

# 403BfinalprojectUber

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## I. Introduction

dependent variable ???mean travel time range=upper-lower/average independent variable???weather / temperature/ holiday/ weekday

## Data clean

```
## import data
rm(list=ls(all=TRUE))
setwd("~/Users/hongyiwang/Desktop/winter/403B/final project")
ub=read.csv("uberdata.csv",header = T)
temp = read.csv("tempre.csv",header = T)
```

## II. Data Exploratory

### II.1 Data Quality

```
### find missing value
sum(is.na(ub))

## [1] 39
sum(is.na(temp$Precip))

## [1] 0
```

There are 39 NAs in total in dataset “ub”, 0 NAs in dataset “temp”.

subsitute NAs, by: if NA, then NA value = average of the the day before and the day after eg. day 10th is NA, X10 = (X9 + X11)/2 both ub[285,] ub[286,] are NA, so we define ub[285,] = ub[284,],ub[286,] = ub[287]

```
### define NAs in ub
index1 = which(is.na(ub$mean),arr.ind = T)
for (i in index1){
  ub$lower[i] = (ub$lower[i-1] + ub$lower[i+1])/2
  ub$upper[i] = (ub$upper[i-1] + ub$upper[i+1])/2
  ub$mean[i] = (ub$upper[i] + ub$lower[i])/2
}
ub[285,2:4] = ub[284,2:4]
ub[286,2:4] = ub[287,2:4]

### define NAs in precip
precip = temp$Precip[1:365]
index2 = levels(precip)
for (i in 1:length(precip)){
  if (precip[i] == 'T') {
```

```

    precip[i] = 0
} else if (precip[i] == ''){
  precip[i] = NA
} else {
  for (j in index2[2:22]) {
    if (precip[i] == j) {
      precip[i] = as.numeric(j)
    }
  }
}
index3 = which(is.na(precip))
## all NAs define as 0
precip[index3] = 0

```

Define dataset:

```

travel = ub$mean
travelrange = ub$upper - ub$lower
holiday = ub$holiady
## we will not use temp so I commented these two lines
#meantemp = (as.numeric(temp$High) + as.numeric(temp$low) )/2
#temprange = as.numeric(temp$High) - as.numeric(temp$low)
# create dummy variable rainfall
rainfall = rep(1,365)
rainfall[precip == 0] = 0

```

## II.2 Dummy Variables Creation

```

weekdayloop = rep(1:7,53)
weekdayloop = weekdayloop[1:365]

Sun = rep(0,365)
Sun[weekdayloop == 1] = 1
Mon = rep(0,365)
Mon[weekdayloop == 2] = 1
Tue = rep(0,365)
Tue[weekdayloop == 3] = 1
Wed = rep(0,365)
Wed[weekdayloop == 4] = 1
Thur = rep(0,365)
Thur[weekdayloop == 5] = 1
Fri = rep(0,365)
Fri[weekdayloop == 6] = 1
Sat = rep(0,365)
Sat[weekdayloop == 7] = 1

#combine dataset
ubdata = data.frame(travel,travelrange,holiday,rainfall,Mon,Tue,Wed,Thur,Fri,Sat,Sun)

```

### II.3 Time Series Classification, Plot and Observation

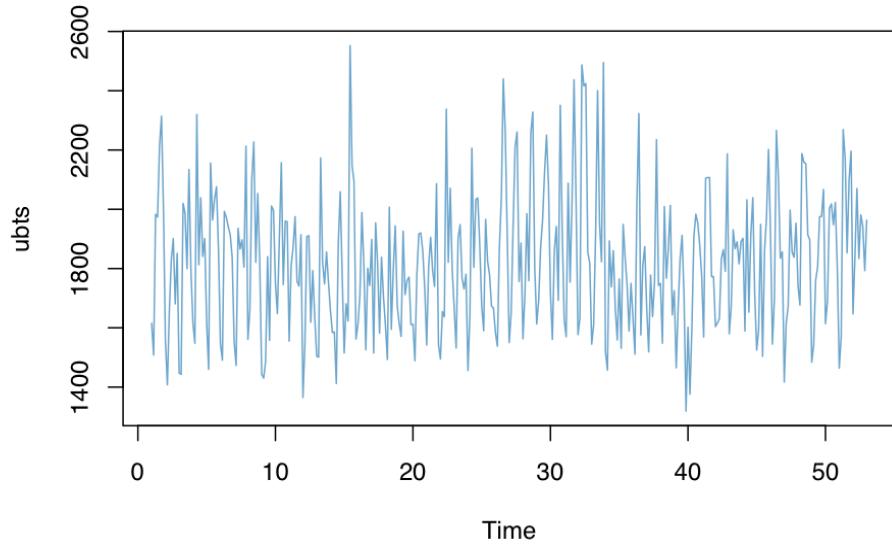
```
inds <- seq(as.Date("2017-04-01"), as.Date("2018-03-01"), by = "day")

#myts <- ts(rnorm(length(inds)),      # random data
#            start = c(2014, as.numeric(format(inds[1], "%j"))),
#            frequency = 365)
#ubts =ts(ubdata$travel,start=c(2017,as.numeric(format(inds[1],"%j"))),freq=365)
ubts =ts(ubdata$travel, start = c(1, 1), freq = 7)
ubrangets =ts(ubdata$travelrange,start=c(2017,as.numeric(format(inds[1],"%j"))),freq=365)
```

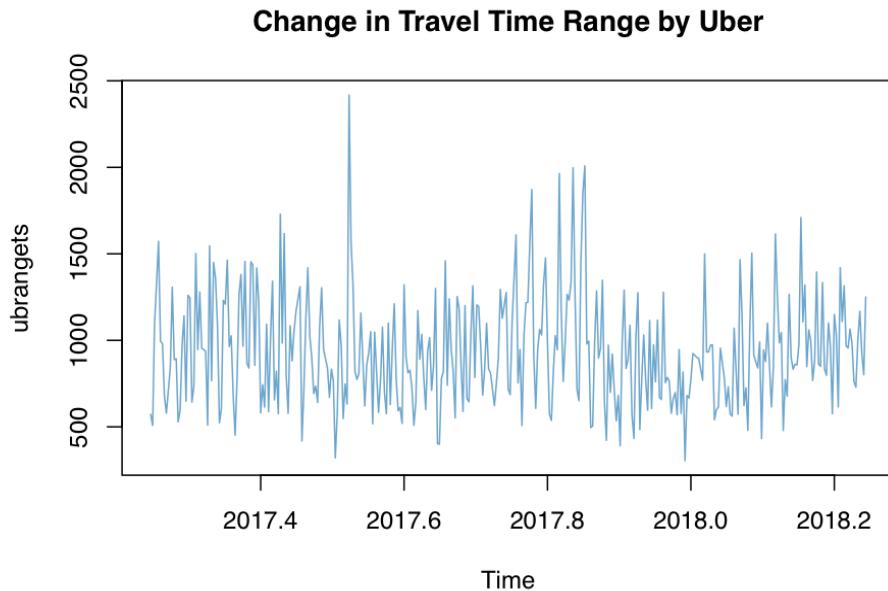
(a) Time Series Plot

```
plot(ubts,col='skyblue3',main="Change in Mean Travel Time by Uber")
```

**Change in Mean Travel Time by Uber**



```
plot(ubrangets,col='skyblue3',main="Change in Travel Time Range by Uber")
```



(b) Covariance Stationary

```
library(tseries)
adf.test(ubts)

## Warning in adf.test(ubts): p-value smaller than printed p-value
##
##  Augmented Dickey-Fuller Test
##
##  data: ubts
##  Dickey-Fuller = -5.5863, Lag order = 7, p-value = 0.01
##  alternative hypothesis: stationary
adf.test(ubrangegets)

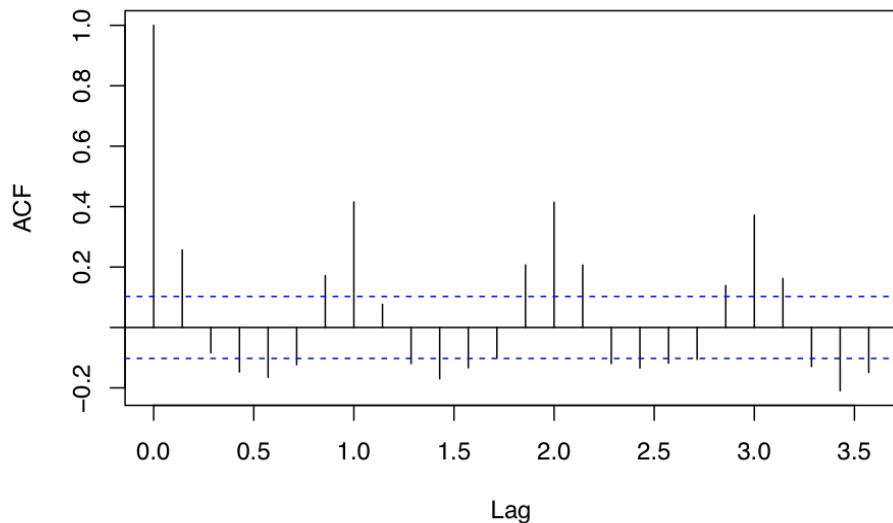
## Warning in adf.test(ubrangegets): p-value smaller than printed p-value
##
##  Augmented Dickey-Fuller Test
##
##  data: ubrangegets
##  Dickey-Fuller = -4.5237, Lag order = 7, p-value = 0.01
##  alternative hypothesis: stationary
##  stationary
```

According to the results, since the p-value is smaller than 0.05, we could conclude that our time series is stationary. There is no need for transforming.

(c) Plot and Discuss the ACF and PACF

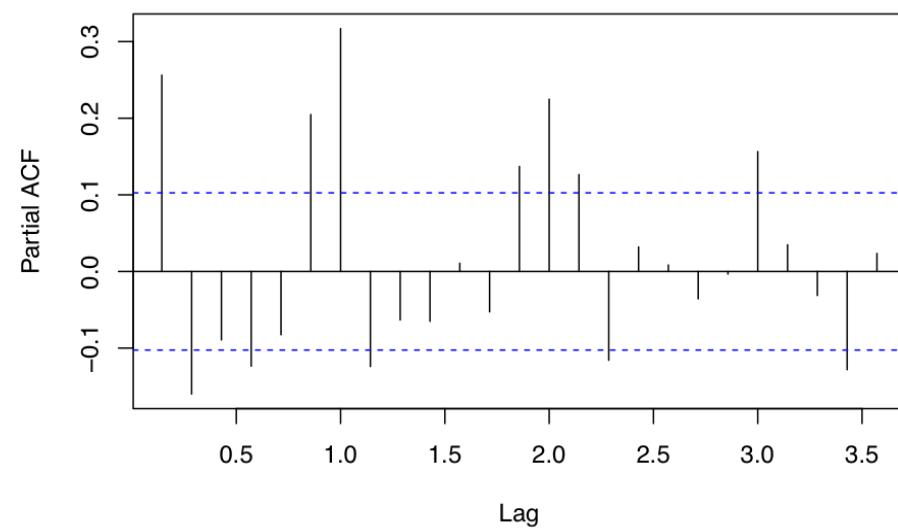
```
acf(ubts)
```

**Series ubts**



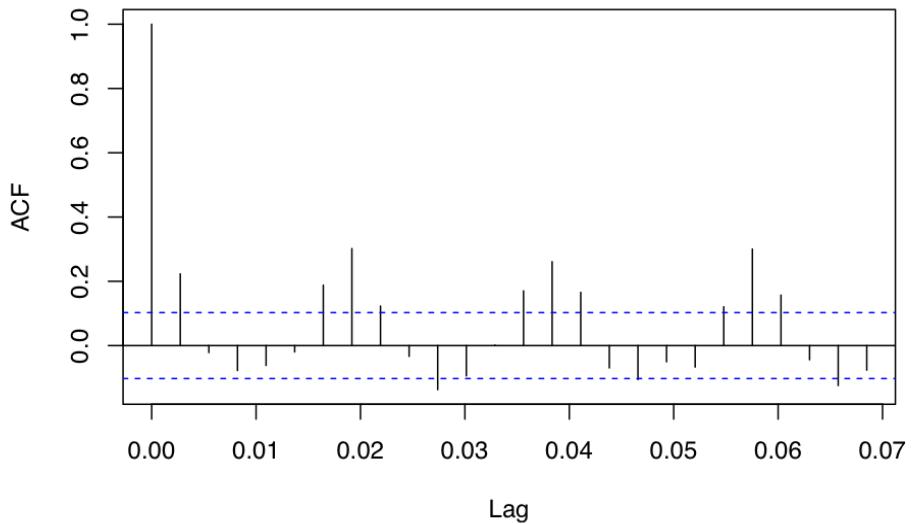
```
pacf(ubts)
```

**Series ubts**



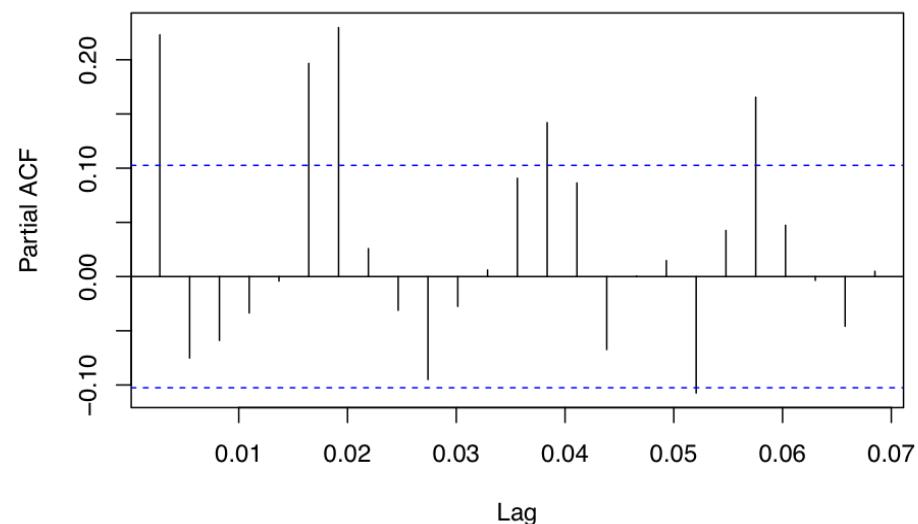
```
acf(ubrangets)
```

**Series ubrangets**



```
pacf(ubrangets)
```

**Series ubrangets**



From the graph, we could notice that there exists a circle of seven days, which is the same as days in a week.

(d) 1st difference

```
ubts.1 <- diff(ubts, 1)
ubts.1.7 <- diff(ubts.1, 7)
summary(ur.df(ubts.1.7, type = "none"))

##
## #####
## # Augmented Dickey-Fuller Test Unit Root Test #
## #####
## 
## Test regression none
##
##
## Call:
## lm(formula = z.diff ~ z.lag.1 - 1 + z.diff.lag)
##
## Residuals:
##     Min      1Q  Median      3Q     Max
## -1035.59 -173.39   -6.52   167.06 1141.80
##
## Coefficients:
##             Estimate Std. Error t value Pr(>|t|)
## z.lag.1    -1.81086   0.08570 -21.130 < 2e-16 ***
## z.diff.lag  0.29152   0.05127  5.686 2.73e-08 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 277.7 on 353 degrees of freedom
## Multiple R-squared:  0.7237, Adjusted R-squared:  0.7221
## F-statistic: 462.2 on 2 and 353 DF,  p-value: < 2.2e-16
##
##
## Value of test-statistic is: -21.1302
##
## Critical values for test statistics:
##      1pct 5pct 10pct
## tau1 -2.58 -1.95 -1.62
```

### III linear regression

(a) pure season dummy

```
m2 =lm(ubts ~ Mon+Tue+Sat+Sun+Thur+Wed)
m3 =lm(ubts ~ Mon+Tue+Sat+Sun+Thur+Fri)
m4 =lm(ubts ~ Mon+Tue+Sat+Fri+Thur+Wed)
m5 =lm(ubts ~ Sun+Tue+Sat+Fri+Thur+Wed)
m6 =lm(ubts ~ Mon+Sun+Sat+Fri+Thur+Wed)
## comparisons
AIC(m2,m3,m4,m5,m6)

##      df      AIC
```

```

## m2 8 4854.371
## m3 8 4854.371
## m4 8 4854.371
## m5 8 4854.371
## m6 8 4854.371
summary(m6)

##
## Call:
## lm(formula = ubts ~ Mon + Sun + Sat + Fri + Thur + Wed)
##
## Residuals:
##   Min     1Q Median     3Q    Max
## -528.64 -110.76 -10.08  84.31 647.36
##
## Coefficients:
##             Estimate Std. Error t value Pr(>|t|)
## (Intercept) 1927.779   25.611  75.271 <2e-16 ***
## Mon        -349.087   36.220 -9.638 <2e-16 ***
## Sun        -313.694   36.048 -8.702 <2e-16 ***
## Sat        -80.139   36.220 -2.213  0.0276 *
## Fri         18.303   36.220  0.505  0.6136
## Thur       -19.558   36.220 -0.540  0.5896
## Wed         6.981   36.220  0.193  0.8473
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 184.7 on 358 degrees of freedom
## Multiple R-squared:  0.3908, Adjusted R-squared:  0.3806
## F-statistic: 38.28 on 6 and 358 DF,  p-value: < 2.2e-16

```

(b) add holiday and rainfall

```

m7=lm(ubts ~ Mon+Tue+Sat+Sun+Thur+Wed+Fri+holiday)
summary(m7)

##
## Call:
## lm(formula = ubts ~ Mon + Tue + Sat + Sun + Thur + Wed + Fri +
##     holiday)
##
## Residuals:
##   Min     1Q Median     3Q    Max
## -528.64 -111.24 -14.24  84.13 647.36
##
## Coefficients: (1 not defined because of singularities)
##             Estimate Std. Error t value Pr(>|t|)
## (Intercept) 1946.082   25.400  76.616 < 2e-16 ***
## Mon        -358.321   36.085 -9.930 < 2e-16 ***
## Tue         -6.212   36.212 -0.172  0.86390
## Sat        -98.442   35.921 -2.740  0.00644 **
## Sun       -329.031   35.769 -9.199 < 2e-16 ***
## Thur       -34.838   35.940 -0.969  0.33303
## Wed        -8.299   35.940 -0.231  0.81751

```

```

## Fri          NA          NA          NA          NA
## holiday     -157.184    59.546   -2.640  0.00866 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 183.2 on 357 degrees of freedom
## Multiple R-squared:  0.4025, Adjusted R-squared:  0.3908
## F-statistic: 34.35 on 7 and 357 DF,  p-value: < 2.2e-16
m8=lm(ubts~Mon+Tue+Sat+Sun+Thur+Wed+Fri+rainfall)
summary(m8)

##
## Call:
## lm(formula = ubts ~ Mon + Tue + Sat + Sun + Thur + Wed + Fri +
##     rainfall)
##
## Residuals:
##      Min    1Q Median    3Q   Max
## -526.86 -109.57  -9.75  86.09 649.14
##
## Coefficients: (1 not defined because of singularities)
##             Estimate Std. Error t value Pr(>|t|)
## (Intercept) 1943.110   25.869  75.112 < 2e-16 ***
## Mon        -366.201   36.263 -10.098 < 2e-16 ***
## Tue         -18.897   36.242  -0.521  0.60240
## Sat        -97.254   36.263  -2.682  0.00766 **
## Sun       -331.358   36.072  -9.186 < 2e-16 ***
## Thur        -37.266   36.242  -1.028  0.30452
## Wed         -9.539   36.298  -0.263  0.79285
## Fri          NA        NA        NA        NA
## rainfall     30.901   37.106   0.833  0.40552
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 184.8 on 357 degrees of freedom
## Multiple R-squared:  0.392, Adjusted R-squared:  0.3801
## F-statistic: 32.88 on 7 and 357 DF,  p-value: < 2.2e-16

```

#### IV Model

##### (a) arima

```

## arma(6,1)+seasonality
fit.a.1 <- arima(ubts, order = c(6, 1, 0), seasonal = list(order = c(7, 1, 0)))
confint(fit.a.1)

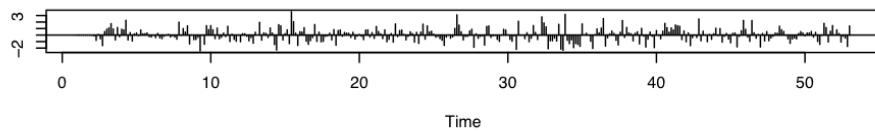
##           2.5 %      97.5 %
## ar1  -0.9448685 -0.73145401
## ar2  -0.8831545 -0.59873374
## ar3  -0.8422949 -0.51000909
## ar4  -0.7712368 -0.43776246
## ar5  -0.6795740 -0.33266517
## ar6  -0.6257342 -0.26331815
## sar1 -1.4274840 -1.04307564

```

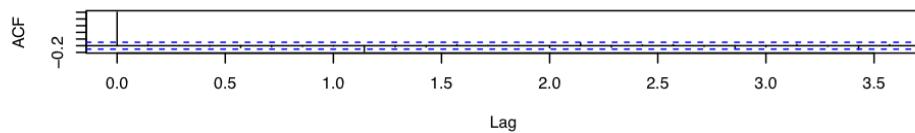
```
## sar2 -1.4596211 -0.86244384  
## sar3 -1.3179197 -0.67176613  
## sar4 -1.1358802 -0.55443401  
## sar5 -0.8762835 -0.38977152  
## sar6 -0.6320074 -0.26717570  
## sar7 -0.1998667  0.03526851
```

```
tsdiag(fit.a.1)
```

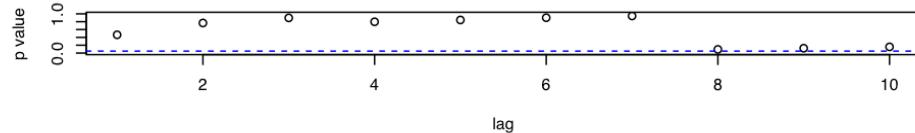
Standardized Residuals



ACF of Residuals



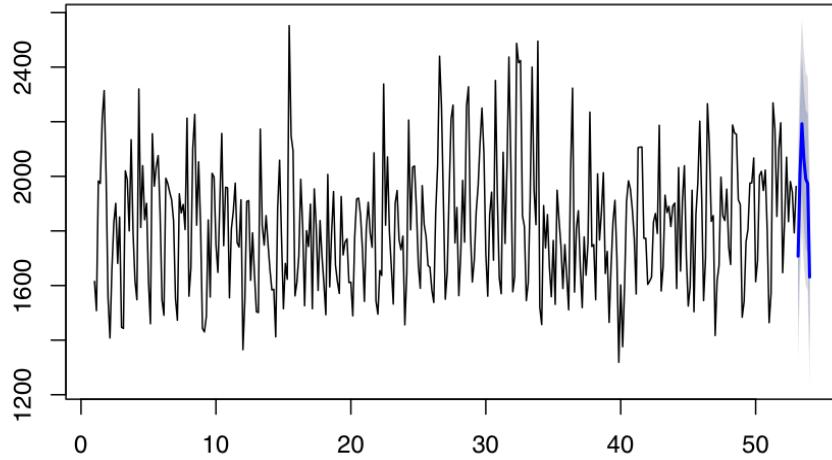
p values for Ljung–Box statistic



```
Box.test(fit.a.1$residuals, lag = 14, type = "Ljung-Box")
```

```
##  
## Box-Ljung test  
##  
## data: fit.a.1$residuals  
## X-squared = 16.016, df = 14, p-value = 0.3124  
plot(forecast(fit.a.1, h = 7))
```

## Forecasts from ARIMA(6,1,0)(7,1,0)[7]



```

summary(fit.a.1, h = 7)

##
## Call:
## arima(x = ubts, order = c(6, 1, 0), seasonal = list(order = c(7, 1, 0)))
##
## Coefficients:
##          ar1      ar2      ar3      ar4      ar5      ar6      sar1
##     -0.8382  -0.7409  -0.6762  -0.6045  -0.5061  -0.4445  -1.2353
##  s.e.    0.0544   0.0726   0.0848   0.0851   0.0885   0.0925   0.0981
##          sar2      sar3      sar4      sar5      sar6      sar7
##     -1.1610  -0.9948  -0.8452  -0.6330  -0.4496  -0.0823
##  s.e.    0.1523   0.1648   0.1483   0.1241   0.0931   0.0600
##
## sigma^2 estimated as 36707:  log likelihood = -2389.09,  aic = 4806.18
##
## Training set error measures:
##          ME      RMSE      MAE       MPE      MAPE      MASE
## Training set 4.498029 189.4795 144.6594 -0.5071734 7.883135 0.6466652
##          ACF1
## Training set 0.03801147

# plot the estimated seasonal factors and interpret your plot

#install.packages("lmtest")
#coeftest(fit.a.1)

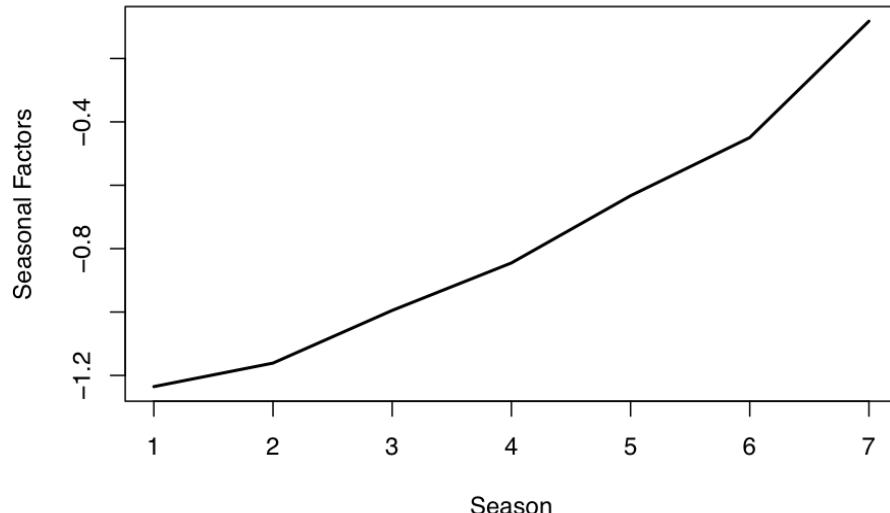
fit.a.1$coef

##          ar1      ar2      ar3      ar4      ar5      ar6
## -0.83816126 -0.74094412 -0.67615200 -0.60449963 -0.50611959 -0.44452618
##          sar1      sar2      sar3      sar4      sar5      sar6
## -1.23527984 -1.16103245 -0.99484291 -0.84515711 -0.63302750 -0.44959156

```

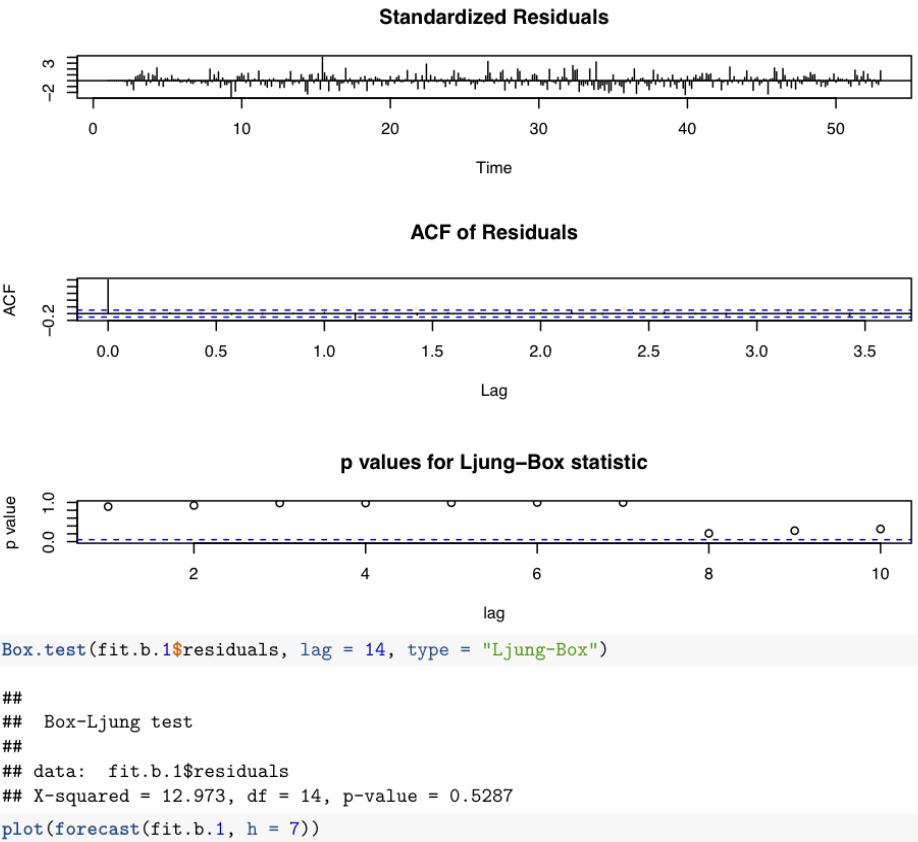
```
##          sar7
## -0.08229908
plot(fit.a.1$coef[7:13],type='l',ylab='Seasonal Factors',xlab="Season",lwd=2, main="Plot of Seasonal Factors")
```

**Plot of Seasonal Factors**

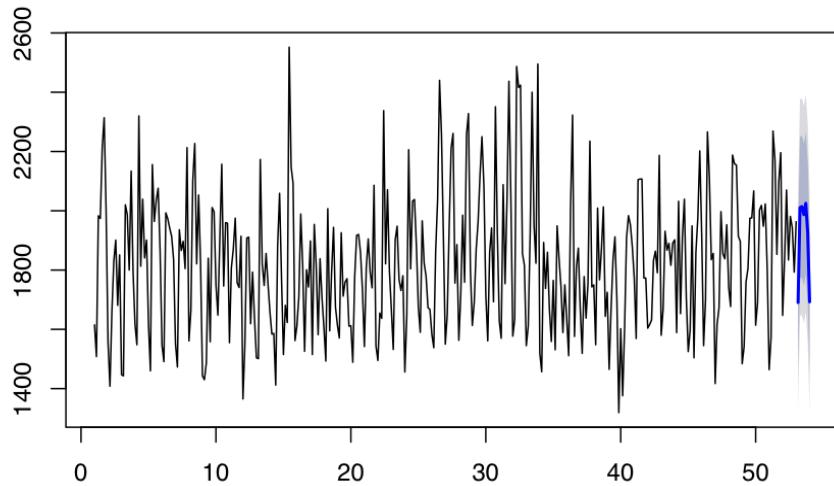


```
fit.b.1 <- arima(ubts, order = c(1, 1, 1), seasonal = list(order = c(0, 1, 1)))
confint(fit.b.1)

##           2.5 %    97.5 %
## ar1   0.009028253  0.2285816
## ma1  -0.984889362 -0.9080785
## sma1 -1.073345165 -0.9266372
tsdiag(fit.b.1)
```



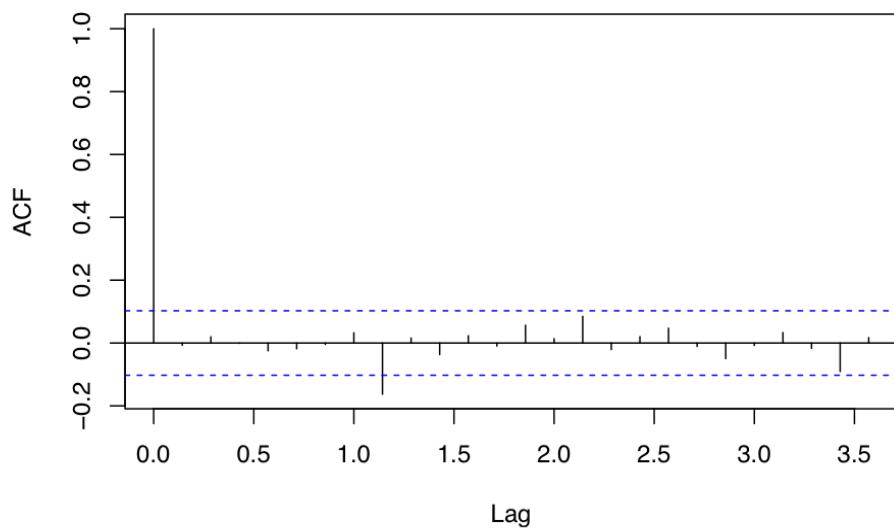
### Forecasts from ARIMA(1,1,1)(0,1,1)[7]



(b) GARCH

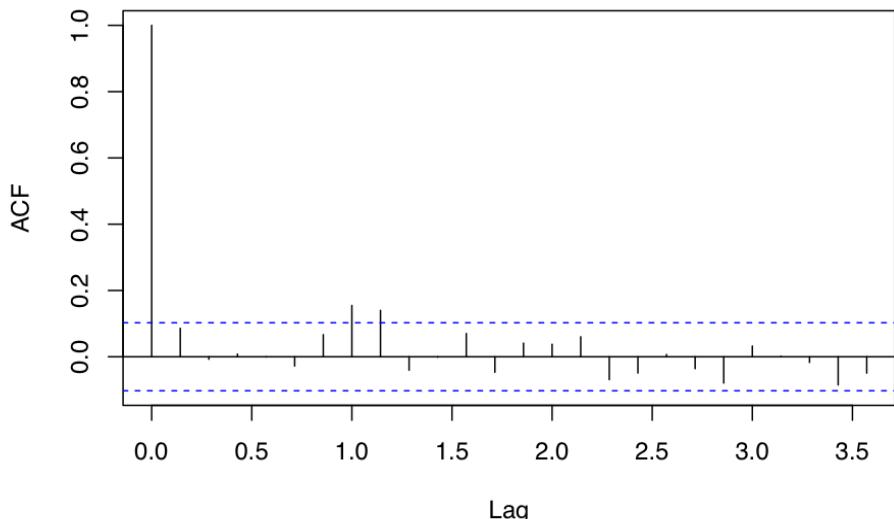
```
## ARCH(1)
res.arimab=fit.b.1$res
acf(res.arimab)
```

**Series res.arimab**



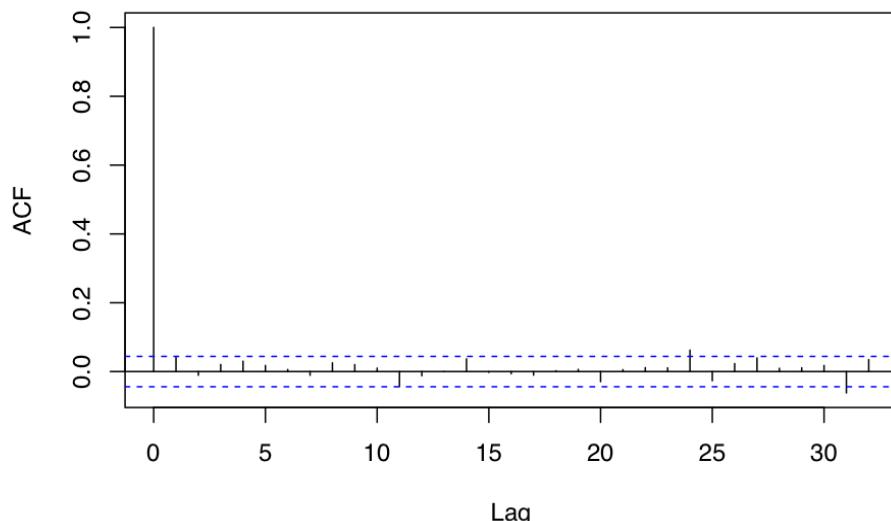
```
acf(res.arimab^2)
```

### Series res.arimab^2



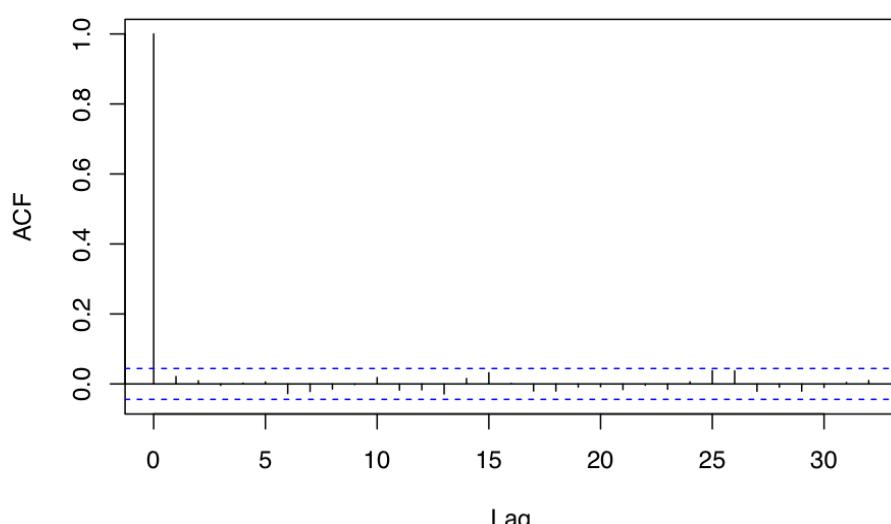
```
squared=res.arimab^2  
m_g<-garchFit(res.arimab~garch(1,3),trace = FALSE)  
res<-m_g@residuals/m_g@sigma.t  
acf(res)
```

**Series res**



```
acf(res^2)
```

**Series res<sup>2</sup>**



```
summary(m_g)
```

```
##
```

```

## Title:
## GARCH Modelling
##
## Call:
## garchFit(formula = res.arimab ~ garch(1, 3), trace = FALSE)
##
## Mean and Variance Equation:
## data ~ garch(1, 3)
## <environment: 0x7fcc73305510>
## [data = fGarch::dem2gbp]
##
## Conditional Distribution:
## norm
##
## Coefficient(s):
##          mu        omega      alpha1      beta1      beta2
## -0.00376519  0.01127952  0.19055243  0.38822250  0.00000001
##          beta3
##  0.37653535
##
## Std. Errors:
## based on Hessian
##
## Error Analysis:
##           Estimate Std. Error t value Pr(>|t|)
## mu     -3.765e-03  8.482e-03 -0.444 0.657123
## omega   1.128e-02  2.826e-03  3.991 6.58e-05 ***
## alpha1  1.906e-01  2.791e-02  6.828 8.63e-12 ***
## beta1   3.882e-01  1.038e-01  3.741 0.000184 ***
## beta2   1.000e-08  1.723e-01  0.000 1.000000
## beta3   3.765e-01  1.279e-01  2.944 0.003240 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Log Likelihood:
## -1099.094    normalized: -0.5567853
##
## Description:
## Fri Mar 22 14:21:20 2019 by user:
##
##
## Standardised Residuals Tests:
##                               Statistic p-Value
## Jarque-Bera Test   R Chi^2 999.1862 0
## Shapiro-Wilk Test  R W 0.9632659 0
## Ljung-Box Test     R Q(10) 9.320895 0.5019442
## Ljung-Box Test     R Q(15) 16.09236 0.37594
## Ljung-Box Test     R Q(20) 18.16239 0.5767117
## Ljung-Box Test     R^2 Q(10) 4.600815 0.9162016
## Ljung-Box Test     R^2 Q(15) 9.874221 0.8275762
## Ljung-Box Test     R^2 Q(20) 11.76969 0.9237748
## LM Arch Test       R TR^2 5.615319 0.9342197
##
## Information Criterion Statistics:

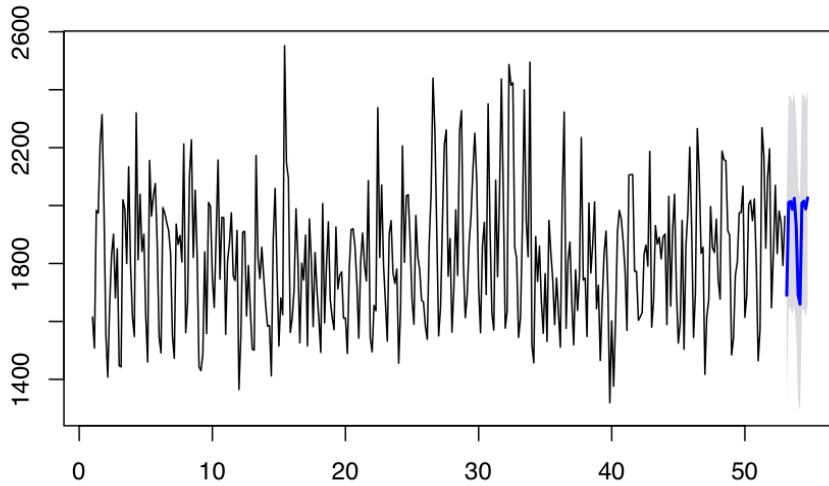
```

```

##      AIC      BIC      SIC      HQIC
## 1.119650 1.136634 1.119631 1.125890
forecast212=forecast(fit.b.1,12,level=95)
plot(forecast212)

```

### Forecasts from ARIMA(1,1,1)(0,1,1)[7]



```
mgg<-garch(res.arimab,order=c(1,3),trace=FALSE)
```

```

## Warning in garch(res.arimab, order = c(1, 3), trace = FALSE): singular
## information
summary(mgg)

```

```

##
## Call:
## garch(x = res.arimab, order = c(1, 3), trace = FALSE)
##
## Model:
## GARCH(1,3)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -2.940830 -0.617130 -0.007312  0.610993  3.865123
##
## Coefficient(s):
##             Estimate Std. Error t value Pr(>|t|)
## a0  2.578e+04      NA      NA      NA
## a1  1.091e-01      NA      NA      NA
## a2  4.466e-10      NA      NA      NA
## a3  8.400e-03      NA      NA      NA
## b1  6.430e-02      NA      NA      NA
##
## Diagnostic Tests:
## Jarque Bera Test

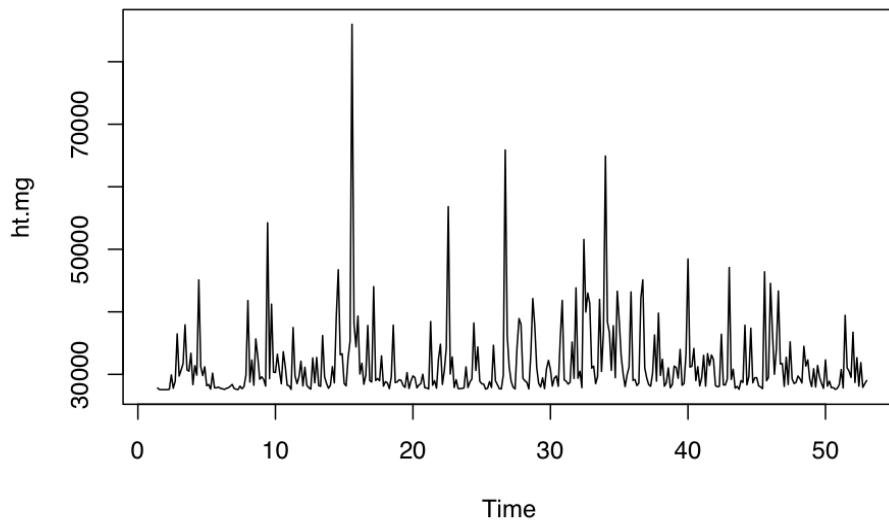
```

```

## 
## data: Residuals
## X-squared = 15, df = 2, p-value = 0.0005531
##
##
## Box-Ljung test
##
## data: Squared.Residuals
## X-squared = 0.0011018, df = 1, p-value = 0.9735
ht.mg =mgg$fit[,1]^2 #use 1st column of fit plot(ht.arch08,main='Conditional variances')
plot(ht.mg,main='Conditional variances')

```

### Conditional variances

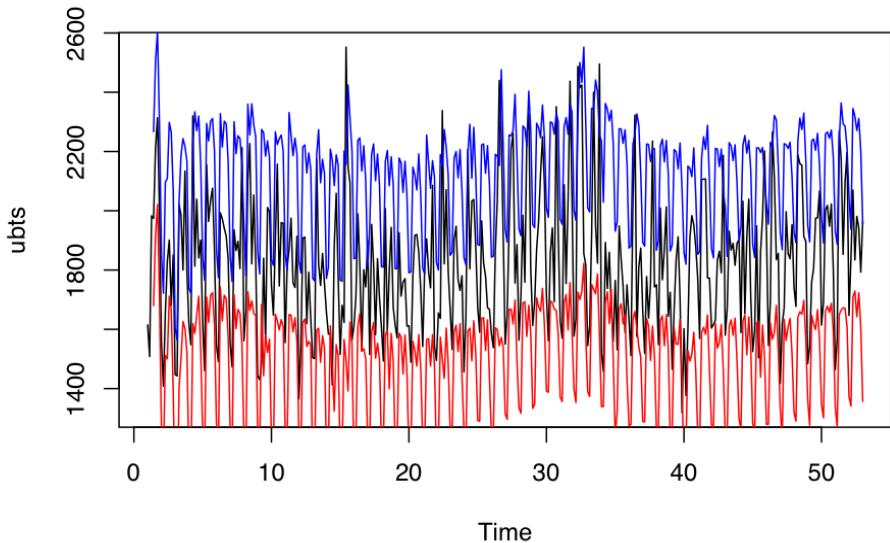


```

res212<-mgg$residuals
fit212=fitted.values(fit.b.1)
low=fit212-1.76*sqrt(ht.mg)
high=fit212+1.76*sqrt(ht.mg)
plot(ubts,type='l',main='Uber,Low,High')
lines(low,col='red')
lines(high,col='blue')

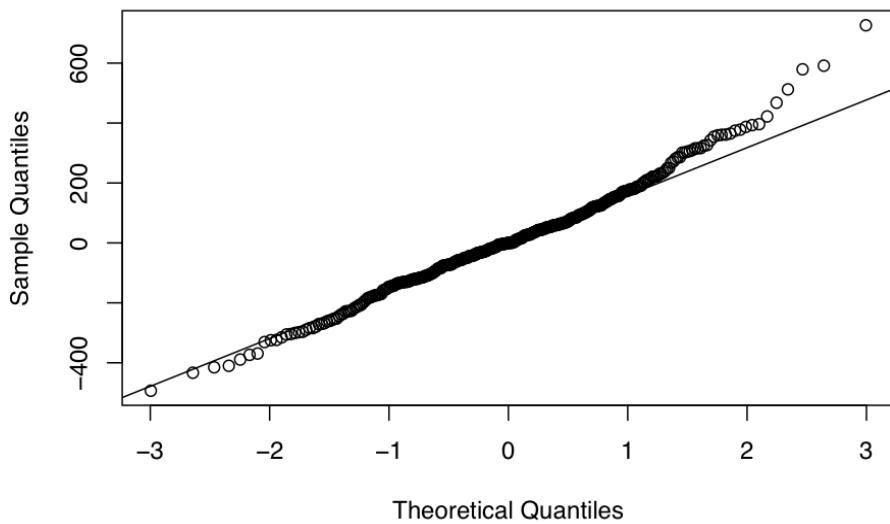
```

### Uber,Low,High



```
## normal distribution  
qqnorm(fit.b.1$residuals,main='ARIMA Residuals')  
qqline(fit.b.1$residuals)
```

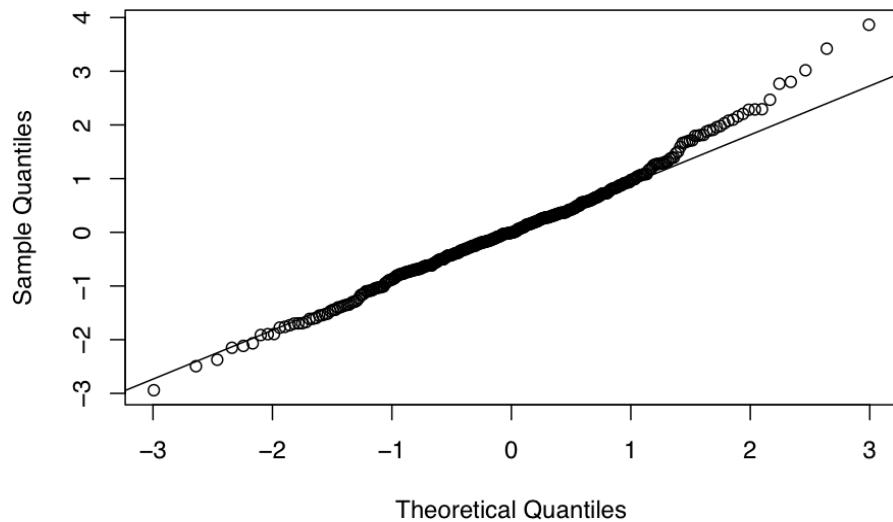
### ARIMA Residuals



```
archres=res212  
qqnorm(archres,main='ARIMA-ARCH Residuals')
```

```
qqline(archres)
```

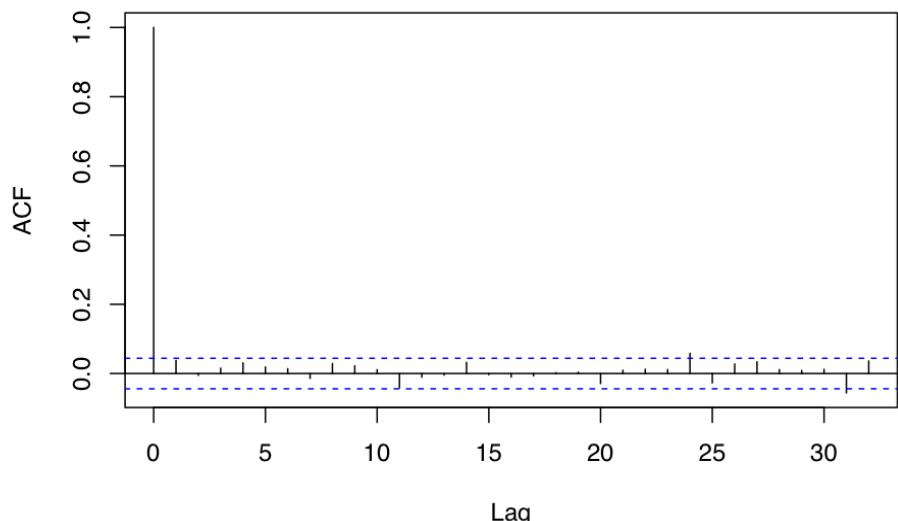
**ARIMA-ARCH Residuals**



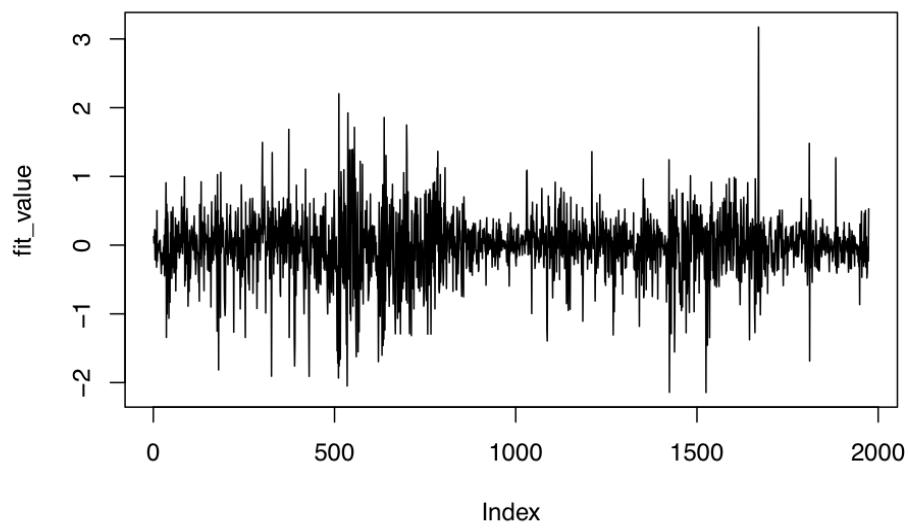
(c) Garch + arima

```
mag<-garchFit(formula =ubts.1~arma(1,1,1)+res.arimab-garch(1,3),cond.dist ="std",trace=FALSE)
resi=mag@residuals/mag@sigma.t   ###Standardised Residuals
acf(resi)
```

### Series resi



```
fit_value<-fitted(mgg)  
plot(fit_value,type="l")
```



```
AIC(mgg,fit.b.1)
```

```
##          df      AIC  
## mgg      5 4793.872  
## fit.b.1  4 4766.726
```

```

BIC(mgg,fit.b.1)

##          df      BIC
## mgg      5      NA
## fit.b.1  4 4782.236

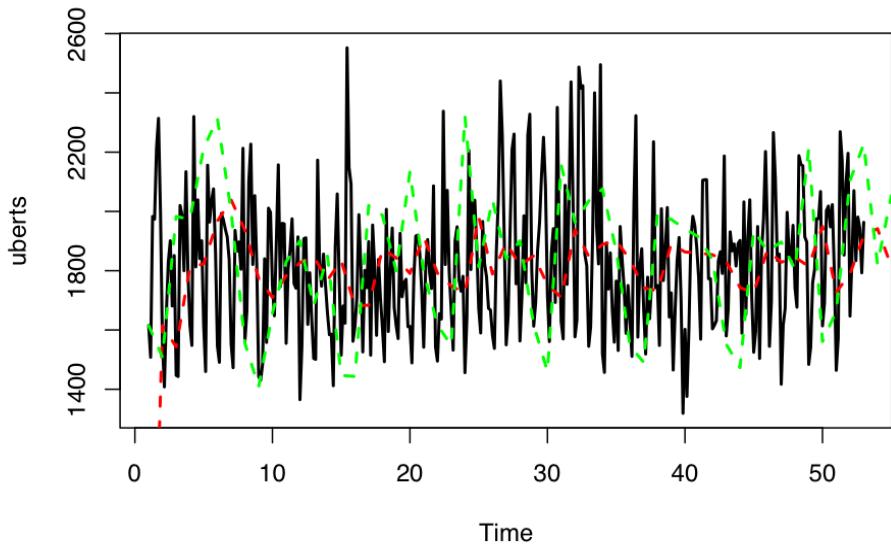
(d) Kalman kilter
library(TSPred)
library(KFAS)

fitkf<-fittestArimaKF(ubts, timeseries.test=NULL, h=12, na.action=na.omit,
level=0.9, filtered = TRUE, initQ=NULL,rank.by="AIC")

pred <- fitkf$pred

#extracting Kalman filtered and smoothed time series from the best fitted model
fs <- KFAS::KFS(fitkf$model,filtering=c("state","mean"),smoothing=c("state","mean"))
f <- fitted(fs, filtered = TRUE) #Kalman filtered time series
s <- fitted(fs) #Kalman smoothed time series
#plotting the time series data
plot(ubts,type='l',lwd=2,xlab="Time",ylab="uberts")
#plotting the Kalman filtered time series
lines(f,col='red',lty=2,lwd=2)
#plotting the Kalman smoothed time series
lines(s,col='green',lty=2,lwd=2)
#plotting predicted values
lines(ts(pred$mean,start = 2017.4),lwd=2,col='blue')
#plotting prediction intervals
lines(ts(pred$upper,start = 2017.4),lwd=2,col='light blue')
lines(ts(pred$lower,start = 2017.4),lwd=2,col='light blue')

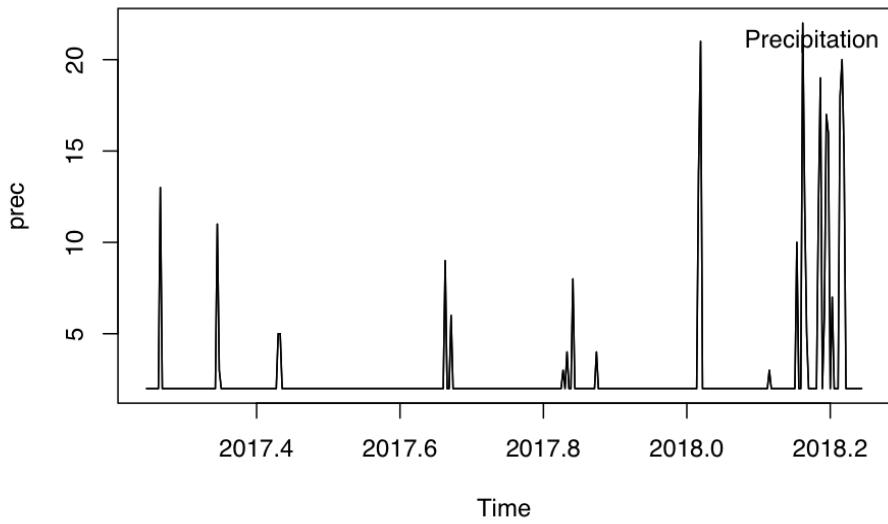
```



#### VAR

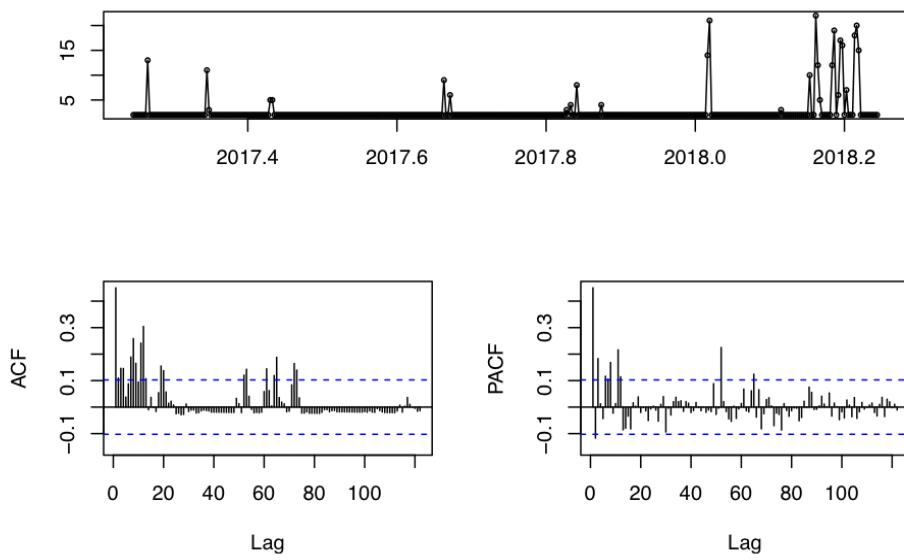
```
# Look at the data
prec =ts(precip,start=c(2017,as.numeric(format(ind[1],"%j"))),freq=365 )
ubts =ts(ubdata$travel,start=c(2017,as.numeric(format(ind[1],"%j"))),freq=365)

plot(prec)
nberShade()
lines(prec,ylab="Precipitation and Mean Travel Time by Uber")
legend("topright",legend=c("Precipitation"),text.col=c("black"),bty="n")
```



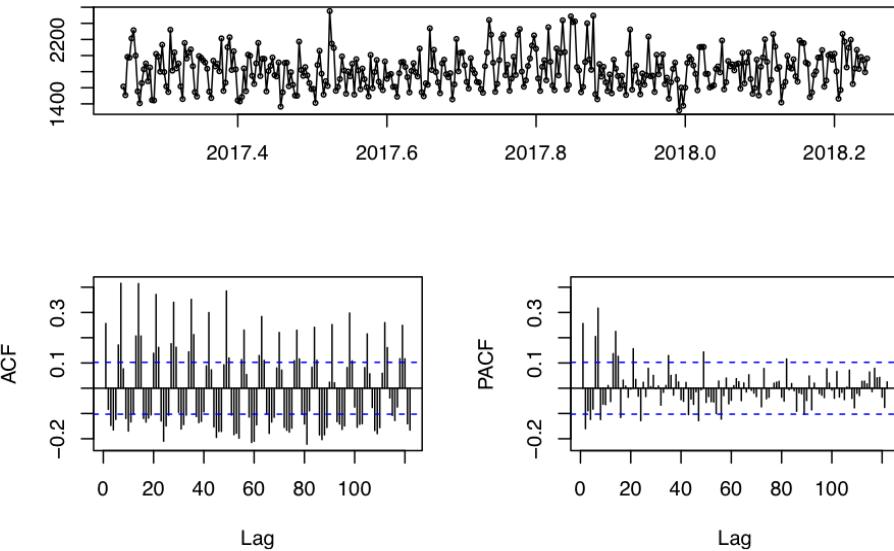
```
# Look at the ACF, PACF, and CCF (cross-correlation function)
tsdisplay(prec,main="Precipitation")
```

**Precipitation**



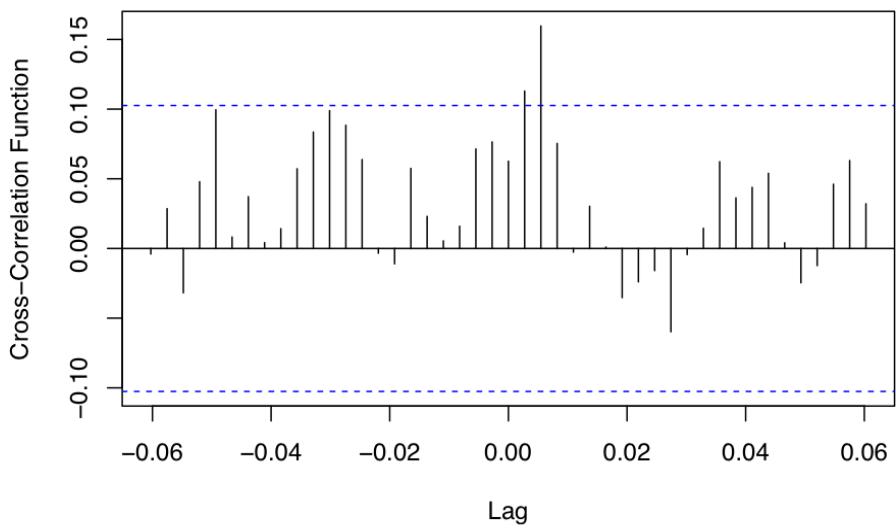
```
tsdisplay(ubts,main="Mean Travel Time by Uber")
```

**Mean Travel Time by Uber**



```
ccf(prec,ubts,ylab="Cross-Correlation Function", main = "Precipitation and Mean Travel Time by Uber CCF")
```

**Precipitation and Mean Travel Time by Uber CCF**



```
# Fit a VAR(p) model to the data  
# Combine the variables into 1 data frame first:
```

```

y=cbind(prec, ubts)
#y_ts=ts.union(starts, comps) # You can also use this function
y_tot=data.frame(y)

# To fit a VAR( $p$ ) model, simply call 'VAR' and set p=value
y_tot<-y_tot[1:365,]
y_model=VAR(y_tot,p=4)
summary(y_model)

## 
## VAR Estimation Results:
## =====
## Endogenous variables: prec, ubts
## Deterministic variables: const
## Sample size: 361
## Log Likelihood: -3291.576
## Roots of the characteristic polynomial:
## 0.6914 0.6914 0.6534 0.5413 0.5413 0.5219 0.5219 0.05785
## Call:
## VAR(y = y_tot, p = 4)
##
## 
## Estimation results for equation prec:
## =====
## prec = prec.l1 + ubts.l1 + prec.l2 + ubts.l2 + prec.l3 + ubts.l3 + prec.l4 + ubts.l4 + const
##
##          Estimate Std. Error t value Pr(>|t|)
## prec.l1  0.5110361  0.0533216  9.584 <2e-16 ***
## ubts.l1  0.0007117  0.0005850  1.217 0.224569
## prec.l2 -0.2050054  0.0592538 -3.460 0.000607 ***
## ubts.l2  0.0010439  0.0006039  1.729 0.084745 .
## prec.l3  0.1682997  0.0589933  2.853 0.004589 **
## ubts.l3 -0.0001164  0.0006050 -0.192 0.847495
## prec.l4  0.0116481  0.0528936  0.220 0.825829
## ubts.l4 -0.0001154  0.0005877 -0.196 0.844420
## const   -1.4166064  1.8763355 -0.755 0.450762
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## 
## Residual standard error: 2.462 on 352 degrees of freedom
## Multiple R-Squared: 0.2537, Adjusted R-squared: 0.2368
## F-statistic: 14.96 on 8 and 352 DF, p-value: < 2.2e-16
##
## 
## Estimation results for equation ubts:
## =====
## ubts = prec.l1 + ubts.l1 + prec.l2 + ubts.l2 + prec.l3 + ubts.l3 + prec.l4 + ubts.l4 + const
##
##          Estimate Std. Error t value Pr(>|t|)
## prec.l1   5.73549   4.81925   1.190  0.23480
## ubts.l1   0.26154   0.05287   4.946 1.17e-06 ***
## prec.l2   5.50192   5.35541   1.027  0.30496
## ubts.l2  -0.15367   0.05458  -2.816  0.00514 **

```

```

##  prec.13   -1.26871   5.33186  -0.238  0.81206
##  ubts.13   -0.06884   0.05468  -1.259  0.20892
##  prec.14    1.32522   4.78056   0.277  0.78178
##  ubts.14   -0.13696   0.05312  -2.578  0.01033 *
##  const     1970.82680  169.58479 11.621 < 2e-16 ***
##
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
##
## Residual standard error: 222.5 on 352 degrees of freedom
## Multiple R-Squared: 0.1224, Adjusted R-squared: 0.1024
## F-statistic: 6.134 on 8 and 352 DF, p-value: 2.061e-07
##
##
##
## Covariance matrix of residuals:
##      prec      ubts
## prec  6.06    16.17
## ubts 16.17  49502.47
##
## Correlation matrix of residuals:
##      prec      ubts
## prec 1.00000 0.02952
## ubts 0.02952 1.00000

# We interpret the coefficients in the usual way, but now have a
# system of equations. For example, for VAR(1) we have:
# y1 = c11 y(1,t-1) + c12 y(2,t-1)
# y2 = c21 y(1,t-1) + c22 y(2,t-1)
# The output from summary are cij, cov, and corr.

# Plot the fit and original data
quartz()
plot(y_model)

```

Diagram of fit and residuals for prec

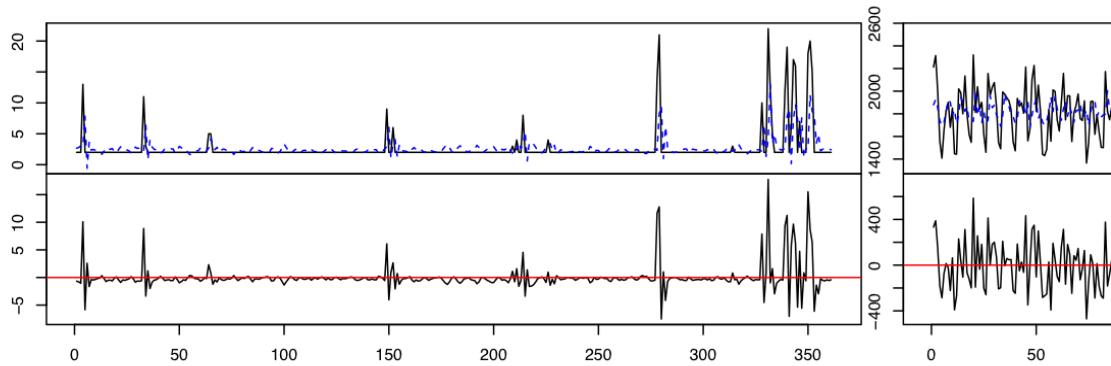
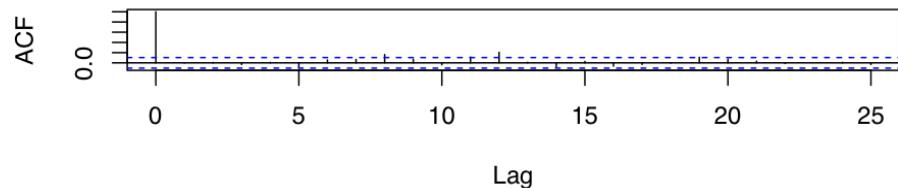


figure margins too large

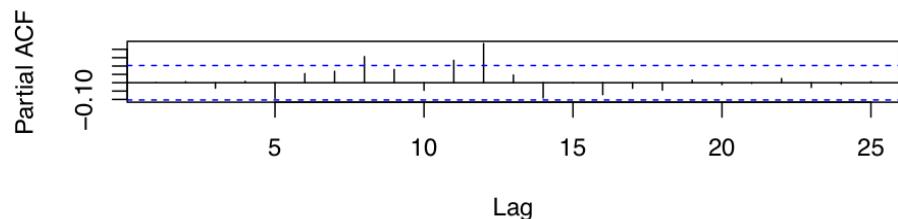
figure margins too

```
# Look at ACF and PACF
par(mfrow=c(2,1))
acf(residuals(y_model)[,1])
pacf(residuals(y_model)[,1])
```

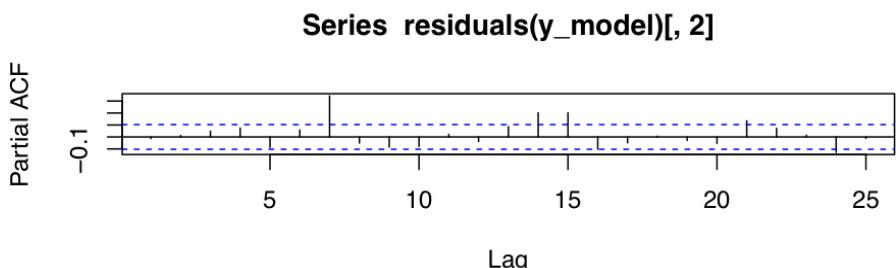
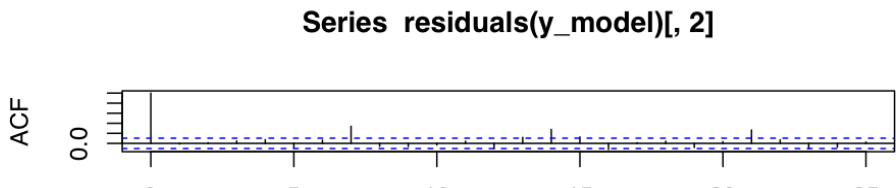
**Series residuals(y\_model)[, 1]**



**Series residuals(y\_model)[, 1]**

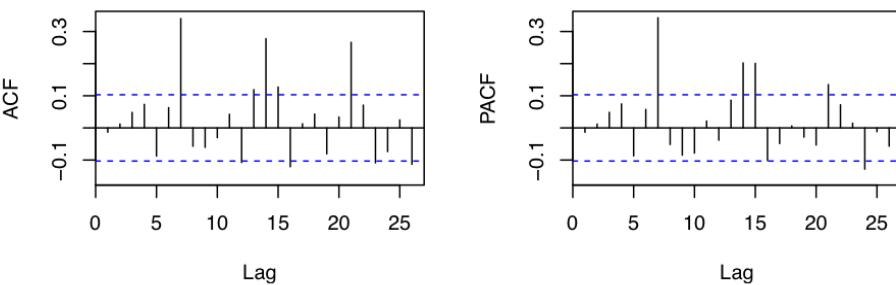
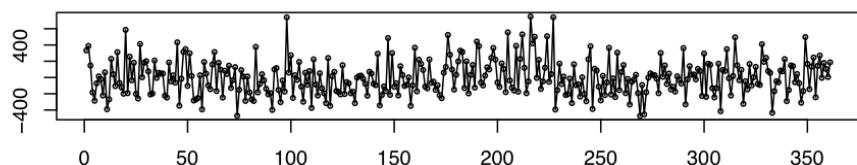


```
par(mfrow=c(2,1))
acf(residuals(y_model)[,2])
pacf(residuals(y_model)[,2])
```



```
tsdisplay(residuals(y_model) [,2], main ="Comps = prec(t-k) + ubts(t-k)")
```

**Comps = prec(t-k) + ubts(t-k)**



```
# Impulse Response Function
irf(y_model)
```

```

## 
## Impulse response coefficients
## $prec
##          prec      ubts
## [1,] 2.46171078 6.5668732
## [2,] 1.26269694 15.8366040
## [3,] 0.15874597 23.9190916
## [4,] 0.26936090 8.1047065
## [5,] 0.37442979 0.5332210
## [6,] 0.18177838 0.1800728
## [7,] 0.06029803 -0.8973371
## [8,] 0.05825468 -0.1811793
## [9,] 0.05121531 0.9365321
## [10,] 0.02705701 1.0885402
## [11,] 0.01571132 0.7191171
##
## $ubts
##          prec      ubts
## [1,] 0.00000000 222.394573
## [2,] 0.15828641 58.164895
## [3,] 0.35443666 -18.054597
## [4,] 0.17065039 -26.065064
## [5,] -0.02865184 -35.777055
## [6,] -0.04541401 -11.540400
## [7,] -0.02492658 6.581525
## [8,] -0.00645062 9.396940
## [9,] 0.01286801 6.986139
## [10,] 0.01852126 1.520397
## [11,] 0.01197190 -2.071960
##
##
## Lower Band, CI= 0.95
## $prec
##          prec      ubts
## [1,] 1.920963850 -18.1110234
## [2,] 0.913228599 -2.1221102
## [3,] -0.115319316 -2.9615759
## [4,] 0.013124991 -14.6782229
## [5,] 0.116912279 -18.4251474
## [6,] -0.052850311 -17.9874519
## [7,] -0.067376825 -7.4500466
## [8,] -0.027850583 -5.4918182
## [9,] -0.031578714 -2.4385456
## [10,] -0.023582371 -1.4055478
## [11,] -0.007781658 -0.7078437
##
## $ubts
##          prec      ubts
## [1,] 0.000000000 203.4322519
## [2,] -0.121632579 30.3515893
## [3,] 0.063504229 -40.6580099
## [4,] -0.115126526 -45.3770259
## [5,] -0.319489233 -56.7162810
## [6,] -0.250214270 -25.5315591

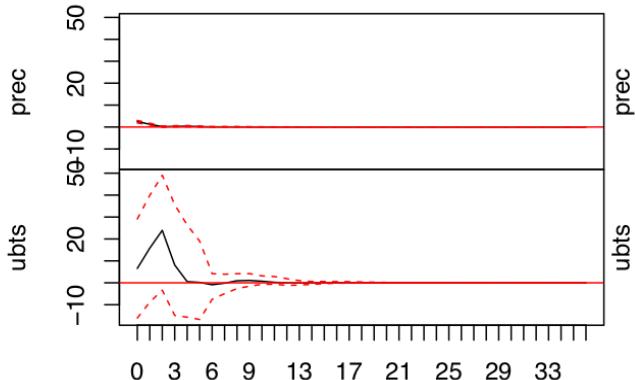
```

```

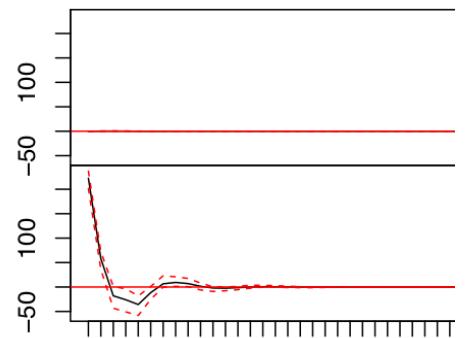
## [7,] -0.122850286 -1.3912408
## [8,] -0.060153115  1.1220402
## [9,] -0.020640745  0.1057883
## [10,] -0.005583414 -2.8958085
## [11,] -0.007420173  -7.3874819
##
##
## Upper Band, CI= 0.95
## $prec
##      prec      ubts
## [1,] 3.07683872 30.391871
## [2,] 1.61989814 39.280888
## [3,] 0.45049211 45.192755
## [4,] 0.55397786 29.923379
## [5,] 0.64808584 23.818441
## [6,] 0.41663756 15.305859
## [7,] 0.18300841 5.153154
## [8,] 0.21091574 3.533929
## [9,] 0.20056683 4.512264
## [10,] 0.13276400 3.755854
## [11,] 0.08553401 2.995258
##
## $ubts
##      prec      ubts
## [1,] 0.00000000 241.9772098
## [2,] 0.40130084 80.7478624
## [3,] 0.63063225 2.9988325
## [4,] 0.46749279 -5.3197409
## [5,] 0.23618693 -12.0477771
## [6,] 0.11684813 0.1174549
## [7,] 0.03232911 16.8463193
## [8,] 0.04890581 22.2931109
## [9,] 0.05963831 15.4225757
## [10,] 0.05555062 7.2496209
## [11,] 0.03973105 0.7041196
#quartz()
#pdf("irf.pdf", width=8, height=8)
plot(irf(y_model, n.ahead=36))

```

Orthogonal Impulse Response from prec



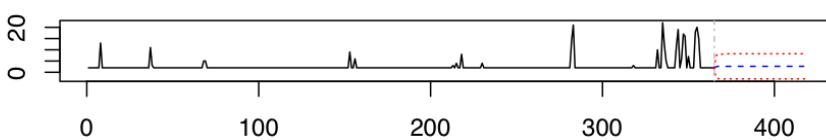
Orthogonal Impulse Response frc



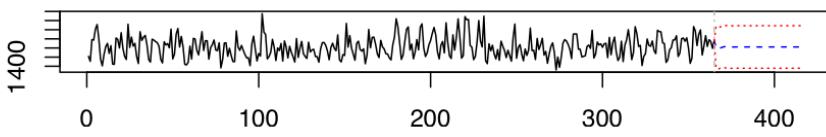
95 % Bootstrap CI, 100 runs

```
#dev.off()  
  
#Forecast  
#holdout_matrix = hold out data  
#var.predict = predict(object=y_model, n.ahead=52, dumvar=holdout_matrix);  
var.predict = predict(object=y_model, n.ahead=52)  
plot(var.predict)
```

**Forecast of series prec**



**Forecast of series ubts**



```

dev.print(device=postscript,"forecast.eps",width=7,height=7, horizontal=FALSE)

## pdf
## 2
#dev.off()

#Granger Test
grangertest(ubts ~ prec, order = 8)

## Granger causality test
##
## Model 1: ubts ~ Lags(ubts, 1:8) + Lags(prec, 1:8)
## Model 2: ubts ~ Lags(ubts, 1:8)
##   Res.Df Df      F Pr(>F)
## 1     340
## 2     348 -8 0.9389 0.4843
#Variance Decomposition (Forecast Error Variance Decomposition)
#plot(fevd(y_model, n.ahead = 5))

#CUSUM Plot
#plot(stability(y_model, type = "Rec-CUSUM"), plot.type="single")

```