G}_s] ds \middle| \mathcal{G}_u\right]$$
$$=E[X_t|\mathcal{G}_u] - \int_0^t E[Y_s|\mathcal{G}_u] ds$$

which confirms it is indeed a martingale.

2.

To match the state equations, we have that

$$a_{i,j}^1 = iK^1, \quad s_{i,j}^1 = i^2$$

and

$$a_{i,j}^2 = rj, \quad s_{i,j}^2 = \frac{r}{K^2}(j^2 + \alpha_{21}ij)$$

and

$$\begin{split} Lf(i,j) &= a_{i,j}^1[f(i+1,j) - f(i,j)] + a_{i,j}^2[f(i,j+1) - f(i,j)] \\ &+ s_{i,j}^1[f(i-1,j) - f(i,j)] + s_{i,j}^2[f(i,j-1) - f(i,j)] \end{split}$$

Then the 2 state equations are consistent with the martingale problem:

$$f(X_t^1, X_t^2) - f(X_0^1, X_0^2) - \int_0^t Lf(X_u^1, X_u^2) du$$

which is a  $\sigma(X^1_s,X^2_s,s\leq t)\text{-martingale}.$ 

3.

We know that

$$F_{X_1 \vee X_2}(x) = (F_{X_i}(x))^2$$

Hence,

$$f_{X_1 \vee X_2}(x) = 2F_{X_i}(x)f_{X_i}(x)$$

$$= 2\frac{e^{-\frac{x^2}{2}}}{\sqrt{2\pi}} \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-\frac{u^2}{2}} du$$

$$= \frac{e^{-\frac{x^2}{2}}}{\pi} \int_{-\infty}^x e^{-\frac{u^2}{2}} du$$

Therefore, we can calculate

$$f_{X_1 \vee X_2}(0) = \frac{1}{\pi} \underbrace{\int_{-\infty}^{0} e^{-\frac{u^2}{2}} du}_{-\frac{1}{\sqrt{2\pi}}} = \frac{1}{\sqrt{2\pi}}$$
$$\int_{-\infty}^{1} f_{X_1 \vee X_2}(x) dx$$

$$\int_{0}^{1} f_{X_{1} \vee X_{2}}(x) dx$$

$$= F_{X_{1} \vee X_{2}}(1) - F_{X_{1} \vee X_{2}}(0)$$

$$= F_{X_{i}}(1)^{2} - F_{X_{i}}(0)^{2}$$

$$= 0.84134475^{2} - 0.5^{2}$$

$$= 0.45786098835$$

4.

Let  $g(B_t, t) = f\left(\frac{\sigma^2}{2}t\right)$ , hence we have that

$$\begin{split} \frac{\partial}{\partial x}g(B,x) &= \frac{\sigma^2}{2}e^{\frac{\sigma^2x}{2}}\left(f'\circ e^{\frac{\sigma^2x}{2}}\right)\\ \frac{\Delta}{2}g(B,s) &= 0 \end{split}$$

and that

$$g(B_t, t) - g(0, 0) - \int_0^t \frac{\Delta}{2} g(B_s, s) + \frac{\partial}{\partial s} g(B_s, s) ds$$

is a  $\sigma(B_u, u \leq t)$ -martingale problem. However, we have that

$$\frac{\Delta}{2}g(B_s, s) + \frac{\partial}{\partial s}g(B_s, s) = \frac{\sigma^2}{2} \exp\left(\frac{\sigma^2 s}{2}\right) f'\left(\exp\left(\frac{\sigma^2 s}{2}\right)\right)$$

Hence, if we let  $X_t = \exp\left(\frac{\sigma^2}{2}t\right)$ , then we get

$$f(X_t) - f(X_0) - \int_0^t \frac{\sigma^2 X_u}{2} f'(X_u) du$$

is a martingale. If there is another solution  $Y_t$  that has  $Y_0 = X_0$ . Let  $f(x) = \ln(x)$  so that  $f'(x) = \frac{1}{x}$  and therefore,  $X_u f'(X_u) = 1$ . Thus,

$$\ln(X_t) - \ln(X_0) - \int_0^t \frac{\sigma^2}{2} du = \ln(Y_t) - \ln(Y_0) - \int_0^t \frac{\sigma^2}{2} du$$

and hence  $\ln(X_t) = \ln(Y_t) \implies X_t = Y_t$  for all t which proves its uniqueness.