Applied Time Series Econometrics

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December 3, 2019

SVAR with long-run restrictions: example

The most known example of the SVAR model with long-run restrictions is from the study by Blanchard and Quah (1989). Their model is a bivariate VAR for ΔQ (growth rate of output) and u_t (unemployment rate).

$$\begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix} \begin{pmatrix} \Delta Q_t \\ u_t \end{pmatrix} = \begin{pmatrix} \lambda_1 \\ \lambda_2 \end{pmatrix} + A_1 \begin{pmatrix} \Delta Q_{t-1} \\ u_{t-1} \end{pmatrix} + \dots + A_p \begin{pmatrix} \Delta Q_{t-p} \\ u_{t-p} \end{pmatrix} + \begin{pmatrix} b_{11} & b_{12} \\ b_{21} & b_{22} \end{pmatrix} \begin{pmatrix} \varepsilon_t^{as} \\ \varepsilon_t^{ad} \end{pmatrix}$$

Assmption: aggregate demand shocks have only transitory effects on the level of output. It means that the long-run effect of ε_t^{ad} on ΔQ is equal to zero:

$$\Psi(1,2) = 0$$

2 variables $\Rightarrow \Psi$ has dimension 2 \times 2.The assumption makes Ψ lower triangular.

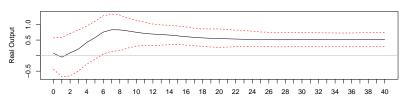
Blanchard and Quah(1989) model

```
var_bq <- VAR(bq_ts, p = 8, type = "const")</pre>
 svar bq = BQ(var bq)
 svar_bq[["B"]] # or svar_bq$B
##
              dQ
## dQ 0.07467792 -0.9296093
## u 0.21980221 0.2082399
svar_bq[["LRIM"]] # or svar_bq$LRIM
##
               dQ
## dQ 0.518604141 0.000000
## u 0.007847413 4.043684
```

IRFs for BQ(1989) model: code

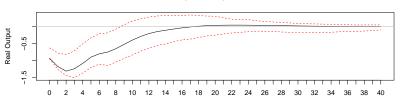
IRFs for BQ(1989) model: graph

Cumulative SVAR Impulse Response from supply shock



95 % Bootstrap CI, 100 runs

Cumulative SVAR Impulse Response from demand shock

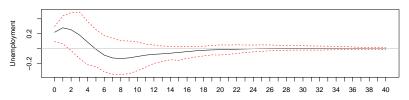


95 % Bootstrap CI, 100 runs

IRFs for BQ(1989) model: code(2)

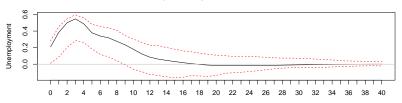
IRFs for BQ(1989) model: graph(2)

SVAR Impulse Response from supply shock



95 % Bootstrap CI, 100 runs

SVAR Impulse Response from demand shock



95 % Bootstrap CI, 100 runs

FEVD for Blanchard-Quah model

```
fevd bq <- fevd(svar bq, n.ahead = 12)
fevd bq
## $dQ
##
                  dQ
## [1,] 0.006411946 0.9935881
##
   [2,] 0.022255946 0.9777441
## [3.] 0.040914220 0.9590858
##
   [4.] 0.054972433 0.9450276
##
   [5.] 0.092693239 0.9073068
## [6,] 0.110173642 0.8898264
## [7.] 0.132653559 0.8673464
## [8.] 0.136643574 0.8633564
## [9.] 0.135503117 0.8644969
   [10,] 0.134572642 0.8654274
##
   [11,] 0.134098229 0.8659018
   [12,] 0.133771206 0.8662288
##
##
```

Sign restrictions (1)

A structural VAR model can be specified as:

$$B_0 y_t = \lambda + B_1 y_{t-1} + \dots + B_p y_{t-p} + \varepsilon_t, \qquad \varepsilon_t \sim iidN(0, \Sigma), \quad (1)$$

where Σ is diagonal. Reduced-form VAR for the model (1) is:

$$y_t = \mu + \Phi_1 y_{t-1} + \dots + \Phi_p y_{t-p} + u_t, \qquad u_t \sim iidN(0, \Omega)$$

Recursive identification implied that:

- B₀ is a lower triangular matrix,
- Σ is identity matrix.

With these assumptions we obtained:

$$\upsilon_t = B_0^{-1} \varepsilon_t$$

$$E(\upsilon_t \upsilon_t') = B_0^{-1} E(\varepsilon_t \varepsilon_t') (B_0^{-1})'$$

$$\Omega = B_0^{-1} (B_0^{-1})'$$

Sign restrictions (2)

Q: What if the structural model was different?

Alternative specification:

$$\tilde{B}_0 y_t = \tilde{\lambda} + \tilde{B}_1 y_{t-1} + \dots + \tilde{B}_p y_{t-p} + \omega_t, \qquad \omega_t \sim iidN(0, I),$$
 (2)

where $\tilde{B}_i = QB_i$, $\tilde{\lambda} = Q\lambda$ and Q is an orthogonal matrix: $Q' = Q^{-1}$ so that QQ' = Q'Q = I.

Write the reduced-form model for (2):

$$QB_0y_t = Q\lambda + QB_1y_{t-1} + \dots + QB_py_{t-p} + \omega_t$$

$$B_0y_t = \lambda + B_1y_{t-1} + \dots + B_py_{t-p} + \underbrace{Q'\omega_t}_{\varepsilon_t}$$

$$y_t = \mu + \Phi_1y_{t-1} + \dots + \Phi_py_{t-p} + u_t,$$

where
$$u_t = (QB_0)^{-1}\omega_t$$
, $\Phi_i = B_0^{-1}B_i$, $\mu = B_0^{-1}\lambda$

Sign restrictions (3)

Find covariance matrix of new reduced-form residuals:

$$E(u_t u_t') = (QB_0)^{-1} E(\omega_t \omega_t') ((QB_0)^{-1})'$$

= $B_0^{-1} Q' Q(B_0^{-1})' = B_0^{-1} (B_0^{-1})' = \Omega$

The baseline and an alternative model has the same reduced-form representation, so they are observationally equivalent with $\varepsilon_t = Q'\omega_t$.

we cannot discriminate ω_t with respect to ε_t .

Sign restrictions: example

In case of two variables the matrix Q can be written in a simple way:

$$Q = \begin{pmatrix} \cos \gamma & -\sin \gamma \\ \sin \gamma & \cos \gamma \end{pmatrix} \quad 0 \le \gamma \le \pi$$

$$\omega_t = \begin{pmatrix} \cos \gamma & -\sin \gamma \\ \sin \gamma & \cos \gamma \end{pmatrix} \begin{pmatrix} \varepsilon_{1t} \\ \varepsilon_{2t} \end{pmatrix} = \begin{pmatrix} \varepsilon_{1t} \cos \gamma - \varepsilon_{2t} \sin \gamma \\ \varepsilon_{1t} \sin \gamma + \varepsilon_{2t} \cos \gamma \end{pmatrix}$$

$$\begin{split} \tilde{B}_{0} &= QB_{0} = \\ &= \begin{pmatrix} \cos\gamma & -\sin\gamma \\ \sin\gamma & \cos\gamma \end{pmatrix} \begin{pmatrix} b_{0,1,1} & 0 \\ b_{0,2,1} & b_{0,2,2} \end{pmatrix} \\ &= \begin{pmatrix} b_{0,1,1}\cos\gamma - b_{0,2,1}\sin\gamma & -b_{0,2,2}\sin\gamma \\ b_{0,1,1}\sin\gamma + b_{0,2,1}\cos\gamma & b_{0,2,2}\cos\gamma \end{pmatrix} \end{split}$$

Sign restrictions: idea(3)

- If $\gamma = 0$, then \tilde{B}_0 is lower triangular (y_1 is "more exogenous" than y_2).
- If $\gamma = \frac{\pi}{2}$, then \tilde{B}_0 is upper triangular (y_2 is "more exogenous" than y_1).
- Other structural shocks are identified by choosing alternative values of γ within the range $0 \le \gamma \le \pi$.
- Sign restriction approach is an approach of generating alternative models using the orthonormal rotation and selecting those models that generate impulse responses consistent with economic theory.

Sign restriction algorithm

- **Step 1.** Estimate a VAR and construct VMA parameters, C_i , compute B_0^{-1} so that $B_0^{-1}(B_0^{-1})' = \Omega$.
- **Step 2.** Draw a value of γ from $0 \le \gamma \le \pi$ and compute $Q, \tilde{B}_0, \tilde{\Psi}_i$.
- **Step 3.** Compute a finite number of impulse responses (for example, IRF for 4 periods for quarterly data).
- **Step 4.** If all IRF have the correct sign, select the model, otherwise discard it.
- **Step 5.** Repeat steps 2 to 4 and generate other models that satisfy the restrictions.

Literature

- Lütkepohl, H., and Krätzig, M. (2004) Applied Time Series Econometrics, New York, Cambridge University Press
- Martin, V., Hurn, S., and HarriS,D. (2013) Econometric Modelling with Time Series: Specification, Estimating and Testing, New York, Cambridge University Press