

IP PARIS

# Asymptotic Guarantees for Learning Generative Models with the Sliced-Wasserstein Distance

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# Minimum Distance Estimation

- Observations  $Y_{1:n} = (Y_1, \dots, Y_n), Y_i \in Y \subset \mathbb{R}^d$ , i.i.d. from  $\mu_{\star} \in \mathcal{P}(Y)$ , with  $\mathcal{P}(Y)$ : set of probability measures on Y.
- A family of distributions on Y parameterized by  $\theta \in \Theta \subset \mathbb{R}^{d_{\theta}}$ :  $\mathcal{M} = \{ \mu_{\theta} \in \mathcal{P}(\mathsf{Y}), \ \theta \in \Theta \}.$
- Purely generative models: We can generate  $m \in \mathbb{N}^*$  i.i.d. samples from  $\mu_{\theta}$ , but the likelihood is intractable.  $\hat{\mu}_{\theta,m}$  is the empirical distribution.

Given  $Y_{1:n}$ , its empirical distribution  $\hat{\mu}_n$  and a distance **D** on  $\mathcal{P}(\mathsf{Y})$ , we perform Minimum Distance Estimation (MDE):

$$\hat{\theta}_n = \operatorname{argmin}_{\theta \in \Theta} \mathbf{D}(\hat{\mu}_n, \mu_\theta)$$
 (1)

or Minimum Expected Distance Estimation (MEDE):

$$\hat{\theta}_{n,m} = \operatorname{argmin}_{\theta \in \Theta} \mathbb{E} \left[ \mathbf{D}(\hat{\mu}_n, \hat{\mu}_{\theta,m}) | Y_{1:n} \right]$$
 (2)

# Optimal Transport (OT) Metrics

For  $p \geq 1$ ,  $\mathcal{P}_p(Y)$ : set of probability measures on Y with finite p'th moment. Let  $\mu, \nu \in \mathcal{P}_p(\mathsf{Y})$ .

Wasserstein distance  $(W_p)$ . Computationally expensive, except in 1d  $(Y \subset \mathbb{R}) \to \text{analytical form.}$ 

#### Sliced-Wasserstein (SW) distance.

 $\mathbb{S}^{d-1}$ : d-dimensional unit sphere,  $\sigma$ : uniform distribution on  $\mathbb{S}^{d-1}$ .

Practical metric based on projections:  $\forall u \in \mathbb{S}^{d-1}, y \in \mathbf{Y}, \ u^{\star}(y) = \langle u, y \rangle$ 

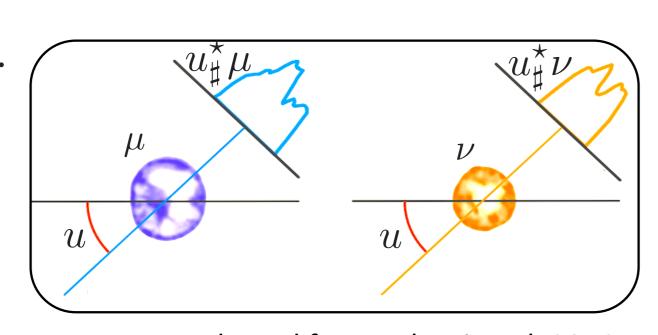


Image adapted from Kolouri et al. 2016

 $\mathbf{SW}_p^p(\mu,\nu) = \int_{\mathbb{S}^{d-1}} \mathbf{W}_p^p(u_{\sharp}^{\star}\mu, u_{\sharp}^{\star}\nu) d\boldsymbol{\sigma}(u)$ 

# Combining MDE and OT

Minimum Wasserstein estimators, defined in (1) and (2) with  $\mathbf{D} = \mathbf{W}_p$ , have asymptotic guarantees [1] but are not practical.

 $\Rightarrow$  With  $\mathbf{D} = \mathbf{SW}_p$  in (1) and (2), we get the minimum (expected) SW estimators (M(E)SWE) of order p.

Recent studies show the empirical success of SW-based estimators on generative modeling, but lack of theoretical guarantees.

 $\Rightarrow$  We investigate the asymptotic properties of these estimators.

# Theoretical Results

The convergence in  $\mathbf{SW}_p$  implies the weak convergence in  $\mathcal{P}(\mathbb{R}^d)$ .

### Key assumptions.

- Continuity: For any  $(\theta_n)_{n\in\mathbb{N}}$  in  $\Theta$  such that  $\lim_{n\to+\infty} \rho_{\Theta}(\theta_n,\theta) = 0$ ,
  - **A1.**  $(\mu_{\theta_n})_{n\in\mathbb{N}}$  converges weakly  $(\xrightarrow{w})$  to  $\mu_{\theta}$ .
  - **A2.**  $\lim_{n\to+\infty} \mathbb{E}[\mathbf{SW}_p(\mu_{\theta_n}, \hat{\mu}_{\theta_n,n})|Y_{1:n}] = 0.$
- Data-generating process:
  - **A3.**  $\lim_{n\to+\infty} \mathbf{SW}_p(\hat{\mu}_n, \mu_{\star}) = 0$ ,  $\mathbb{P}$ -almost surely.
- Bounded sets: For some  $\epsilon > 0$ ,
  - **A4.**  $\Theta_{\epsilon}^{\star} = \{\theta \in \Theta : \mathbf{SW}_{p}(\mu_{\star}, \mu_{\theta}) \leq \epsilon_{\star} + \epsilon\}, \text{ with } \epsilon_{\star} =$  $\inf_{\theta \in \Theta} \mathbf{SW}_p(\mu_{\star}, \mu_{\theta})$ , is bounded.
  - **A5.**  $\Theta_{\epsilon,n} = \{\theta \in \Theta : \mathbf{SW}_p(\hat{\mu}_n, \mu_\theta) \leq \epsilon_n + \epsilon\}, \text{ with } \epsilon_n = \mathbf{SW}_p(\hat{\mu}_n, \mu_\theta) \leq \epsilon_n + \epsilon\}$  $\inf_{\theta\in\Theta} \mathbf{SW}_p(\hat{\mu}_n,\mu_\theta)$ , is bounded almost surely.

## Existence and consistency of MSWE

Assume A1, A3, A4. Then, there exists E with  $\mathbb{P}(\mathsf{E}) = 1$  such that, for all  $\omega \in \mathsf{E}$ ,

$$\lim_{n \to +\infty} \inf_{\theta \in \Theta} \mathbf{SW}_p(\hat{\mu}_n(\omega), \mu_{\theta}) = \inf_{\theta \in \Theta} \mathbf{SW}_p(\mu_{\star}, \mu_{\theta}),$$

 $\limsup \operatorname{argmin}_{\theta \in \Theta} \mathbf{SW}_p(\hat{\mu}_n(\omega), \mu_{\theta}) \subset \operatorname{argmin}_{\theta \in \Theta} \mathbf{SW}_p(\mu_{\star}, \mu_{\theta})$  $n \rightarrow +\infty$ 

Besides, for all  $\omega \in \mathsf{E}$ , there exists  $n(\omega)$  such that, for all  $n \geq n(\omega)$ ,  $\operatorname{argmin}_{\theta \in \Theta} \mathbf{SW}_p(\hat{\mu}_n(\omega), \mu_{\theta})$  is non-empty.

Guarantees for MESWE. Existence and consistency (with A1 to A4), convergence to MSWE as  $m \to \infty$  (A1, A2, A5).

# Central limit theorem for MSWE with p=1

Consider A1, A3, A4,  $\mu_{\star} = \mu_{\theta_{\star}}$  (with  $\theta_{\star} \in \Theta$  well-separated) and  $H: \theta \mapsto \int_{\mathbb{S}^{d-1}} \int_{\mathbb{R}} |G_{\star}(u,t) - \langle \theta, D_{\star}(u,t) \rangle | dt d\boldsymbol{\sigma}(u), \text{ with }$ 

- $\sqrt{n}(\hat{F}_n F_{\theta_{\star}}) \xrightarrow{w} G_{\star}$ , where  $\hat{F}_n$  and  $F_{\theta_{\star}}$  contain the CDFs of the projected  $\hat{\mu}_n$  and  $\mu_{\theta_*}$
- $D_{\star}(u,\cdot)$ : the "derivative" of  $F_{\theta}(u,\cdot)$  in  $\theta_{\star}$

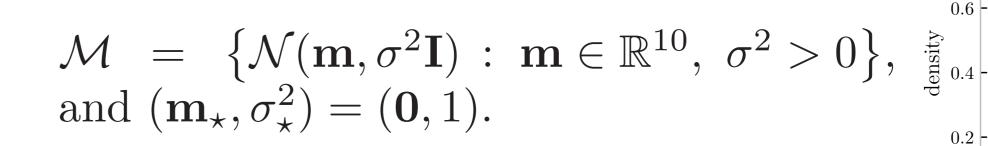
 $\sqrt{n} \inf_{\theta \in \Theta} \mathbf{SW}_1(\hat{\mu}_n, \mu_\theta) \xrightarrow{w} \inf_{\theta \in \Theta} H(\theta),$ Then,

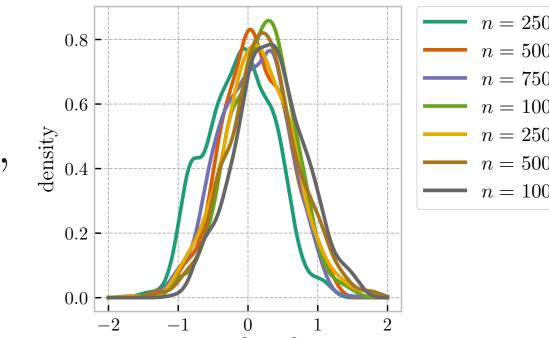
 $\sqrt{n}(\hat{\theta}_n - \theta_\star) \xrightarrow{w} \operatorname{argmin}_{\theta \in \Theta} H(\theta), \text{ as } n \to +\infty$ 

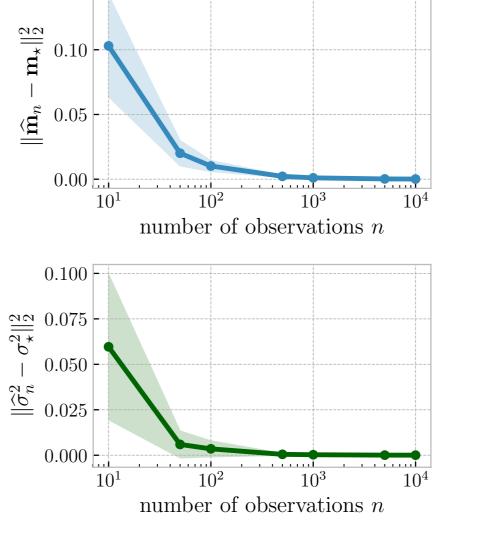
 $\Rightarrow$  Convergence rate of  $\sqrt{n}$  independent of the dimension

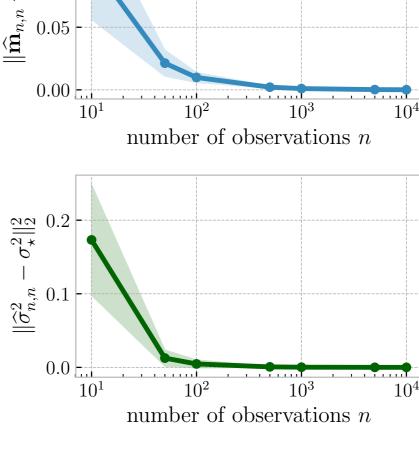
# Numerical Experiments

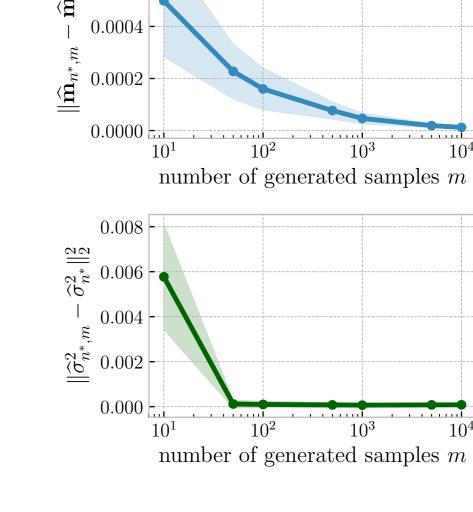
• Multivariate Gaussians.











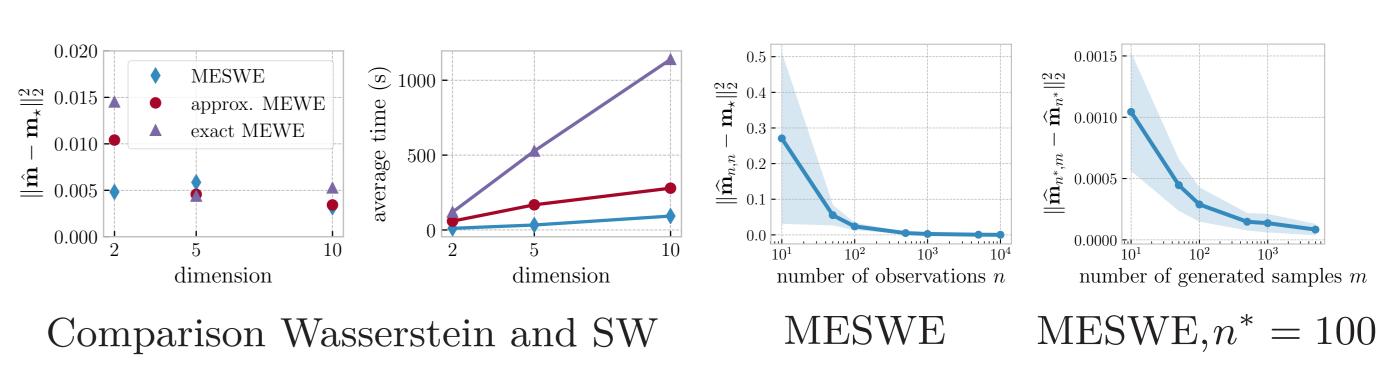
MSWE vs. n

MESWE vs. n = m

MESWE, n = 2000 vs. m

# • Multivariate elliptically contoured stable distributions.

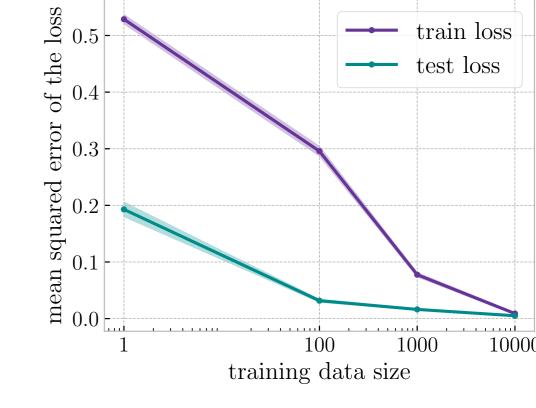
 $\mathcal{M} = \{ \mathcal{E} \alpha \mathcal{S}_c(\mathbf{I}, \mathbf{m}) : \mathbf{m} \in \mathbb{R}^d \} \text{ with } \alpha = 1.8, \text{ and } \mathbf{m}_{\star} = 2.$ 



#### • High-dimensional real data.

We train the Sliced-Wasserstein Generator [2] (based on MESWE), on MNIST.

We plot the mean-squared error between the training/test loss obtained for (n, m) $(\text{from } (1,1) \text{ to } (10\ 000,60)) \text{ and for }$  $(n^*, m^*) = (60\,000, 200).$ 



#### Main References

- [1] E. Bernton, P. E. Jacob, M. Gerber, C. P. Robert. On parameter estimation with the Wasserstein distance. Information and Inference: A Journal of the IMA, Jan 2019.
- [2] I. Deshpande, Z. Zhang, A. G. Schwing. Generative modeling using the sliced Wasserstein distance. CVPR 2018.