

The t - and F-distributions

Case when σ^2 is not known

An outline proof of this results is as follows:

1. We first prove that S^2 is an unbiased estimator for σ^2 .
(Theorem 4) $E[S^2] = \sigma^2$
2. We then derive the distribution of S^2 under the assumption that $Y_i \sim N(\mu, \sigma^2)$.
(Lemma 4, Definition 2.5, Theorem 5) $\frac{(n-1)S^2}{\sigma^2} \sim \chi^2_{n-1}$
 $S^2 \perp \bar{Y}$
3. We find the distribution of $\frac{\bar{Y} - \mu}{S/\sqrt{n}} \sim t_{n-1}$
(Definition 2.6, Theorem 6)

Definition 2.6

Suppose $Z \sim N(0,1)$ and $X \sim \chi_n^2$ independently, and let

$$T = \frac{Z}{\sqrt{\frac{X}{n}}},$$

then T is said to have a t -distribution with n degrees of freedom and we write

$$X \sim (t_n).$$

$$f(x) = \frac{\Gamma(\frac{n+1}{2})}{\sqrt{n\pi} \Gamma(\frac{n}{2})} \left(1 + \frac{x^2}{n}\right)^{-\frac{(n+1)}{2}} \quad \text{for } -\infty < x < \infty$$

$$E(X) = 0 \quad \text{if } n > 1$$

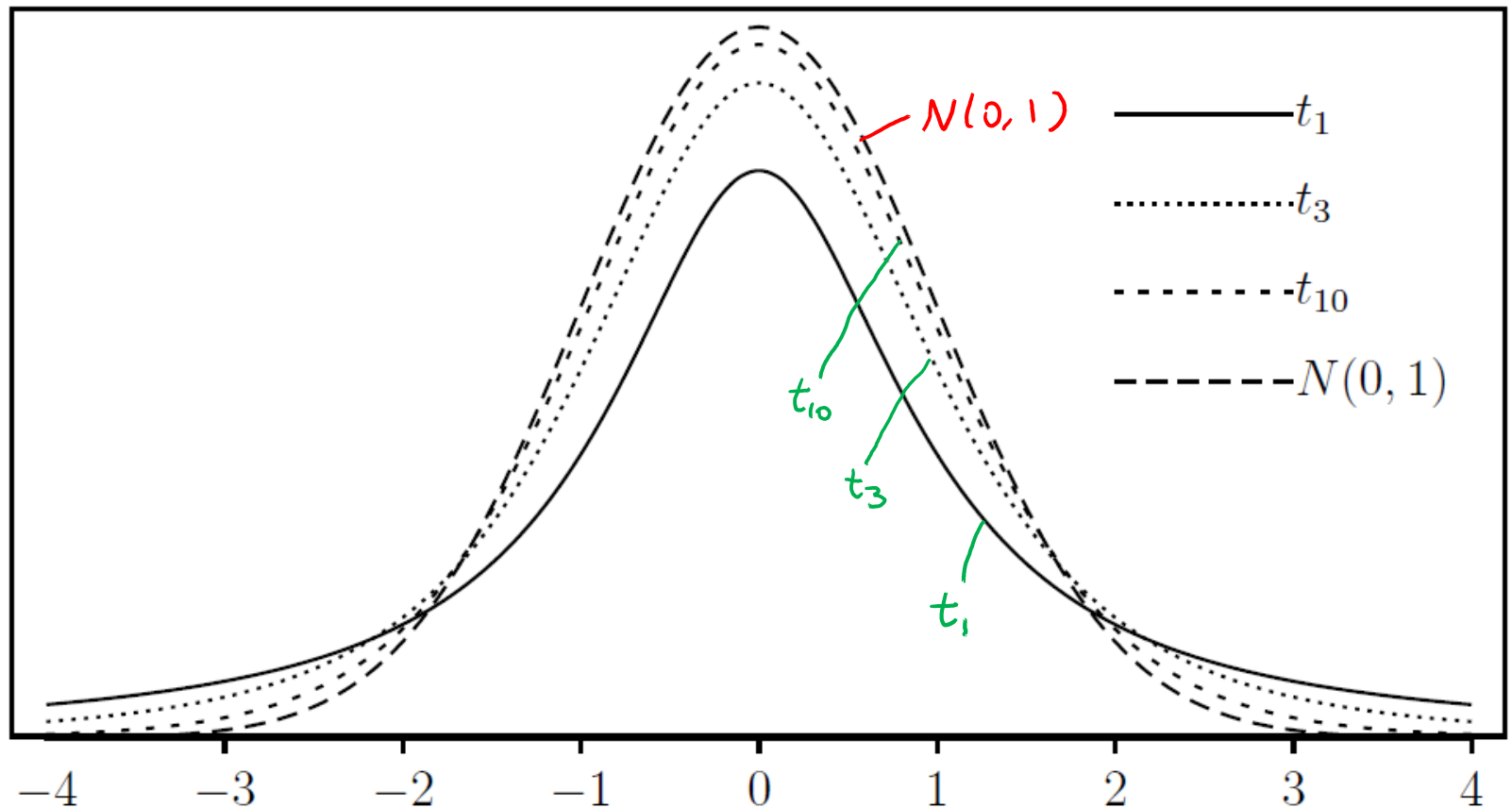
$$\text{var}(X) = \frac{n}{n-2} \quad \text{if } n > 2$$

MGF does not exist

The t -distribution

Symmetric
flatter tails than normal
(heavy tails)

t distributions



As $n \rightarrow \infty$, $t_n \rightarrow N$

Theorem 6

Suppose Y_1, Y_2, \dots, Y_n are i.i.d. $N(\mu, \sigma^2)$ random variables.

Then

$$T = \frac{\bar{Y} - \mu}{S/\sqrt{n}} \sim t_{n-1}.$$

We will show that T can be written as
a ratio of a $N(0,1)$ random variable
and a $\frac{\chi^2_{n-1}}{n-1}$ random variable.

Proof of Theorem 6

$$T = \frac{\frac{\bar{Y} - \mu}{\frac{s}{\sqrt{n}}}}{\frac{s/\sqrt{n}}{s/\sqrt{n}}} \sim N(0,1)$$

$$= \frac{\bar{Y} - \mu}{\frac{s}{\sqrt{n}}} \cdot \frac{s/\sqrt{n}}{s/\sqrt{n}}$$

$$= \frac{\bar{Y} - \mu}{\frac{s}{\sqrt{n}}} \cdot \frac{s}{s}$$

$$= \frac{\bar{Y} - \mu}{\frac{s}{\sqrt{n}}} / \sqrt{\frac{s^2}{\sigma^2}}$$

$$= \frac{\bar{Y} - \mu}{\frac{s}{\sqrt{n}}} / \sqrt{\frac{s^2 (n-1)}{\sigma^2 (n-1)}} \quad \begin{matrix} \nearrow N(0,1) & \nearrow \chi^2_{n-1} \end{matrix}$$

$$= Z / \sqrt{\frac{\chi^2}{n-1}}$$

$$\sim t_{n-1}$$

$$\text{Let } Z = \frac{\bar{Y} - \mu}{\frac{s}{\sqrt{n}}} \sim N(0,1)$$

$$\chi = \frac{s^2 (n-1)}{\sigma^2} \sim \chi^2_{n-1}$$

Definition 2.7

Let W_1 and W_2 be independent χ^2 -distributed random variables with ν_1 and ν_2 degrees of freedom respectively, then

$$F = \frac{W_1/\nu_1}{W_2/\nu_2} \sim \bar{F}_{\nu_1, \nu_2}$$

is said to have an F-distribution with ν_1 numerator degrees of freedom and ν_2 denominator degrees of freedom.

$$f(x) = \frac{\Gamma(\frac{\nu_1 + \nu_2}{2})}{\Gamma(\frac{\nu_1}{2}) \Gamma(\frac{\nu_2}{2})} \frac{\nu_1^{\frac{\nu_1}{2}} \nu_2^{\frac{\nu_2}{2}} x^{\frac{\nu_1}{2} - 1}}{(\nu_1 x + \nu_2)^{(\frac{\nu_1 + \nu_2}{2})}}, \quad x > 0$$

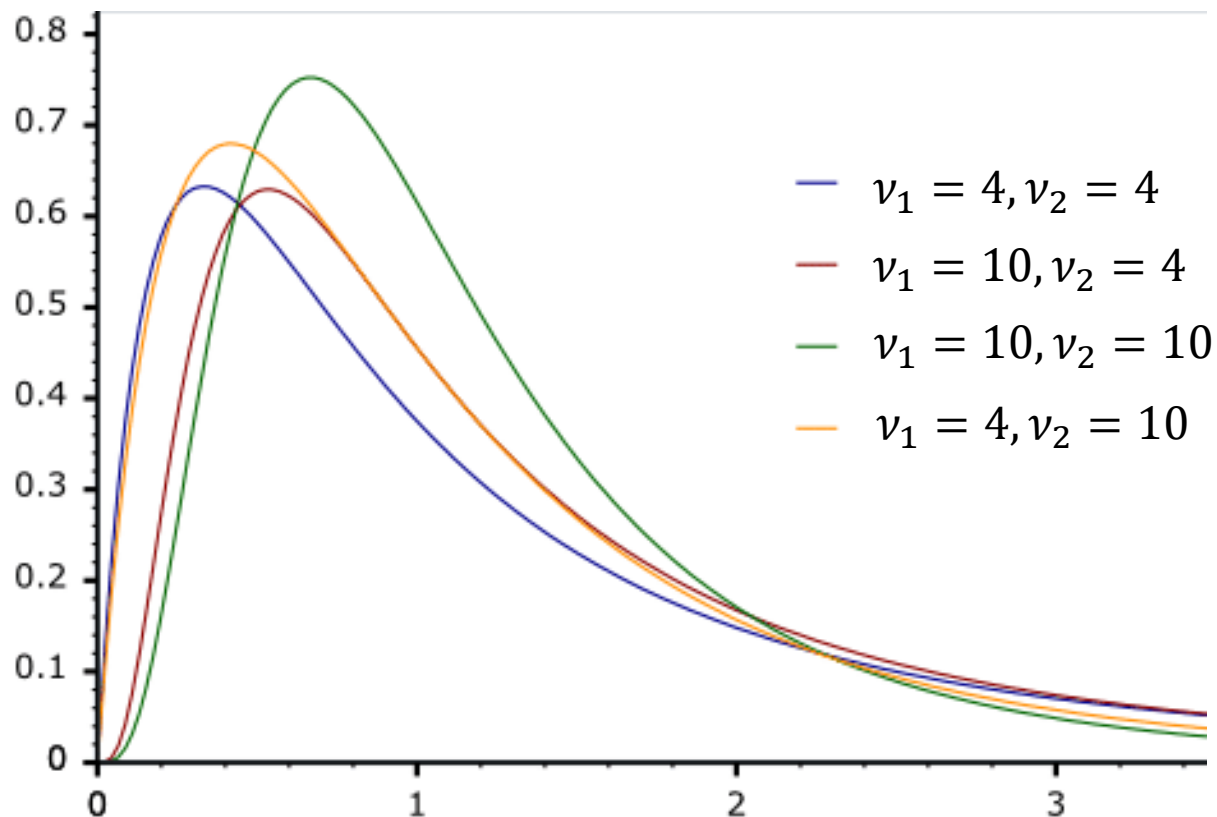
$$E(X) = \frac{\nu_2}{\nu_2 - 2} \quad \text{if } \nu_2 > 2$$

If $X \sim t_n$, then $X^2 \sim F_{1, n}$.

$$\text{var}(X) = \frac{2 \nu_2^2 (\nu_1 + \nu_2 - 2)}{\nu_1 (\nu_2 - 2)^2 (\nu_2 - 4)} \quad \text{if } \nu_2 > 4$$

MGF does not exist

The F-distribution



skewed to the right

Example 2.4

Consider two independent random samples, with sample size n_1 and n_2 , from normal distributions with variance σ_1^2 and σ_2^2 respectively. Let S_1^2 and S_2^2 be the sample variance of the corresponding samples. Show that

$$\frac{S_1^2/\sigma_1^2}{S_2^2/\sigma_2^2}$$

has an F-distribution.

$$\frac{\frac{S_1^2}{\sigma_1^2}}{\frac{S_2^2}{\sigma_2^2}} = \frac{\boxed{\frac{S_1^2}{\sigma_1^2} (n_1-1)}}{\boxed{\frac{S_2^2}{\sigma_2^2} (n_2-1)}} = \frac{\frac{W_1}{n_1-1}}{\frac{W_2}{n_2-1}} \sim F_{n_1-1, n_2-1}$$

$\nearrow \chi_{n_1-1}^2$
 $\searrow \chi_{n_2-1}^2$

$$\text{where } W_1 = \frac{(n_1-1)S_1^2}{\sigma_1^2} \sim \chi_{n_1-1}^2 \text{ and}$$

$$W_2 = \frac{(n_2-1)S_2^2}{\sigma_2^2} \sim \chi_{n_2-1}^2 \text{ are independent}$$

Example 2.5

Let S_1^2 denote the sample variance of a random sample of size 10 from Population I and let S_2^2 denote the sample variance for a random sample of size 8 from population II. The variance of Population I is three times the variance of Population II. Find the two numbers a and b such that $P(a \leq S_1^2/S_2^2 \leq b) = 0.90$, assuming S_1^2 is independent of S_2^2 .

$$F = \frac{\frac{S_1^2}{\sigma_1^2}}{\frac{S_2^2}{\sigma_2^2}} = \frac{\frac{S_1^2}{3\cancel{\sigma_2^2}}}{\frac{S_2^2}{\cancel{\sigma_2^2}}} = \frac{S_1^2}{3S_2^2} \sim F_{n_1-1, n_2-1} = F_{9,7}$$

$n_1 = 10, n_2 = 8, \sigma_1^2 = 3\sigma_2^2$



$$P\left(\frac{S_1^2}{S_2^2} < a\right) = 0.05$$

$$P\left(\frac{S_1^2}{S_2^2} > b\right) = 0.05$$

Example 2.5

$$0.05 = P\left(\frac{S_1^2}{S_2^2} < a\right)$$

$$= P\left(\frac{S_1^2}{3S_2^2} < \frac{a}{3}\right)$$

$$= P\left(F < \frac{a}{3}\right) \text{ where } F \sim F_{9,7}$$

Use R `qf(0.05, 9, 7)`

$$0.05 = P(F < 0.304)$$

$$\Rightarrow \frac{a}{3} = 0.304$$

$$a = 3(0.304)$$

$$= 0.912$$

$$0.05 = P\left(\frac{S_1^2}{S_2^2} > b\right)$$

$$= 1 - P\left(\frac{S_1^2}{S_2^2} \leq b\right)$$

$$= 1 - P\left(\frac{S_1^2}{3S_2^2} \leq \frac{b}{3}\right)$$

$$= 1 - P\left(F \leq \frac{b}{3}\right)$$

$$0.95 = P\left(F \leq \frac{b}{3}\right)$$

Using R `qf(0.95, 9, 7)`

$$0.95 = P(F \leq 3.677)$$

$$\Rightarrow \frac{b}{3} = 3.677$$

$$b = 11.03$$

$$\therefore P\left(0.912 \leq \frac{S_1^2}{S_2^2} \leq 11.03\right) = 0.90$$