

STATS 2107
Statistical Modelling and Inference II
Tutorial 2
Solutions

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1. a. If $X \sim \chi_k^2$, show that $E(X) = k$ and $\text{Var}(X) = 2k$. **Hint: Use MGFs.**

Solutions:

$$\begin{aligned} E[X] &= \left. \frac{d}{dt} M_X(t) \right|_{t=0} \\ &= \left. \frac{d}{dt} (1 - 2t)^{-k/2} \right|_{t=0} \\ &= \left. k(1 - 2t)^{-k/2-1} \right|_{t=0} \\ &= k. \end{aligned}$$

$$\begin{aligned} E[X^2] &= \left. \frac{d^2}{dt^2} M_X(t) \right|_{t=0} \\ &= \left. \frac{d}{dt} k(1 - 2t)^{-k/2-1} \right|_{t=0} \\ &= \left. k(k+2)(1 - 2t)^{-k/2-2} \right|_{t=0} \\ &= k(k+2). \end{aligned}$$

Now

$$\begin{aligned} \text{Var}(X) &= E[X^2] - E[X]^2 \\ &= k(k+2) - k^2 \\ &= 2k. \end{aligned}$$

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- b. Suppose $X_1 \sim \chi_{k_1}^2$ and $X_2 \sim \chi_{k_2}^2$ independently. Find the distribution of $X_1 + X_2$.

Solutions:

Use the fact that the MGF of a sum of independent random variables is equal to the product of the individual MGFs.

$$\begin{aligned} M_{X_1+X_2}(t) &= M_{X_1}(t) \times M_{X_2}(t) \\ &= (1 - 2t)^{-k_1/2} \times (1 - 2t)^{-k_2/2} \\ &= (1 - 2t)^{-(k_1+k_2)/2}. \end{aligned}$$

By observation and the uniqueness of MGFs, we recognise this as the MGF of a χ^2 distribution with $k_1 + k_2$ degrees of freedom

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2. A random sample of 500 hospital records shows that the length of stay in one of South Australia's hospitals had a (sample) mean 5.4 days and (population) standard deviation 3.1 days.
- a. A health agency hypothesizes that the average length of stay is 5 days. Do the data support this hypothesis? You may use $\alpha = 0.05$.
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Solutions:

Let μ be the true mean length of stay in the hospital. Then we are testing

$$H_0 : \mu = 5 \quad \text{against} \quad H_a : \mu \neq 5.$$

The test statistic is

$$z = \frac{\bar{y} - \mu_0}{\frac{\sigma}{\sqrt{n}}} = \frac{5.4 - 5}{\frac{3.1}{\sqrt{500}}} \approx 2.89.$$

The critical value is $z_{\alpha/2} = z_{0.025} \approx 1.96$. Hence there is sufficient evidence to reject H_0 since $z > 1.96$.

- b. For the hypothesis test in part a, and using the significance level $\alpha = 0.05$, find β for $\mu = 5.5$.
Hint: first calculate the power using the formula from the lectures.
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Solutions:

We want to use the formula for the two-sided hypothesis test that:

$$\begin{aligned} 1 - \beta &= \text{Power} \\ &= \Phi\left(-z_{\alpha/2}; \frac{\mu - \mu_0}{\sigma/\sqrt{n}}, 1\right) + 1 - \Phi\left(z_{\alpha/2}; \frac{\mu - \mu_0}{\sigma/\sqrt{n}}, 1\right), \end{aligned}$$

Where Φ is the cdf of the normal distribution. In this situation, $\alpha = 0.05$, $\mu = 5.5$, $\mu_0 = 5$, $\sigma = 3.1$, and $n = 500$. Using this (and R) we get that

$$\begin{aligned} 1 - \beta &= 0.9501796 \\ \therefore \beta &\approx 0.05. \end{aligned}$$

- c. How large should the sample size be if we require that $\alpha = 0.01$ and $\beta = 0.05$, assuming $\mu = 5.5$?
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Solutions:

This is a two-sided z-test. Using the formula from lecture, we have

$$n = \frac{(z_{\alpha/2} + z_{\beta})^2 \sigma^2}{(\mu_a - \mu_0)^2} = 684.7765.$$

Hence, we need a sample of size 685 to provide the desired levels.

3. A study is to be conducted to investigate the amount of toxic chemicals in freshwater lakes. A common measure of toxicity for any pollutant is LC50 (lethal concentration killing 50% of test species), which is the concentration of the pollutant that will kill half of the test species in a given amount of time (usually 96 hours for fish species). In many studies, the natural logarithm of LC50 measurements, $\log(\text{LC50})$, are normally distributed. For copper, the variance of $\log(\text{LC50})$ measurements is around 0.4 mg/L (milligrams per litre) on fish species A and around 0.8 mg/L for fish species B.

- a. Suppose 10 samples were collected for species A. Find the probability that the sample mean of $\log(\text{LC50})$ will differ from the population mean by no more than 0.5.

Solutions:

Let \bar{X}_A denote the sample mean of the species A sample, and μ_A denote the population mean for this species. We have $\sigma_A^2 = 0.4$ and sample size $n = 10$.

$$\begin{aligned} P(|\bar{X}_A - \mu_A| \leq 0.5) &= P\left(\left|\frac{\bar{X}_A - \mu_A}{\frac{\sigma_A}{\sqrt{n}}}\right| \leq \frac{0.5}{\frac{\sigma_A}{\sqrt{n}}}\right) \\ &= P(|Z| \leq 2.5) \\ &= P(-2.5 \leq Z \leq 2.5) \\ &= 1 - 2P(Z > 2.5) \\ &= 1 - 2(0.00621) \\ &= 0.9876 \end{aligned}$$

- b. If we want the sample mean (for species A) to differ from the population by no more than 0.5 with probability 0.95, how many samples do we need to collect?

Solutions:

We want

$$\begin{aligned} P(|\bar{X}_A - \mu_A| \leq 0.5) &= P\left(|Z| \leq \frac{0.5}{\sqrt{\frac{0.4}{n}}}\right) \\ &= 1 - P\left(|Z| > \frac{0.5}{\sqrt{\frac{0.4}{n}}}\right) \\ &= P\left(-\frac{0.5}{\sqrt{0.4}}\sqrt{n} \leq Z \leq \frac{0.5}{\sqrt{0.4}}\sqrt{n}\right) \\ &= 0.95 \end{aligned}$$

We know that $P(|Z| > 1.96) = 0.05$. Hence it follows that $\frac{0.5}{\sqrt{\frac{0.4}{n}}} = 1.96$, which implies $n = \frac{0.4}{\left(\frac{0.5}{1.96}\right)^2} = 6.15$. Hence, we need to collect 7 samples.

- c. Assuming the population mean for both species is the same, what is the probability that the sample mean of species A will exceed the sample mean of species B by at least 1 mg/L, if we collected 10 measurements from each species?

Solutions:

Let \bar{Y}_B denote the sample mean of the species B sample. We have $E(\bar{X}_A - \bar{Y}_B) = \mu_A - \mu_B = 0$. Also, since \bar{X}_A and \bar{Y}_B are independent, then $\text{Var}(\bar{X}_A - \bar{Y}_B) = \frac{\sigma_A^2 + \sigma_B^2}{n} = \frac{0.4 + 0.8}{10} = 0.12$. Hence,

$$\begin{aligned} P(\bar{X}_A - \bar{Y}_B \geq 1) &= P\left(\frac{\bar{X}_A - \bar{Y}_B}{\sqrt{0.12}} \geq \frac{1}{\sqrt{0.12}}\right) \\ &= P\left(Z \geq \frac{1}{\sqrt{0.12}}\right) \\ &= P(Z \geq 2.89) \\ &= 0.0019 \end{aligned}$$

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4. Suppose X_1, X_2, \dots, X_m and Y_1, Y_2, \dots, Y_n are independent random samples, with $X_i \sim N(\mu_1, \sigma_1^2)$ and $Y_j \sim N(\mu_2, \sigma_2^2)$ for $i = 1, 2, \dots, m$ and $j = 1, 2, \dots, n$.

a. State $E(\bar{X} - \bar{Y})$.

Solutions:

$$E(\bar{X} - \bar{Y}) = E(\bar{X}) - E(\bar{Y}) = \mu_1 - \mu_2$$

b. State $\text{Var}(\bar{X} - \bar{Y})$.

Solutions:

As \bar{X} and \bar{Y} are independent,

$$\text{Var}(\bar{X} - \bar{Y}) = \text{Var}(\bar{X}) + \text{Var}(\bar{Y}) = \frac{\sigma_1^2}{m} + \frac{\sigma_2^2}{n}.$$

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- c. What is the sample size needed so that $(\bar{X} - \bar{Y})$ will be within k units of $(\mu_1 - \mu_2)$ with probability $1 - \alpha$? You may assume $m = n$.

Solutions:

It is required that $P(|(\bar{X} - \bar{Y}) - (\mu_1 - \mu_2)| \leq k) = 1 - \alpha$. Observe that $\bar{X} - \bar{Y} \sim N(\mu_1 - \mu_2, \frac{\sigma_1^2}{m} + \frac{\sigma_2^2}{n})$. Taking $n = m$, we have

$$\begin{aligned} 1 - \alpha &= P(|(\bar{X} - \bar{Y}) - (\mu_1 - \mu_2)| \leq k) \\ &= P\left(\left|\frac{(\bar{X} - \bar{Y}) - (\mu_1 - \mu_2)}{\sqrt{\frac{\sigma_1^2 + \sigma_2^2}{n}}}\right| \leq \frac{k}{\sqrt{\frac{\sigma_1^2 + \sigma_2^2}{n}}}\right) \\ &= P\left(|Z| \leq \frac{k}{\sqrt{\frac{\sigma_1^2 + \sigma_2^2}{n}}}\right). \end{aligned}$$

Equivalently, this is

$$P\left(|Z| \geq \frac{k}{\sqrt{\frac{\sigma_1^2 + \sigma_2^2}{n}}}\right) = \alpha.$$

Since $Z \sim N(0, 1)$, it follows that

$$\frac{k}{\sqrt{\frac{\sigma_1^2 + \sigma_2^2}{n}}} = z_{\alpha/2},$$

and hence

$$n = \frac{(\sigma_1^2 + \sigma_2^2) z_{\alpha/2}^2}{k^2}.$$
