

STA642 hw2
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Q2

(a)

$$\begin{aligned}
 y_t &= x_t + \nu_t \\
 &= \phi x_{t-1} + \epsilon_t + \nu_t \\
 &= \phi(y_{t-1} - \nu_{t-1}) + \epsilon_t + \nu_t \\
 &= \phi y_{t-1} + \eta_t \\
 &\text{where } \eta = \epsilon_t + \nu_t - \phi \nu_{t-1}
 \end{aligned}$$

(b)

Note that $Var(\eta_t) = Var(\epsilon_t) + Var(\nu_t) + \phi^2 Var(\nu_{t-1})$ and $C(\eta_t, \eta_{t-1}) = C(\epsilon_t + \nu_t - \phi \nu_{t-1}, \epsilon_{t-1} + \nu_{t-1} - \phi \nu_{t-2}) = -\phi\omega + \text{bunch of 0's (by iid)}$. Thus we have $Corr(\eta_t, \eta_{t-1}) = \frac{-\phi\omega}{\omega(1+\phi^2)+\nu}$.

(c)

Note that $Var(y_t) = s + \omega = q$. For $C(y_t, y_{t+k})$:

$$\begin{aligned}
 C(y_t, y_{t+k}) &= C(x_t + \nu_t, \phi(x_t + \nu_t) + \sum_{i=1}^{k-1} \phi^i (\epsilon_{t+k-i} + \nu_{t+k-i} - \phi \nu_{t+k-i-1})) \\
 &= C(x_t, \phi^k x_t) \text{ by iid} \\
 &= \phi^k s
 \end{aligned}$$

Thus $Corr(y_t, y_{t+k}) = \frac{s}{q} \phi^k$.

Note that the correlation is positively related to the signal:noise ratio, this is intuitive. If ω is large relatively to s (small s/q), then more variation will be added to the AR(1) model, x_t , to form y_t . Then the auto-correlation of y_t will be small relatively. Note that ϕ^k is the lag- k auto-correlation of x_t .

(d)

I think y_t is not AR(1) because η_t is not iid. y_t is Markov since $y_t = \phi y_{t-1} + \eta_t$ only depends on y_{t-1} .

Q3

For 2 by 2 matrix case:

(a)

$$AG = \begin{bmatrix} 1 & 0 \\ 0 & \phi_2 \end{bmatrix} \begin{bmatrix} \phi_1 & \phi_2 \\ 1 & 0 \end{bmatrix} = \begin{bmatrix} \phi_1 & \phi_2 \\ \phi_2 & 0 \end{bmatrix}$$

AG is symmetric.

(b)

(i)

Suppose $|\phi_p| \neq 0$, then $\|A\| = \det\left(\begin{bmatrix} 1 & 0 \\ 0 & \phi_2 \end{bmatrix}\right) = \phi_2 \neq 0$, thus A is non-singular.

(ii)

$$G'A = \begin{bmatrix} \phi_1 & 1 \\ \phi_2 & 0 \end{bmatrix} \begin{bmatrix} 1 & 0 \\ 0 & \phi_2 \end{bmatrix} = \begin{bmatrix} \phi_1 & \phi_2 \\ \phi_2 & 0 \end{bmatrix} = AG$$

(iii)

$$AF = \begin{bmatrix} 1 & 0 \\ 0 & \phi_2 \end{bmatrix} \begin{bmatrix} 1 \\ 0 \end{bmatrix} = \begin{bmatrix} 1 \\ 0 \end{bmatrix} = F$$

(c)

It suffices to show that $G'z_{t-1} + F\epsilon_t = Ax_t$. Apply (b), Note that since A is non-singular, A^{-1} exists.

$$\begin{aligned} G'z_{t-1} + F\epsilon_t &= \\ &= AGA^{-1}z_{t-1} + AF\epsilon_t \\ &= A(GA^{-1}z_{t-1} + F\epsilon_t) \\ &= A(GA^{-1}Ax_{t-1} + F\epsilon_t) \\ &= A(Gx_{t-1} + F\epsilon_t) \\ &= Ax_t \end{aligned}$$

(d)

$$z_t = Ax_t = \begin{bmatrix} 1 & 0 \\ 0 & \phi_2 \end{bmatrix} \begin{bmatrix} y_t \\ y_{t-1} \end{bmatrix} = \begin{bmatrix} y_t \\ \phi_2 y_{t-1} \end{bmatrix}$$

The first element of z_t is y_t , the current state. For the rest of the elements, the elements are the weighted sum of $y_{t-1}, \dots, y_{t-p+1}$ with weights of ϕ 's. Note that G and G' are similar, thus x_t and z_t form the same forecast function.