

# Metropolis-Hastings part

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## MAIN FUNCTION

Here we write a Metropolis-Hastings algorithm which is sampled from a Beta distribution. we should use the Detail Balance equation here:

$$\pi(\phi)T(\phi \rightarrow \phi') = \pi(\phi')T(\phi' \rightarrow \phi)$$

, which also can be written as

$$\frac{T(\phi \rightarrow \phi')}{T(\phi' \rightarrow \phi)} = \frac{\pi(\phi')}{\pi(\phi)}$$

- Firstly, we choose a start value:  $\phi$  at random.
- Then, we separate the transition in two sub-steps: the proposal distribution ( $Q(\phi \rightarrow \phi')$ ) and the acceptance probability ( $A(\phi \rightarrow \phi')$ ).

$$T(\phi \rightarrow \phi') = Q(\phi \rightarrow \phi')A(\phi \rightarrow \phi')$$

in the part below, we write the proposal function of the form

$$\phi_{prop}|\phi_{old} \sim \text{Beta}(c\phi_{old}, c(1 - \phi_{old}))$$

, which is the probability of proposing a state  $\phi'|\phi$ .

And we create the posterior funtion which is

$$\frac{P(\phi')}{P(\phi_{old})} = \frac{\text{dbeta}(\phi', a, b)}{\text{dbeta}(\phi_{old}, a, b)}$$

where P follows the Beta distribution, and (a,b) = (6,4).

```
proposalfunction <- function(c, phi_old){  
  return (rbeta(1, c*phi_old, c*(1-phi_old)))  
}  
  
posterior <- function(c, phi_old,trueA,trueB){  
  return(dbeta(proposalfunction(c, phi_old),trueA,trueB)/dbeta(phi_old,trueA,trueB))  
}
```

- Next, we write the Metropolis-Hastings algorithm which is sampled from a Beta distribution. and set the acceptance probability as:

$$A(\phi \rightarrow \phi') = \min(1, \frac{\pi(\phi')Q(\phi' \rightarrow \phi)}{\pi(\phi)Q(\phi \rightarrow \phi')})$$

If state is accepted, set the current value to  $\phi'$ ; otherwise, set it to the startvalue  $\phi$ .

\*Final, we repeat this process until generate the required number of samples we need.

```
beta_metropolis <- function(c,startvalue,iterations){  
  trueA = 6  
  trueB = 4  
  chain <- rep(0,iterations)  
  chain[1] = startvalue
```

```

for (i in 1:iterations){
  phi <- proposalfunction(c, startvalue)
  proposal <- dbeta(startvalue, c*phi, c*(1-phi))/dbeta(phi, c*startvalue, c*(1-startvalue))
  #post <- posterior(c,startvalue,trueA,trueB)
  post <- dbeta(phi,trueA,trueB)/dbeta(startvalue,trueA,trueB)
  probab <- min(1,post*proposal)

  if(runif(1) < probab){
    startvalue <- phi
    chain[i] <- phi
  } else {
    chain[i] <- startvalue
  }
}
return(chain)
}

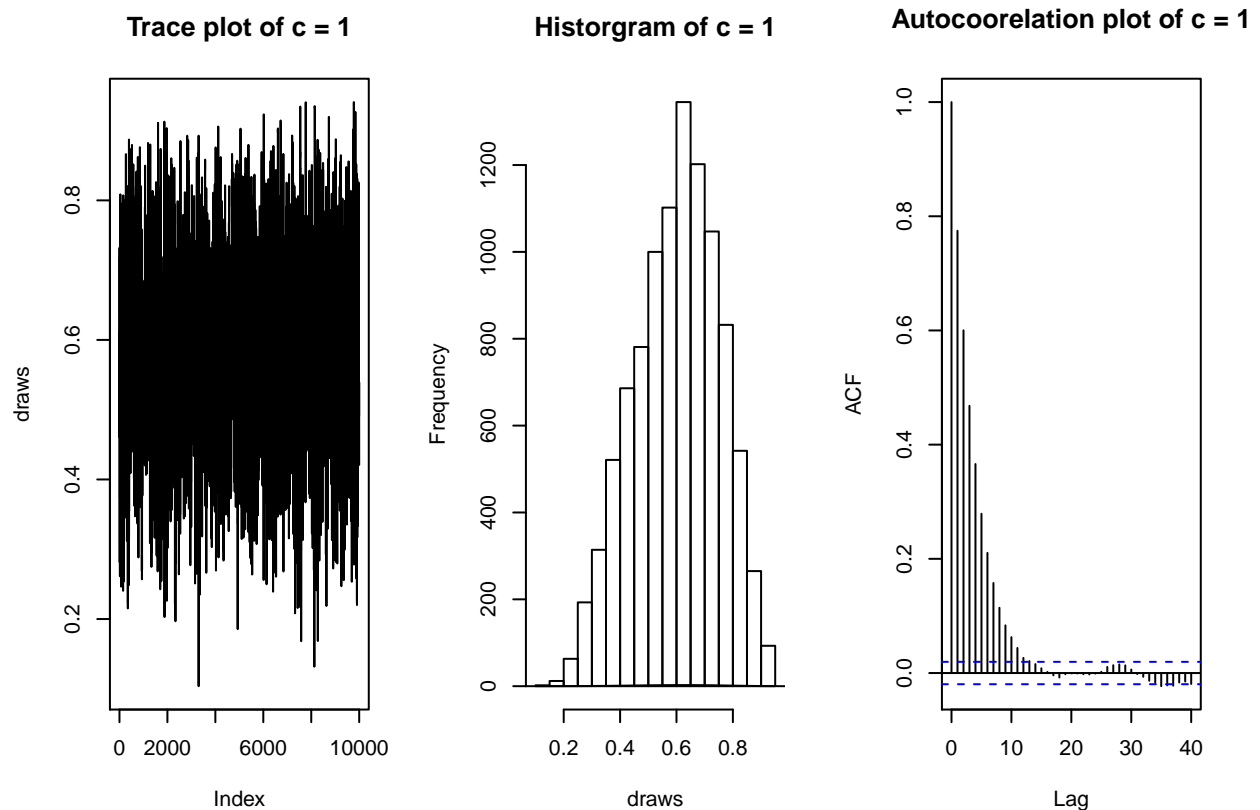
```

by test, we give a sample with  $c = 1$ , iterations = 10000 and start value = random uniform(1):

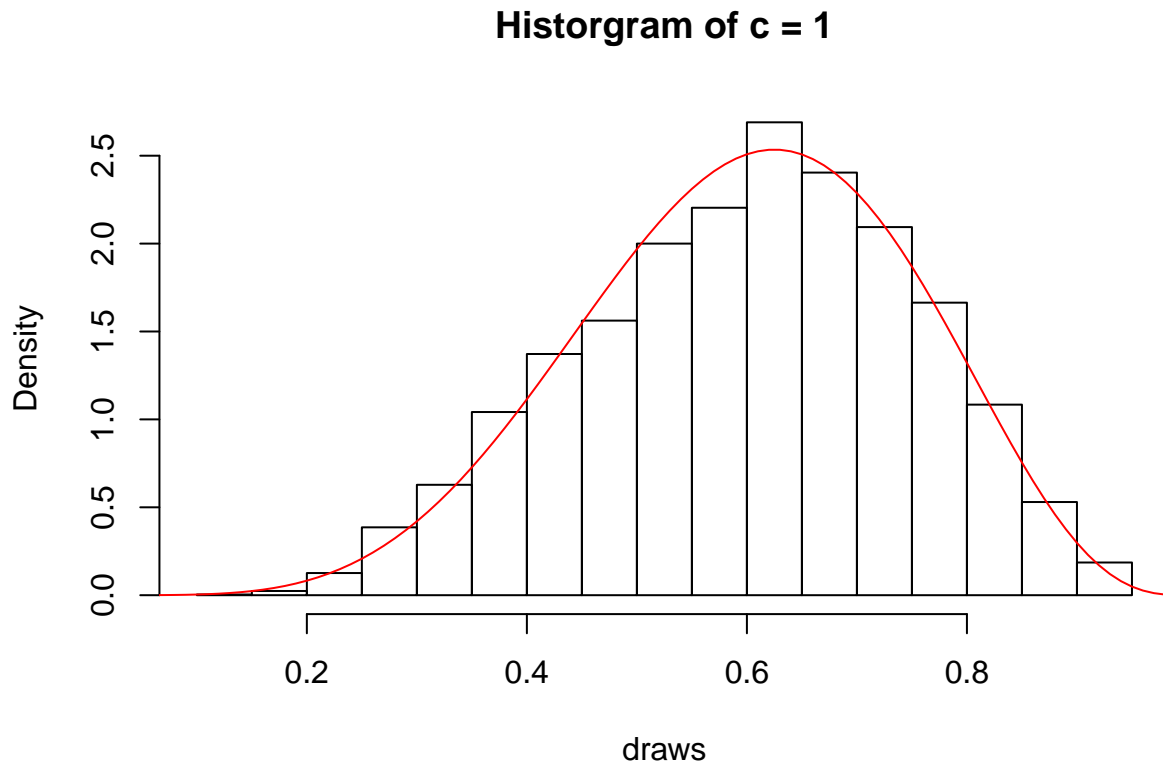
```

draws <- beta_metroplis(1,runif(1),10000)
par(mfrow = c(1,3))
plot(draws,type = "l",main = "Trace plot of c = 1")
hist(draws, main = "Histogram of c = 1")
x <- seq(0,1,0.01)
lines(x,dbeta(x,6,4))
acf(draws, main = "Autocoorelation plot of c = 1")

```



```
par(mfrow = c(1,1))
x <- seq(0,1,0.01)
hist(draws, main = "Histogram of c = 1",freq = FALSE)
lines(x,dbeta(x,6,4),col = "red")
```



```
ks.test(jitter(draws,0.0001),rbeta(10000,6,4))
```

```
##
## Two-sample Kolmogorov-Smirnov test
##
## data: jitter(draws, 1e-04) and rbeta(10000, 6, 4)
## D = 0.0201, p-value = 0.03519
## alternative hypothesis: two-sided
```

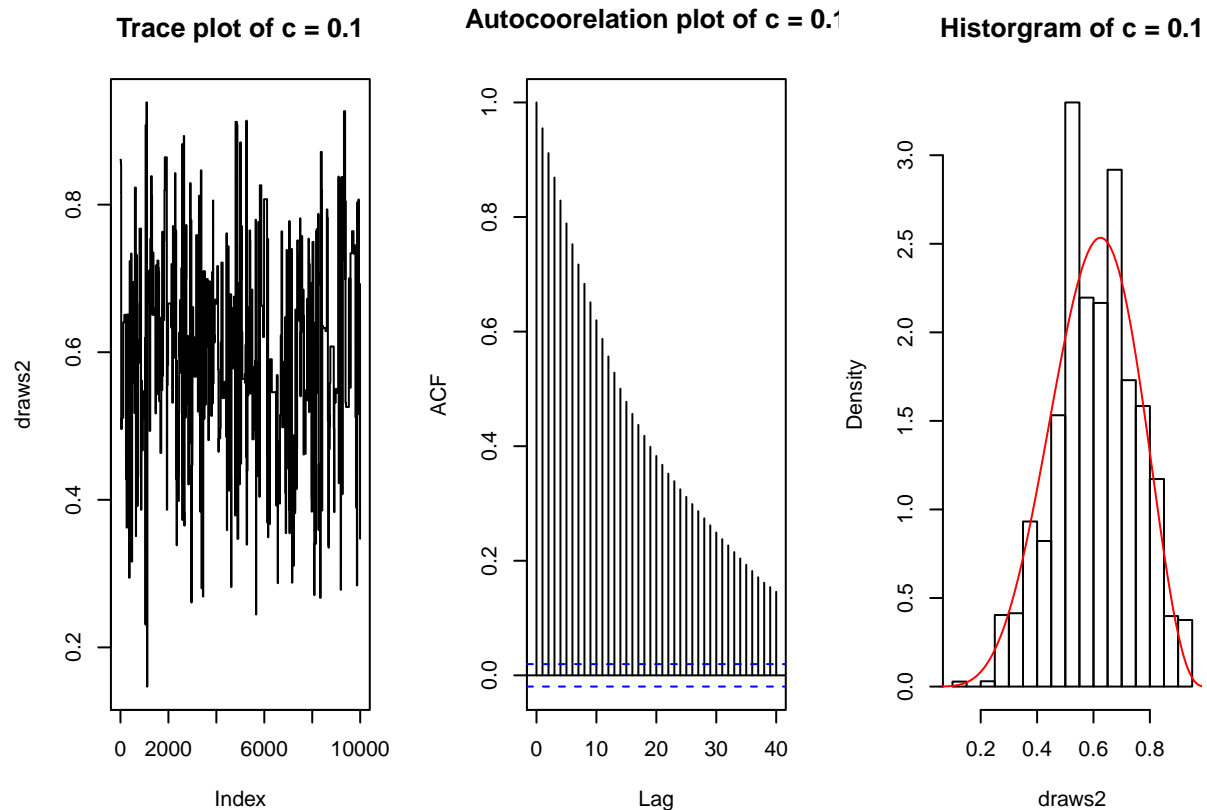
From the Kolmogorov-Smirnov statistic test, we can see the p-value equals 0.0009305, which is much smaller than  $p = 0.01$ , then we reject the null hypothesis, the data does not follow the beta distribution.

## Re-run

$c = 0.1$ :

```
draws2 <- beta_metroplis(0.1,runif(1),10000)
par(mfrow = c(1,3))
plot(draws2,type = "l",main = "Trace plot of c = 0.1")
acf(draws2, main = "Autocoorelation plot of c = 0.1")
hist(draws2, main = "Histogram of c = 0.1",freq = FALSE)
```

```
x <- seq(0,1,0.01)
lines(x,dbeta(x,6,4),col = "red")
```



From the acf plot, we could see there are lots of lags, so it should be tested for 50,000 draws before thinning and burn-in. After we thin about 50 times, test for 10,000 draws, the p-value is much better but still small, so it does not work.

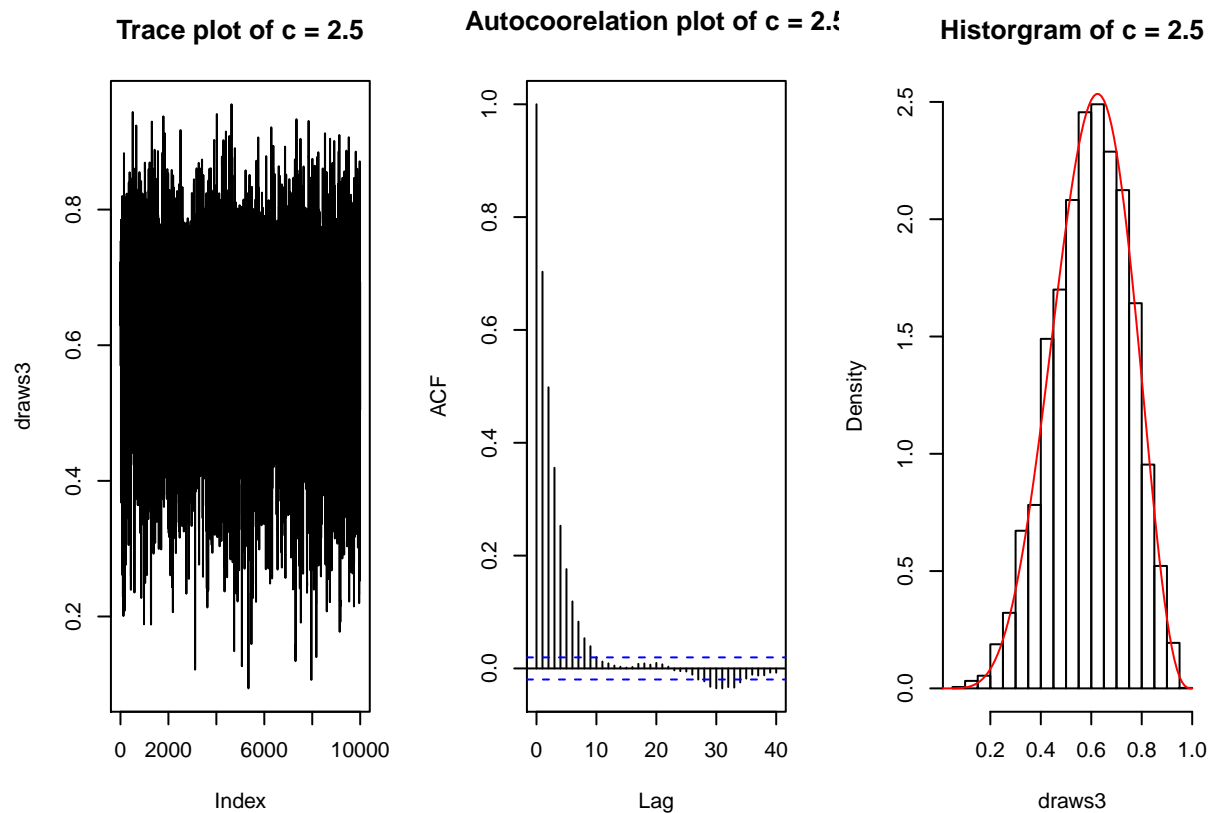
```
ks.test(jitter(draws2,0.0001),rbeta(10000,6,4))
```

```
##
## Two-sample Kolmogorov-Smirnov test
##
## data: jitter(draws2, 1e-04) and rbeta(10000, 6, 4)
## D = 0.0619, p-value < 2.2e-16
## alternative hypothesis: two-sided
```

From the Kolmogorov-Smirnov statistic test, we can see the p-value equals 6.142e-11, which is much smaller than  $p = 0.01$ , then we reject the null hypothesis, the data does not follow the beta distribution.

$c = 2.5$ :

```
draws3 <- beta_metroplis(2.5,runif(1),10000)
par(mfrow = c(1,3))
plot(draws3,type = "l",main = "Trace plot of c = 2.5")
acf(draws3, main = "Autocoorelation plot of c = 2.5")
hist(draws3, main = "Histogram of c = 2.5",freq = FALSE)
x <- seq(0,1,0.01)
lines(x,dbeta(x,6,4),col = "red")
```



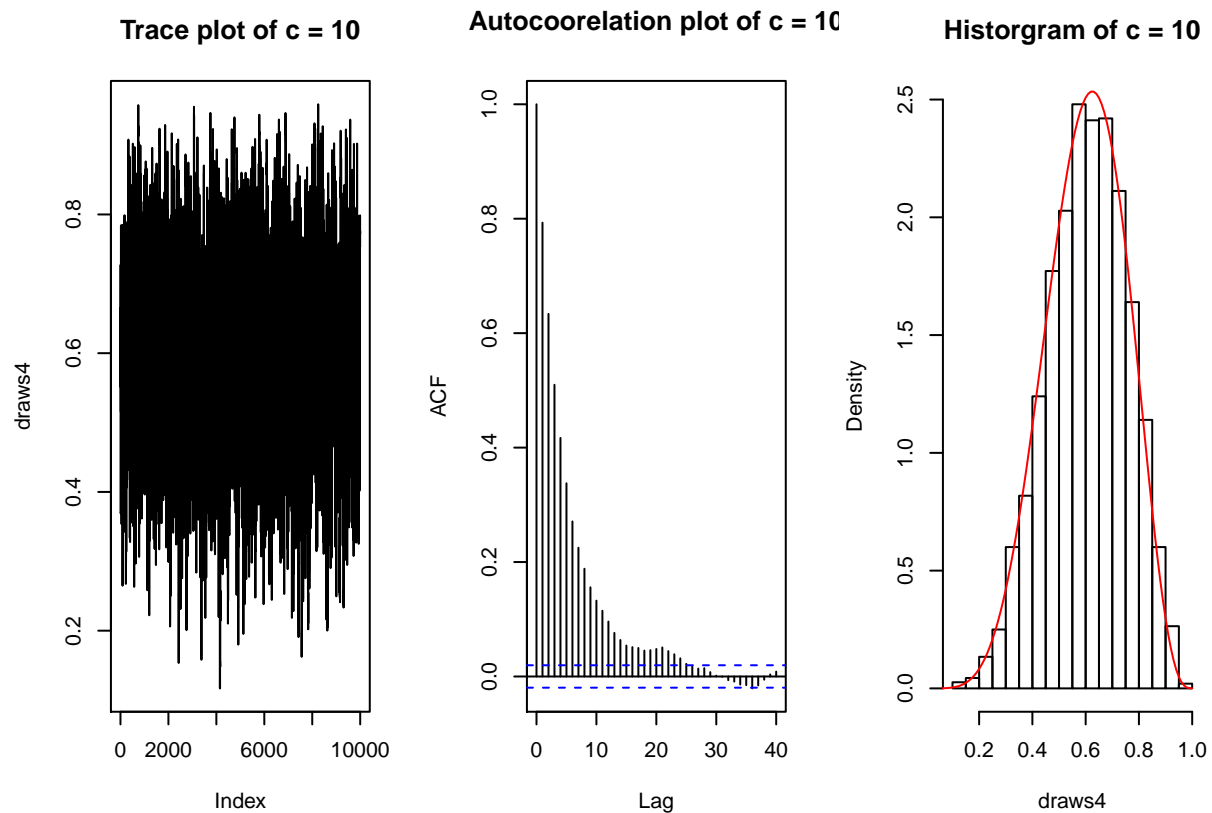
Test for 50,000 draws before thinning and burn-in. After we thin 5 times, test for 10,000 draws, the p-value of k-s test is much better.

```
ks.test(jitter(draws3,0.0001),rbeta(10000,6,4))
```

```
##
## Two-sample Kolmogorov-Smirnov test
##
## data: jitter(draws3, 1e-04) and rbeta(10000, 6, 4)
## D = 0.0149, p-value = 0.2169
## alternative hypothesis: two-sided
```

c = 10:

```
draws4 <- beta_metroplis(10,runif(1),10000)
par(mfrow = c(1,3))
plot(draws4,type = "l",main = "Trace plot of c = 10")
acf(draws4, main = "Autocorelation plot of c = 10")
hist(draws4, main = "Histogram of c = 10",freq = FALSE)
x <- seq(0,1,0.01)
lines(x,dbeta(x,6,4),col = "red")
```



Test for 50,000 draws before thinning and burn-in. After we thin 5 times, test for 10,000 draws, the p-value of k-s test is much better.

```
ks.test(jitter(draws4,0.0001),rbeta(10000,6,4))
```

```
##
## Two-sample Kolmogorov-Smirnov test
##
## data: jitter(draws4, 1e-04) and rbeta(10000, 6, 4)
## D = 0.0281, p-value = 0.0007444
## alternative hypothesis: two-sided
```

From the Kolmogorov-Smirnov statistic test, we can see the p-value equals 0.0007444, which is larger than  $p = 0.05$ , then we cannot reject the null hypothesis, the data follows the beta distribution.

## Conclusion:

By re-run the sampler with  $c = 0.1$ ,  $c = 2.5$  and  $c = 1$ , we can see when  $c$  is getting larger, it meets the target distribution better and have larger p-value, but when  $c$  is too large, the autocorrelation would be worse.

1. Compare their plot, as  $c$  becomes larger, the more efficient it meets target distribution.
2. Through the acf plot, we can see that when  $c = 10$ , it is better than  $c = 0.1$  but worse than  $c = 2.5$ . So we should choose a fitable  $c$  which cannot be too large or too small.
3. By comparing their histogram, we can see as  $c$  becomes larger, it meets the target distribution more.
4. Through the Kolmogorov-Smirnov statistic test, the p-value is getting larger as  $c$  gets larger.

So by comparing this values of  $c$ ,  $c = 2.5$  is most effective at drawing from the target distribution since it could meets the target distribution enough well and has a better autocorrelation plot than  $c = 10$ .