Part 4 MC and multivariate distribution

Dr. Nguyen Quang Huy

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To simulate (X_1, X_2, \dots, X_n) with distribution function $F(x_1, x_2, \dots, x_n)$:

- Calculate $F_1(x_1)$ the distribution function of X_1 ; simulate $U_1 \sim \mathcal{U}(0,1)$ and return $X_1 = F_1^{\leftarrow}(U_1)$.
- ② Calculate $F_{2|1}(x_2|x_1)$, the conditional distribution function of X_2 given X_1 ; simulate $U_2 \sim \mathcal{U}(0,1)$ and return $X_2 = F_{2|1}^{\leftarrow}(U_2)$.
- **3** Calculate $F_{3|2,1}(x_3|x_2,x_1)$, the conditional distribution function of X_3 given X_2 and X_1 ; simulate $U_3 \sim \mathcal{U}(0,1)$ and return $X_3 = F_{3|2,1}^{\leftarrow}(U_3)$.

. . .

Calculate $F_{n|n-1,...,1}(x_n|x_{n-1},...,x_1)$, the conditional distribution function of X_n given $X_{n-1},...,X_1$; simulate $U_n \sim \mathcal{U}(0,1)$ and return $X_n = F_{n|(n-1),...,1}^{\leftarrow}(U_n)$.

Example: Given (X,Y) with multivariate distribution function $F(x,y):(0,\infty)^2\to [0,1]$ and

$$F(x,y) = 1 - \frac{1}{x+1} - \frac{1}{y+1} + \frac{1}{1+x+y}$$

- Calculate the distribution F_X of X and F_Y of Y?
- \bigcirc X and Y are independent?
- 3 Calculate the conditional distribution $F_{Y|X}(y|x)$

$$F_{Y|X}(y|x) = \mathbb{P}(Y \le y|X = x) = \frac{\partial F(x,y)/\partial x}{f_X(x)}$$

3 Simulate X based on F_X then simulate Y based on $F_{Y|X}(y|x)$

1 Distribution F_X of X and F_Y of Y

1 Distribution F_X of X and F_Y of Y

$$F_X(x) = F(x, \infty) = 1 - \frac{1}{x+1} - \frac{1}{1+\infty} + \frac{1}{1+x+\infty}$$

= $1 - \frac{1}{x+1}$

Similarly,

$$F_Y(y) = 1 - \frac{1}{v+1}$$

Density functions of X and Y

$$f_X(x) = \frac{1}{(x+1)^2}$$

 $f_Y(y) = \frac{1}{(y+1)^2}$

 \bigcirc X and Y are independent

 $oldsymbol{2}$ X and Y are independent if and only if

$$F(x,y) = \mathbb{P}(X \le x, Y \le y)$$

$$= \mathbb{P}(X \le x) \mathbb{P}(Y \le y)$$

$$= F_X(x) F_Y(y) \quad \forall x, y$$

However

$$F(x,y) = 1 - \frac{1}{x+1} - \frac{1}{y+1} + \frac{1}{1+x+y}$$
$$F_X(x)F_Y(y) = \left(1 - \frac{1}{x+1}\right)\left(1 - \frac{1}{y+1}\right)$$

 $\rightarrow X$ and Y are NOT independent.

3 Calculate the conditional distribution $F_{Y|X}(y|x)$

$$F_{Y|X}(y|x) = \frac{\partial F(x,y)/\partial x}{f_X(x)}$$

We have

$$\frac{\partial F(x,y)}{\partial x} = \frac{1}{(1+x)^2} - \frac{1}{(1+x+y)^2}$$

then

$$F_{Y|X}(y|x) = \left(\frac{1}{(1+x)^2} - \frac{1}{(1+x+y)^2}\right)(1+x)^2$$
$$= 1 - \frac{(1+x)^2}{(1+x+y)^2}$$

• Calculate the inverse function of $F_X(x)$ and $F_{Y|X}(y|x)$ (a function of y given x)

• Calculate the inverse function of $F_X(x)$ and $F_{Y|X}(y|x)$ (a function of y given x)

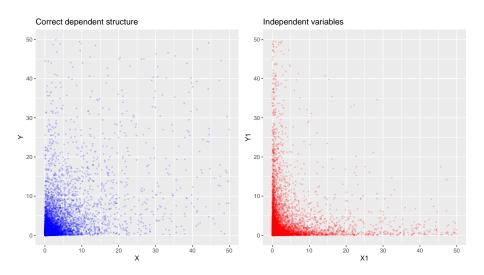
$$F_X^{\leftarrow}(t) = \frac{1}{1-t} - 1$$

$$F_{Y|X}^{\leftarrow}(t|x) = (1+x) \left(\frac{1}{\sqrt{1-t}} - 1\right)$$

To simulate (X, Y)

- Simulate $U_1 \sim \mathcal{U}(0,1)$, return $X = \frac{1}{1-U_1}-1$
- Simulate $U_2 \sim \mathcal{U}(0,1)$, return $Y = (1+X_1)\left(\frac{1}{\sqrt{1-U_2}}-1\right)$

```
N<-10<sup>3</sup>
u1 < -runif(N, 0, 1)
u2 < -runif(N.0.1)
X<-(1)/(1-u1)-1
Y < -(1+X)*(1/(1-u^2)^0.5-1)
# TNDEPENDENT X1 a.n.d. Y1
X1 < -(1)/(1-u1)-1
Y1 < -(1)/(1-u2)-1
dat<-data.frame("X"=X,"Y"=Y,"X1"=X1,"Y1"=Y1)
p1<-ggplot(dat,aes(X,Y))+geom_point(cex=0.5,alpha=0.5,color="l
  ggtitle("Correct dependent structure")
p2<-ggplot(dat,aes(X1,Y1))+geom_point(cex=0.5,alpha=0.5,color=
  ggtitle("Independent variable")
grid.arrange(p1, p2, ncol=2)
```

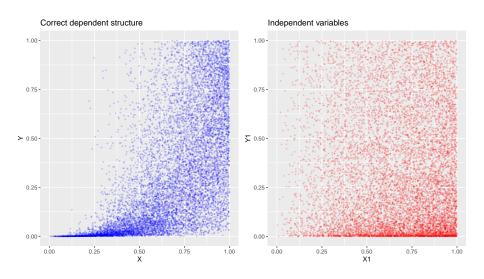


Given (X,Y) with multivariate distribution function $F(x,y):(0,1)^2\to [0,1]$ and

$$F(x,y) = (x^{-2\theta} + y^{-\theta/2} - 1)^{-1/\theta}$$

with $\theta > 0$

- Calculate $F_X(x)$, $F_{Y|X}(y|x)$
- 2 Calculate $F_X^{\leftarrow}(x)$, $F_{Y|X}^{\leftarrow}(y|x)$
- **3** With $\theta = ...$, simulate (X, Y)
- Calculate $\mathbb{P}(X + Y > 0.8)$ with $\theta = ...$



- In some cases, the simulation of dependent random variables does not need the calculation of conditional distribution, multivariate normal random variable is an example.
- Suppose that $\mathbf{N} = (N_1, N_2, \cdots, N_k)'$ is a k-dimensional normal random vector with mean μ and covariance matrix Σ

$$\mu = \begin{pmatrix} \mu_1 \\ \mu_2 \\ \dots \\ \mu_k \end{pmatrix} ; \Sigma = \begin{pmatrix} \sigma_{11} & \sigma_{12} & \cdots & \sigma_{1k} \\ \sigma_{21} & \sigma_{22} & \cdots & \sigma_{2k} \\ \dots & \dots & \dots \\ \sigma_{k1} & \sigma_{k2} & \cdots & \sigma_{kk} \end{pmatrix}$$

• We can prove that $\mathbf{N} = \mu + \mathbf{CZ}$ where $\mathbf{Z} = (Z_1, Z_2, \cdots, Z_k)'$ is the vector of k i.i.d standard normal r.v and matrix \mathbf{C} is the Cholesky decomposition of Σ : $\Sigma = \mathbf{CC}'$

Given $\mathbf{N} \sim \mathcal{N}(\mu, \Sigma)$ where

$$\mu = \left(egin{array}{c} 1 \ 1 \end{array}
ight) \; ; \; \Sigma = \left(egin{array}{cc} 4 & 4 \ 4 & 9 \end{array}
ight)$$

- ullet Find a lower triangular matrix ${f C}$ such that $\Sigma = {f CC}'$
- Using stochastic representation $N = \mu + CZ$ to simulate N.

n < -1000

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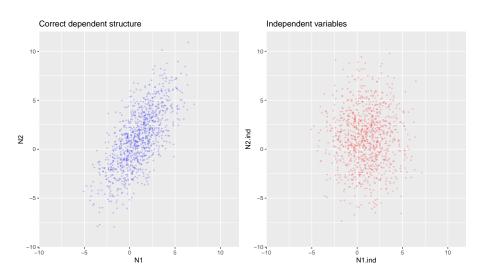
Z<-matrix(rnorm(2*n,0,1),2,n) C<-matrix(c(2,2,0,5^0.5),2,2) mu<-matrix(rep(c(1,1),n),2,n)

```
N < -mu + C % * % Z
# INDEPENDENT X1 and Y1
C1 \leftarrow matrix(c(2,0,0,3),2,2)
N1<-mu+C1 %*% Z
dat<-data.frame("N1"=N[1,],"N2"=N[2,],"N1.ind"=N1[1,],"N2.ind
p1<-ggplot(dat,aes(N1,N2))+geom_point(cex=0.5,alpha=0.2,color=
  ggtitle("Correct dependent structure")+xlim(c(-9,11))+ylim(c
p2<-ggplot(dat,aes(N1.ind,N2.ind))+geom point(cex=0.5,alpha=0
  ggtitle("Independent variables")+xlim(c(-9,11))+ylim(c(-9,1)
grid.arrange(p1, p2, ncol=2)
```

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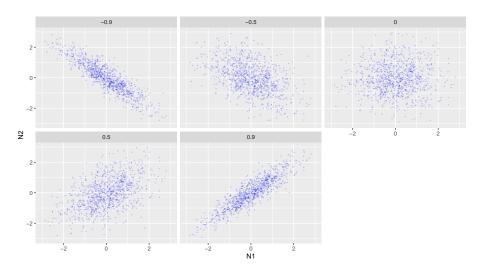
• The function *chol(M)* return the **transpose** of the cholesky decomposition matrix *M*.

```
Sigma<-matrix(c(4,4,4,9),2,2)
t(chol(Sigma))
```

```
## [,1] [,2]
## [1,] 2 0.000000
## [2,] 2 2.236068
```

• Simulate $N=(N_1,N_2)$ where N_1 and N_2 are standard normal r.v and correlation between N_1 and N_2 is ρ ; $\rho=-0.9,-0.5,0,0.5,0.9$ respectively

```
rho < -c(-0.9, -0.5, 0, 0.5, 0.9)
n<-1000
N \leftarrow matrix(0,2,5*n)
Z \leftarrow matrix(rnorm(2*n,0,1),2,n)
for (i in 1:5){
  C<-matrix(c(1,rho[i],rho[i],1),2,2)</pre>
  N[,((i-1)*n+1):(i*n)] < -t(chol(C)) %*% Z
}
dat < -data.frame("N1"=N[1,],"N2"=N[2,],
                  gr=(((1:(5*n))-1) \%/\% n + 1))
dat<-mutate(dat,cor=as.factor(rho[gr]))</pre>
dat%>%ggplot(aes(N1,N2))+
  geom point(cex=0.2,alpha=0.2,col="blue")+
  xlim(c(-3,3))+
  ylim(c(-3,3))+facet_wrap(\sim cor)
```



Bivariate Brownian motion

A 2-dimensional Brownian motion $B_t = (B_t^{(1)}, B_t^{(2)})$ is defined as follows:

- $B_0 = (0,0)$
- ② $(B_{t_2}^{(i)}-B_{t_1}^{(i)})$ and $(B_{s_2}^{(j)}-B_{s_1}^{(j)})$ are independent for all $1\leq i,j\leq 2$ and $s_1\leq s2\leq t_1\leq t_2$
- **3** $(B_{t_2}^{(1)} B_{t_1}^{(1)}, B_{t_2}^{(2)} B_{t_1}^{(2)})$ is a 2-dimensional normal random vector with mean μ and covariance matrix Σ

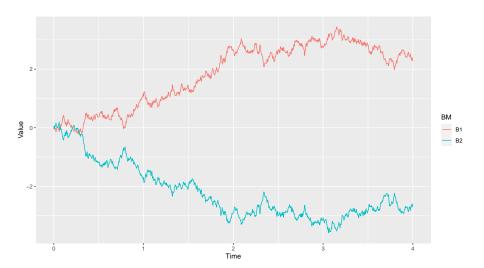
$$\mu = \left(egin{array}{c} 0 \ 0 \end{array}
ight) \; ; \; \Sigma = \left(egin{array}{cc} (t_2-t_1) &
ho(t_2-t_1) \
ho(t_2-t_1) & (t_2-t_1) \end{array}
ight)$$

where ρ is the correlation between these Brownian motions.

Bivariate Brownian motion

```
Simulate a 2-dimensional Brownian motion from 0 to T with T = 4, n =
1000 and \rho = -0.9, 0, 0.9, respectively
T<-4
n < -1000
delta<-T/n
d < -2
rho < -0.9
Bt \leftarrow matrix(0,d,(n+1))
C<-t(chol(matrix(c(delta,rho*delta,rho*delta,delta),d,d)))
for (i in 2:(n+1)){
  dBt < -C%*%rnorm(d,0,1)
  Bt[,i] \leftarrow Bt[,(i-1)] + dBt
}
dat<-data.frame("B1"=Bt[1,],"B2"=Bt[2,],Time=seq(0,T,T/n))
dat <- gather (dat, BM, Value, B1: paste0("B", d))
dat%>%ggplot(aes(Time, Value, group=BM, col=BM))+geom line()
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```

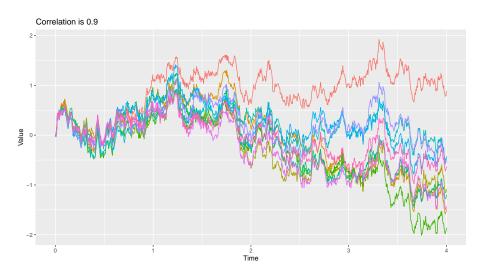
Bivariate Brownian motion



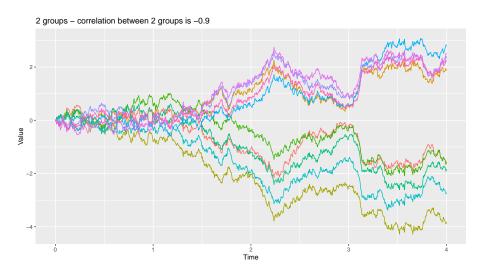
Multivariate Brownian motion

```
Simulate a d-dimensional Brownian motion from 0 to T with d = 10, T =
4, n = 1000 and \rho_{ii} = 0.9 for all 1 \le i, j \le k and i \ne j
d < -10
rho < -0.9
Bt \leftarrow matrix(0,d,(n+1))
Sigma <- matrix (rho,d,d)
Sigma < - delta * (Sigma + diag(1-rho,d,d))
C<-t(chol(Sigma))</pre>
for (i in 2:(n+1)){
  dBt < -C%*%rnorm(d.0.1)
  Bt[,i] \leftarrow Bt[,(i-1)] + dBt
}
dat<-as.data.frame(t(Bt))</pre>
names(dat) <- paste0("B",1:d)
dat<-dat%>%mutate(Time=seq(0,T,T/n))%>%
  gather(dat,BM,Value,B1:paste0("B",d))
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```

Multivariate Brownian motion



Multivariate Brownian motion



In one factor (One Brownian motion) short rate models (Vasicek, CIR) the continuously compounded spot rates R(t,T) is a function of r_t

$$P(t,T) = A(t,T) \times exp(-B(t,T)r_t)$$

$$R(t,T) = \frac{-log(P(t,T))}{(T-t)}$$

$$= a(t,T) + r_tb(t,T)$$

where a(t, T) and b(t, T) are deterministic functions. With $T_1 < T_2$

$$cor(R(t, T_1), R(t, T_2)) = cor(a(t, T_1) + r_t b(t, T_1), a(t, T_2) + r_t b(t, T_2))$$

= 1 $\forall T_1 < T_2$

 \rightarrow the price of 1-month ZC bond and the price of 30-years ZC bond are perfectly correlated!!!

The two-factor version of Vasicek short rate model (G2 model):

$$r_t = x_t + y_t$$

$$x_t = \kappa_x(\theta_x - x_t)dt + \sigma_x dB_t^{(1)}$$

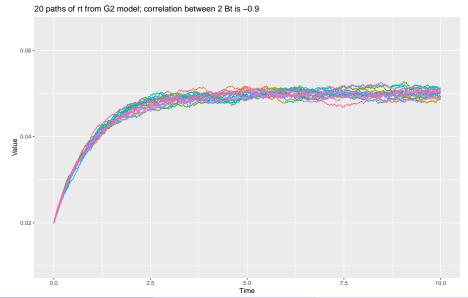
$$y_t = \kappa_y(\theta_y - y_t)dt + \sigma_y dB_t^{(2)}$$

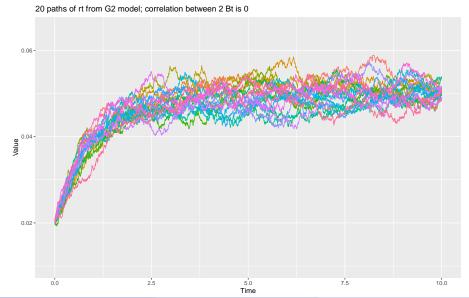
where $(B_t^{(1)}, B_t^{(2)})$ is a 2-dimensional Brownian motion with correlation ρ

• Simulate r_t from G2 model with $ho = -0.9, 0, 0.9, \ T = 10, \ n = 2500$ and

$$x_0 = 1\%, \kappa_x = 1, \theta_x = 2.5\%, \sigma_x = 0.25\%$$

 $y_0 = 1\%, \kappa_y = 1, \theta_y = 2.5\%, \sigma_y = 0.25\%$





20 paths of rt from G2 model; correlation between 2 Bt is 0.9 0.06 -Value 0.02 2.5 0.0 10.0 5.0 7.5 Time