

**Examination in School of Mathematical Sciences**  
**Semester 2, 2018**

**104959 STATS 7107 Statistical Modelling and Inference PG**

Official Reading Time: 10 mins  
Writing Time: 180 mins  
Total Duration: 190 mins

**NUMBER OF QUESTIONS: 8      TOTAL MARKS: 100**

**Instructions**

- Attempt all questions.
- Begin each answer on a new page.
- Examination materials must not be removed from the examination room.

**Materials**

- 1 Blue book is provided.
- Formulae sheets are provided.
- Calculators without remote communications capability are allowed.
- English and foreign-language dictionaries may be used.

**DO NOT COMMENCE WRITING UNTIL INSTRUCTED TO DO SO.**

1. Let  $Y_1, Y_2, \dots, Y_n$  be independent and identically distributed (*i.i.d.*) random variables with probability density function  $f(y; \theta)$  for a real scalar parameter  $\theta \in \Theta$ , where  $\Theta$  denotes the parameter space.

Let  $T = T(Y_1, Y_2, \dots, Y_n)$  be an estimator for  $\theta$ .

- (a) Define the *mean squared error*,  $\text{MSE}_T(\theta)$ , of  $T$ . [1 marks]

- (b) Define the *bias*,  $b_T(\theta)$ , of  $T$ . [1 marks]

- (c) Prove that

$$\text{MSE}_T(\theta) = \text{Var}(T) + b_T(\theta)^2.$$

[3 marks]

- (d) Suppose  $Y_1, Y_2, \dots, Y_n$  are independent identically distributed (*i.i.d.*)  $N(\mu, \sigma^2)$  random variables and let

$$\bar{Y} = \frac{1}{n} \sum_{i=1}^n Y_i$$

be an estimator for  $\mu$ . Calculate  $\text{MSE}_{\bar{Y}}(\mu)$ .

[4 marks]

[Total: 9]

2.

- (a) Carefully define the  $t$ -distribution with  $k$  degrees of freedom. [3 marks]

- (b) Suppose that  $Y_1, Y_2, \dots, Y_n$  are *i.i.d.*  $N(\mu, \sigma^2)$ , then prove that

$$\frac{\bar{Y} - \mu}{S/\sqrt{n}} \sim t_{n-1}.$$

You may assume that  $\bar{Y}$  and  $S^2$  are independent.

[3 marks]

- (c) Let  $Z \sim N(0, 1)$ . Show that the moment generating function of  $Z^2$  is

$$M_{Z^2}(t) = (1 - 2t)^{-\frac{1}{2}}, \quad t < 1/2.$$

[4 marks]

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- (d) Suppose  $Z_1, Z_2, \dots, Z_k$  are independent and identically distributed  $N(0, 1)$  random variable and let

$$X = \sum_{i=1}^k Z_i^2.$$

Show that the moment generating function of  $X$  is

$$M_X(t) = (1 - 2t)^{-\frac{k}{2}}, \quad t < 1/2.$$

[3 marks]

- (e) Hence, or otherwise, show that if

$$X \sim \chi_k^2,$$

then

$$E[X] = k \text{ and } \text{Var}(X) = 2k.$$

[5 marks]

[Total: 18]

3. Consider the simple linear regression model

$$Y_i = \beta_0 + \beta_1 x_i + \epsilon_i,$$

where  $\epsilon_1, \epsilon_2, \dots, \epsilon_n$  are independent with  $\epsilon_i \sim N(0, \sigma^2)$  for  $i = 1, 2, \dots, n$ .

- (a) Consider

$$S_{xy} = \sum_{i=1}^n (y_i - \bar{y})(x_i - \bar{x})$$

Show that  $S_{xy}$  can be written as

$$S_{xy} = \sum_{i=1}^n (x_i - \bar{x})y_i.$$

[3 marks]

- (b) Given that

$$\hat{\beta}_1 = \frac{S_{xy}}{S_{xx}},$$

find the constants  $a_1, a_2, \dots, a_n$ , such that

$$\hat{\beta}_1 = \sum_{i=1}^n a_i Y_i.$$

[2 marks]

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(c) Prove that

$$E[\hat{\beta}_1] = \beta_1 \text{ and } \text{Var}(\hat{\beta}_1) = \frac{\sigma^2}{S_{xx}},$$

where

$$S_{xx} = \sum_{i=1}^n (x_i - \bar{x})x_i.$$

[6 marks]

[Total: 11]

4. An analysis of the effect of displacement (`displ`) and drive type (`drv`) on the city fuel efficiency (`cty`) for 38 popular models of car was performed in R. The commands and output are given in Appendix A.

The displacement of a car is the volume of the cylinders, while the drive is the type of drive, in this case we have just three levels - front-wheel drive, rear-wheel drive and four-wheel drive.

Three models are fitted:

- `cty` on `displ` (Model 1 - identical regression)
- `cty` on `displ` and `drv` (Model 2 - parallel regression)
- `cty` on `displ` and `drv` with interaction (Model 3 - separate regression)

- (a) Consider the scatterplot of city fuel efficiency against displacement given in Figure 1. Describe the relationship. [3 marks]
- (b) Consider the separate regression model. Write down the line of best fit for the relationship between displacement and city fuel efficiency for rear-wheel drive cars. [2 marks]
- (c) Test for a statistically significant interaction term in the separate regression model at the 5% significance level. Remember to include the null and alternative hypotheses, the value of the test statistic, the P-value and your conclusion. [4 marks]
- (d) Using the Akaike's Information Criterion, which model fits the data the best? Justify your answer. [2 marks]

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- (e) Assess the assumptions of the linear model used in the separate regression model. The plots given in Figure 2 may be used where appropriate. [4 marks]

[Total: 15]

5. Suppose  $y_1, y_2, \dots, y_n$  are independent Poisson observations with parameter  $\lambda$ ,  $\lambda > 0$ . That is, for  $i = 1, 2, \dots, n$ ,

$$f(y_i; \lambda) = \frac{e^{-\lambda} \lambda^{y_i}}{y_i!}, \quad y_i > 0.$$

- (a) Write down the likelihood. [1 marks]
- (b) Write down the log-likelihood. [1 marks]
- (c) Find the maximum likelihood estimate of  $\lambda$ ,  $\hat{\lambda}$ . [3 marks]
- (d) Find the Fisher information. [3 marks]
- (e) Let  $\phi = \log(\lambda)$ . Write down the maximum likelihood estimate,  $\hat{\phi}$ . [1 marks]

[Total: 9]

6. Haemophilia is a X-chromosome linked, recessive disorder. Suppose a woman has a haemophilic brother, her father is normal, and her mother is a carrier. Let

$$\theta = \begin{cases} 1 & \text{if the woman is a carrier,} \\ 0 & \text{otherwise.} \end{cases}$$

It follows from genetic considerations that the prior distribution is

$$p(\theta) = \begin{cases} \frac{1}{2} & \text{if } \theta = 1, \\ \frac{1}{2} & \text{if } \theta = 0. \end{cases}$$

- (a) Suppose the woman has two sons, of which neither have haemophilia. Find the probability the woman is a carrier. [5 marks]

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- (b) Suppose the woman has a third son. Given that the first two sons are not haemophiliacs, what is the probability that the third son is not a haemophiliac? [3 marks]

[Total: 8]

7. Consider the multiple regression model

$$\mathbf{Y} = X\boldsymbol{\beta} + \boldsymbol{\varepsilon},$$

where  $\mathbf{Y}$  is an  $n \times 1$  vector of response random variables,  $X$  is an  $n \times p$  design matrix,  $\boldsymbol{\beta}$  is a  $p \times 1$  vector of regression parameters and  $\boldsymbol{\varepsilon}$  is an  $n \times 1$  vector of random errors with  $\varepsilon_i \sim N(0, \sigma^2)$ ,  $i = 1, \dots, n$ .

- (a) State the necessary and sufficient condition on  $X$  for the least squares estimate

$$\hat{\boldsymbol{\beta}} = (X^T X)^{-1} X^T \mathbf{y}$$

to be uniquely identified.

[1 marks]

- (b) Prove that  $\hat{\boldsymbol{\beta}}$  uniquely minimises the sum of squares  $Q(\boldsymbol{\beta}) = \|\mathbf{y} - X\boldsymbol{\beta}\|^2$ .

[14 marks]

[Total: 15]

8. To investigate the effect of Vitamin C on tooth growth in Guinea Pigs, 60 Guinea Pigs were given doses of Vitamin C. Each animal received one of three dose levels of vitamin C (0.5, 1, and 2 mg/day), denoted by dose, by one of two delivery methods, orange juice or ascorbic acid (a form of vitamin C and coded as VC): denoted by supp. The response is the length of odontoblasts (cells responsible for tooth growth) denoted len.

An analysis was performed in R and the output is given in Appendix C.

- (a) Describe what the code in the section **Clean data** is doing. [2 marks]
- (b) Is the experiment balanced? Justify your answer. [2 marks]
- (c) Using the interaction plot in Figure 3, describe the relationship between length of odontoblasts and dose; and between length of odontoblasts and delivery method. Does an interaction between dose and supplementary appear to be present? [3 marks]

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- (d) From the output, is an interaction term necessary? Justify your conclusion with reference to the output. [2 marks]
- (e) Using the interaction model, calculate the predicted mean length of odontoblasts for Guinea Pigs on a low dose of Vitamin C given as orange juice. [2 marks]
- (f) Using the interaction model, calculate the predicted mean length of odontoblasts for Guinea Pigs on a high dose of Vitamin C given as ascorbic acid. [2 marks]
- (g) Using the normal QQ-plots given in Figure 4, is the assumption of normality of length for each treatment reasonable? Justify your conclusion. [2 marks]

[Total: 15]

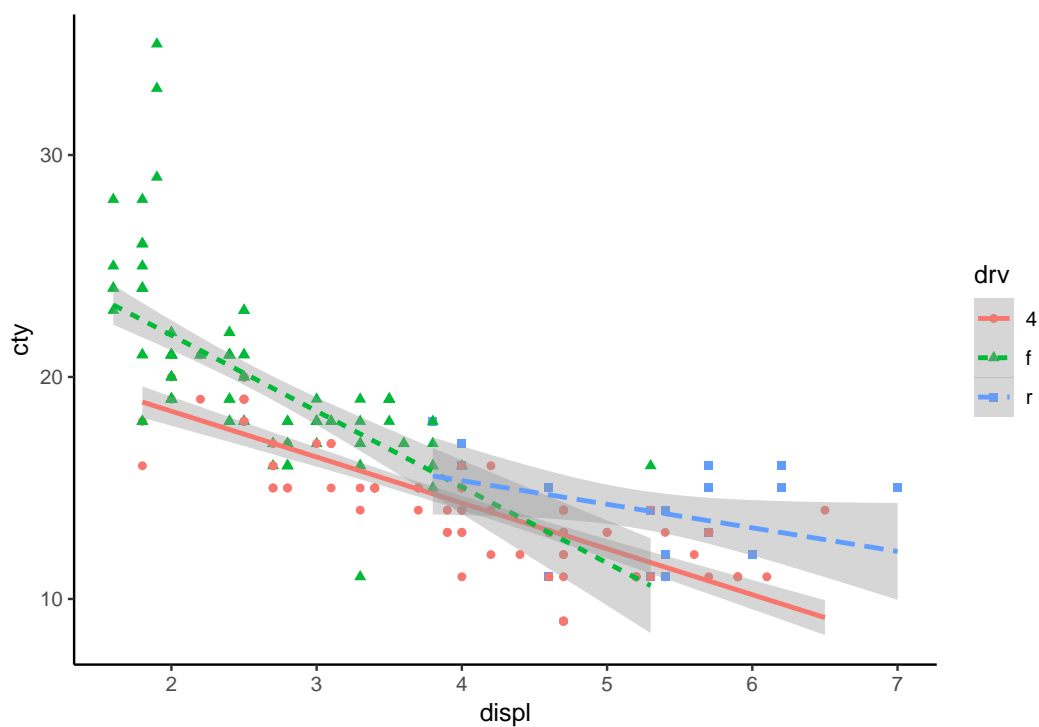


Figure 1: Scatterplot of Fuel efficiency against displacement for the MPG dataset. Colour and shape of points indicates drive (type).

## Appendix A

### Load the data

```
library(tidyverse)
data(mpg)
theme_set(theme_classic())
```

### Visualise data

```
mpg %>%
  ggplot(aes(displ, cty, col = drv, shape = drv)) +
  geom_point() +
  geom_smooth(method = "lm", aes(linetype = drv))
```

### Fit models



```
identical <- lm(cty ~ displ, data = mpg)
parallel <- lm(cty ~ displ + drv, data = mpg)
separate <- lm(cty ~ displ * drv, data = mpg)
```

## Model Coefficients

```
summary(separate)
```

```
##
## Call:
## lm(formula = cty ~ displ * drv, data = mpg)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -6.4363 -1.2957 -0.0863  1.1203 12.7768
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  22.5914     0.8136  27.768 < 2e-16 ***
## displ       -2.0663     0.1958 -10.554 < 2e-16 ***
## drvf         6.1284     1.1632   5.269 3.18e-07 ***
## drvr        -3.0124     3.1043  -0.970 0.332872
## displ:drvf  -1.3529     0.3696  -3.661 0.000313 ***
## displ:drvr   1.0039     0.6048   1.660 0.098285 .
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 2.252 on 228 degrees of freedom
## Multiple R-squared:  0.7261, Adjusted R-squared:  0.7201
## F-statistic: 120.9 on 5 and 228 DF, p-value: < 2.2e-16
```

```
anova(separate)
```

```
## Analysis of Variance Table
##
## Response: cty
##           Df Sum Sq Mean Sq F value    Pr(>F)
## displ      1 2691.06 2691.06  530.7574 < 2.2e-16 ***
## drv        2  277.99  138.99  27.4136 2.144e-11 ***
## displ:drv   2   95.28   47.64   9.3963 0.0001199 ***
```

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```
## Residuals 228 1156.01    5.07
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
anova(separate, parallel)
```

```
## Analysis of Variance Table
##
## Model 1: cty ~ displ * drv
## Model 2: cty ~ displ + drv
##   Res.Df    RSS Df Sum of Sq    F    Pr(>F)
## 1      228 1156.0
## 2      230 1251.3 -2    -95.283 9.3963 0.0001199 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

## Model Selection

```
AIC(identical, parallel, separate)
```

```
##           df      AIC
## identical  3 1109.336
## parallel   5 1066.391
## separate   7 1051.857
```

## Assumption checking

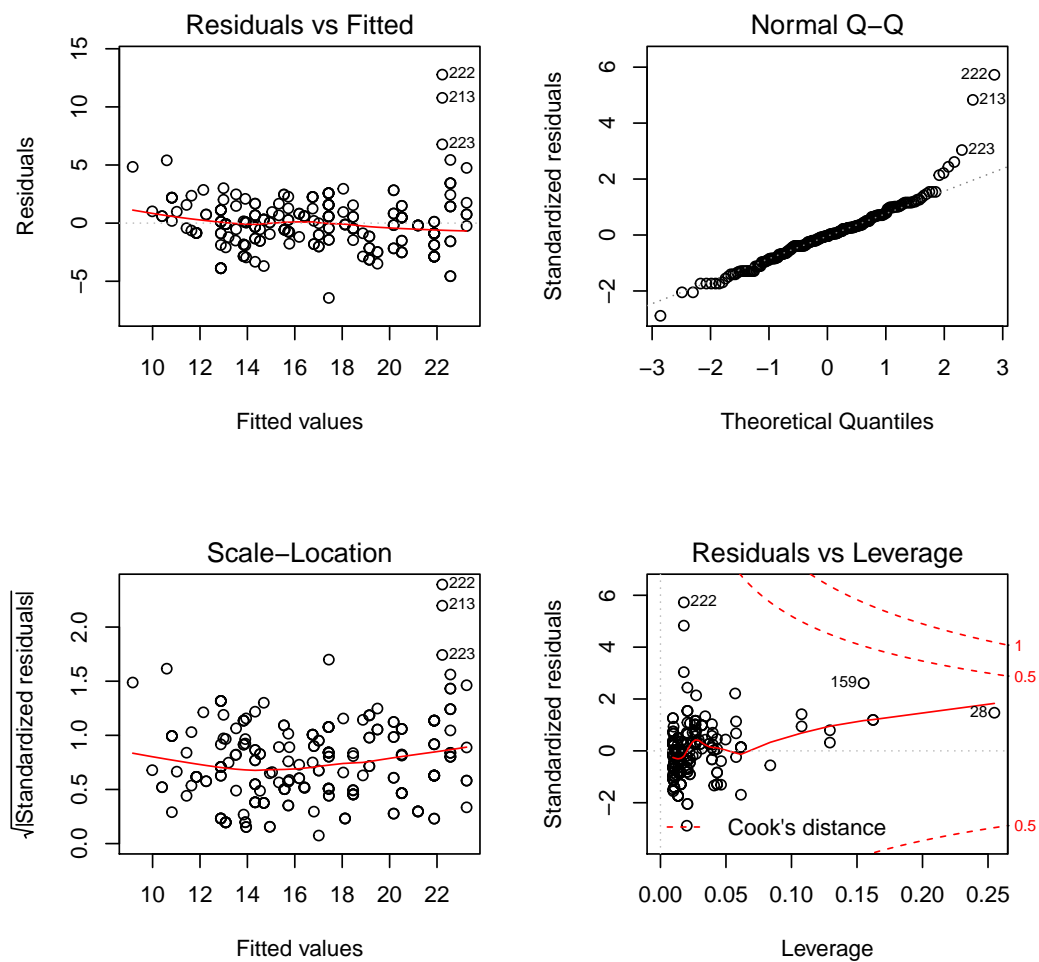


Figure 2: Assumption plots of the separate model for the MPG dataset.

## Appendix B

Distribution	Probability mass function / probability density function	Expectation	Variance
Binomial	$p(x) = \binom{n}{x} p^x (1-p)^{n-x}$ for $x = 0, 1, 2, \dots, n$	$np$	$np(1-p)$
Geometric	$p(x) = p(1-p)^{x-1}$ for $x = 1, 2, \dots$	$\frac{1}{p}$	$\frac{1-p}{p^2}$
Poisson	$p(x) = \frac{e^{-\lambda} \lambda^x}{x!}$ for $x = 0, 1, 2, \dots$	$\lambda$	$\lambda$
Uniform	$f(x) = \frac{1}{b-a}$ for $a < x < b$	$\frac{a+b}{2}$	$\frac{(b-a)^2}{12}$
Exponential	$f(x) = \lambda e^{-\lambda x}$ for $x > 0$	$\frac{1}{\lambda}$	$\frac{1}{\lambda^2}$
Gamma	$f(x) = \frac{\lambda^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-\lambda x}$ for $x > 0$	$\frac{\alpha}{\lambda}$	$\frac{\alpha}{\lambda^2}$
Normal	$f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-(1/2\sigma^2)(x-\mu)^2}$ for $-\infty < x < \infty$	$\mu$	$\sigma^2$
Beta	$f(x) = \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} \theta^{\alpha-1} (1-\theta)^{\beta-1}$ for $0 < \theta < 1$	$\frac{\alpha}{\alpha+\beta}$	$\frac{\alpha\beta}{(\alpha+\beta)^2(\alpha+\beta+1)}$

## Appendix C

### Load data

```
library(tidyverse)
data("ToothGrowth")
ToothGrowth <- as_tibble(ToothGrowth)
ToothGrowth
```

```
## # A tibble: 60 x 3
##       len supp  dose
##   <dbl> <fct> <dbl>
## 1   4.2 VC     0.5
## 2  11.5 VC     0.5
## 3   7.3 VC     0.5
## 4   5.8 VC     0.5
## 5   6.4 VC     0.5
## 6  10   VC     0.5
## 7  11.2 VC     0.5
## 8  11.2 VC     0.5
## 9   5.2 VC     0.5
## 10   7   VC     0.5
## # ... with 50 more rows
```

### Clean data

```
ToothGrowth <-
  ToothGrowth %>%
  mutate(dose = case_when(
    dose == 0.5 ~ "low",
    dose == 1 ~ "medium",
    TRUE ~ "high"
  ))
ToothGrowth$dose <- factor(ToothGrowth$dose,
  levels = c("low", "medium", "high"))
```

### Descriptive analysis

```
ToothGrowth %>%
  count(dose, supp) %>%
  spread(supp, n)
```

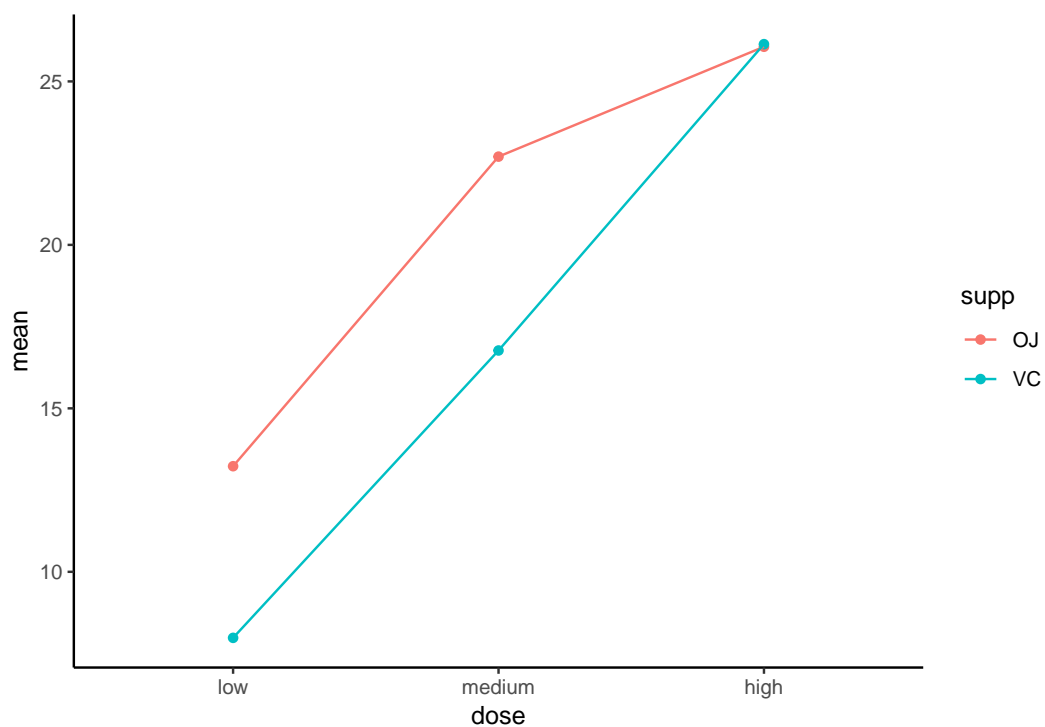


Figure 3: Interaction plot of mean length for each dose and type of supplement.

```
## # A tibble: 3 x 3
##   dose      OJ    VC
##   <fct> <int> <int>
## 1 low      10    10
## 2 medium   10    10
## 3 high     10    10
```

## Two-way ANOVA

```
ToothGrowth_M1 <- lm(len ~ dose * supp, data = ToothGrowth)
ToothGrowth_M2 <- lm(len ~ dose + supp, data = ToothGrowth)
anova(ToothGrowth_M1)
```

```
## Analysis of Variance Table
##
## Response: len
##          Df Sum Sq Mean Sq F value    Pr(>F)
## dose      2 2426.43  1213.22   92.000 < 2.2e-16 ***
## supp      1  205.35   205.35   15.572 0.0002312 ***
## dose:supp  2  108.32    54.16    4.107 0.0218603 *
```

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```
## Residuals 54 712.11 13.19
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
anova(ToothGrowth_M1, ToothGrowth_M2)
```

```
## Analysis of Variance Table
##
## Model 1: len ~ dose * supp
## Model 2: len ~ dose + supp
##   Res.Df    RSS Df Sum of Sq    F Pr(>F)
## 1      54 712.11
## 2      56 820.43 -2   -108.32 4.107 0.02186 *
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
summary(ToothGrowth_M1)
```

```
##
## Call:
## lm(formula = len ~ dose * supp, data = ToothGrowth)
##
## Residuals:
##   Min       1Q   Median       3Q      Max
## -8.20  -2.72  -0.27   2.65   8.27
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)      13.230      1.148  11.521 3.60e-16 ***
## dosemedium        9.470      1.624   5.831 3.18e-07 ***
## dosehigh        12.830      1.624   7.900 1.43e-10 ***
## suppVC          -5.250      1.624  -3.233 0.00209 **
## dosemedium:suppVC -0.680      2.297  -0.296 0.76831
## dosehigh:suppVC    5.330      2.297   2.321 0.02411 *
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 3.631 on 54 degrees of freedom
## Multiple R-squared: 0.7937, Adjusted R-squared: 0.7746
## F-statistic: 41.56 on 5 and 54 DF, p-value: < 2.2e-16
```

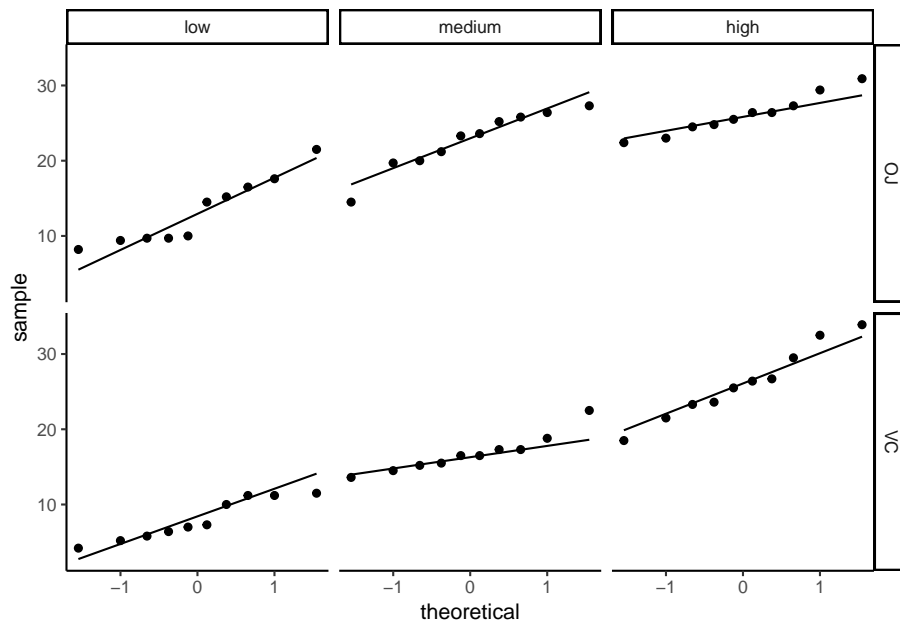


Figure 4: Normal QQ-plots of length for each treatment.

```
summary(ToothGrowth_M2)
```

```
##
## Call:
## lm(formula = len ~ dose + supp, data = ToothGrowth)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -7.085 -2.751 -0.800  2.446  9.650
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  12.4550     0.9883  12.603 < 2e-16 ***
## dosemedium    9.1300     1.2104   7.543 4.38e-10 ***
## dosehigh     15.4950     1.2104  12.802 < 2e-16 ***
## suppVC       -3.7000     0.9883  -3.744 0.000429 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 3.828 on 56 degrees of freedom
## Multiple R-squared:  0.7623, Adjusted R-squared:  0.7496
## F-statistic: 59.88 on 3 and 56 DF,  p-value: < 2.2e-16
```