Simulation of Markov genealogy processes

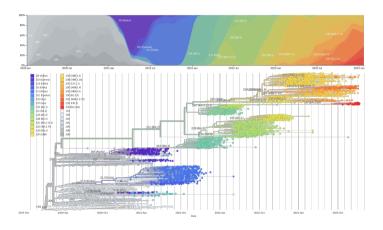
Aaron A. King

February 27, 2024

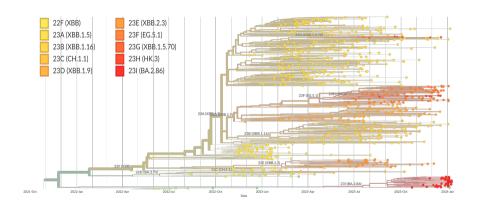
Outline

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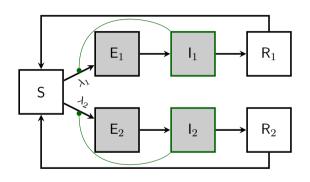
Objectives



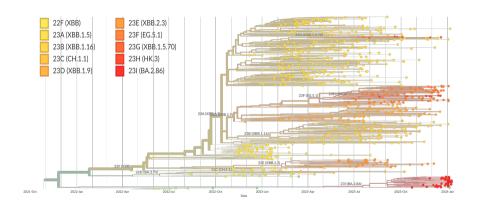
nextstrain.org (?)



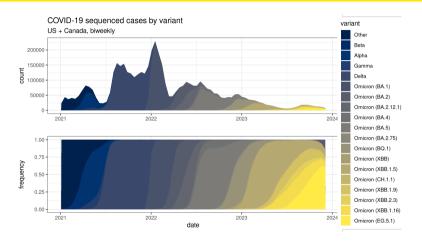
nextstrain.org(?)



$$\lambda_1 = \beta_1 \frac{I_1}{N} \qquad \lambda_2 = \beta_2 \frac{I_2}{N}$$



nextstrain.org(?)

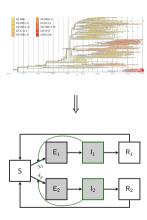


What is phylodynamics?

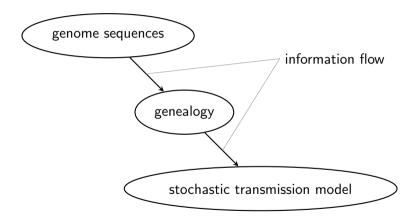
Broadly: Phylodynamics is

Phylodynamics is the project of inferring determinants of epidemic spread using genomic data from pathogen samples.

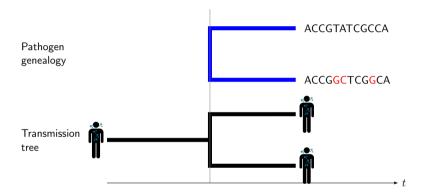
In this talk:
Phylodynamics means using
genomic data
to infer
stochastic dynamic transmission models.



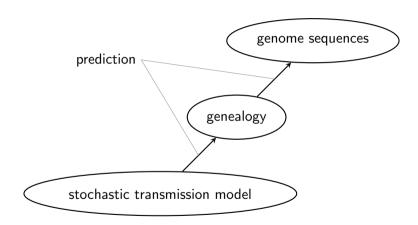
Core problems of phylodynamics



Core problems of phylodynamics



Core problems of phylodynamics

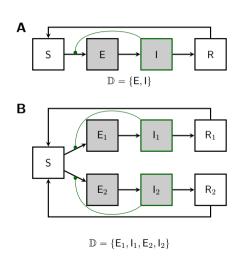


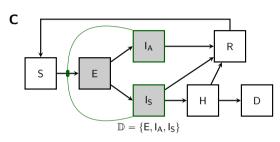
Overview

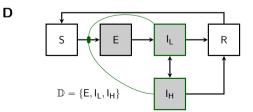
- We show how a given population process induces a unique genealogy process.
- Pruning and obscuration project a genealogy onto observable data.
- We derive the exact likelihood as the solution to a nonlinear filtering problem
- This equation can be solved by standard Monte Carlo methods.

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- Non-explosive Markov jump process, $\mathbf{X}_t \in \mathbb{X}$, $t \in \mathbb{R}_+$: the *population process*.
- Initial-state distribution, p_0 :

$$\mathsf{Prob}\left[\mathbf{X}_0 \in \mathcal{E}\right] = \int_{\mathcal{E}} p_0(x) \, \mathrm{d}x$$

• Jump rates: $\alpha(t, x, x') = \text{rate of jump } x \to x'$

$$\alpha(t, x, x') \ge 0, \qquad \int_{\mathbb{X}} \alpha(t, x, x') \, \mathrm{d}x' < \infty$$

Multiple events at each jump are allowed.

Kolmogorov forward equation (KFE):

lf

$$\frac{\partial w}{\partial t}(t,x) = \int w(t,x') \,\alpha(t,x',x) \,\mathrm{d}x' - \int w(t,x) \,\alpha(t,x,x') \,\mathrm{d}x'$$

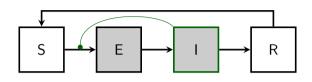
and

$$w(0,x) = p_0(x)$$

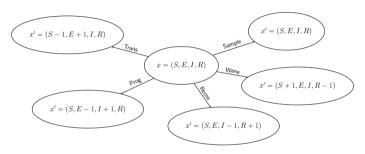
then

$$\int_{\mathcal{E}} w(t, x) \, \mathrm{d}x = \mathsf{Prob} \left[\mathbf{X}_t \in \mathcal{E} \right].$$

KFE is sometimes called the *master equation* for \mathbf{X}_t .

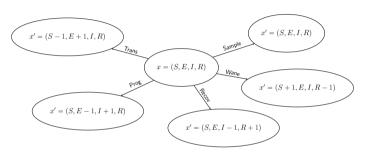


$$\frac{\partial w}{\partial t}(t,x) = \int w(t,x') \,\alpha(t,x',x) \,\mathrm{d}x' - \int w(t,x) \,\alpha(t,x,x') \,\mathrm{d}x'$$



 $\mathbb{U} = \{\mathsf{Trans}, \mathsf{Prog}, \mathsf{Recov}, \mathsf{Wane}, \mathsf{Sample}\}$

$$\frac{\partial w}{\partial t}(t,x) = \sum_{u \in \mathbb{U}} \left\{ \int w(t,x') \, \alpha_u(t,x',x) \, \mathrm{d}x' - \int w(t,x) \, \alpha_u(t,x,x') \, \mathrm{d}x' \right\}$$



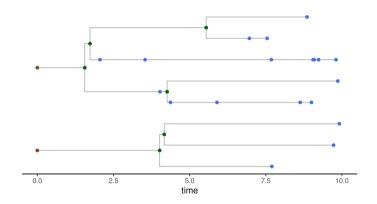
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$$\frac{\partial w}{\partial t}(t, S, E, I, R) = \frac{\beta(t) \, (S+1) \, I}{N} \, w(t, S+1, E-1, I, R) - \frac{\beta(t) \, S \, I}{N} \, w(t, S, E, I, R) + \sigma \, (E+1) \, w(t, S, E+1, I-1, R) - \sigma \, E \, w(t, S, E, I, R) + \frac{\beta(t) \, S \, I}{N} \, w(t$$

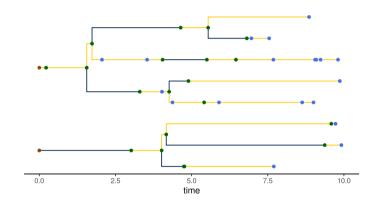
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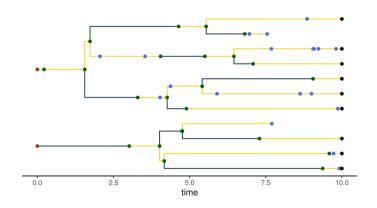
What is a genealogy?

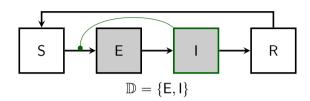


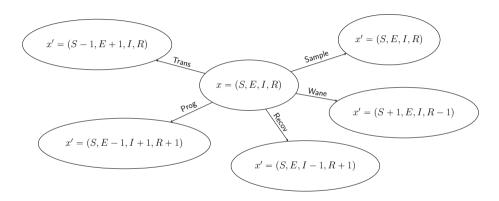
What is a genealogy?



What is a genealogy?







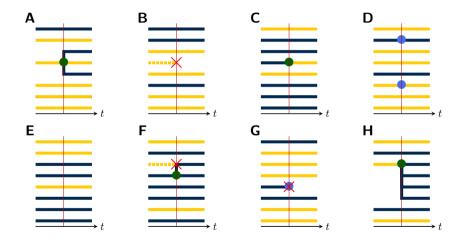
 $\mathbb{U} = \{\mathsf{Trans}, \mathsf{Prog}, \mathsf{Recov}, \mathsf{Wane}, \mathsf{Sample}\}$

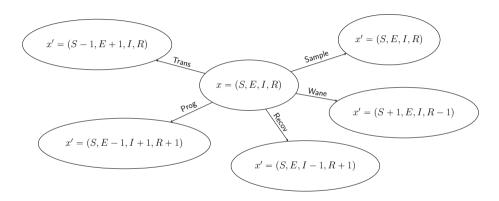
If we write

$$\alpha(t, x, x') = \sum_{u \in \mathbb{U}} \alpha_u(t, x, x'),$$

the KFE becomes

$$\frac{\partial w}{\partial t}(t, x) = \sum_{u} \int w(t, x') \alpha_u(t, x', x) dx' - \sum_{u} \int w(t, x) \alpha_u(t, x, x') dx'$$





 $\mathbb{U} = \{\mathsf{Trans}, \mathsf{Prog}, \mathsf{Recov}, \mathsf{Wane}, \mathsf{Sample}\}$

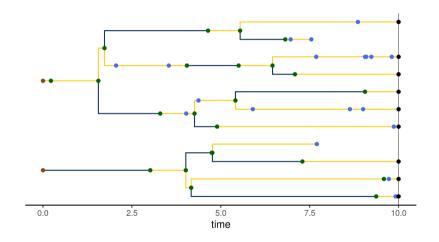
A population process induces a genealogy process

- ullet G_t is a stochastic process on the space of genealogies.
- The map $X \mapsto G$ is random.
- **Key assumption:** Lineages within a deme are *exchangeable*. There is no more structure than is implied by the population process.
- Simulation code on github.com/kingaa/phylopomp
- Animations at https://kingaa.github.io/manuals/phylopomp/vignettes/

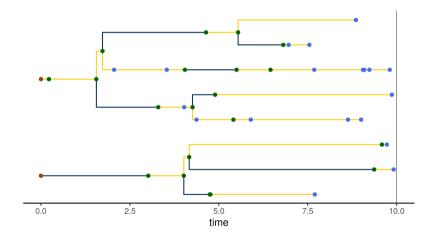
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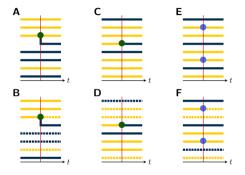
Full genealogy



Pruned genealogy

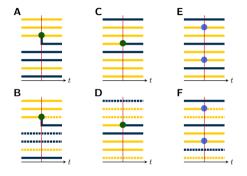


Local structure of a pruned genealogy



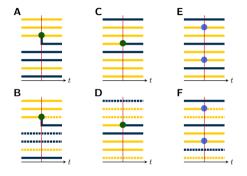
Top row shows the *unpruned genealogy* in neighborhood of an event. Bottom row shows the corresponding *pruned genealogy*.

Local structure of a pruned genealogy



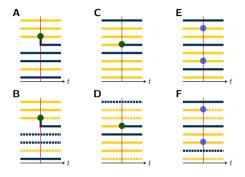
For $x \in \mathbb{X}$, $i \in \mathbb{D}$, $n_i(x)$ is the *occupancy* of deme i when the system is in state x. In panel A $n = (n_{\mathsf{blue}}, n_{\mathsf{yellow}}) = (4, 4)$; in panel C n = (3, 5);

Local structure of a pruned genealogy



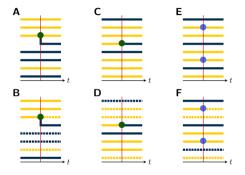
For $u \in \mathbb{U}$, $i \in \mathbb{D}$, r_i^u is the *production* of event u in deme i. In panel A, $r = (r_{\text{blue}}, r_{\text{yellow}}) = (1, 1)$; in panel E, r = (0, 2).

Local structure of a pruned genealogy



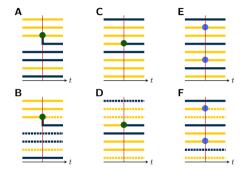
The *lineage count*, $\ell_i(t)$, is the number of unpruned lineages in deme i at time t. In this case, for all panels, $\ell=(2,2)$.

Local structure of a pruned genealogy



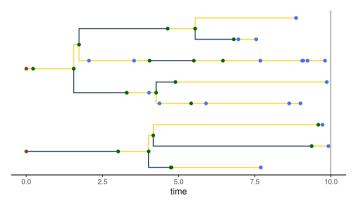
The saturation, s_i , is the number of unpruned lineages in deme i descending from the event. In panels B and D, s = (1,0); in panel F, s = (0,1).

Local structure of a pruned genealogy



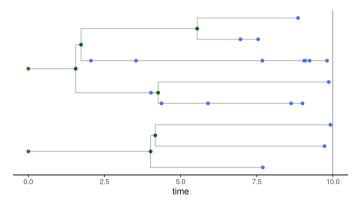
Obviously, $s_i \leqslant r_i \leqslant n_i$ and $s_i \leqslant \ell_i \leqslant n_i$.

Pruned genealogy



A pruned genealogy is specified by two functions of time, (Y,Z): Z_t gives the local topological structure; Y_t gives the local coloring.

Obscured genealogy



An obscured genealogy is specified by $(\mathrm{T},\mathrm{Z}).$

Binomial ratio

For $n, r, \ell, s \in \mathbb{Z}_+^{\mathbb{D}}$, define the *binomial ratio*

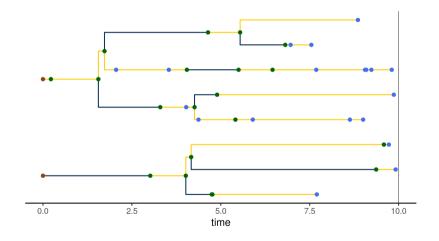
$$\begin{pmatrix} n & \ell \\ r & s \end{pmatrix} := \begin{cases} \prod_{i \in \mathbb{D}} \frac{\binom{n_i - \ell_i}{r_i - s_i}}{\binom{n_i}{r_i}}, & \text{if } \forall i \ n_i \geqslant \{\ell_i, r_i\} \geqslant s_i \geqslant 0, \\ 0, & \text{otherwise}. \end{cases}$$

Observe that $\binom{n}{r} \stackrel{\ell}{s} \in [0,1]$. Moreover,

$$\sum_{s \in \mathbb{Z}_{-}^{\mathbb{D}}} \binom{n}{r} \binom{\ell}{s} \binom{\ell}{s} = 1.$$

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Suppose that P=(Y,Z) is a given pruned genealogy with depth T. Define

$$\phi_u(x, y, y') \coloneqq \begin{pmatrix} n(x) & \ell(y') \\ r^u & s(y, y') \end{pmatrix} Q_u(y, y').$$

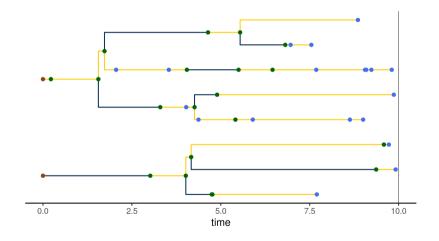
Here, Q=1 if the local structure of P is compatible with an event of type u at that time; Q=0 otherwise.

If w = w(t, x) satisfies the initial condition $w(0, x) = p_0(x)$ and the filter equation

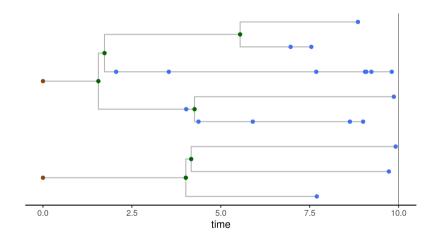
$$\begin{split} \frac{\partial w}{\partial t}(t,x) &= \sum_{u} \int w(t,x') \, \alpha_u(t,x',x) \, \phi_u(x,\widetilde{\mathbf{Y}}_t,\mathbf{Y}_t) \, \mathrm{d}x' - \sum_{u} \int w(t,x) \, \alpha_u(t,x,x') \, \mathrm{d}x' \\ &+ \sum_{e \in \mathsf{ev}(\mathbf{P})} \delta(t,e) \, \left\{ \sum_{u} \int w(t,x') \, \alpha_u(t,x',x) \, \phi_u(x,\widetilde{\mathbf{Y}}_t,\mathbf{Y}_t) \, \mathrm{d}x' - w(t,x) \right\}, \end{split}$$

then the likelihood of P is

$$\mathcal{L}(\mathbf{P}) = \int w(\mathbf{T}, x) \, \mathrm{d}x.$$



Theorem: likelihood of an obscured genealogy



Theorem: likelihood of an obscured genealogy

Let (T, Z), be a given obscured genealogy. Then there are probability kernels π and q such that if

$$\beta_u(t, x, x', y, y') = \alpha_u(t, x, x') \pi_u(t, x, x', y, y'), \qquad \psi_u(t, x, x', y, y') = \frac{\phi_u(x', y, y')}{\pi_u(t, x, x', y, y')},$$

and if w = w(t, x, y) satisfies the initial condition $w(0, x, y) = p_0(x) q(x, y)$ and the filter equation

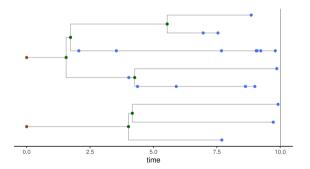
$$\frac{\partial w}{\partial t} = \sum_{uy'} \int w(t, x', y') \,\beta_u(t, x', x, y', y) \,\psi_u(t, x', x, y', y) \,\mathrm{d}x' - \sum_{uy'} \int w(t, x, y) \,\beta_u(t, x, x', y, y') \,\mathrm{d}x'$$

$$\frac{\partial t}{\partial t} = \sum_{uy'} \int w(t, x, y) \beta_u(t, x, x, y, y) \psi_u(t, x, x, y, y) dx - \sum_{uy'} \int w(t, x, y) \beta_u(t, x, x, y, y) dx + \sum_{e \in \text{ev}(\mathbf{Z})} \delta(t, e) \left\{ \sum_{uy'} \int w(t, x', y') \beta_u(t, x', x, y, y') \psi_u(t, x', x, y', y) dx' - w(t, x, y) \right\},$$

Theorem: likelihood of an obscured genealogy

then the likelihood of Z is

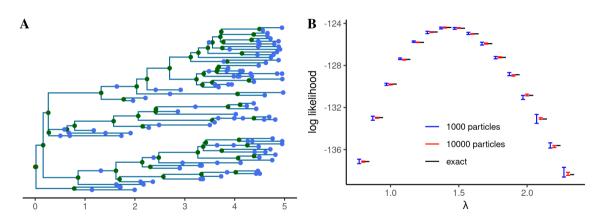
$$\mathcal{L}(\mathbf{Z}) = \sum_{y} \int w(T, x, y) \, \mathrm{d}x.$$



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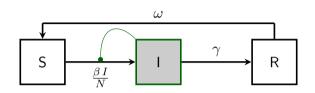
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Linear birth-death model



Uniform sampling. Exact likelihoood is available in closed form.

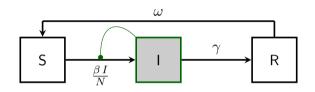
SIRS model



Between genealogical events:

$$\frac{\partial w}{\partial t} = \frac{\beta (S+1) (I-1)}{N} \left(1 - \frac{\binom{\ell}{2}}{\binom{I}{2}} \right) w(t, S+1, I-1, R) + \gamma (I+1) w(t, S, I+1, R-1) + \omega (R+1) w(t, S-1, I, R+1) - \left(\frac{\beta S I}{N} + \gamma I + \omega R + \psi I \right) w(t, S, I, R).$$

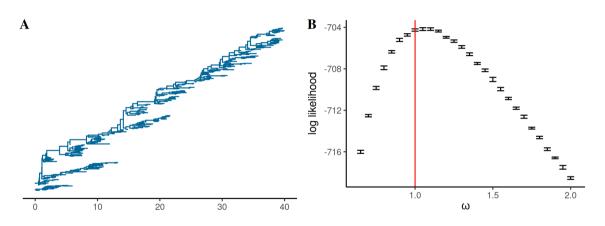
SIRS model



At genealogical events:

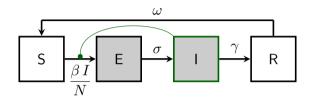
$$w(t,S,I,R) = \begin{cases} \frac{2\,\beta\,(S+1)}{I\,N}\,\widetilde{w}(t,S+1,I-1,R), & \text{branch point at }t,\\ \\ \psi\,\widetilde{w}(t,S,I,R), & \text{internal sample at }t,\\ \\ \psi\,\,(I-\ell)\,\,\widetilde{w}(t,S,I,R), & \text{terminal sample at }t. \end{cases}$$

SIRS model

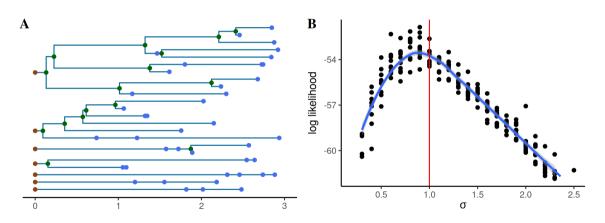


Uniform sampling. One deme only.

SEIRS model



SEIRS model



Concluding remarks

- The theory *corrects* and *strictly extends* all existing likelihood-based phylodynamic methods (e.g., ???????).
- It eliminates the need for large population-size and small sample-fraction assumptions, as well as any dependence on linearization.
- All computations can be carried out forward in time.
- This greatly expands the class of models that can be entertained.
- Great flexibility in sampling model
- Applications in disease ecology and beyond
- The unstructured case can be found in King et al. (2022).

Outstanding challenges

- Choice of importance-sampling kernel
- Borrowing information from future is allowed.
- Phylogenetic uncertainty
- Efficient algorithms
- Reassortment and recombination

Summary

- A discretely structured Markov population process uniquely induces a genealogy-valued Markov process.
- The likelihood of an observed genealogy satisfies a nonlinear filtering equation, which can be efficiently computed via Feynman-Kaç (sequential Monte Carlo) algorithms.
- In principle, these results liberate us to entertain models that more closely match our biological questions, without less hindrance from inference methodology.

References

- King AA, Lin Q, Ionides EL (2022). "Markov genealogy processes." *Theor Popul Biol*, **143**, 77–91. doi: 10.1016/j.tpb.2021.11.003.
- King AA, Nguyen D, Ionides EL (2016). "Statistical inference for partially observed Markov processes via the R package pomp." *J Stat Softw*, **69**(12), 1–43. doi: 10.18637/jss.v069.i12.

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- Produced with R version 4.3.2 and **pomp** version 5.6.
- Compiled on February 27, 2024.

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