# Markov Chain Basic Concepts

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#### P2P System Evolution: Brief Overview

- Unstructured Approaches
  - random connections
  - search mechanisms: TTL-enhanced flooding
  - examples: Gnutella, Kazaa, BitTorrent,...
  - pro: simplicity, cons: limited scalability
- Structured Approaches
  - Distributed Hash Tables, Delaunay Overlay
  - examples: Kademlia, implementation in eMule/Bittorrent
  - pro: routing efficiency, cons: consistency mainteinance in presence of churn
- Stochastic Approaches
  - Gossip, Random Walk
  - simple, based on a well known mathematical theory
  - exploited also for distributed crawling of social networks



# P2P System: Stochastic Approaches

- unstructured/structured approaches presented in the Course "P2P Systems" of the Master Degree in Computer Science
- here we focus on stochastics appoches
- mathematical background
  - Markov Chains, Random Walk
- applications:
  - Distributed sampling for extimating network parameters
  - Epidemic information diffusion
  - Distributed computation af aggregate values

#### Outline

- Basic Definitions
- 2 Examples
- 3 It's All Just Matrix Theory?
- 4 The Basic Theorem

#### Markov Chain: Basic Characteristics

- Markov Chain: describes a system whose states change over time
  - discrete time stochastic process
- Changes are governed by a probability distribution.
- The next state only depends upon the current system state
  - the path to the present state is not relevant
- Class of random process useful in different areas
  - developments in theory and applications in recent decades.

## Markov Chain Specification

A sequence of Random Variables  $\{X_0, X_1, \dots X_n\}$ :  $X_i$  describes the state of the system at time i

To specify a Markov Chain we need:

- a finite or countable set of states S
  - $S = \{1, 2, \dots, N\}$  for some finite N
  - the value of the random variables  $X_i$  are taken from S
- an initial distribution  $\pi_0$ ,  $\pi_0(i) = \mathbb{P}\{X_0 = i\}$ : probability that the Markov Chain starts out in state i
- the probability transition rules

#### The Probability Matrix

- The probability matrix  $P = (P_{ij})$  specifies the transition rules
- if the size of S is N, P is a  $N \times N$  stochastic matrix
  - each entry is non negative
  - the sum of each row is 1
- $P_{ij}$  is a conditional probability: defines the probability that the chain jumps to state j, at time n + 1, given that it is in state i at time n,

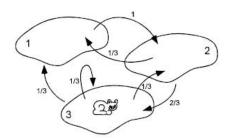
$$P_{ij} = P(X_{n+1} = j \mid X_n = i)$$

• We assume Time Homogeneity: the probability does not depend on the time n, but probability only on depends state i and i

### Time Homogeneity

- "where I go next given that I am in state s at time x is equal to where I go next given that I am in the same state s at time  $y \neq x$ "
- every time the chain is in state s, the probability of jumping to another state is the same
- the same probability over time
- we will assume time homogeneity in the following

### The Markov Frog



A frog hopping among lily pads

• State Space:  $S = \{1, 2, 3\}$  represent the pads

• Initial Distribution:

- $\pi_0 = (1/2, 1/4, 1/4)$
- Probability Transition Matrix:

$$P = \begin{pmatrix} 0 & 1 & 0 \\ 1/3 & 0 & 2/3 \\ 1/3 & 1/3 & 1/3 \end{pmatrix}$$

• defines the probabilities of jumping from one state to another one

# The Markov Frog

- the frog chooses its initial position  $X_0$  according to the initial distribution  $\pi_0$
- to this purpose, the frog can ask its computer to generate a uniformly distributed random number  $U_0$  in the interval I = [0,1] and then taking

$$X_0 = \begin{cases} 1 & if \quad 0 \le U_0 \le 1/2 \\ 2 & if \quad 1/2 < U_0 \le 3/4 \\ 3 & if \quad 3/4 < U_0 \le 1 \end{cases}$$

- for instance, if  $U_0 = 0.8419$  then  $X_0 = 3$ 
  - the frog starts on the third lily pad

# The Markov Frog

Basic Definitions

- Evolution of the Markov Chain: the frog chooses a lily pad to jump
  - state after the first jump = value of the random variable  $X_1$

- the frog starts from lily pad 3 so look at the probability distribution in row 3 of P, namely (1/3,1/3,1/3)
- again, generate a uniformly distributed random number  $U_1$  in the interval I = [0, 1] then takes

$$X_1 = \begin{cases} 1 & if & 0 \le U_1 \le 1/3 \\ 2 & if & 1/3 < U_1 \le 2/3 \\ 3 & if & 2/3 < U_1 \le 1 \end{cases}$$

- if  $U_1 = 0.1234$ , then  $X_1 = 1$ 
  - the frog jumps from the lily pad 3 to lily pad 1
  - $X_1 = 1$  there is no choice for the value of  $X_2$ , it must be 2
    - and so on....

### The Markov Property

• in the previous example

$$\mathbb{P}\{X_3 = j \mid X_2 = 2, X_1 = 1, X_0 = 3\} = \mathbb{P}\{X_3 = j \mid X_2 = 2\} \ \forall j$$

It's All Just Matrix Theory?

- the only information relevant to the distribution to  $X_3$  is the information that  $X_2 = 2$
- $X_0 = 3$  and  $X_1 = 1$  may be ignored!

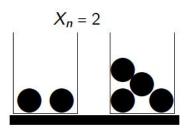
#### Definition

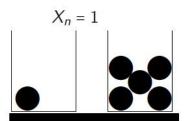
A stochastic process  $X_0, X_1, \dots$  satisfies the Markov Property if  $\mathbb{P}\{X_{n+1}=i_{n+1}\mid X_n=i_n,X_{n-1}=i_{n-1},\ldots,X_0=i_0\}=$  $\mathbb{P}\{X_{n+1} = i_{n+1} \mid X_n = i_n\}$ 

#### Markov Chains: Ehrenfest Chain

There is a total of 6 balls in two urns, 4 in the first and 2 in the second. We pick one of the 6 balls at random and move it to the other urn.

 $X_n$  number of balls in the first urn, after the nth move.





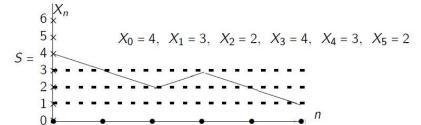
#### Markov Chains: Ehrenfest Chain

$$\mathbb{P}(X_0 = 4) = 1$$

$$\mathbb{P}(X_1 = j) = \begin{cases} 4/6 & j = 3\\ 2/6 & j = 5\\ 0 & otherwise \end{cases}$$

$$\mathbb{P}(X_{n+1} = l \mid X_n = j) = \begin{cases} j/6 & l = j - 1\\ (6 - j)/6 & l = j + 1\\ 0 & otherwise \end{cases}$$

In 5 unites of time  $X_0, \ldots X_5$  might follow the following path:



#### Ehrenfast Chain: Probability Transition Matrix

$$P = \begin{pmatrix} 0 & 1 & 0 & 0 & 0 & 0 & 0 \\ 1/6 & 0 & 5/6 & 0 & 0 & 0 & 0 \\ 0 & 2/6 & 0 & 4/6 & 0 & 0 & 0 \\ 0 & 0 & 3/6 & 0 & 3/6 & 0 & 0 \\ 0 & 0 & 0 & 4/6 & 0 & 2/6 & 0 \\ 0 & 0 & 0 & 0 & 5/6 & 0 & 1/6 \\ 0 & 0 & 0 & 0 & 0 & 1 & 0 \end{pmatrix}$$

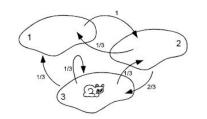
$$0 \xrightarrow{1/6} 1 \xrightarrow{2/6} 2 \xrightarrow{3/6} 3 \xrightarrow{4/6} 4 \xrightarrow{5/6} 5 \xrightarrow{5/6} 6$$

#### Markov Chains and Random Walks

- given a graph and a starting point
- select one of its neighbours at random, and move to this neighbor
- then select a neighbour of the new point point at random, and move to it, and so on ...
- the (random) sequence of points selected in this way is a random walk on the graph.
- there is a strict relation between Random Walk and Markov chains
  - time-reversible Markov chains can be viewed as random walks on undirected graphs
  - we will see this later ...

#### Random Walks and Complex Networks

- what is the connection between Random Walk and P2P/Social Networks?
- a simple way to analyse complex graph
  - gather information about all nodes and links at a single server
  - exploit classical graph theory
  - unfeasible solution due to the huge dimension of the graphs
- an alternative solution
  - compute a node sample through a random walk
  - exploit the node sample to estimate network properties



- what is the probability  $p_{ij}^n$  that, given the chain in state i, it will be in state j, n step after?
- if we start on lily pad 3, what is the probability of being on lily pad 1, after 2 steps?

$$p_{31}^2 = p_{31}p_{11} + p_{32}p_{21} + p_{33}p_{31} = \frac{1}{3}0 + \frac{1}{3}\frac{1}{3} + \frac{1}{3}\frac{1}{3} = \frac{1}{9} + \frac{1}{9} = \frac{2}{9}$$

- This is the dot product third row-first column of P
  - this returns the 1,3-entry of the product of P with itself.

The generalization of the previous observation leads to the following theorem.

#### Theorem

Let P be the transition matrix of a Markov chain. The ij-th entry of the matrix  $P^n$  gives the probability that the Markov chain, starting in state  $s_i$ , will be in state  $s_j$  after n steps.

The power of the transition matrix gives interesting information about the evolution of the process

$$P^{1} = \begin{array}{c} pad_{1} & pad_{2} & pad_{3} \\ pad_{1} & 0 & 1 & 0 \\ 0.333 & 0 & 0.666 \\ 0.333 & 0.333 & 0.333 \\ pad_{1} & pad_{2} & pad_{3} \\ \end{array}$$
 
$$P^{2} = \begin{array}{c} pad_{1} \\ pad_{2} \\ pad_{3} \\ \end{array} \begin{pmatrix} 0.333 & 0 & 0.666 \\ 0.221 & 0.554 & 0.221 \\ 0.221 & 0.443 & 0.332 \\ \end{array}$$
 
$$\begin{array}{c} pad_{1} \\ pad_{2} \\ pad_{3} \\ \end{array} \begin{pmatrix} 0.258 & 0.295 & 0.442 \\ 0.244 & 0.404 & 0.342 \\ 0.244 & 0.392 & 0.355 \\ \end{array}$$
 
$$\begin{array}{c} pad_{1} \\ pad_{2} \\ pad_{3} \\ \end{array} \begin{pmatrix} 0.246 & 0.368 & 0.371 \\ 0.244 & 0.369 & 0.367 \\ 0.245 & 0.369 & 0.367 \\ \end{array}$$

$$P^{5} = \begin{array}{c} pad_{1} & pad_{2} & pad_{3} \\ pad_{1} & 0.241 & 0.363 & 0.362 \\ pad_{2} & 0.239 & 0.361 & 0.360 \\ 0.240 & 0.361 & 0.361 \\ pad_{1} & pad_{2} & pad_{3} \\ \end{array}$$

$$P^{6} = \begin{array}{c} pad_{1} & 0.231 & 0.249 & 0.248 \\ 0.230 & 0.247 & 0.246 \\ 0.230 & 0.247 & 0.247 \\ pad_{1} & pad_{2} & pad_{3} \\ \end{array}$$

$$P^{7} = \begin{array}{c} pad_{1} & 0.213 & 0.322 & 0.321 \\ pad_{2} & 0.212 & 0.320 & 0.320 \\ pad_{3} & 0.212 & 0.321 & 0.320 \\ \end{array}$$

- The probabilities of the three lily pads are .2, .3, and .3 no matter where the frog starts at the first step.
- the long range predictions are independent from the starting state
- the columns are about identical because the chain forgets the initial state
- but ... this does not happen for all chains ....
  - we will show the conditions to be satisfied by the chain to guarantee this behaviour

• consider  $\pi_0$  is a vector of N component defining,  $\forall$  state i, the probability that the Markov chain is initially at i

$$\pi_0(i) = \mathbb{P}\{X_0 = i\}, i=1...N$$

•  $\pi_n(j)$  defines the probability that the Markov chain is at j, after n steps.

$$\pi_n = \{\pi_n(1), \dots, \pi_n(N)\}\$$

#### Theorem

Let P be the transition matrix of a Markov chain, and let  $\pi_0$  be the probability vector which represents the starting distribution. Then the probability that the chain is in state  $s_i$  after n steps is the ith entry of  $\pi^n = \pi_0 P^n$ 

•  $\pi_n(j)$  defines,  $\forall$  state j, the probability that the Markov chain is at j, after n steps.

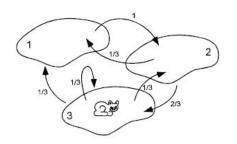
$$\pi_n = \{\pi_n(1), \dots, \pi_n(N)\}$$

• thus, we have

$$\pi_1 = \pi_0 P$$
 $\pi_2 = \pi_1 P = \pi_0 P^2$ 
 $\pi_3 = \pi_2 P = \pi_0 P^3$ 

• in general

$$\pi_n = \pi_0 P^n$$



- let us suppose  $\pi_0 = (1/3, 1/3, 1/3)$
- we want to compute the state distribution after 3 frog jumps

• 
$$\pi^3 = \pi_0 P^3 =$$

$$(\frac{1}{3}, \frac{1}{3}, \frac{1}{3}) \begin{pmatrix} 0.258 & 0.295 & 0.442 \\ 0.244 & 0.404 & 0.342 \\ 0.244 & 0.392 & 0.355 \end{pmatrix}$$

#### No, not just matrix theory...

In principle, matrix algebra can give us the answer to any question about the probabilistic behaviour of a Markov chain, but this is not viable, in practice

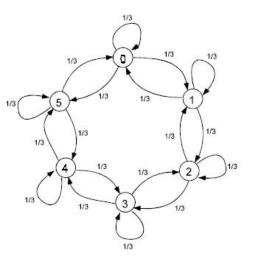
#### Consider a P2P system

- Markov chains describe random walk on P2P overlays
  - states corresponds to peer
  - transition correspond to link overlays
- the size of the matrix is huge, matrix algebra may be not exploited

The same for social graph analysis



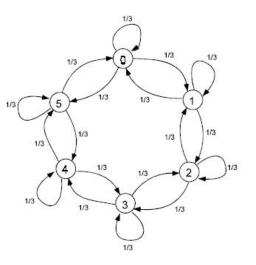
#### A Random Walk on a Clock



- a simplified clock with 6 number 0,1,2,3,4,5
- from each state we can move clockwise, counter clockwise, or stay in place with the same probability
- the transition matrix is

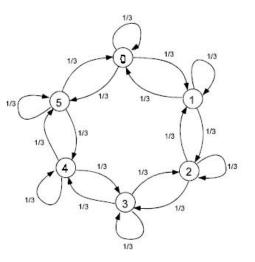
$$P(i,j) = \begin{cases} 1/3 & \text{if } j=i-1 \text{ mod } 6\\ 1/3 & \text{if } j=i\\ 1/3 & \text{if } j=i+1 \text{ mod } 6 \end{cases}$$

#### A Random Walk on a Clock



- suppose we start out at  $X_0 = 2$ , that is  $\pi_0 = (0, 0, 1, 0, 0, 0)$
- $\pi_1 = (0, 1/3, 1/3, 1/3, 0, 0)$ •  $\pi_2 = (1/9, 2/9, 1/3, 2/9, 1/9, 0)$
- $\pi_3 = (3/27, 6/27, 7/27, 6/27, 3/27, 2/27, 1$
- notice that the probability is spreading out away its initial concentration on the initial state 2

#### A Random Walk on a Clock



- now guess what is the state of the random walk at time 10000
- an intuitive answer:  $X_{10000}$  is uniformly distributed over the 6 states
- this can be proven formally  $\pi_n \to (\frac{1}{6}, \frac{1}{6}, \frac{1}{6}, \frac{1}{6}, \frac{1}{6}, \frac{1}{6})$
- after 10000 steps, the random walk has forgotten that it started out in state 2

#### Markov Chains: The Basic Theorem

#### Theorem (Basic Limit Theorem)

Any irreducible, aperiodic, Markov chain defined on a set of state S and with a stochastic transition matrix P has a unique stationary distribution  $\pi$  with all its component positive. Furthermore, let  $P^n$  be the n-th power of P, then  $\lim_{n\to+\infty} P_{i,j}^n = \pi(j), \forall i,j\in S$ 

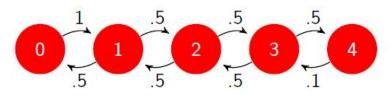
to understand the theorem, we need to define the words

- irreducible
- aperiodic
- stationary distribution

Before we shall see some examples where the limit does not exist

#### Random Walk with Reflecting Boundary(Periodic)

#### Periodic chain with period 2



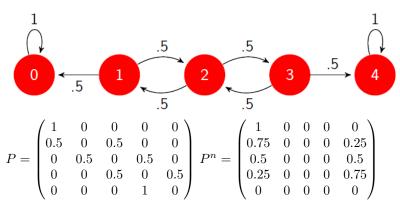
$$P = \begin{pmatrix} 0 & 1 & 0 & 0 & 0 \\ 0.5 & 0 & 0.5 & 0 & 0 \\ 0 & 0.5 & 0 & 0.5 & 0 \\ 0 & 0 & 0.5 & 0 & 0.5 \\ 0 & 0 & 0 & 1 & 0 \end{pmatrix} P^{2n} = \begin{pmatrix} 0.25 & 0 & 0 & 0.5 & 0.25 \\ 0 & 0.5 & 0 & 0.5 & 0 \\ 0.25 & 0 & 0.5 & 0 & 0.25 \\ 0 & 0.5 & 0 & 0.5 & 0 \\ 0.25 & 0 & 0.5 & 0 & 0.25 \end{pmatrix}$$

$$P^{2n+1} = \begin{pmatrix} 0 & 0.5 & 0 & 0.5 & 0 \\ 0.25 & 0 & 0.5 & 0 & 0.25 \\ 0 & 0.5 & 0 & 0.5 & 0 \\ .25 & 0 & 0.5 & 0 & 0.25 \\ 0 & 0.5 & 0 & 0.5 & 0 \end{pmatrix}$$

This means at every state i, if  $P^n(i,i) > 0$ , then n is even.

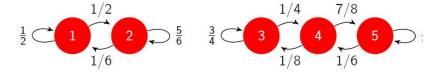


#### Random Walk With Absorbing Boundary



- rows are not identical, i.e, the chain does not forget the initial state
- states 1, 2, 3 are transient:  $p_{i,1}^n = p_{i,2}^n = p_{i,3}^n = 0 \ \forall i$ , i.e. after passing a large amount of time the chain will stop visiting states 1, 2, 3.

#### Reducible Chains



$$P = \begin{pmatrix} 1/2 & 1/2 & 0 & 0 & 0 \\ 1/6 & 5/6 & 0 & 0 & 0 \\ 0 & 0 & 3/4 & 1/4 & 0 \\ 0 & 0 & 1/8 & 2/3 & 5/24 \\ 0 & 0 & 0 & 1/6 & 1/6 \end{pmatrix}$$

$$P^{n} = \begin{pmatrix} 0.25 & 0.75 & 0 & 0 & 0\\ 0.25 & 0.75 & 0 & 0 & 0\\ 0 & 0 & 0.182 & 0.364 & 0.455\\ 0 & 0 & 0.182 & 0.364 & 0.455\\ 0 & 0 & 0.182 & .364 & 0.455 \end{pmatrix}$$

# Stationary Distribution

a distribution which satisfies

$$\pi = \pi P$$

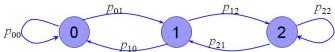
- every finite state Markov chain has at least one stationary distribution
- A trivial example with infinitely many stationary distribution: when P is the identity matrix, in which case all distributions are stationary
- The stationary distribution is the one obtained by computing the power of the transition matrix

## Markov Chain Irreducibility

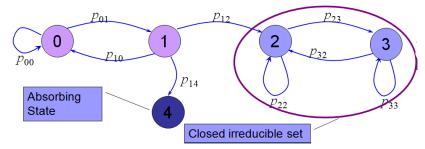
- State j is reachable from state i if probability to go from i to i in n > 0 steps is greater then 0
- A subset X of the state space S is closed if  $p_{ij} = 0, \forall i \in S$ and  $i \notin S$
- A closed set of states is *irreducible* if any state  $j \in S$  is reachable from any state  $i \in S$
- A Markov Chain is *irreducible* if the state space S is irreducible

### Markov Chain Irreducibility

• Irreducible Markov Chain



• Reducible Markov Chain



#### Markov Chain Periodicity

#### Definition

Given a Markov chain  $X_0, X_1, \ldots$ , the period  $p_i$  of a state i is the greatest common divisor (gcd)

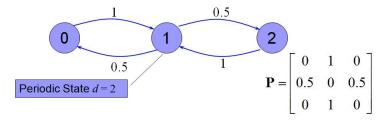
$$p_i = \{ \gcd n \text{ such that } P^n(i, i) > 0 \}$$

It's All Just Matrix Theory?

#### Definition

An irreducible Markov chain is aperiodic if its period is 1, and periodic otherwise.

### Markov Chain Periodicity



- If the period of state i is  $\neq 1$ , the chain returns to state i at regular times,
- If the period of state i is 1, the chain returns to state i can occur at irregular times
- A simple sufficient (but not necessary) condition for an irreducible chain to be aperiodic is that there exist a state i such that P(i,i) > 0

#### Markov Chain Recurrence

- Hitting Time:  $T_{ij} = min\{k > 0 : X_0 = i, X_k = j\}$
- Recurrence Time: is the first time that the chain returns to state i
- Let  $\rho_i$  the probability that the chain will come back to state i, given it starts from state i
- A stete is recurrent if  $\rho = 1$ , otherwise it is transient.

#### Recurrent and Transient States

