

Rank versus Inequality—Does Gender Composition Matter?*

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Abstract

This study investigates the influence of gender composition on allocation decisions involving a rank–inequality tradeoff. In a real-time online experiment, participants chose to either alleviate inequality by relinquishing their current relative rank or exacerbate inequality while maintaining their current rank. Two essential features of the experiment are: 1) participants’ relative rank is the outcome of their real-effort performance and luck; 2) participants’ genders are naturally revealed by gender-specific nicknames. We found that female participants are more reluctant to relinquish their current relative rank when the persons ranked below and above them are of the opposite gender. This tendency was less pronounced in the male participants.

JEL Classification: C91, D63

Keywords: Gender composition, Positional concerns, Preferences for Redistribution, Last-Place Aversion, Perception of luck

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1 Introduction

Imagine a situation where you have to choose to either: 1) relinquish your current relative rank or position to reduce economic inequalities around you, or 2) maintain your current rank, which exacerbates inequalities; your absolute monetary payoffs are unaffected in both cases. Although drastically simplified, this situation captures the tradeoff relationship between concerns for relative rank and social preferences for distributional outcomes. For instance, such a tradeoff could affect the political support for redistribution schemes: Supporting more redistribution than the status quo may lead some supporters to relinquish their current relative rank while reducing economic inequalities. [Kuziemko et al. \(2014\)](#) provide survey evidence that some workers whose current hourly wage is slightly above the minimum wage stand against the idea of increasing the minimum wage. Increasing the minimum wage can reduce economic inequalities without making any harms to them, but it involves the workers relinquishing their current relative rank from ‘above the minimum wage’ to ‘at the minimum wage’. Our research questions regarding this hypothetical situation are as follows: How is your decision affected if you know the gender of the persons affected? Would your decisions differ if you obtain a relatively lower/higher rank? To answer our research questions, we examined the effect of gender composition on participants’ allocation decisions in a plausible environment where both individual performance and luck jointly determine payoff-relevant rank.¹

Our question stems from two actively studied research agendas: utility from a relative rank and preferences for redistribution. Although muted in most economic theories, concerns for relative rank drive our daily decisions and affect our well-being to a non-negligible extent.² In the two-dimensional space of incomes and favors toward progressive tax–redistribution policies, the standard economic models à la [Meltzer and Richard \(1981\)](#) predict that the preferences for redistribution should be negatively aligned with income levels. However, we sometimes observe "off-diagonal" preferences for redistribution: Some low-income citizens vote against redistribution, while some high-income citizens vote for it.

¹Another straightforward question is whether men and women act differently when trading rank versus inequality. [Kuziemko et al. \(2014\)](#) as well as [Martinangeli and Windsteiger \(2021\)](#) however do not report statistically significant differences in that regard.

²The concerns for relative standing were recognized by early philosophers such as [Veblen \(1899\)](#), [Keynes \(1930\)](#), and [Russell \(1930\)](#). Following [Duesenberry \(1949\)](#), some theoretical studies focus on the resource allocation to positional goods ([Frank, 1985](#); [Hirsch, 1995](#); [Hopkins and Kornienko, 2004](#)) and discuss possible policy interventions ([Frank, 2008](#)) and optimal taxation ([Aronsson and Johansson-Stenman, 2010, 2018](#)). For a broader review of the literature on relative rank and happiness, see [Clark et al. \(2008\)](#).

Among several plausible ways of attempting to answer "why the poor vote for the rich,"³ we are particularly interested in the role of utility from a relative rank.

The broad question regarding the relationship between concerns for relative position and preferences for redistribution has been examined in various ways. People may be averse to redistribution even when it is costless to themselves because they want to maintain their relative standing (Kuziemko et al., 2014), because they prefer to avoid decisions that reverse the existing rankings (Xie et al., 2017), or because they take into account the potential loss from other dimensions associated with changes in the redistribution policy (Kim, 2019; Gallice and Grillo, 2019). However, to the best of our knowledge, little is known about the role of gender composition of persons adjacent in rank to the decision maker. Gender composition could affect the decision maker's well-being via preferences for rank and redistribution. On the one hand, individuals could be habituated to care more about their intra-gender rank, as in many real-life reference groups consisting of either men or women only, such as sports teams, circles of friends, and professional networks in sectors dominated by one gender. In addition, partnering or marriage decisions involve only intra-gender comparisons. Given this context, people may care more about their relative position within the same-gender group. On the other hand, evidence from dictator games (Ben-Ner et al., 2004; Sharma, 2015) suggests that preferences for redistribution might depend on the gender of the decision maker's counterparts. We contribute to the literature by experimentally investigating the effect of gender composition on allocation decisions when facing a rank–inequality trade-off.

We faced two empirical challenges, which was the reasoning behind using a real-time online experiment to answer our research questions. One challenge was to make gender information salient without imposing an emphasis on it. The second challenge was to allow for within-participant variation in rank over rounds while ensuring that ranks should not be entirely random. In our experimental setting, the gender information of other participants

³Low-income citizens with the prospect of upward mobility may believe that redistribution will prevent them from prospering in the future (Benabou, 2000). Biased perceptions in individuals' evaluations of their own relative positions (Cruces et al., 2013) or heterogeneous beliefs about intergenerational mobility (Alesina et al., 2018) and about immigrants (Alesina et al., 2022) could affect preferences for redistribution. The public may be sensitive to the media coverage of household-income inequality, rather than the inequality itself (Kelly and Enns, 2010). Doubts about the political system and the government's efficiency may prevent the poor from supporting liberal parties or policies (Houtman et al., 2008). Economic inequality could have an informational value for social decisions (Corneo and Grüner, 2000). The social and cultural conservatism applied to a specific region (Frank, 2004) or religion (Guiso et al., 2003) may drive opposition to the political party which advocates for economic redistribution. Meanwhile, some high-income citizens' support for redistribution can be explained when values are viewed as luxury goods (Enke et al., 2023).

was revealed innocuously by displaying the (gender-identifiable) nicknames self-chosen for anonymity, and the luck factor and the participant’s performance co-determined the relative rank of the participant in approximately equal weights. Considering that both luck and effort affect an individual’s relative standing in diverse real-life circumstances, the inclusion of the luck factor in the experimental setup adds a significant dimension. We believe that both experimental features contribute to the methods in experimental economics, as the revelation of innocuous gender information can be used in many other experimental studies,⁴ and the adjustment of the random scaling parameter can allow us to examine another underexplored aspect regarding the preferences for redistribution—the (perceived) attribution of luck and effort for economic achievement.⁵

The experiment consisted of two main parts. Prior to the first part, the participants were asked to provide basic demographic information, including gender. In the first part, participants individually worked on a real-effort task, knowing that better performance would more likely lead to a higher payoff. The second part consisted of five decision rounds. In each decision round, the participants were randomly divided into groups of six and informed about their relative rank within the group. Their rank was based on their performance in the first part of the experiment, scaled by a luck factor (dice roll of a computer). The participants decide to allocate additional money to either a participant one rank above or one rank below⁶ while knowing the gender (via the gender-identifiable nicknames) of all group members. After each decision round, the participants were randomly reshuffled to form new groups, chose new nicknames, and were informed about new ranks. After the five rounds of the second part, the participants filled out the post-experiment survey, which included their perceptions about the role of luck and their performance.

We found that female participants were more reluctant to relinquish their current rela-

⁴Making gender salient without an unnecessary emphasis is easier said than done, as there is a tradeoff relationship. Sometimes concerns for demand effects make the gender salience weak: For example, [Baranski et al. \(2023\)](#) use colors (pink for female participants and blue for male participants) as cues for gender information in a three-person bargaining experiment where no explanation was given about what the colors meant. [Croson et al. \(2008\)](#) composed the single-gender sessions by recruiting subjects from fraternities and sororities, while the both-gender sessions were recruited from a service fraternity and a business fraternity. The opposite case also exists: For example, [Solnick \(2001\)](#) uses real first names (and tries to exclude participants with unrevealing first names) to examine gender differences in the ultimatum game.

⁵By adjusting the support of the luck scaling factor, we could build different institutions varying from a complete meritocracy (when luck plays no role) to a complete ex-ante egalitarian (when luck dominates effort).

⁶Following usage in the existing literature, we use the term ‘redistribution’ to refer to the act of allocating additional money to lower-ranked individuals. We acknowledge that this differs from conventional redistribution, which typically involves transferring resources from one person to another. In our case, no participant loses income as a result of this allocation. Nonetheless, we retain the term due to its prevalence in related studies and its descriptive value in capturing the resulting changes in income distribution.

tive rank—and thereby alleviate inequality—in mixed-gender environments where both the rank-below and the rank-above were male, compared to same-gender environments where both were female. A similar pattern was observed for male participants, but it was not as strong as that of females. We did not find support for last-place aversion (LPA, [Kuziemko et al., 2014](#)), the reluctance to allocate downward when ranked second from the bottom. A well-established finding is that low-income females are more likely to support redistributive policies ([Mengel and Weidenholzer, 2023](#)), highlighting the importance of the gender composition of decision makers. Our findings complement this by suggesting that the gender composition of redistribution beneficiaries may also shape support for such policies.

While our experimental setting does not have a direct real-world analog, the behavioral patterns we uncover, such as gender-conditioned preferences for status maintenance or redistribution, may shape public attitudes toward redistributive policies. For instance, in many countries, parental leave benefits and pension bonuses for parents disproportionately benefit one gender. Understanding how such identity markers interact with rank-based preferences may offer insight into the perceived fairness and political support for such policies.

The remainder of this paper is organized as follows. The following subsection provides an overview of the relevant literature. Section 2 outlines the experimental design, hypotheses, and procedures. Section 3 reports the results of the experiment, and Section 4 concludes.

1.1 Literature Review

Our experimental design adopts the encryption task used in [Erkal et al. \(2011\)](#) and the money allocation task used in [Kuziemko et al. \(2014\)](#), henceforth KBRN). The encryption task involves repeatedly decoding a random sequence of letters into corresponding numbers using a decoding table during a prespecified period. Our experimental setup was similar to theirs in the sense that an individual’s earnings were determined by the relative rank of the task performance, rather than by the absolute level of performance. The primary difference was that we considered a luck factor adjustment so that each participant’s rank could vary substantially by decision round, while their relative level of the absolute performance would vary less.

Although our research is not focused on reexamining KBRN’s LPA, it is worth discussing as our experimental design allows us to examine it in the second part. Motivated by the fact that some workers whose hourly wage is just above the minimum wage rate do not support the idea of raising the minimum wage, KBRN conducted an experiment where each

participant in a group of six was randomly assigned a payoff-relevant rank and asked to decide whether to allocate additional money to either a person ranked just above or a person ranked just below. KBRN found that participants with the second rank from the bottom are distinctively reluctant to allocate money to the person in the last place, calling this effect LPA. A replication of their experiment ([Camerer et al., 2016](#)) reported the opposite effect, questioning the validity of the observations. [Martinangeli and Windsteiger \(2021\)](#) replicated and extended KBRN, claiming that the way of assigning rank could matter.⁷ They found that when ranks were randomly updated in each decision round, as in KBRN, LPA is rejected. However, when considering a treatment where the randomly assigned ranks are maintained over the entire decision rounds, they find support for LPA. This suggests that randomly updated ranks may not be perceived as participants' actual ranks. We appreciate the way of assigning the same rank for the entire experiment, as it renders the participants some sense that the rank is endowed, but it does not allow for the observations of how the participants' decisions vary with rank. Our experimental design embraces the advantages of the design considered in KBRN and [Martinangeli and Windsteiger \(2021\)](#): By taking the performance in the real-effort task as a base, each participant's rank is determined by a rank of the adjusted performance that is scaled by a luck factor. In this way, we can provide the sense that the participant's "average" rank is endowed, yet the participant's rank varies by decision round.

Laboratory experiments on gender differences form another important strand of the related literature.⁸ We focus on reviewing a subset of studies that help us form conjectures on the effect of gender composition on redistribution. Overall, the results are mixed: Some studies suggest that women tend to allocate money to men at a lower rank, while others suggest the opposite. [Ben-Ner et al. \(2004\)](#) reported that in a dictator game context, women give less to other women when the gender of participants is known. [Dalmia and Filiz-Ozbay \(2021\)](#) reported that behindness aversion is distinct among women, and [Dasgupta et al.](#)

⁷Next to adding gender-specific nicknames, the second stage of our experiment follows the design of KBRN except for two differences. First, unlike KBRN, we use a real-effort task to determine participants' ranks, while KBRN assign ranks randomly in each round. Second, in our design, the amount that the decision maker must allocate to the player ranked above or below is smaller relative to the differences in initial endowments. In contrast, KBRN's design allows for rank reversals. [Martinangeli and Windsteiger \(2021\)](#) replicate KBRN's original design and extend it in three ways: (i) they use a smaller allocation amount to eliminate rank reversals, as we do; (ii) they introduce fixed ranks, instead of reshuffling them in each round; and (iii) they combine both modifications. Both studies investigate LPA. While KBRN find evidence for LPA, [Martinangeli and Windsteiger \(2021\)](#) cannot confirm this finding in their replication treatment. They however find evidence for LPA in their treatment involving no rank reversals and fixed ranks.

⁸For an extensive review on gender differences in economics experiments, consult [Croson and Gneezy \(2009\)](#) and [Niederle \(2017\)](#).

(2019) also found evidence of stronger behindness aversion among women. Regarding our experimental setting, these findings suggest that a female participant might allocate less to a woman of a lower rank. On the other hand, women give more to others in a dictator game (Eckel and Grossman, 2001; Bilén et al., 2021) and in a variant of the dictator game called the solidarity game (Selten and Ockenfels, 1998). In addition, women tend to favor redistributive policies more than men do (Alesina and Giuliano, 2011; Funk and Gathmann, 2015). These observations may indicate that women are more inequality-averse than men; hence, a female participant would allocate more to a participant at a lower rank, regardless of gender.

More nuanced still, our experimental setting provides noisy feedback on performance, which may interact with gender differences in how feedback is interpreted. Shastry et al. (2020) find that women, when receiving negative feedback in competitive settings, are more likely than men to attribute their lower standing to a lack of ability—even when the outcome is actually driven by luck. In our experiment, this perception could shape reactions to low rank in two contrasting ways. A participant who believes her low rank reflects poor performance may be less inclined to defend her position, and thus more willing to allocate downward. Conversely, a participant who believes she performed well but attributes her low rank to bad luck may see her current position as undeserved and either be less willing to sacrifice it further—or, alternatively, feel a sense of solidarity with others disadvantaged by luck and choose to redistribute downward.⁹

It is also worth noting that women tend to perform better in competitive environments when surrounded by other women (Gneezy et al., 2003), and exert more effort in such contexts (Mago and Razzolini, 2019). Although these findings are not directly transferable to our setting, where participants worked individually and did not know how their performance would translate into payoffs, they suggest that intra-gender reference points may influence behavior. If competitive motivation among women is partly shaped by comparisons with other women, a female participant may be especially reluctant to allocate downward when it reduces her standing relative to same-gender peers.

These considerations highlight the importance of understanding how perceptions of luck, performance, and rank interact with gender. Our design allows us to explore these dynamics, which were not directly addressed in previous studies.

⁹This additional interpretation was suggested by an anonymous reviewer, and we thank them for raising it.

2 Experimental Design, Hypotheses, and Procedure

2.1 Design

We adopted a within-subject design, where the primary source of variation was the gender composition of the adjacent participants. The aim of the experiment was to observe: 1) how gender composition affects the allocation decision; 2) how such effects vary in the participant’s current rank; and 3) how the perception of luck affects the allocation decision.¹⁰

Prior to the main experiment, the participants were asked to provide basic information, such as gender, age, field of study, and previous experience of online experiments, without leaving any questions unanswered. Since our primary manipulation was to make the group members’ genders salient, while avoiding the experimenter demand effect (Zizzo, 2010) for the participants to become aware that the revelation of the gender is crucial, we did not emphasize any part of the questionnaire.¹¹

The experiment consisted of two parts. In Part 1, the participants worked on an encryption task (Erkal et al., 2011) for five minutes. Per trial, each participant was given a sequence of five random letters and a decoding table showing how letters map into numbers. Their task was to decode the sequence of letters into corresponding numbers in the table. For example, if the text sequence YDDYO is given, and a decoding table states that D, O, and Y correspond to 3, 6, and 8, respectively, then the correct answer is 83386. A new decoding table was provided for each text sequence. The participants were told that the more correct answers they gave, the more money they were likely to earn. We also informed the participants that final earnings would be based on their decisions, others’ decisions, and some luck, but we only provided further details of how their earnings would be determined in Part 2. Previous studies reported no significant gender differences in encryption task performance (Erkal et al., 2011).

Part 2 of the experiment comprised five decision-making rounds of the money allocation task à la Kuziemko et al. (2014). Figure 1 illustrates how Part 2 proceeded.

At the beginning of each round, the participants were randomly assigned to a group of six, and each picked their nickname from a provided gender-specific list of alternatives. We explained that picking a nickname is a way to make every participant remain anonymous, yet clearly distinguish other participants in the group by their nicknames. After they

¹⁰The complete experimental instructions can be found in Appendix A.

¹¹In addition, since the first part of the experiment—the individual real-effort task—had nothing to do with gender, we believe that the experimenter demand effect, if any, is negligible.

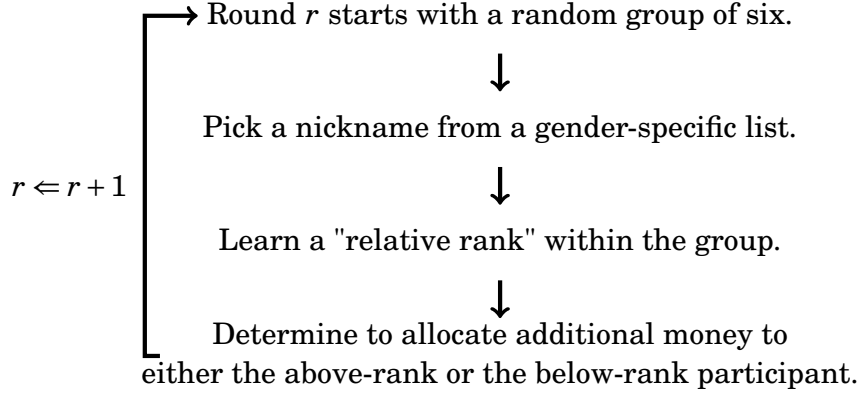


Figure 1: Flowchart of Part 2

selected their nicknames, a server computer calculated and displayed their *relative rank* within the group. The rank was determined by their performance in Part 1, scaled by the outcomes of the computerized dice rolls. Specifically, the rank of participant i , R_i , is determined by $\#\{j \in G_i | P_j \times [D_j + 2] \geq P_i \times [D_i + 2]\}$, where G_i is the group for participant i , P_j is the performance measure (i.e., the number of correct answers in Part 1) of participant j , and D_j is the random outcome of a computerized dice roll for participant j . In case of a tie, the participant whose participant number (a unique number internally assigned to each participant) was smaller received a better rank. We instructed the participants that the computer randomly rolled a dice for each participant, calculated the "adjusted performance," which is the number of correct answers in Part 1 multiplied by [dice roll+2], and showed the relative rank of the adjusted performances within the group. For example, if participant 1 gave 10 correct answers in Part 1, her adjusted performance could vary from 30 ($= 10 \times [1 + 2]$) to 80 ($= 10 \times [6 + 2]$). The participants received different amounts of money according to the ranking: The person obtaining rank r , $r \in \{1, 2, \dots, 6\}$, received $17 - 2r$ euros. As the participants were only informed about their relative rank, they could infer neither the realization of the dice roll nor the other group members' actual performance measures.

After every participant observed the rank of all group members, hence the initial money allocations, they were asked to allocate two additional euros to one of the two adjacent participants in ranking.¹² For example, a third-ranked participant was required to allocate two additional euros to either a person ranked second or fourth. For notational simplicity, we refer to the person who is at one rank above the participant as Rank-Above and the

¹²Participants were not allowed to decline the allocation of the additional two euros. Our experimental design aimed to capture a rank-inequality tradeoff, and with the inclusion of the option for no allocation, efficiency concerns also could become relevant for participants.

person who is at one rank below the participant as Rank-Below. The participant ranked first (last) had to allocate money to the Rank-Below (Rank-Above). Since the base payoff of each rank increases by two euros, allocating the additional two euros to the Rank-Below implies either that the participant sacrifices her relative rank (if ranks are counted from the bottom), or that she allows the person below to catch up and share the same rank (if ranks are counted from the top). Participants were informed that the allocation decision of only one randomly selected participant would be implemented. Final payoffs were not displayed during Part 2, such that other participants' decisions remained unknown.

For every new round, the participants repeated the procedures mentioned above. All participants were reshuffled and assigned to a new group of six. The computer rolled new dice for the new round, independent of previous outcomes. The new ranking was then determined based on their performance in Part 1, their new group members' performance, and new dice rolls. Once all five rounds ended, the computer randomly selected one round and one group member's allocation decision for the actual payment. Until the end of the experiment, the participants did not receive any feedback on the selected outcome.

There are several points worth discussing. We made the participants' gender salient by asking them to select one fictitious name from a list of popular names.¹³ Thus, it was crucial to make the participants unambiguously grasp the gender of other participants. For this purpose, we carefully selected a list of first names. First, we collected 100 boy names and 100 girl names that were popular for babies born in the 1990s and 2000s, excluding known unisex names. Most participants were born in the 1990s and 2000s, so the name list contains contemporary names from their perspectives.¹⁴ We then asked 80 Amazon Mechanical Turk (mTurk) participants to classify the 200 names as male, female, or unsure/unisex and used 60 names for each gender on which the survey participants formed a near-consensus as boy and girl names.¹⁵ We filtered survey respondents to obtain 29 German-speaking,

¹³Chakraborty and Serra (2023) also asked experiment participants to choose a nickname from a list of gender-specific names. However, their primary purpose was to prevent confounding bias of race, nationality, or ethnicity associated with the participant's actual name, so they used distinctively white-sound names taken from Bertrand and Mullainathan (2004), and the chosen nickname was used throughout the entire experiment.

¹⁴An undesirable side effect is that participants might attribute good or bad personal impressions to certain first names. However, we claim that such name effects are orthogonal to rank and gender composition and thus do not affect our primary observations in aggregate.

¹⁵The following boy names are used: Aaron, Adam, Albert, Anthony, Arthur, Christopher, David, Derek, Edward, Ethan, Frank, Fred, Hugo, James, Jason, Patrick, Peter, Richard, Robert, Ronald, Steven, Thomas, Walter, Xavier, Benjamin, Brandon, Bruno, Charles, Christian, Daniel, Diego, Eric, Gregor, Isaac, Ivan, Jacob, John, Jonathan, Kevin, Leonard, Manuel, Marco, Matthew, Nicholas, Norman, Paul, Phillip, Ralph, Samuel, Scott, Stanley, Victor, Vincent, William, Alexander, Andrew, Brian, George, Henry, and Oscar. The following girl names are used: Alyssa, Elena, Helena, Isabella, Jasmine, Jennifer, Lisa, Luisa, Monica, Sarah,

37 European, and 14 Asian respondents. This procedure aimed to match the proportion of survey respondents with that of participants in the other experiments that used the same participant pool.¹⁶ Also, to avoid potential confusion, every experiment participant received different name options per round; therefore, the participants were not concerned about possible duplication within a round.

We carefully calibrated the range of the random scaling factor from 3 to 8 (dice roll+2), which determined the adjusted performance. If the random scaling factor varied too much compared to the performance variance among participants, then the actual performance would not matter for the relative rank. In contrast, if the random scaling factor varied too little, the relative rank across rounds would also vary too little. Ideally, we wanted to ensure that an increase in the actual performance rank would on average raise the adjusted performance rank by 0.5.¹⁷

The main experiment was followed by a post-experiment survey, which included the participant’s guess of the actual rank in performance, perceptions about the luck factor, and a general view about the inequality caused by luck.¹⁸

2.2 Hypotheses

We had three testable null hypotheses regarding how (perception of) ranks and gender composition affect allocation decisions. When previous findings from related studies are mixed, we hypothesize that there are no effects.

Hypothesis 1. *There is no last-place aversion. Fifth-ranked participants are equally as likely to allocate downward as participants holding ranks 2, 3, and 4.*

Hypothesis 1 was based on the previous findings on LPA. Using random assignment of ranks in every round, KBRN provided supporting evidence for LPA, rejecting this hypothesis. [Camerer et al. \(2016\)](#) and [Martinangeli and Windsteiger \(2021\)](#) both replicated the same experimental setup and provided evidence against LPA. Our experimental setting differed in a crucial dimension: The existence of the performance factor in the determination

Sylvia, Alicia, Angelica, Christina, Jessica, Julia, Katherina, Lara, Laura, Marisa, Olivia, Rebecca, Regina, Samantha, Sandra, Selina, Teresa, Claudia, Cynthia, Elizabeth, Ella, Eva, Lena, Linda, Lydia, Melissa, Mia, Michaela, Natasha, Rose, Sabrina, Sophia, Ariana, Bianca, Cassandra, Clara, Daniela, Diana, Emily, Emma, Erica, Lily, Maya, Vanessa, Alina, Amanda, Gloria, Maria, Stella, and Victoria.

¹⁶See [Online Appendix](#) Section A for mTurk instructions and other details.

¹⁷See [Online Appendix](#) Section B for more details.

¹⁸Before answering the survey, participants learned their adjusted performance rank in five different groups. This noisy feedback on performance is thus incorporated in their perceptions.

of ranks is very consequential on individual behavior. It potentially affects the system of attitudes towards ranks and preferences for alterations of the distribution. However, we believe that most real-world rankings, such as in school, the workplace, or sports, reflect a mix of both individual effort and random factors. By incorporating both performance and luck into the determination of rank in our experiment, we aim to better mirror these real-world environments. This improves the external validity of our findings, particularly with regard to how individuals respond to rank positions that may not be fully deserved.

Hypothesis 2. *The genders of the Rank-Above and Rank-Below do not affect an individual's redistribution decision.*

Hypothesis 2 states our null hypothesis regarding the effect of gender composition. As described in the literature review section, our conjectures regarding the effect of gender composition on allocation decisions are mixed.

Hypothesis 3. *The tendency to allocate downward may be influenced by individual perceptions of performance and luck. However, gender differences in allocation behavior are not explained by differences in these perceptions.*

Hypothesis 3 examines whether participants' perceived "deservingness" based on their beliefs about performance and the role of luck affects their allocation behavior. Specifically, participants who believe their low rank reflects bad luck might either (1) allocate upward to preserve their undeservedly low rank or (2) allocate downward in solidarity with others they perceive as similarly disadvantaged. Given that gender differences in confidence and attribution are known (Niederle and Vesterlund, 2007; Shastry et al., 2020), we test whether these perceptions explain the observed gender differences in allocation behavior. This helps determine whether our main finding is robust to an important class of alternative explanations. To the best of our knowledge, no previous studies report how the differences in individual perceptions are interacting with the effect of the gender composition in making allocation decisions.

2.3 Procedures

The experimental sessions were conducted in November 2021 and January 2022 and included participants recruited from the Mannheim Laboratory for Experimental Economics (mLab) of the University of Mannheim. Invitations were sent to members of the partici-

pant pool who had already provided gender information in the recruitment system.¹⁹ Instructions were provided in English. Owing to COVID-19 restrictions, we did not bring participants to the laboratory. Instead, we invited them to join an online meeting to receive instructions from the experimenter, distributed a link for participating in the online experiment, and paid them via either PayPal or bank transfer. Six sessions with 18 participants (nine females and nine males) were conducted, hence a total of 108 participants participated in the experiment. We used an interactive online platform called LIONESS (Live Interactive Online Experimental Server Software, [Arechar et al., 2018](#)). After the participants joined an online meeting, the experimenter disabled their profile photos and asked them to turn off their webcam. Furthermore, the participants were asked to edit their displayed names to two letters they arbitrarily chose so that their identities, and hence decisions, remained anonymous to the experimenter as well as the other participants. The participants were asked to carefully read the instructions for each part of the experiment and pass a comprehension quiz.

The average payment per participant was 10.33 euros. Payments were made via online transfers after receiving the personal payment code that was generated at the end of the experiment. Each session lasted less than 40 minutes.

3 Results

In Part 1, on average, each participant correctly decoded 24.79 text sequences (SD 3.93). As the minimum and maximum performances were 15 and 36, respectively, there is a nonzero probability that the highest-performing participant could be ranked lower than the lowest-performing participant. The correlation between the ranks based on adjusted performance and those based on actual performance was 0.4065 ($p < 0.001$), which means that the luck factors perturbed the relative rank of the group members to some extent, while the rank of actual performance significantly affected the relative rank.

As intended, there were no differences based on gender in the participants' performance. The average number of correctly decoded text sequences was 24.76 (SD 3.78) for females and 24.81 (SD 4.11) for males, and the difference was not statistically significant (Kolmogorov-Smirnov (KS) test, $p = 0.60$; Mann-Whitney (MW) test, $p = 0.90$). The distribution of ranks based only on actual performance was not statistically different by gender (KS, $p = 0.90$; MW, $p = 0.81$) either. Thus, we claim that our findings are not driven by gender differences

¹⁹Using that information, we invited only subjects who identify as either male or female.

in the ability to perform in Part 1.

The randomly determined groups of six in Part 2 were fairly balanced in terms of the overall gender composition: 37.78% of groups consisted of exactly three males and females. 90% of groups were composed of at least two and at most four females.²⁰

3.1 The effects of rank and gender composition

Figure 2 shows the proportion of decisions allocating additional money to the Rank-Below, collated by rank. Trivial decisions from participants holding rank 1 and rank 6 were excluded, as they had only one option to allocate money. For all ranks, the majority of participants (80.83%) preferred to sacrifice their relative rank by allocating money to the Rank-Below.²¹ Thus, participants tended to prioritize inequality over rank concerns. This tendency was slightly more pronounced for females, as 83.05% of female participants allocated downward, and 78.69% of males did so. However, the difference was not statistically significant ($p = 0.51$). Another notable finding was that we did not observe LPA. Although statistically insignificant, we observed the opposite pattern: The participants at rank 5 allocated additional money to the Rank-Below more often than participants at any other rank. These results support our Hypothesis 1. A potential explanation for the absence of LPA lies in the noisy performance feedback and participants' perceptions of their rank. If participants holding low ranks attributed their results to bad luck, they might resort to solidarity with participants ranked below, who could also be seen as disadvantaged by luck. This behavioral tendency may align with previous findings of greater support for redistribution when inequality is perceived to result from chance rather than merit (Alesina and Angeletos, 2005; Almås et al., 2020; Fong, 2007).

Result 1. *Fifth-ranked participants allocated downward as much as other participants.*

Unlike rank, we find significant effects of the gender composition of the Rank-Above and Rank-Below on allocation decisions. From the perspective of the member holding rank $r \in \{2, 3, 4, 5\}$, when both members at rank $r - 1$ and $r + 1$ are of the same gender as member r , we call it the *same-gender* situation (SG, 75 observations). When both Rank-Above and

²⁰Simulations (random permutation of 9 males and 9 females to form three groups of six for 1,000,000 iterations) show that the empirical CDF for the number of females per group is close to the true CDF under random allocation. In particular, the probability of a group with three females is 37.97%, and that with two to four females is 86.88%.

²¹We do not find any order effects: The share of subjects allocating downward in the first round (84.72%) is almost the same as in the last round (83.33%), and the estimated coefficient of the round number is statistically insignificant for all regressions we performed.

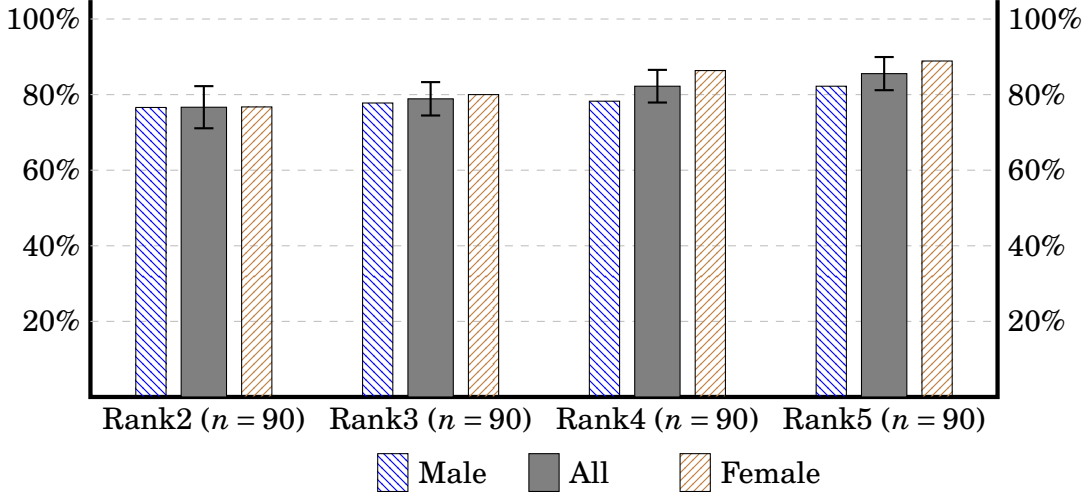


Figure 2: Proportions of allocating downward, by rank

Error bars represent standard errors from an OLS regression (no controls) clustered at the individual level. Results of OLS and Probit regressions can be found in Appendix B, Table B1.

Rank-Below members are of the opposite gender from member r , we call it the *mixed-gender* situation (MG, 88 observations). When the Rank-Above is of the same gender as member r , and the Rank-Below is of the different gender from member r , we call it the *same-above* situation (SA, 99 observations). The remaining situation is analogously called the *same-below* situation (SB, 98 observations).²²

Figure 3 shows the proportion of decisions allocating additional money to the Rank-Below, collated by gender composition. Compared to other situations, there was a distinct tendency to allocate money to the Rank-Below in the SG (t-test²³, $p = 0.005$). The proportion of downward allocation in the MG was the lowest, and it was not statistically different from that in the SA ($p = 0.895$). The proportion of downward allocation in the SB was slightly larger than that in the MG, but the difference was not statistically significant ($p = 0.151$) either. These observations imply that the participants were more inclined to allocate

²²Since the participants are asked to consider two adjacent participants, we believe that it is appropriate not to differentiate the situations further with respect to the gender composition of participants at remote ranks. Although the contexts are somewhat different, Fisman et al. (2020) reported that experimental participants are locally competitive, preferring to decrease the income level directly above them, but not two positions above. This evidence suggests that the gender composition of the participants two or more ranks away would not lead to noticeable differences. Indeed, the number of females in a group has no statistically significant effect ($p=0.35$) on allocation behavior when controlling for the genders of the decision maker and the adjacent players. This holds also true when only considering female ($p=0.82$) or male ($p=0.76$) decision makers.

²³Unless otherwise stated, we have reported the p value of an estimated coefficient of a linear regression model controlling for ranks. The standard errors for the test statistics are clustered at the individual level. For regression results, including Probit regression results that are not reported in the main context, see Appendix B.

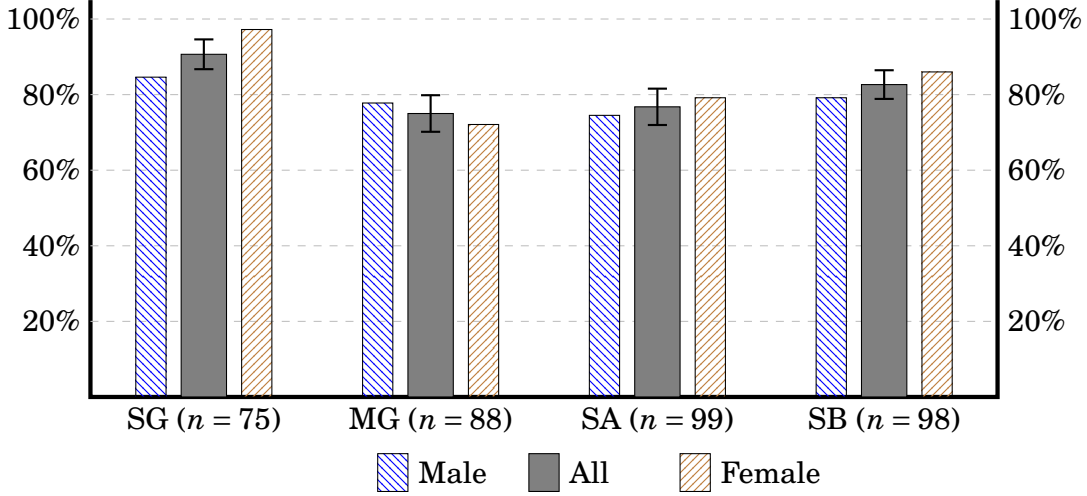


Figure 3: Proportions of allocating downward, by gender composition

Error bars represent standard errors from an OLS regression (no controls) clustered at the individual level. SG/MG=rank $r - 1$ and $r + 1$ have the same/different gender. SA/SB=the rank above/below has same gender. Results of OLS and Probit regressions can be found in Appendix B, Table B2.

the money to the Rank-Below, especially when the recipient of the money and the Rank-Above shared their gender. Moreover, this tendency of allocating more to the Rank-Below was driven by the behavior of female participants. In the SG, almost every female (97%) allocated money to the female Rank-Below, while 85% of males allocated it to the male Rank-Below ($p = 0.126$).²⁴

Result 2. *Participants allocated downward more in the same-gender situation than in the mixed-gender situation. This tendency was driven by females.*

To further examine the significant differences between the SG and MG, we checked the proportions of downward allocation by rank only for these two situations. We found suggestive evidence that participants at ranks 2 and 5 allocated downward less often in the MG. Figure 4 shows the proportions of downward allocation in the SG and MG, collated by rank. The participants holding ranks 2 and 5 allocated additional money to the Rank-Below significantly more frequently when the adjacent participants were of the same gender ($p = 0.001$). This difference was more distinctive for females ($p = 0.014$) than for males ($p = 0.048$),²⁵

²⁴This observation might be due to in-group bias based on gender: Rudman and Goodwin (2004) reported that women tend to be more favorable to other women than men to other men. Conversely, given that the patterns of the overall observations are weak but similar for men, this observation might indicate that women are more sensitive to the experimental context, as reported in Croson and Gneezy (2009) and Miller and Ubeda (2012).

²⁵When separately examining the effect for: 1) the male participants ranked fifth, 2) the male participants

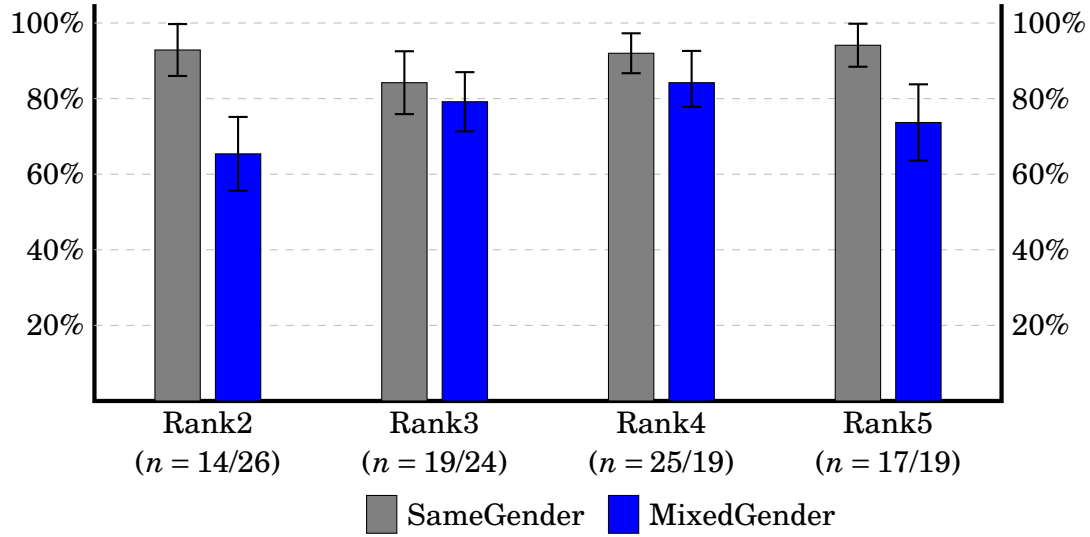


Figure 4: Proportions of decisions allocating to the Rank-Below in the SG and MG, by rank. Error bars represent standard errors from an OLS regression (no controls) clustered at the individual level. Results of OLS and Probit regressions can be found in Appendix B, Table B4.

and many second-ranked and fifth-ranked female participants were reluctant to bring the male Rank-Below up to the same rank. Although this finding should be interpreted with caution given potential limitations in subgroup-level statistical power, one possible interpretation for participants ranked fifth is that their reluctance to allocate downward may reflect mixed-gender LPA. The reluctance of participants ranked second to allocate downward is also observed by [Martinangeli and Windsteiger \(2021\)](#) in their treatment involving fixed ranks and no rank reversals.²⁶ A possible explanation for this pattern in the MG is that participants want to avoid rankings that are dominated by other genders. If a second-ranked female allocated downward, two men would end up in the first place and (a shared) second place.²⁷

ranked second, 3) the female participants ranked fifth, and 4) the female participants ranked second, the effect is statistically significant only for the last case ($p = 0.027$). However, there are only a few (16–20) observations per case, so we pooled the observations for ranks 2 and 5.

²⁶[Xie et al. \(2017\)](#) show that even third-party dictators have an aversion to change a pre-existing ranking. [Martinangeli and Windsteiger \(2021\)](#) thus argue that allocations involving rank reversals would not allow for a clean test of last-place aversion.

²⁷Some may argue that the second-ranked participants' reluctance to allocate downward might be related to first-place loving ([Gill et al., 2019](#)), but neither mixed-gender first-place loving nor within-gender first-place loving can clearly explain our observation in the MG. If a female participant cares for a relative standing with respect to male participants, she ranks second, and her decision is irrelevant to the first-place loving. If a female participant cares for a relative standing among female participants, she ranks first, but neither of her decision affects her within-gender rank.

3.2 Perceptions of Performance and Allocation Decisions

Participants were generally overconfident about their actual performance rank in their post-experiment survey, with the average difference between the guessed performance rank and the actual performance rank among the 18 participants per session being equal to 1.68 (SD 5.58). This difference was statistically significant (t-test, $p = 0.002$). Male participants were more overconfident than female participants regarding their actual performance rank.²⁸ Figure 5 shows the empirical cumulative distribution functions of the actual performance rank by session and perceived performance rank. Male participants' guesses first-order stochastically dominate their actual ranks, implying that the men were overconfident, regardless of their actual performance. Female participants were also overconfident, but the magnitude was much smaller, with some high-performing females less confident about their relative performance rank. On average, male participants guessed that their relative performance rank was 2.5 ranks (SD 5.40) better than their actual rank, while female participants guessed that their rank was 0.85 ranks (SD 5.67) better. Although statistically insignificant, the average overconfidence of males is greater than females ($p = 0.054$, controlling for actual performance rank).

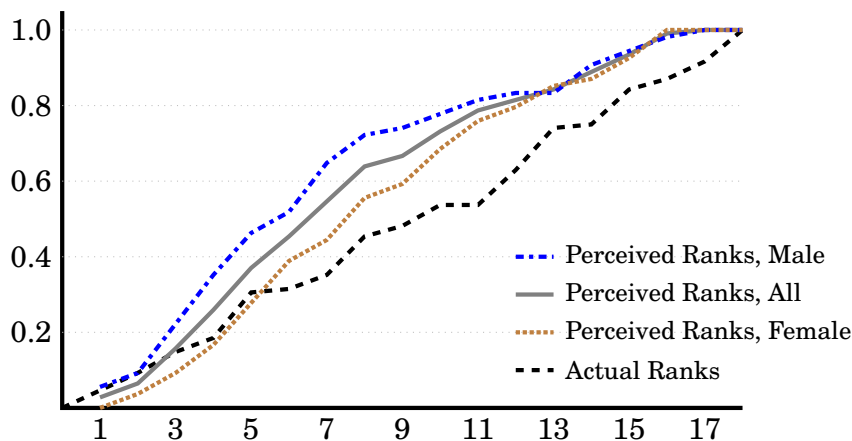


Figure 5: Empirical cumulative distribution functions for relative ranks

Moreover, we also notice that the participants whose performance was worse tend to overestimate their relative performance in a more significant manner: For each one-rank decrease²⁹ in actual performance, the gap between actual and perceived rank increased by

²⁸This finding is in line with Niederle and Vesterlund (2007), who found that men are more overconfident about their relative performance on a math task, although there is no significant gender difference in actual performances.

²⁹Since the participant with the best performance is ranked first (1) and the one with the worst performance

0.74 on average ($p < 0.001$). Combined with the gender difference in overconfidence, this suggests that low-performing male participants were more likely to perceive their monetary allocation as worse than deserved. This motivates us to examine the behavior of fifth-ranked males in particular. If they felt unfairly disadvantaged by bad luck, they might either (1) avoid allocating downward in an effort to preserve their already low but undeserved rank, or (2) allocate downward out of solidarity with others they believe also suffered from bad luck. We find no evidence supporting the second interpretation. Allocation decisions by fifth-ranked male participants were not significantly different from those of males ranked 2nd to 4th. In fact, though statistically insignificant ($p = 0.261$), fifth-ranked males in the MG allocated downward less often, suggesting that the mutual solidarity narrative may not apply, at least for males in our setting.

3.3 Perceptions of Luck and Allocation Decisions

During the post-experiment survey, we asked the participants to evaluate the following three statements on a 1–5 Likert scale:

1. "I feel that, overall, the random adjustment of the performance places me worse ranks than my actual ranks."
2. "I feel that the effect of random adjustment on the relative rank is stronger than that of actual performance."
3. "A society should accept inequality caused by differences in luck."

The first question was used to check the validity of the survey answers. If the participants strongly agreed with the statement, they would report their performance rank better than the rank from Part 2; thus, the difference between the actual performance rank and the perceived performance rank would increase. Confirming this narrative, we found a strong and positive correlation between them ($\rho=0.2160$, $p < 0.001$).

The second question asked about the perceived importance of the luck factor. We examined whether participants who felt that the luck factor's impact was strong behaved differently from those who did not. We found that male participants agreed that the luck factor's impact was stronger than the actual performance (KS, $p = 0.017$) more often than female participants, echoing the above-reported higher overconfidence of males. Moreover,

is ranked last (18), a decrease in rank means that the rank increases. We are aware of this potential confusion, so we clarified its meaning to the experiment participants.

we found that the perceived importance of the luck factor does not significantly affect allocation decisions in any way.³⁰

Regarding the last question, we found that the participants who strongly agreed with the statement tended to allocate additional money to the Rank-Below less often ($p = 0.061$). From the perspective of the decision maker, allocating downward involves giving up one's rank to mitigate inequality. Our preferred interpretation for this observation is that those who embrace luck-driven inequality would be less inequality-averse and put relatively more weight on their concern for relative rank. However, this association between the attitude of accepting inequality and the tendency of downward allocation does not alter the effects of gender composition we report in Result 2. Results of OLS and Probit regressions in Appendix B, Table B6 show no noticeable differences in the estimated coefficients on gender composition with or without the post-experimental measures.

Result 3. *The participants' beliefs about how much their performance-based rank was higher than the announced rank was not associated with their allocation decisions. Participants with a general attitude of accepting inequality caused by differences in luck tended to allocate downward less often. However, after controlling for the perceptions, the main finding we report in Result 2 holds.*

4 Conclusions

In our experiment, we added a gender composition factor to the existing studies on relative rank and distributional outcomes. We hypothesized that decision makers' weightings between concerns for their relative position and equality of outcomes might depend on the gender of the potentially affected persons. Thus, the gender composition of the reference groups could affect allocation decisions. However, there is no clear indication of whether people are relatively more or less sensitive to rank concerns when affected persons are of the same or opposite gender. This uncertainty motivated the present experiment.

We let the participants encounter different payoff-relevant ranks, innocuously provided the gender information of others, and asked them to allocate additional money to a person at one rank below or above. The crucial variation was whether the Rank-Below was of the same gender. When the Rank-Below was of the different gender, allocating downward did not alter the decision maker's within-gender rank. The experiment participants allocated

³⁰This also holds true when controlling for the interaction of the perceived importance of the luck factor with the rank obtained.

downward more often when the Rank-Below was of the same gender. This tendency is more distinctive when the Rank-Above also was of the same gender. As this occurred primarily among women, we interpreted that solidarity is higher among women. Allocating money to a female Rank-Below minimizes inequality within the group of women, whereas allocating money to a female Rank-Above exacerbates income differences.

Meanwhile, we found no support for LPA. The participants at the fifth rank (one rank above the bottom) allocated money downward as much as the participants at other ranks. Combined with the findings on same-gender solidarity, these results indicate that gender composition is pertinent in allocation decisions that involve a rank-inequality tradeoff, potentially more than the overall rank.

Future research is needed to investigate the many underexplored questions on the trade-off between concerns for relative rank and distributional outcomes. For the current study, the participant's performance and luck were similarly attributed to determine the observed ranks, and this may have limited participants' sense of rank ownership. It would be worth investigating the effects of different rank-determining schemes, ranging from pure luck to pure performance. Future work could explore how varying the strength of rank signals affects redistribution preferences. The effects of group size, variation in ranks over rounds, and varying degrees of income inequality relative to the magnitude of redistribution could also be investigated. Another possible direction of future research is to utilize our experimental design—using nicknames to innocuously reveal some characteristic information about the subjects—for investigating the effect of individual characteristics other than gender, e.g., immigration status and ethnicity, on the preferences for redistribution.

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A Appendix - Experimental Instructions

Thank you for your participation. Before moving on to the main experiment, please complete a short survey.

[Ask participants' gender, age, a field of study, and experience of online experiments.]

In this experiment, your earnings will be based on your decisions, others' decisions, and some luck. It is important that you fully understand the following instructions. Please read them carefully.

Overview

This experiment mainly consists of two parts. In Part 1 of the experiment, which we call WORKING, you will work on a task for 5 minutes. In Part 2 of the experiment, which we call DECISIONS, you will report your decisions under a given situation. Your earnings in this experiment will depend on both—your task performance and random factors. More details follow.

Part 1: WORKING

You will perform a so-called Encryption Task. Specifically, you will read sequences of random letters. These have to be **decoded into numbers** with the given decoding table. For each text sequence, the decoding table changes. Below is an example of such a sequence.

Encryption task
Decode as many text sequences into numbers as possible
Correctly decoded text sequences: 0

Decoding table

letter	B	V	K	D	C	L	O	A	Y	F
corresponds to number	0	1	2	3	4	5	6	7	8	9

Text sequence: YDDYO

Your answer (enter only numbers):

Submit

Remaining time: 04:56

In the example, you can see the text sequence *YDDYO*. The decoding table tells you that B=0, V=1, and so on. This means that you have to decode *YDDYO* into 83386 and enter this numeric value into the answer field.

You will have **5 minutes** to correctly decode as many text sequences into numbers as possible. Note that each letter must be decoded correctly. After entering the decoded text sequence, hit the submit button. Subsequently, irrespective of whether the text sequence was decoded correctly or not, a new text sequence and a new decoding table will appear. You will see the remaining time in the upper right corner of the screen. **The more correct answers, the more money you will earn.** We will provide you more detail in the second part of the experiment.

Quiz Please answer the following two questions. If you answer one or more questions wrongly, you have to re-take the quiz.

- Q1 What is the aim of the task? (A) Write meaningful words from the given letters. (B) Decode the letters into numbers. (C) Reorder the letters according to the alphabet. (D) Assign a value to each letter and calculate the total value.
- Q2 Which of the following statements is **false**? (A) For each text sequence, the decoding table changes. (B) The answer only contains numbers. (C) Regardless of whether you decode the sequence correctly, a new text sequence will appear. (D) You have as much time as you want to decode the text sequences.

[participants perform the real-effort task for 5 minutes.]

Part 2: DECISIONS

This part of the experiment consists of 5 rounds. At the beginning of each decision round, we want you to pick your nickname among the list of alternatives. This way, you and others are completely anonymous, yet you can clearly distinguish other participants by their nicknames. (Everyone receives different options, so do not worry about possible duplication within a decision round. If you see the same names in other rounds, it is the result of pure coincidence, not implying the same person picking the same name twice.)

In each round, after you pick your nickname, a computer randomly assigns you to a group of six and displays your and other group members' **relative rank**. The rank is determined by your performance within Part 1 (WORKING) **and the unknown outcomes of a dice roll**. Specifically, the rank is determined as follows.

1. After the computer rolls a dice, the "adjusted performance" is determined. For simplicity, let N denote the number of correct answers in Part 1.
 - If your dice roll is 1, then your adjusted performance is $3 \times N$.

- If your dice roll is 2, then your adjusted performance is $4 \times N$.
- If your dice roll is 3, then your adjusted performance is $5 \times N$.
- If your dice roll is 4, then your adjusted performance is $6 \times N$.
- If your dice roll is 5, then your adjusted performance is $7 \times N$.
- If your dice roll is 6, then your adjusted performance is $8 \times N$.

So, for example, if you have 10 correct answers in Part 1, your adjusted performance could vary from 30 to 80.

2. The relative rank shows the rank of the adjusted performance, not the rank of the actual performance.
3. You will not learn the realizations of dice rolls.

For illustration, suppose that your rank is higher than the rank of another member. A higher rank could mean (1) that your performance in Part 1 (WORKING) was better than her, or (2) that your dice roll outcome was higher than her. When your adjusted performance happens to be the same as another member's, a computer randomly determines who receives the higher rank.

In each round, you and your group members will receive different amounts of money according to the ranking. The person ranked first will receive 15€, the person ranked second will receive 13€, the person ranked third will receive 11€, and so on.

After you observed the initial money amounts, everyone will be asked to choose who should get more money. The additional money comes from an **external source** and does not take away from the amount of money you have. The choices you make are private and will not be shown to anyone playing the game.

Once everyone in your group has decided, the computer will randomly select one group member and award the additional money according to the selected member's decision. At that point, the computer records everyone's updated earnings, but it will not show your final earnings from the round.

For every new round, all participants are reshuffled, and you are randomly assigned to a new group of six. The computer rolls new dices for the new round, which are independent of the previous outcomes. The new ranking is then determined based on your performance within Part 1 (Working), your new group members' performance, and the new dice rolls.

Payment

At the end of the experiment, the server computer will randomly select one of the five rounds with equal probability. Your final earnings in that selected round will be your actual payment. Since each round is equally likely selected, it is in your best interest to take every round equally seriously.

Quiz

Please answer the following two questions. If you answer one or more questions wrongly, you have to re-take the quiz.

- Q1 Suppose Adam and Bea are in the same group. Adam solved 10 tasks and Bea solved 11 tasks within Part 1 (WORKING). Which of the following statements is **false**? (A) The "adjusted performance" of Adam could range from 30 to 80. (B) Bea will hold a better rank than Adam in any case. (C) If their dice rolls are equal, Bea will hold a better rank than Adam. (D) If Bea holds a better rank than Adam, she will initially receive a higher amount of money.
- Q2 Which of the following statements is **true**? (A) Your nickname will remain the same during all five rounds. (B) At the end of each round, the choices of all group members are implemented. (C) Your group members will remain the same during all five rounds. (D) The additional money you are asked to give to another group member comes from an external source.

[Participants make redistribution decisions for 5 rounds. See Figure A1 for illustration.]

Allocate additional money.

There are 2€ from an external source. You can assign the additional money to either a person at 1 rank above or below you.

(If you are ranked first or last, you will have only one option.)

When your decision is selected at the end of the experiment, the additional money goes to the person you chose.
Of course, if someone else decides to give you the money, then you could get the additional money.

Rank	Name	Earning
1	Adam	15€
2	Julia	13€
(you) 3	Diana	11€
4	Patrick	9€
5	Richard	7€
6	Katherina	5€

Choose to whom you want to allocate the additional money.

Figure A1: Screenshot of Part 2

[participant filled out post-experimental survey questions.]

Survey

1. There were 18 (including you) in the experiment you just finished. Please give your best guess about your actual rank of the performance. (Write an integer between 1 and 18.)
2. Please specify how you agree on the following statement: "I feel that, overall, the random adjustment of the performance places me worse ranks than my actual ranks."
3. Please specify how you agree on the following statement: "I feel that the effect of random adjustment on the relative rank is stronger than that of actual performance."
4. Please specify how you agree on the following statement: "A society should accept inequality caused by differences in luck."

B Appendix - Regression results

	OLS			Probit		
	(1)	(2)	(3)	(4)	(5)	(6)
	Pooled	Females	Males	Pooled	Females	Males
rank2	-0.0889 (0.0669)	-0.121 (0.101)	-0.0563 (0.0919)	-0.333 (0.249)	-0.490 (0.393)	-0.198 (0.324)
rank3	-0.0667 (0.0482)	-0.0889 (0.0740)	-0.0444 (0.0649)	-0.258 (0.192)	-0.379 (0.323)	-0.159 (0.237)
rank4	-0.0333 (0.0542)	-0.0253 (0.0604)	-0.0396 (0.0884)	-0.137 (0.223)	-0.124 (0.296)	-0.143 (0.321)
_cons	0.856*** (0.0467)	0.889*** (0.0531)	0.822*** (0.0778)	1.061*** (0.204)	1.221*** (0.278)	0.924** (0.296)
<i>N</i>	360	177	183	360	177	183

The dependent variable is a dummy indicating whether a participant allocated downward. Observations holding ranks 1 and 6 are omitted. Standard errors clustered at the individual level are in parentheses.

* $p < 0.05$, ** $p < 0.01$, *** $p < 0.001$

Table B1: Rank effects

	OLS			Probit		
	(1) Pooled	(2) Females	(3) Males	(4) Pooled	(5) Females	(6) Males
SG	0.157** (0.0568)	0.251** (0.0909)	0.0684 (0.0663)	0.646* (0.259)	1.329* (0.516)	0.255 (0.266)
SA	0.0177 (0.065)	0.0707 (0.109)	-0.0327 (0.0755)	0.0567 (0.208)	0.227 (0.344)	-0.106 (0.24)
BS	0.0765 (0.0533)	0.139 (0.0795)	0.0139 (0.0742)	0.266 (0.182)	0.495 (0.26)	0.0475 (0.253)
_cons	0.750*** (0.0513)	0.721*** (0.0847)	0.778*** (0.0601)	0.674*** (0.161)	0.586* (0.25)	0.765*** (0.2)
<i>N</i>	360	177	183	360	177	183

The dependent variable is a dummy indicating whether a participant allocated downward. Observations holding ranks 1 and 6 are omitted. Standard errors clustered at the individual level are in parentheses.

* $p < 0.05$, ** $p < 0.01$, *** $p < 0.001$

Table B2: Effects of gender composition

	OLS			Probit		
	(1) Pooled	(2) Females	(3) Males	(4) Pooled	(5) Females	(6) Males
rank2	-0.0849 (0.0632)	-0.109 (0.0921)	-0.0561 (0.0910)	-0.322 (0.239)	-0.468 (0.377)	-0.194 (0.319)
rank3	-0.0737 (0.0486)	-0.0884 (0.0744)	-0.0527 (0.0676)	-0.299 (0.197)	-0.429 (0.329)	-0.188 (0.249)
rank4	-0.0474 (0.0556)	-0.0401 (0.0612)	-0.0505 (0.0911)	-0.202 (0.231)	-0.191 (0.305)	-0.182 (0.327)
SG	0.152** (0.0551)	0.240** (0.0876)	0.0685 (0.0660)	0.636* (0.257)	1.313** (0.501)	0.263 (0.271)
SA	0.0092 (0.0652)	0.0557 (0.110)	-0.0364 (0.0764)	0.0289 (0.209)	0.161 (0.353)	-0.110 (0.238)
SB	0.0746 (0.0518)	0.129 (0.0759)	0.0168 (0.0736)	0.262 (0.175)	0.460 (0.251)	0.0643 (0.243)
_cons	0.805*** (0.0552)	0.789*** (0.0786)	0.818*** (0.0812)	0.898*** (0.208)	0.907** (0.297)	0.904** (0.290)
<i>N</i>	360	177	183	360	177	183

The dependent variable is a dummy indicating whether a participant allocated downward. SG, SB, and SA respectively refer to a *same-gender*, *same-above*, and *same-below* situation. Observations holding ranks 1 and 6 are omitted. Standard errors clustered at the individual level are in parentheses.

* $p < 0.05$, ** $p < 0.01$, *** $p < 0.001$

Table B3: Effects of rank and gender composition

	OLS		Probit	
	(1) SG	(2) MG	(3) SG	(4) MG
rank2	-0.0126 (0.0920)	-0.0830 (0.148)	-0.0995 (0.707)	-0.238 (0.422)
rank3	-0.0991 (0.0754)	0.0548 (0.131)	-0.562 (0.436)	0.179 (0.415)
rank4	-0.0212 (0.0815)	0.105 (0.141)	-0.160 (0.617)	0.370 (0.493)
_cons	0.941*** (0.0594)	0.737*** (0.104)	1.565** (0.496)	0.634* (0.312)
<i>N</i>	75	88	75	88

The dependent variable is a dummy indicating whether a participant allocated downward. SG and MG respectively refer to a *same-gender* and *mixed-gender* situation. Observations holding ranks 1 and 6 are omitted. Standard errors clustered at the individual level are in parentheses.

* $p < 0.05$, ** $p < 0.01$, *** $p < 0.001$

Table B4: Rank effects in SG and MG

	OLS			Probit		
	(1)	(2)	(3)	(4)	(5)	(6)
rank2	-0.0888 (0.0672)	-0.0871 (0.0657)	-0.0948 (0.0676)	-0.333 (0.249)	-0.319 (0.244)	-0.363 (0.257)
rank3	-0.0667 (0.0483)	-0.0612 (0.0471)	-0.0737 (0.0490)	-0.258 (0.192)	-0.232 (0.186)	-0.281 (0.198)
rank4	-0.0334 (0.0537)	-0.0272 (0.0517)	-0.0241 (0.0531)	-0.137 (0.221)	-0.105 (0.213)	-0.103 (0.221)
PL	-0.0012 (0.0250)			-0.0060 (0.0942)		
IL		0.0325 (0.0297)			0.116 (0.109)	
SL			-0.0487 (0.0258)			-0.183 (0.0940)
_cons	0.860*** (0.0890)	0.734*** (0.107)	0.982*** (0.0706)	1.081** (0.353)	0.628 (0.376)	1.558*** (0.321)
<i>N</i>	360	360	360	360	360	360

The dependent variable is a dummy indicating whether a participant allocated downward. PL (Perception of Luck), IL (Importance of Luck), and SL (Society embracing Luck) respectively refer to the Likert-scale answers on survey questions 2–4. Observations holding ranks 1 and 6 are omitted. Standard errors clustered at the individual level are in parentheses.

* $p < 0.05$, ** $p < 0.01$, *** $p < 0.001$

Table B5: Effects of post-experimental perceptions

	OLS				Probit			
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
rank2	-0.0849 (0.0635)	-0.0833 (0.0619)	-0.0901 (0.0641)	-0.0880 (0.0631)	-0.3220 (0.2396)	-0.3091 (0.2322)	-0.3549 (0.2495)	-0.3357 (0.2405)
rank3	-0.0737 (0.0486)	-0.0684 (0.0471)	-0.0792 (0.0493)	-0.0744 (0.0476)	-0.2988 (0.1968)	-0.2736 (0.1880)	-0.3227 (0.2043)	-0.2963 (0.1928)
rank4	-0.0474 (0.0552)	-0.0413 (0.0528)	-0.0376 (0.0545)	-0.0334 (0.0522)	-0.2025 (0.2278)	-0.1676 (0.2170)	-0.1671 (0.2286)	-0.1372 (0.2141)
SG	0.1515** (0.0553)	0.1482** (0.0545)	0.1516** (0.0542)	0.1486** (0.0542)	0.6359* (0.2556)	0.6111* (0.2503)	0.6655* (0.2611)	0.6459* (0.2542)
SA	0.0092 (0.0648)	0.0100 (0.0647)	0.0167 (0.0633)	0.0181 (0.0624)	0.0298 (0.2070)	0.0299 (0.2083)	0.0671 (0.2052)	0.0730 (0.2023)
SB	0.0746 (0.0515)	0.0758 (0.0515)	0.0706 (0.0505)	0.0726 (0.0501)	0.2633 (0.1732)	0.2698 (0.1754)	0.2497 (0.1727)	0.2626 (0.1717)
PL	-0.0000 (0.0244)			-0.0075 (0.0230)	-0.0053 (0.0938)			-0.0371 (0.0940)
IL		0.0307 (0.0294)		0.0260 (0.0304)		0.1053 (0.1101)		0.0942 (0.1152)
SL			-0.0473 (0.0254)	-0.0447 (0.0264)			-0.1853 (0.0947)	-0.1793 (0.0974)
_cons	0.806*** (0.0983)	0.691*** (0.1086)	0.926*** (0.0758)	0.847*** (0.1536)	0.915* (0.3684)	0.503 (0.3815)	1.389*** (0.3233)	1.137* (0.5610)
<i>N</i>	360	360	360	360	360	360	360	360

The dependent variable is a dummy indicating whether a participant allocated downward. SG, SB, and SA respectively refer to a *same-gender*, *same-above*, and *same-below* situation. PL (Perception of Luck), IL (Importance of Luck), and SL (Society embracing Luck) respectively refer to the Likert-scale answers on survey questions 2–4. Observations holding ranks 1 and 6 are omitted. Standard errors clustered at the individual level are in parentheses.

* $p < 0.05$, ** $p < 0.01$, *** $p < 0.001$

Table B6: Effects of rank, gender composition, and post-experimental perceptions