

STATS 2107  
Statistical Modelling and Inference II  
Tutorial 4

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Semester 2 2022

1. (a) Consider regression data

$$(x_1, y_1), (x_2, y_2), \dots, (x_n, y_n)$$

and let

$$S_{xx} = \sum_{i=1}^n (x_i - \bar{x})^2, S_{xy} = \sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y}), S_{yy} = \sum_{i=1}^n (y_i - \bar{y})^2.$$

Prove that

$$S_{xy} = \sum_{i=1}^n x_i y_i - n \bar{x} \bar{y}.$$

- (b) Consider independent random variables  $Y_1, Y_2, \dots, Y_n$  with

$$E(Y_i) = \beta_0 + \beta_1 x_i \quad \text{and} \quad \text{Var}(Y_i) = \sigma^2.$$

Let

$$\hat{E}_i = Y_i - (\hat{\beta}_0 + \hat{\beta}_1 x_i).$$

Prove that

$$E \left[ \sum_{i=1}^n \{Y_i - (\hat{\beta}_0 + \hat{\beta}_1 x_i)\}^2 \right] = (n-2)\sigma^2.$$

Hence deduce that  $S_e^2$  is an unbiased estimator for  $\sigma^2$ .

2. Suppose  $X$  is an  $n \times p$  matrix with linearly independent columns and let

$$\mathbf{H} = \mathbf{X}(\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top.$$

In lectures, we have stated that the matrices  $\mathbf{H}$  and  $(\mathbf{I} - \mathbf{H})$  are symmetric and idempotent. We will prove these properties here.

- (a) Show that  $\mathbf{H}$  is symmetric, *i.e.*,  $\mathbf{H} = \mathbf{H}^\top$ .
- (b) Show that  $\mathbf{H}$  is idempotent, that is,  $\mathbf{H}^2 = \mathbf{H}$ .
- (c) Show that  $(\mathbf{I} - \mathbf{H})$  is symmetric and idempotent, where  $\mathbf{I}$  is the  $n \times n$  identity matrix.

3. Consider the multiple regression model

$$\mathbf{M} : \mathbf{Y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon},$$

where  $Y_i \sim N(\eta_i, \sigma^2)$  independently for  $i = 1, 2, \dots, n$  and  $E[\mathbf{Y}] = \boldsymbol{\eta} = \mathbf{X}\boldsymbol{\beta}$ . The vector of residuals is defined by  $\hat{\mathbf{E}} = \mathbf{Y} - \mathbf{X}\hat{\boldsymbol{\beta}}$ . In lectures, we have stated some properties of  $\hat{\mathbf{E}}$ . We will look at these here.

- (a) Prove that  $\hat{\mathbf{E}} = (\mathbf{I} - \mathbf{H})\mathbf{Y}$ , where  $\mathbf{H} = \mathbf{X}(\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top$ .

- (b) Prove that  $E(\hat{\mathbf{E}}) = \mathbf{0}$  (and hence  $E(\hat{E}_i) = 0$  for  $i = 1, 2, \dots, n$ ).
- (c) Prove that:  $\text{Var}(\hat{E}_i) = \sigma^2(1 - h_{ii})$ , where  $h_{ii}$  is the  $(i, i)$ th element of  $H$ . **Hint: Let  $(I - H)$  have rows  $\mathbf{a}_1^T, \mathbf{a}_2^T, \dots, \mathbf{a}_n^T$ .**
4. Linearise the following equations:
- (a)  $Y = \alpha\beta^x$
  - (b)  $Y = \alpha e^{\frac{\beta}{x}}$
  - (c)  $Y = \alpha + \frac{\beta}{x}$
  - (d)  $Y = \frac{\alpha}{\beta + x}$
  - (e)  $Y = \alpha + \beta x^n$
  - (f)  $Y = \frac{1}{\alpha + \beta e^{-x}}$
  - (g)  $Y = e^{-\alpha x_1} e^{-\frac{\beta}{x_2}}$

*The following question is optional:*

5. In the proof of Theorem 11, we have used the result that if  $\mathbf{X}$  is a random variable with  $E(\mathbf{X}) = \boldsymbol{\mu}$  and  $\text{Var}(\mathbf{X}) = \boldsymbol{\Sigma}$ , then

$$E(\mathbf{X}^\top \mathbf{A} \mathbf{X}) = \text{tr}(\mathbf{A} \boldsymbol{\Sigma}) + \boldsymbol{\mu}^\top \mathbf{A} \boldsymbol{\mu}.$$

Prove this result.