

Problem 1

a.

$$\begin{aligned} m_1 = \bar{X} = \mu'_1, \quad \mu'_1 &= \mathbb{E} \left[\frac{\theta x^{\theta-1}}{3^\theta} I_{(0,3)}(x) \right] \\ &= \int_0^3 x \frac{\theta x^{\theta-1}}{3^\theta} dx = \frac{\theta}{3^\theta} \int_0^3 x^\theta dx \\ &= \frac{\theta}{3^\theta} \frac{x^{\theta+1}}{\theta+1} \Big|_0^3 = \frac{3\theta}{\theta+1} \end{aligned}$$

$$\theta = \frac{-\mu'_1}{\mu'_1 - 3} = \frac{\bar{X}}{3 - \bar{X}}$$

This answer is intuitive because X_i has the range $(0, 3)$; therefore \bar{X} must be in this range as well. The denominator must therefore always be a positive number, and the range of θ becomes $(0, \infty)$, as given in the problem statement.

b.

$$\begin{aligned} L(\theta|x) &= \prod_{i=1}^n \frac{\theta x_i^{\theta-1}}{3^\theta} I_{(0,3)}(x) \\ \mathcal{L}(\theta|x) &= \sum_{i=1}^n \log \theta + (\theta - 1) \log x_i I_{(0,3)}(x) - \theta \log 3 \\ \frac{d\mathcal{L}(\theta|x)}{d\theta} &= \sum_{i=1}^n \frac{1}{\theta} + \log x_i I_{(0,3)}(x) - \log 3 \end{aligned}$$

Solve for a maximum by setting the derivative of the log-likelihood function equal to zero:

$$\begin{aligned} \sum_{i=1}^n \frac{1}{\theta} &= - \sum_{i=1}^n \log \frac{x}{3} \\ \frac{n}{\theta} &= - \sum_{i=1}^n \log \frac{x}{3} \\ \theta &= \frac{-n}{\sum_{i=1}^n \log \frac{x}{3}} \end{aligned}$$

Check the second derivative to ensure that this is a global maximum:

$$\frac{d^2 \mathcal{L}(\theta|x)}{d\theta^2} = \sum_{i=1}^n -\frac{1}{\theta^2} = -\frac{n}{\theta^2}$$

The second derivative is always negative, so the point in question must be a global maximum. This answer is intuitive, because $x < 3$; therefore $x/3 < 1$; therefore $\log x/3 < 0$. The sum of negative numbers is negative in the denominator, while $-n$ is negative in the numerator. Therefore, θ will be in the range $(0, \infty)$, as given in the problem statement.

Problem 2

a.

$$\begin{aligned} m_1 = \bar{X} = \mu'_1, \quad \mu'_1 &= \mathbb{E} \left[\frac{3x_i^2}{\theta^3} I_{(0,\theta]}(x) \right] \\ &= \int_0^\theta x \frac{3x^2}{\theta^3} dx = \frac{3}{\theta^3} \int_0^\theta x^3 dx \\ &= \frac{3}{\theta^3} \frac{x^4}{4} \Big|_0^\theta = \frac{3\theta}{4} \\ \theta &= \frac{4\bar{X}}{3} \end{aligned}$$

b.

$$\begin{aligned} L(\theta|x) &= \prod_{i=1}^n \frac{3x_i^2}{\theta^3} I_{(0,\theta]}(x) \\ &= \begin{cases} \frac{3^n}{\theta^{3n}} \prod_{i=1}^n x_i^2 & \theta \geq x_{(n)} \\ 0 & \theta < x_{(n)} \end{cases} \\ \frac{dL(\theta|x)}{d\theta} &= \frac{-n3^{n+1}}{\theta^{3n+1}} \prod_{i=1}^n x_i^2 \end{aligned}$$

This derivative is always negative, but never reaches zero. Since the derivative is negative, L is always decreasing on $(x_{(n)}, \infty)$. Therefore, the maximum value of the likelihood function is attained when $\theta = x_{(n)}$.

c.

$$\begin{aligned} F(x|\theta) &= \int \frac{3x^2}{\theta^3} dx \\ &= \frac{x^3}{\theta^3} = P(X < x) \end{aligned}$$

The probability that all of n X_i are less than x is equal to this expression raised to the n power. Let $Y = X_{(n)}$, then

$$\begin{aligned} P(Y < y) &= \frac{y^{3n}}{\theta^{3n}} = F(y|\theta) \\ f(y|\theta) &= \frac{d}{dx} \frac{y^{3n}}{\theta^{3n}} = \frac{3ny^{3n-1}}{\theta^{3n}} \\ E[f(y|\theta)] &= \int_0^\theta y \frac{3ny^{3n-1}}{\theta^{3n}} dx \\ &= \frac{3n}{3n+1} \frac{y^{3n+1}}{\theta^{3n}} \Big|_0^\theta = \frac{3n}{3n+1} \theta \end{aligned}$$

As n approaches infinity, $E(X_{(n)}) \rightarrow \theta$.

Problem 3

0.0.1 a.

$$\begin{aligned} p_2 \text{ observed} &= \frac{5}{25} = \theta(1 - \theta) \\ \theta^2 - \theta + 1/5 &= 0 \\ \theta &= \frac{\sqrt{5} \pm 1}{2\sqrt{5}} \end{aligned}$$

Since $\theta \in (0, 1/2)$, only the lower of the two roots is applicable, so

$$\theta = 0.276.$$

b.

$$\begin{aligned} L(\theta|n_1 = 11, n_2 = 5, n_3 = 9) &= \frac{25!}{11!5!9!} (\theta)^{11} (1 - \theta)^5 ((1 - \theta)^2)^9 \\ &= C\theta^{16} (1 - \theta)^{23} \end{aligned}$$

where $C = 25!/(11!5!9!) = 8923714800$.

$$\begin{aligned} \mathcal{L}(\theta) &= \log C + 16 \log \theta + 23 \log (1 - \theta) \\ \frac{d\mathcal{L}(\theta)}{d\theta} &= \frac{16}{\theta} + \frac{23}{\theta - 1} \end{aligned}$$

Set this derivative equal to zero to solve for maximum:

$$\begin{aligned} \frac{16}{\theta} &= \frac{23}{1 - \theta} \\ \theta &= \frac{16}{39} = 0.410 \end{aligned}$$

Check the second derivative to ensure that this is a global maximum:

$$\frac{d^2 \mathcal{L}(\theta)}{d\theta^2} = -\frac{16}{\theta^2} - \frac{23}{(\theta - 1)^2}$$

Since the second derivative is always negative, $\theta = 0.410$ is a global maximum of the likelihood function.

Problem 4

a.

$$\begin{aligned} m_1 &= \bar{X}, \quad \mu'_1 = \frac{1 + \theta x}{2} I_{[-1, 1]}(x) \\ &= \frac{1}{2} \int_{-1}^1 x(1 + \theta x) dx \\ &= \frac{1}{2} \left(\frac{x^2}{2} + \frac{\theta}{3} x^3 \right) \Big|_{-1}^1 = \frac{1}{2} \left(\frac{2\theta}{3} \right) \\ &= \frac{\theta}{3} \\ \theta &= 3\bar{X} \end{aligned}$$

b.

$$L(\theta|x) = \frac{1 + \theta x_1}{2}$$

The range of both x and θ are $[-1, 1]$. If $x > 0$, then likelihood increases with increasing θ , so maximum likelihood is when $\theta = 1$. If $x < 0$, likelihood decreases with increasing θ , so maximum likelihood is when $\theta = -1$. If $x = 0$, likelihood is a constant $1/2$ for all θ . Therefore the maximum likelihood expression is

$$\theta = \begin{cases} 1 & x > 0 \\ [-1, 1] & x = 0 \\ -1 & x < 0 \end{cases}$$

c.

Problem 5

$$m_1 = \bar{X} = 0.5166$$

$$m_2 = \frac{1}{n} \sum_{i=1}^n x_i^2 = 1.3113$$

$$\begin{aligned} \mu'_i &= \mathbb{E} \left[\frac{1}{2\delta} I_{[\gamma-\delta, \gamma+\delta]}(u) \right] \\ &= \int_{\gamma-\delta}^{\gamma+\delta} \frac{1}{2\delta} u du \\ &= \frac{1}{4\delta} u^2 \Big|_{\gamma-\delta}^{\gamma+\delta} = \frac{1}{4\delta} (4\gamma\delta) \\ &= \gamma \end{aligned}$$

$$\begin{aligned} \mu'_2 &= \mathbb{E} \left[\left(\frac{1}{2\delta} \right)^2 I_{[\gamma-\delta, \gamma+\delta]}(u) \right] \\ &= \int_{\gamma-\delta}^{\gamma+\delta} \frac{1}{4\delta^2} u du \\ &= \frac{1}{8\delta^2} (4\gamma\delta) \\ &= \frac{\gamma}{2\delta} \end{aligned}$$

$$\gamma = m_1 = 0.517$$

$$\delta = \frac{\gamma}{2m_2} = \frac{0.5166}{2 \cdot 1.3113} = 0.197$$