## Lecture 5

# Classification II

ISLR 4, ESL 4



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# Bayesian Classification

#### Bayesian Methods

Bayes' formula

Probability of the input, given the output, i.e. class density

Posterior (probability of the input) 
$$\rightarrow$$
  $Pr(Y \mid X) = \frac{Pr(X \mid Y) Pr(Y)}{Pr(X)}$  Prior probability of the output

Pr(X) is a normalizing constant that only depends on the input data and often need not be computed

#### Bayesian classification for K classes

use Bayes' formula to determine posterior density per class  $Pr(Y = k \mid X = x)$ 

$$p_k(x) = \Pr(Y = k \mid X = x) = \frac{\pi_k f_k(x)}{\sum_{\ell=1}^K \pi_\ell f_\ell(x)}$$
 class density

- we compute  $p_k(x)$  by estimating the class prior probabilities  $\pi_k$  and the class densities  $f_k(X)$
- we estimate the prior class probabilities from data,  $\pi_k = \frac{1}{n} \sum_{i=1}^n I(y_i = k)$
- we somehow determine the probability density for point x for a class k
- we then classify each point to its most probable class

# Linear Discriminant Analysis

#### Model assumptions

every class is Gaussian-distributed

$$f_k(x) = \frac{1}{\sqrt{2\pi}\sigma_k} \exp\left(-\frac{1}{2\sigma_k^2}(x - \mu_k)^2\right)$$

all classes have the same variance

$$\sigma_1^2=\sigma_2^2=\cdots=\sigma_k^2=\sigma^2$$

The Bayesian classifier now becomes

$$p_k(x) = \frac{\pi_k f_k(x)}{\sum_{\ell=1}^K \pi_\ell f_\ell(x)} = \frac{\pi_k \frac{1}{\sqrt{2\pi}\sigma_k} \exp\left(-\frac{1}{2\sigma_k^2} (x - \mu_k)^2\right)}{\sum_{l=1}^K \pi_l \frac{1}{\sqrt{2\pi}\sigma^2} \exp\left(-\frac{1}{2\sigma^2} (x - \mu_l)^2\right)}$$

the logarithm of the numerator

$$-\frac{x^2}{2\sigma^2} + x \cdot \frac{\mu_k}{\sigma^2} - \frac{\mu_k^2}{2\sigma^2} + \log \pi_k - \log(\sqrt{2\pi}\sigma)$$

# Linear Discriminant Analysis

The Bayes-optimal choice is to classify x to the class with the largest discriminant

• the discriminant of a class k is the log-probability that cancels in the log odds

$$\log\left(\frac{p_k(x)}{p_l(x)}\right) = \delta_k(x) - \delta_l(x)$$

where

$$\delta_k(x) = x \cdot \frac{\mu_k}{\sigma^2} - \frac{\mu_k^2}{2\sigma^2} + \log \pi_k$$

is the log-numerator from previous slide with the class-independent terms removed

# **Example** Linear Discriminant Analysis

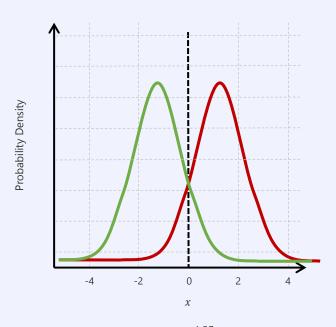
If  $\pi_1 = \pi_2$  we classify an observation x to class 1 if

$$2x(\mu_1 - \mu_2) > \mu_1^2 - \mu_2^2$$

• the Bayes decision boundary is the set of points for which both discriminants are equal, i.e.

$$x = \frac{\mu_1^2 - \mu_2^2}{2(\mu_1 - \mu_2)} = \frac{\mu_1 + \mu_2}{2}$$

- the figure shows two 1D normal density functions.
- the dashed line represents the Bayes decision boundary, at which an observation is equally likely to belong to either class



$$\mu_1 = -1.25$$
 $\mu_2 = 1.25$ 
 $\sigma_1 = \sigma_2 = 1$ 

(ESL 4.4.1)

# Fitting Univariate LDA Models

In general, we do not know the underlying class densities

we estimate these using the finite training sample

$$\hat{\mu}_k = \frac{1}{n_k} \sum_{i:y_i = k} x_i$$

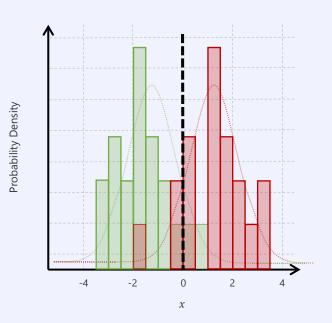
$$\hat{\sigma}^2 = \frac{1}{n - K} \sum_{k=1}^K \sum_{i:y_i = k} (x_i - \hat{\mu}_k)^2$$

$$\pi = n_k/n$$

• we assign x to the class with the largest fitted discriminant

$$\hat{\delta}_k(x) = x \cdot \frac{\hat{\mu}_k}{\hat{\sigma}^2} - \frac{\hat{\mu}_k^2}{2\hat{\sigma}^2} + \log \hat{\pi}_k$$

note that the discriminants are linear (!)



LDA fit over 20 samples per class, fitted decision boundary in dashed black. Bayes error 10.6%, LDA test error 11.1%

(ESL 4.4.1)

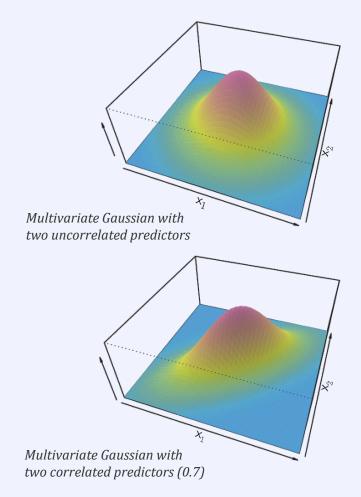
#### Model assumptions

- each class is a multivariate Gaussian
- the covariance matrix is the same for all classes

$$f_{k(x)} = \frac{1}{(2\pi)^{\frac{p}{2}} |\mathbf{\Sigma}|^{\frac{1}{2}}} \exp\left(-\frac{1}{2}(x - \mu_k)^T \mathbf{\Sigma}^{-1} (x - \mu_k)\right)$$

$$\delta_k(x) = x^T \mathbf{\Sigma}^{-1} \mu_k - \frac{1}{2} \mu_k^T \mathbf{\Sigma}^{-1} \mu_k + \log \pi_k$$

•  $\Sigma$  is the  $p \times p$  covariance matrix of the inputs  $\Sigma = \text{Cov}(x)$ 



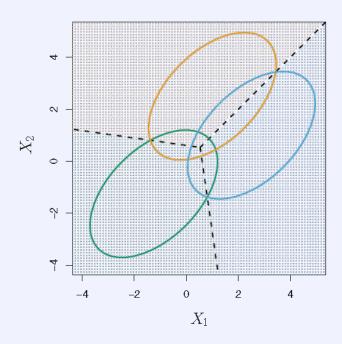
(ESL 4.4.3)

#### Model assumptions

- each class is a multivariate Gaussian
- the covariance matrix is the same for all classes

$$f_{k(x)} = \frac{1}{(2\pi)^{\frac{p}{2}} |\mathbf{\Sigma}|^{\frac{1}{2}}} \exp\left(-\frac{1}{2}(x - \mu_k)^T \mathbf{\Sigma}^{-1} (x - \mu_k)\right)$$
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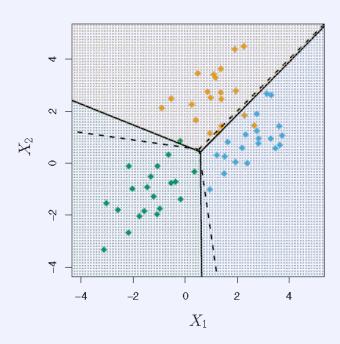
2D synthetic data example with three classes. Ellipses contain 95% of the class probability mass, the Bayes decision boundaries are dashed

#### Model assumptions

- each class is a multivariate Gaussian
- the covariance matrix is the same for all classes

$$f_{k(x)} = \frac{1}{(2\pi)^{\frac{p}{2}} |\mathbf{\Sigma}|^{\frac{1}{2}}} \exp\left(-\frac{1}{2}(x - \mu_k)^T \mathbf{\Sigma}^{-1}(x - \mu_k)\right)$$
$$\delta_k(x) = x^T \mathbf{\Sigma}^{-1} \mu_k - \frac{1}{2} \mu_k^T \mathbf{\Sigma}^{-1} \mu_k + \log \pi_k$$

- **\Sigma** is the  $p \times p$  covariance matrix of the inputs  $\Sigma = \text{Cov}(x)$
- model is fitted using sample estimates similar to the 1D case
- $\mu$  easy, but  $\Sigma$  is the hardest to estimate



LDA fit of data set comprising 20 samples from each class, decision boundary in black

#### Example default with balance and student as inputs

- training error for LDA is 2.75%
- data is highly unbalanced, we have only 3,33% positives
- the No-only classifier has an error of already only 3,33%

#### Sensitivity Sens = $TP/(TP + FN) = TP/P^*$

fraction of correctly predicted positives

Specificity Spec = 
$$TN/(TN + FP) = TN/N^*$$

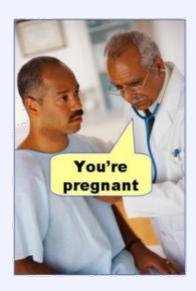
- fraction of correctly predicted negatives
- No Sens =  $\frac{0}{333} = 0\%$ , Spec=  $\frac{9,667}{9,667} = 100\%$
- LDA Sens= $\frac{81}{333}$  = 24.3%, Spec= $\frac{9,644}{9,667}$  = 99.8%
- LDA approximates the Bayes classifier, minimizing error on all observations

#### LDA Model Results

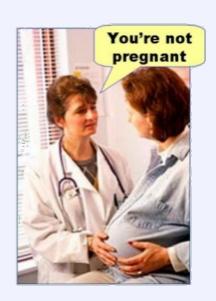
	True Default Status		
Prediction	No (—)	Yes (+)	Total
No (—)	9,644	252	9,896
Yes (+)	23	81	104
Total	9,667	333	10,000

		True Default Status		
	Prediction	No (—)	Yes (+)	Total
	No (—)	TN	FN	N
	Yes (+)	7 FP	TP	P
	Total	N*	$P^*$	n
Type-1 error Confusion matrix alse positive		Т	Type-2 error alse negative	

# Types of Errors – a handy guide



**Type I error** (false positive)



**Type II error** (false negative)

√ la

Biasing the classifier trades sensitivity for specificity

$$\log((p_k(x))/(p_l(x))) = \delta_k(x) - \delta_l(x)$$

- move the decision threshold between class **no** or **yes** from  $Pr(\text{default} = \text{yes} \mid X = x) = 0.5$
- we can increase sensitivity by choosing  $Pr(default = yes \mid X = x) < 0.5$  as this assigns more points to class **yes**
- for Pr(default = yes | X = x) < 0.2
  - Sens = 195/333 = 58.6%
  - Spec = 9,432/9,667 = 97.6%
  - $\blacksquare$  Error = 373/10,000 = 3.73%

#### For a threshold of 0.5 we get Sens = 24.3%, Spec = 99.8%, Error=2.75%

	True Default Status		
Prediction	No (—)	Yes (+)	Total
No (—)	9,644	252	9,896
Yes (+)	23	81	104
Total	9,667	333	10,000

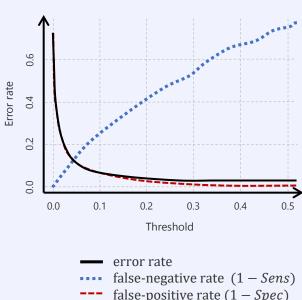
	True Default Status		
Prediction	No (—)	Yes (+)	Total
No (—)	9,432	138	9,570
Yes (+)	235	195	430
Total	9,667	333	10,000

While for a threshold of 0.2 we have Sens = 58.6%, Spec = 97.6%, Error=3.73%

Biasing the classifier trades sensitivity for specificity

$$\log((p_k(x))/(p_l(x))) = \delta_k(x) - \delta_l(x)$$

- move the decision threshold between class **no** or **yes** from Pr(**default** = yes | X = x ) = 0.5
- we can increase sensitivity by choosing  $Pr(default = ves \mid X = x) < 0.5$ as this assigns more points to class yes
- for Pr(default = yes | X = x ) < 0.2
  - Sens = 195/333 = 58.6%
  - $\blacksquare$  Spec = 9,432/9,667 = 97.6%
  - Error = 373/10,000 = 3.73%
- error rates change smoothly when we move the threshold



false-positive rate (1 - Spec)

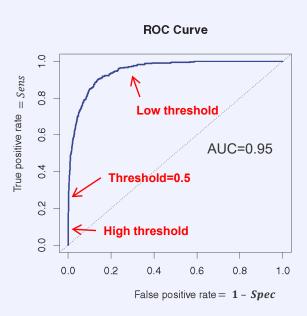
## ROC Curves

#### Receiver-Operating Characteristic (ROC) curves plot Sens against 1 - Spec for all thresholds

- Area Under the ROC-Curve (AUC) measures the quality of a classifier independent of the choice of that threshold
- optimally Spec = Sens = 1 for any threshold (AUC = 1)
- random classifier performs on the diagonal (AUC = 0.5)
- if the ROC curve goes below the diagonal, we can improve accuracy by inverting the classifier

#### ROC curves are **not influenced by imbalance** of the data

balance only affects locations of a threshold along the curve



# Quadratic Discriminant Analysis (QDA)

We give up the assumption that the covariances of all classes are all the same

For QDA we have

$$f_{k(x)} = \frac{1}{(2\pi)^{\frac{p}{2}} |\mathbf{\Sigma}_k|^{\frac{1}{2}}} \exp\left(-\frac{1}{2}(x - \mu_k)^T \mathbf{\Sigma}_k^{-1} (x - \mu_k)\right)$$
$$\delta_k(x) = -\frac{1}{2} x^T \mathbf{\Sigma}_k^{-1} x + x^T \mathbf{\Sigma}_k^{-1} \mu_k - \frac{1}{2} \mu_k^T \mathbf{\Sigma}_k^{-1} \mu_k + \log \pi_k$$

$$f_{k(x)} = \frac{1}{(2\pi)^{\frac{p}{2}} |\mathbf{\Sigma}|^{\frac{1}{2}}} \exp\left(-\frac{1}{2}(x - \mu_k)^T \mathbf{\Sigma}^{-1} (x - \mu_k)\right)$$
$$\delta_k(x) = x^T \mathbf{\Sigma}^{-1} \mu_k - \frac{1}{2} \mu_k^T \mathbf{\Sigma}^{-1} \mu_k + \log \pi_k$$

# Quadratic Discriminant Analysis (QDA)

We give up the assumption that the covariances of all classes are all the same

For ODA we have

$$f_{k(x)} = \frac{1}{(2\pi)^{\frac{p}{2}}|\Sigma_{k}|^{\frac{1}{2}}} \exp\left(-\frac{1}{2}(x - \mu_{k})^{T} \Sigma_{k}^{-1}(x - \mu_{k})\right) \qquad f_{k(x)} = \frac{1}{(2\pi)^{\frac{p}{2}}|\Sigma|^{\frac{1}{2}}} \exp\left(-\frac{1}{2}(x - \mu_{k})^{T} \Sigma^{-1}(x - \mu_{k})\right)$$

$$\delta_{k}(x) = -\frac{1}{2} x^{T} \Sigma_{k}^{-1} x + x^{T} \Sigma_{k}^{-1} \mu_{k} - \frac{1}{2} \mu_{k}^{T} \Sigma_{k}^{-1} \mu_{k} + \log \pi_{k}$$

$$\delta_{k}(x) = x^{T} \Sigma^{-1} \mu_{k} - \frac{1}{2} \mu_{k}^{T} \Sigma^{-1} \mu_{k} + \log \pi_{k}$$

$$f_{k(x)} = \frac{1}{(2\pi)^{\frac{p}{2}}|\mathbf{\Sigma}|^{\frac{1}{2}}} \exp\left(-\frac{1}{2}(x - \mu_k)^T \mathbf{\Sigma}^{-1}(x - \mu_k)\right)$$
$$\delta_k(x) = x^T \mathbf{\Sigma}^{-1} \mu_k - \frac{1}{2}\mu_k^T \mathbf{\Sigma}^{-1} \mu_k + \log \pi_k$$

# Quadratic Discriminant Analysis (QDA)

In QDA every class has its own covariance matrix

$$f_{k(x)} = \frac{1}{(2\pi)^{\frac{p}{2}} |\Sigma_k|^{\frac{1}{2}}} \exp\left(-\frac{1}{2}(x - \mu_k)^T \Sigma_k^{-1} (x - \mu_k)\right)$$
$$\delta_k(x) = -\frac{1}{2} x^T \Sigma_k^{-1} x + x^T \Sigma_k^{-1} \mu_k - \frac{1}{2} \mu_k^T \Sigma_k^{-1} \mu_k + \log \pi_k$$

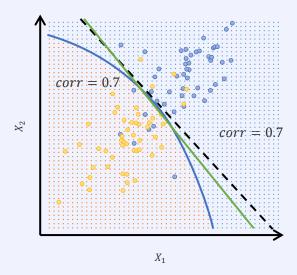
- class boundaries are now quadratic curves
- we fit a different covariance matrix estimate per class
- LDA has (2K + p + 1)p/2 parameters,
- QDA has Kp(p+3)/2 parameters

#### Example

- for p = 4, K = 2, LDA has 18 parameters, QDA has 28 parameters
- for p = 8, K = 2, LDA has 52 parameters, QDA has 88 parameters

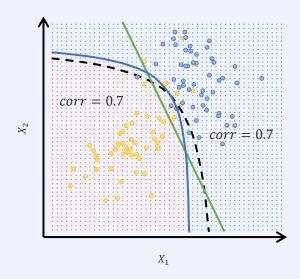
# Example LDA vs. QDA

# Two-class problem with $\Sigma_1 = \Sigma_2$ QDA overtrains



- Bayes decision boundary
- LDA decision boundary
- QDA decision boundary

#### Two-class problem with $\Sigma_1 \neq \Sigma_2$ LDA overtrains



- Bayes decision boundary
- LDA decision boundary
- QDA decision boundary

V (ESL 4.4.3)





Again, we use sample estimates

$$\hat{\mu}_k = \frac{1}{n_k} \sum_{i: y_i = k} x_i$$

$$\widehat{\boldsymbol{\Sigma}} = \frac{1}{n-K} \sum_{k=1}^{K} \sum_{i:y_i=k} (x_i - \hat{\mu}_k) (x_i - \hat{\mu}_k)^T$$

$$\widehat{\boldsymbol{\Sigma}}_k = \frac{1}{n_k - K} \sum_{i: y_i = k} (x_i - \widehat{\mu}_k) (x_i - \widehat{\mu}_k)^T$$

 $\bullet \quad \pi_k = n_k/n$ 

To simplify calculation we use the eigenvalue decomposition of the covariance matrices

$$\widehat{\boldsymbol{\Sigma}}_k = \boldsymbol{U}_k \boldsymbol{D}_k \boldsymbol{U}_k^T$$

- $\mathbf{u}_k$  is a  $p \times p$  orthonormal matrix
- $\mathbf{p}_k$  is a diagonal matrix of decreasing positive eigenvalues  $d_{kl}$

The main terms in the discriminants,

$$\delta_k(x) = -\frac{1}{2}\log|\widehat{\mathbf{\Sigma}}_k| - \frac{1}{2}(x - \mu_k)^T \widehat{\mathbf{\Sigma}}_k^{-1}(x - \mu_k) + \log \pi_k$$

then turn into

$$\log |\widehat{\mathbf{\Sigma}}_k| = \sum_{l} \log d_{kl}$$
$$(x - \hat{\mu}_k)^T \widehat{\mathbf{\Sigma}}_k^{-1} (x - \hat{\mu}_k) = \left[ \mathbf{U}_k^T (x - \hat{\mu}_k) \right]^T D_k^{-1} \left[ \mathbf{U}_k^T (x - \hat{\mu}_k) \right]$$

The LDA estimator

- Step 1: Normalize X to spherical covariance  $X^* \leftarrow D^{-1/2}U^TX$
- Step 2: Classify to the closest class centroid in the transformed space, where distance is weighted by the class prior probabilities  $\pi_k$

# Comparing Different Classifiers

We now know four classifiers: k-NN, LDA, QDA and logistic regression

when should we use which?

Logistic regression and LDA are surprisingly closely related

univariate binary setting

$$p_2(x) = 1 - p_1(x)$$

log-odds for LDA are

$$\log \frac{p_1(x)}{1 - p_1(x)} = c_0 + c_1 x$$

(difference of two linear discriminants)

while for logistic regression

$$\log \frac{p_1(x)}{1 - p_1(x)} = \beta_0 + \beta_1 x$$

#### Similar, but different

- $\beta_0$  and  $\beta_1$  are maximum likelihood estimates
- $c_0$  and  $c_1$  are estimated from sample mean and variance of Gaussian distribution
- relationship extends to multivariate data: LR and LDA often give similar results but not always!
- LDA makes stronger assumptions

We now know four classifiers: k-NN, LDA, QDA and logistic regression

when should we use which?

k-NN is nonparametric and tends to work better for strongly nonlinear settings

it does not allow for inference, i.e. we do not get a model that we can learn from

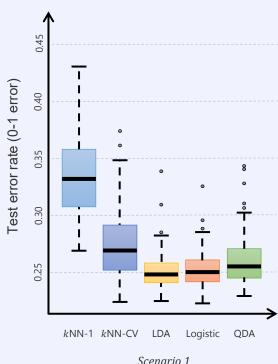
QDA is a compromise between LDA and k-NN

#### Scenario 1

- 100 random training data sets, p = 2 predictors, K = 2 classes
- 20 observations per class
- observations in different classes uncorrelated normal variables with different means and the same variance (spherical Gaussian)
- this matches the LDA assumptions of LDA

#### **Observations**

- LDA works very well
- logistic regression assumes a linear decision boundary, performs only slightly worse than LDA
- k-NN overtrains, as does QDA

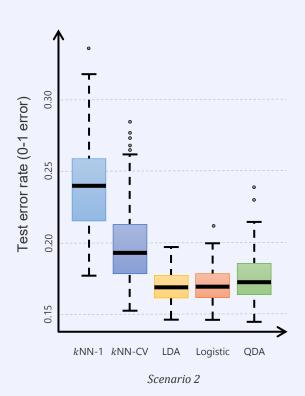


#### Scenario 2

- 100 random training data sets, p = 2 predictors, K = 2 classes
- like scenario 1, but predictors in each class now have a correlation of -0.5 (elliptical multivariate Gaussian)

#### Observations

relative performances are similar to scenario 1



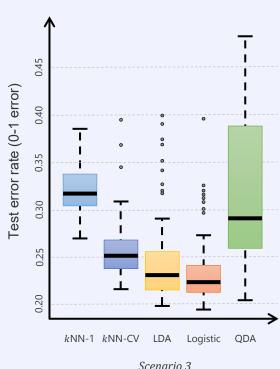
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#### Scenario 3

- 100 random training data sets, p = 2 predictors, K = 2 classes
- $X_1$  and  $X_2$  are generated using a **t**-distribution
- more extreme points than with a Gaussian
- decision boundary is linear, but, setup violates LDA assumption

#### **Observations**

- logistic regression performs best
- QDA deteriorates because of non-normality of the data

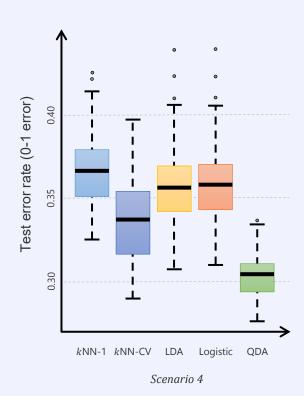


#### Scenario 4

- 100 random training data sets, p = 2 predictors, K = 2 classes
- class 1: normal distribution with correlation 0.5 to predictors
- class 2: normal distribution with correlation -0.5 to predictors
- assumptions of QDA are met (but not LDA!)

#### **Observations**

QDA outperforms all other methods

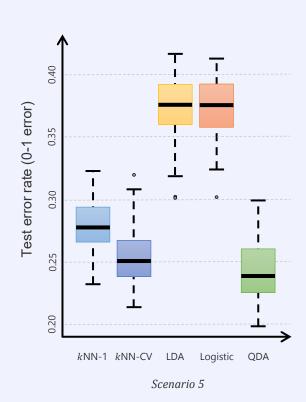


#### Scenario 5

- 100 random training data sets, p = 2 predictors, K = 2 classes
- two normal distributions with uncorrelated predictors
- inputs  $X_1^2$ ,  $X_2^2$  and  $X_1X_2$ , not  $X_1$  and  $X_2$
- the decision boundary is quadratic

#### **Observations**

- QDA performs best
- kNN (CV) follows closely
- the linear methods all perform poorly

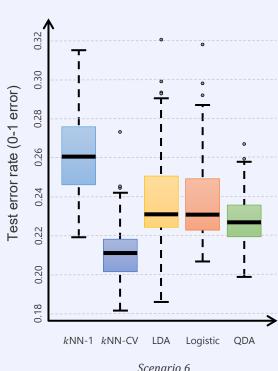


#### Scenario 6

- 100 random training data sets, p = 2 predictors, K = 2 classes
- like 5, but responses sampled from a complicated linear function

#### **Observations**

- even QDA cannot model data well
- k-NN-1 overtrains
- k-NN (CV) outperforms all parametric approaches
- smoothness must be chosen carefully



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