

# Sample Questions for Module 2

- Module 2 is about sampling distributions.
- Main aim is to be familiar with the  $t$ ,  $\chi^2$ , and  $F$ -distributions, and be able to derive CI based on a pivotal quantity.
- Important results: sampling distribution of the sample mean and the sample variance (can use these without proof in the exam)

# Sample Question 4

Consider two independent samples of size  $n_1$  and  $n_2$ :

$$Y_{i1} \sim i.i.d. N(\mu_1, \sigma_1^2), \quad i = 1, 2, \dots, n_1$$

$$Y_{j2} \sim i.i.d. N(\mu_2, \sigma_2^2), \quad j = 1, 2, \dots, n_2$$

and writing  $\sigma_2^2 = k\sigma_1^2$ , where  $k > 0$  is a constant.

Let  $\bar{Y}_1$  and  $\bar{Y}_2$  be their samples means and  $S_1^2$  and  $S_2^2$  be their sample variances.

a) Show that  $Z = \frac{(\bar{Y}_1 - \bar{Y}_2) - (\mu_1 - \mu_2)}{\sigma_1 \sqrt{\frac{1}{n_1} + \frac{k}{n_2}}}$  has a standard normal distribution.

b) Show that  $W = \frac{(n_1 - 1)S_1^2 + (n_2 - 1)S_2^2/k}{\sigma_1^2}$  has a  $\chi^2$ -distribution.

c) Show that  $T = \frac{(\bar{Y}_1 - \bar{Y}_2) - (\mu_1 - \mu_2)}{s_p \sqrt{\frac{1}{n_1} + \frac{k}{n_2}}}$  has a  $t$ -distribution,

$$\text{where } s_p^2 = \frac{\sigma_1^2}{n_1 + n_2 - 2} W.$$

d) Using the results of part c), construct a  $100(1 - \alpha)\%$  confidence interval for  $\mu_1 - \mu_2$ .

# Sample Question 4 Solution

(a)  $\bar{Y}_1 \sim N(\mu_1, \frac{\sigma_1^2}{n_1})$  independent of  $\bar{Y}_2 \sim N(\mu_2, \frac{\sigma_2^2}{n_2})$ .

By the property of linear combination of independent normal random variables,

$$\bar{Y}_1 - \bar{Y}_2 \sim N(\mu_1 - \mu_2, \frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}).$$

$$Z = \frac{(\bar{Y}_1 - \bar{Y}_2) - E(\bar{Y}_1 - \bar{Y}_2)}{\sqrt{\text{var}(\bar{Y}_1 - \bar{Y}_2)}} \sim N(0, 1)$$

$$= \frac{(\bar{Y}_1 - \bar{Y}_2) - (\mu_1 - \mu_2)}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}}$$

$$= \frac{(\bar{Y}_1 - \bar{Y}_2) - (\mu_1 - \mu_2)}{\sigma_1 \sqrt{\frac{1}{n_1} + \frac{k}{n_2}}}$$

(Write  $\sigma_2^2 = k\sigma_1^2$ )

$$\therefore Z \sim N(0, 1).$$

# Sample Question 4 Solution

(b) We know  $U_1 = \frac{(n_1-1)S_1^2}{\sigma_1^2} \sim \chi^2_{n_1-1}$  independent of

$$U_2 = \frac{(n_2-1)S_2^2}{\sigma_2^2} \sim \chi^2_{n_2-1}.$$

$$W = U_1 + U_2 \sim \chi^2_{n_1+n_2-2}$$

$$= \frac{(n_1-1)S_1^2}{\sigma_1^2} + \frac{(n_2-1)S_2^2}{\sigma_2^2}$$

$$= \frac{(n_1-1)S_1^2}{\sigma_1^2} + \frac{(n_2-1)S_2^2}{k\sigma_1^2}$$

$$= \frac{(n_1-1)S_1^2 + (n_2-1)S_2^2\left(\frac{1}{k}\right)}{\sigma_1^2} \sim \chi^2_{n_1+n_2-2}$$

# Sample Question 4 Solution

(c) From parts a and c, we have  $Z \sim N(0,1)$  independent of  $W \sim \chi^2_{n_1+n_2-2}$ .

By the definition of the t-distribution, we have

$$T = \frac{Z}{\sqrt{\frac{W}{n_1+n_2-2}}} \sim t_{n_1+n_2-2}$$

$$= \frac{\bar{Y}_1 - \bar{Y}_2 - (\mu_1 - \mu_2)}{\sigma_1 \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}} \cdot \frac{1}{\sqrt{\frac{W}{n_1+n_2-2}}}$$

$$= \frac{\bar{Y}_1 - \bar{Y}_2 - (\mu_1 - \mu_2)}{\underbrace{\sqrt{\sigma_1^2 \left(\frac{1}{n_1} + \frac{1}{n_2}\right)} \left(\frac{W}{n_1+n_2-2}\right)}_{S_p^2}}$$

$$\begin{aligned} &= \frac{\bar{Y}_1 - \bar{Y}_2 - (\mu_1 - \mu_2)}{\sqrt{S_p^2 \left(\frac{1}{n_1} + \frac{1}{n_2}\right)}} \\ &= \frac{\bar{Y}_1 - \bar{Y}_2 - (\mu_1 - \mu_2)}{S_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}} \\ &\sim t_{n_1+n_2-2} \end{aligned}$$

# Sample Question 4 Solution

$$\begin{aligned} d) \quad 1 - \alpha &= P\left(-t_{n_1+n_2-2, \frac{\alpha}{2}} < T < t_{n_1+n_2-2, \frac{\alpha}{2}}\right) \\ &= P\left(-t_{n_1+n_2-2, \frac{\alpha}{2}} < \frac{(\bar{Y}_1 - \bar{Y}_2) - (\mu_1 - \mu_2)}{S_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}} < t_{n_1+n_2-2, \frac{\alpha}{2}}\right) \\ &= P\left((\bar{Y}_1 - \bar{Y}_2) - t_{n_1+n_2-2, \frac{\alpha}{2}} S_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}} < \mu_1 - \mu_2 < \right. \\ &\quad \left. (\bar{Y}_1 - \bar{Y}_2) + t_{n_1+n_2-2, \frac{\alpha}{2}} S_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}\right). \end{aligned}$$

$\therefore$  The required CI for  $\mu_1 - \mu_2$  is

$$(\bar{Y}_1 - \bar{Y}_2) \pm t_{n_1+n_2-2, \frac{\alpha}{2}} S_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}$$

# Sample Question 5

Suppose  $Y_1, Y_2, \dots, Y_n$  is a sample of size  $n$  from a  $\text{Beta}(\theta, 1)$ . Consider the quantity  $X = -2\theta \sum_{i=1}^n \log Y_i$ .

- a) Find the distribution of  $X$  and show that  $X$  is a pivotal quantity.
- b) Use the above pivotal quantity to construct a symmetric  $100(1 - \alpha)\%$  confidence interval for  $\theta$ .
- c) Suppose we have observed the following values for  $Y$ :  
0.72, 0.48, 0.80, 0.45, 0.67  
Using the results of part b), give a 95% confidence interval for  $\theta$ .

# Sample Question 5 Solution

$$Y_1, Y_2, \dots, Y_n \sim \text{iid Beta}(\theta, 1)$$

$$f_{Y_i}(y) = \theta y^{\theta-1} \quad \text{for } 0 < y < 1$$

$$X = -2\theta \sum_{i=1}^n \log Y_i$$

(a) Let  $X_i = -2\theta \log Y_i$ , then  $Y_i = e^{-\frac{X_i}{2\theta}}$  (inverse transformation)

$$\frac{dY_i}{dX_i} = -\frac{1}{2\theta} e^{-\frac{X_i}{2\theta}}$$

$$f_{X_i}(x) = \cancel{\theta} \left( \cancel{e^{-\frac{x}{2\theta}}} \right)^{\cancel{\theta-1}} \left| -\frac{1}{\cancel{2\theta}} \cancel{e^{-\frac{x}{2\theta}}} \right| \quad (\text{transformation formula})$$

$$= \frac{1}{2} e^{-\frac{x}{2}} \quad \text{for } x > 0$$

$$\therefore X_i \sim \text{iid Exp}\left(\frac{1}{2}\right)$$

$$\therefore X = \sum_{i=1}^n X_i \sim \text{Gamma}\left(n, \frac{1}{2}\right) \quad \text{or } \chi^2_{2n}.$$



# Sample Question 5 Solution

Since  $X$  is a function  $\theta$  and the sample  $Y_1, Y_2, \dots, Y_n$ ,  $\theta$  is the only unknown parameter in  $X$ , and the distribution of  $X$  does not depend on  $\theta$ , so  $X$  is a pivotal quantity of  $\theta$ .

$$\begin{aligned} \text{(b)} \quad 1 - \alpha &= P\left(\chi_{2n, 1-\frac{\alpha}{2}}^2 < X < \chi_{2n, \frac{\alpha}{2}}^2\right) \\ &= P\left(\chi_{2n, 1-\frac{\alpha}{2}}^2 < -2\theta \underbrace{\sum_{i=1}^n \log Y_i}_{\text{less than 0}} < \chi_{2n, \frac{\alpha}{2}}^2\right) \\ &= P\left(\frac{\chi_{2n, 1-\frac{\alpha}{2}}^2}{-2 \sum_{i=1}^n \log Y_i} < \theta < \frac{\chi_{2n, \frac{\alpha}{2}}^2}{-2 \sum_{i=1}^n \log Y_i}\right) \end{aligned}$$

$$\therefore CI = \left( \frac{\chi_{2n, 1-\frac{\alpha}{2}}^2}{-2 \sum_{i=1}^n \log Y_i}, \frac{\chi_{2n, \frac{\alpha}{2}}^2}{-2 \sum_{i=1}^n \log Y_i} \right)$$

# Sample Question 5 Solution

$$\begin{aligned} c) \quad -2 \sum_{i=1}^n \log \gamma_i &= -2 (\log 0.72 + \log 0.48 + \log 0.80 + \log 0.45 + \log 0.67) \\ &= 4.9692 \end{aligned}$$

$$n = 5,$$

$$\chi^2_{10, 0.025} = 20.4832, \quad \text{qchisq}(0.975, 10)$$

$$\chi^2_{10, 0.975} = 3.2470, \quad \text{qchisq}(0.025, 10)$$

$$CI = \left( \frac{3.2470}{4.9692}, \frac{20.4832}{4.9692} \right)$$

$$\approx (0.653, 4.122)$$