

Power

Power function

Suppose that H_0 is false in the Example 1.10, i.e. $\mu \neq \mu_0$.

Let $\mu_1 =$ true value of μ ($\mu_1 \neq \mu_0$)

Power = $1 - \beta = P(\text{reject } H_0 \mid H_0 \text{ false}) = P(|Z^*| \geq Z_{\frac{\alpha}{2}} \mid \mu \neq \mu_0)$

$$Z^* = \frac{\bar{Y} - \mu_0}{\frac{\sigma}{\sqrt{n}}} = \frac{\bar{Y} - \mu_1 + \mu_1 - \mu_0}{\frac{\sigma}{\sqrt{n}}} = \boxed{\frac{\bar{Y} - \mu_1}{\frac{\sigma}{\sqrt{n}}}} + \boxed{\frac{\mu_1 - \mu_0}{\frac{\sigma}{\sqrt{n}}}}$$

$\sim N(0,1)$ by Lemma 3 constant

$$\begin{aligned} E[Z^*] &= E\left[\frac{\bar{Y} - \mu_1}{\frac{\sigma}{\sqrt{n}}} + \frac{\mu_1 - \mu_0}{\frac{\sigma}{\sqrt{n}}}\right] \\ &= E\left[\frac{\bar{Y} - \mu_1}{\frac{\sigma}{\sqrt{n}}}\right] + \frac{\mu_1 - \mu_0}{\frac{\sigma}{\sqrt{n}}} \\ &= 0 + \frac{\mu_1 - \mu_0}{\frac{\sigma}{\sqrt{n}}} \\ &= \frac{\mu_1 - \mu_0}{\frac{\sigma}{\sqrt{n}}} \end{aligned}$$

$$\therefore Z^* \sim N\left(\frac{\mu_1 - \mu_0}{\frac{\sigma}{\sqrt{n}}}, 1\right)$$

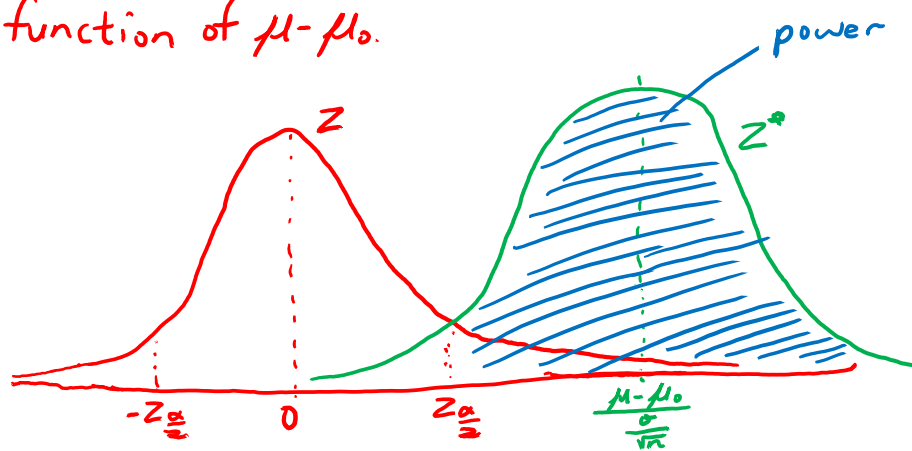
$$\text{var}(Z^*) = \text{var}\left(\frac{\bar{Y} - \mu_1}{\frac{\sigma}{\sqrt{n}}}\right) = 1$$

Power function (cont.)

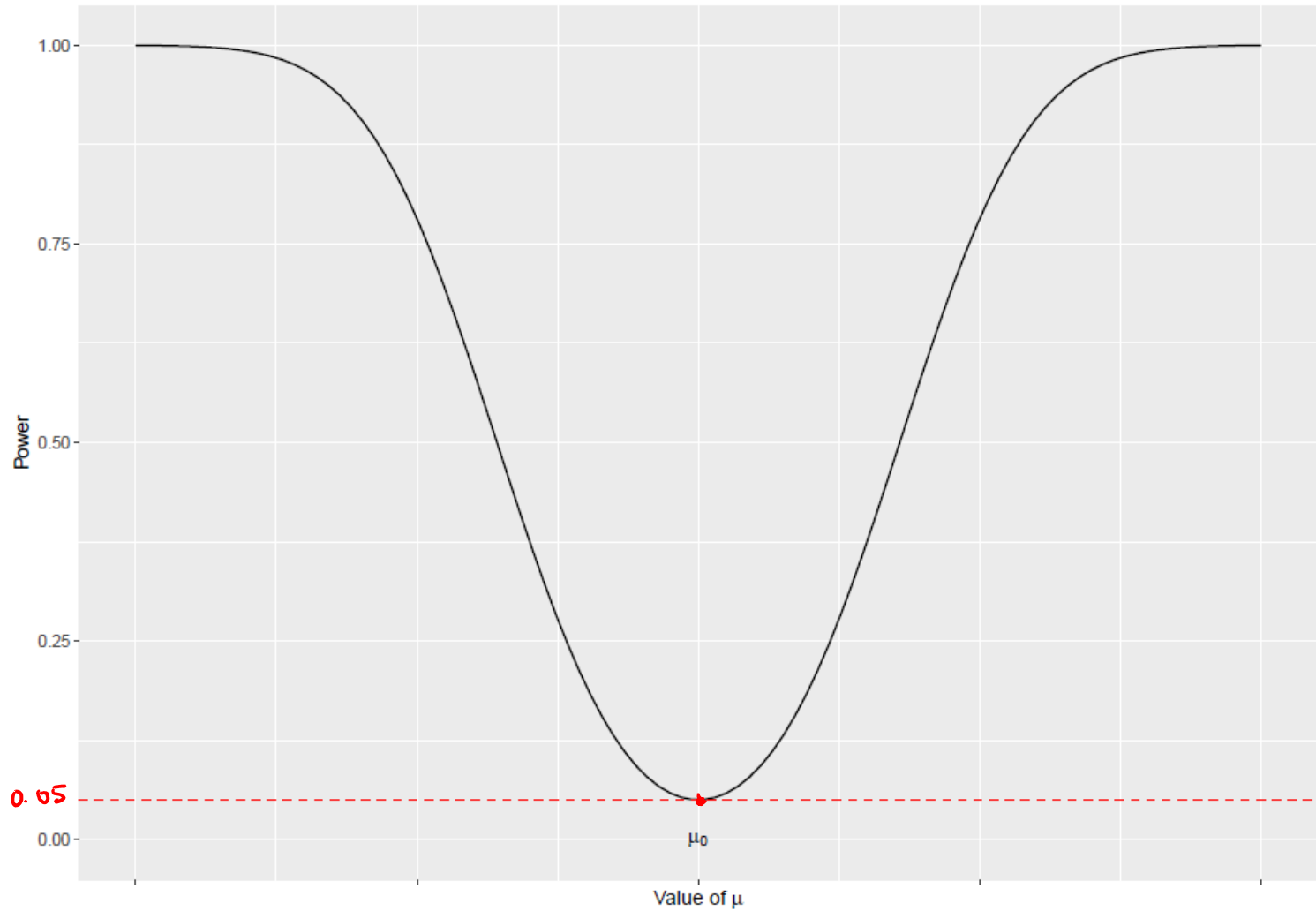
$$\begin{aligned}
 \text{Power} &= P(|Z^*| \geq Z_{\frac{\alpha}{2}} | \mu \neq \mu_0) \\
 &= P(Z^* \leq -Z_{\frac{\alpha}{2}} | \mu \neq \mu_0) + P(Z^* \geq Z_{\frac{\alpha}{2}} | \mu \neq \mu_0) \\
 &= \Phi\left(-Z_{\frac{\alpha}{2}}; \frac{\mu - \mu_0}{\frac{\sigma}{\sqrt{n}}}, 1\right) + 1 - P(Z^* \leq Z_{\frac{\alpha}{2}} | \mu \neq \mu_0) \\
 &= \Phi\left(-Z_{\frac{\alpha}{2}}; \frac{\mu - \mu_0}{\frac{\sigma}{\sqrt{n}}}, 1\right) + 1 - \Phi\left(Z_{\frac{\alpha}{2}}; \frac{\mu - \mu_0}{\frac{\sigma}{\sqrt{n}}}, 1\right)
 \end{aligned}$$

$\Phi(x; \mu, \sigma^2)$ is the traditional notation for the cdf of $N(\mu, \sigma^2)$.
 That is, $\Phi(x; \mu, \sigma^2) = P(X \leq x) = \int_{-\infty}^x \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2\sigma^2}(t-\mu)^2} dt$,
 where $X \sim N(\mu, \sigma^2)$.

In this case, the power is a function of $\mu - \mu_0$.



Power function (cont.)



One-sided tests

Consider y_1, y_2, \dots, y_n are i.i.d. $N(\mu, \sigma^2)$ observations with σ^2 known.

We can test the one-sided hypothesis

$$H_0: \mu \leq \mu_0.$$

$$H_a: \mu > \mu_0.$$

The test statistic is

$$Z = \frac{\bar{Y} - \mu_0}{\sigma/\sqrt{n}}$$

The rule is

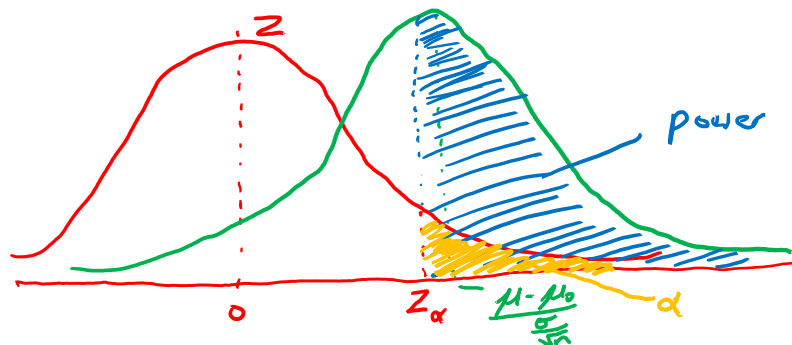
$$\text{Reject } H_0 \text{ if } \underline{z \geq z_\alpha}.$$

One-sided power function

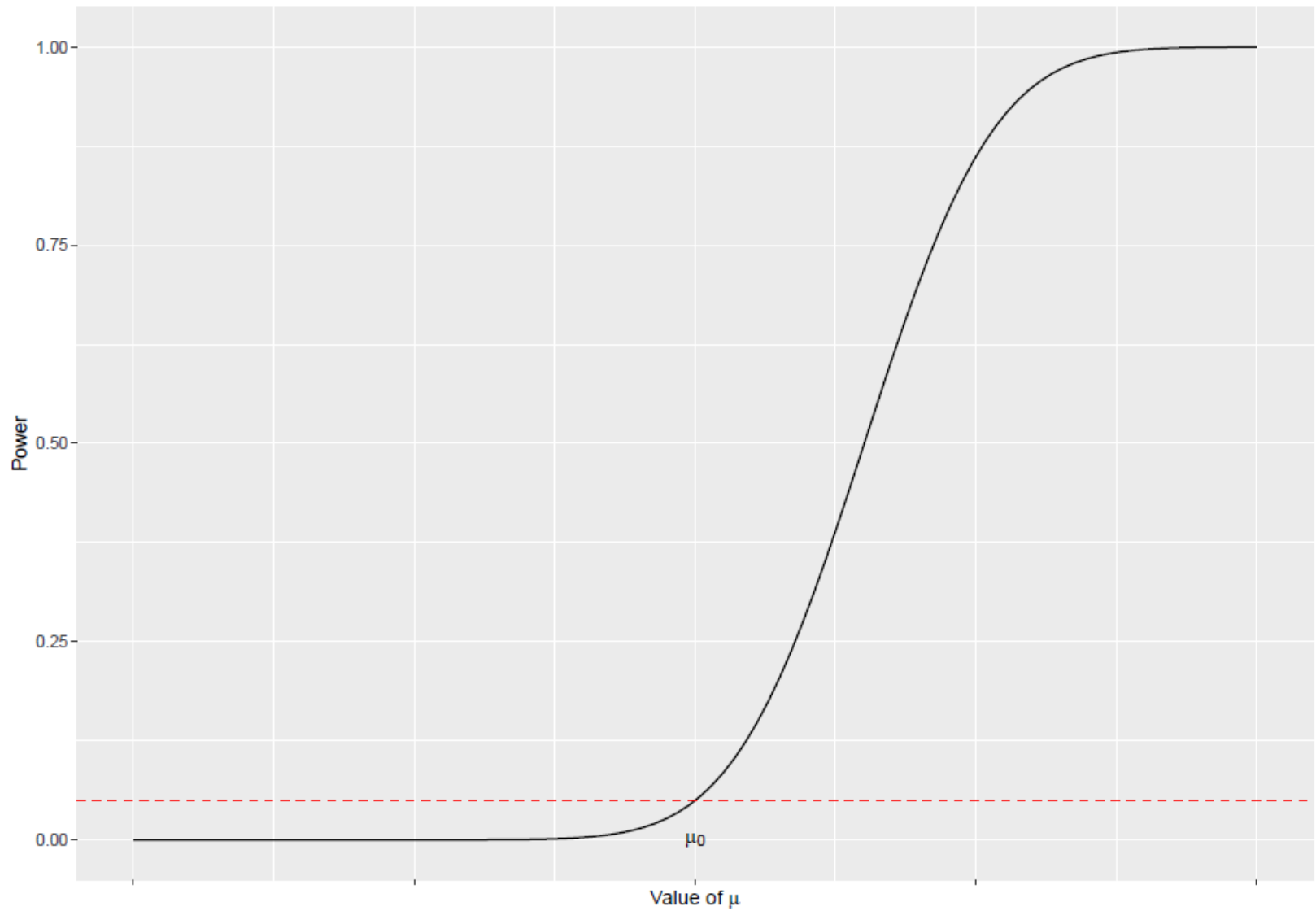
Following the same argument as the two-sided test, we have:

If H_0 is false (i.e. $\mu > \mu_0$), then $Z \sim N\left(\frac{\mu - \mu_0}{\sigma/\sqrt{n}}, 1\right)$.

$$\begin{aligned}\text{Power} &= \mathcal{P}(\text{reject } H_0 \mid H_0 \text{ false}) \\ &= \mathcal{P}(Z^* \geq z_\alpha \mid \mu \neq \mu_0) \\ &= 1 - \mathcal{P}(Z^* < z_\alpha \mid \mu \neq \mu_0) \\ &= 1 - \Phi\left(z_\alpha; \frac{\mu - \mu_0}{\sigma/\sqrt{n}}, 1\right)\end{aligned}$$



One-sided power function



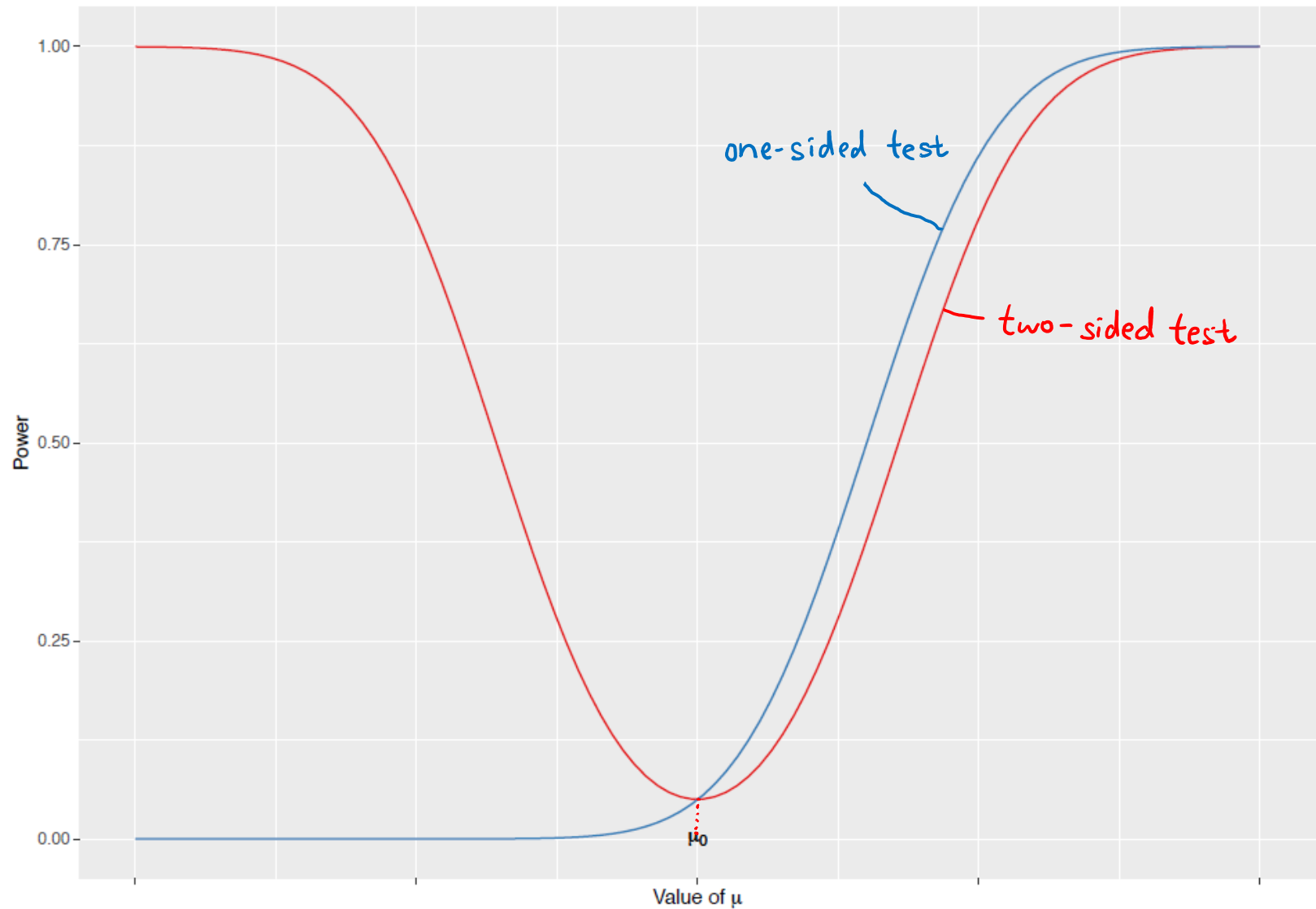
One-sided vs two-sided

For one-sided test, the type I error probability is not a single number, i.e.

$$P(\text{reject } H_0 | H_0 \text{ true}) \leq \alpha$$

$$P(Z^* \geq Z_\alpha | H_0 \text{ true}) = \begin{cases} \alpha & \text{if } \mu = \mu_0 \\ < \alpha & \text{if } \mu \neq \mu_0 \end{cases}$$

One-sided vs two-sided



One-sided vs two-sided

One-sided test are often subject to an abuse:

Suppose we begin with a problem where a two-sided test is appropriate, but then observed a test statistic such that

$$z_{\alpha} \leq z_{obs} \leq z_{\alpha/2}$$

One-sided vs two-sided simulation

If we use a one-sided test when in fact a two-sided test should be used, the type I error probability will double.

Let's do a simulation to see this:

1. Generate 100,000 replications of $n = 100$ observations from $N(5, 4)$.
2. Perform one-sided and two-sided tests with $\mu_0 = 5$.
3. Compare the proportion of times where we commit a type I error (i.e. rejecting H_0 in this case).

$$Y_i \sim \text{i.i.d. } N(5, 4)$$

$$H_0: \mu = \mu_0 = 5$$

One-sided vs two-sided simulation

Function to perform the Z-test

```
perform_z_test <- function(sample, null, cheat = FALSE, known_var){  
  # Get sigma  
  sigma <- sqrt(known_var)  
  # Get number of obs  
  n <- length(sample)  
  # Calculate sample mean  
  sample_mean <- mean(sample)  
  # Calculate test statistic  
  Z <- (sample_mean - null) / (sigma / sqrt(n))  
  # Decide if to reject or retain  
  if(abs(Z) > qnorm(0.025, lower.tail = FALSE)){  
    return("reject")  
  }  
  if(cheat){ #(incorrectly)using the one-sided test  
    if(abs(Z) > qnorm(0.05, lower.tail = FALSE)){  
      return("reject")  
    }  
  }  
  return("fail to reject")  
}
```

One-sided vs two-sided simulation

Generate the samples

```
library(tidyverse)
sim_data <- 100000 %>%
  rerun(rnorm(100, 5, 2))
```

Run the proper test

```
sim_data %>%
  map_chr(perform_z_test, null = 5, known_var = 4) %>%
  table() %>%
  prop.table()
```

Run the wrong test

```
sim_data %>%
  map_chr(perform_z_test, null = 5, known_var = 4, cheat =
    TRUE) %>%
  table() %>%
  prop.table()
```