STATS 2107

Statistical Modelling and Inference II Tutorial 1 Solutions

Jono Tuke, Matt Ryan

Semester 2 2022

1. Suppose Y_1, Y_2, \ldots, Y_n are i.i.d. $N(\mu, \sigma^2)$ random variables and let \bar{Y} denote the sample mean. Using moment generating functions (MGFs), prove that $\bar{Y} \sim N(\mu, \sigma^2/n)$.

Solutions:

First consider the moment generating function for a general sample mean:

$$M_{\bar{Y}}(t) = E[e^{t\bar{Y}}]$$

$$= E[e^{t(Y_1/n + Y_2/n + \dots + Y_n/n)}]$$

$$= E[e^{tY_1/n + tY_2/n + \dots + tY_n/n}]$$

$$= E[e^{tY_1/n}]E[e^{tY_2/n}] \dots E[e^{tY_n/n}] \quad \text{by independence}$$

$$= M_{Y_1}(t/n)M_{Y_2}(t/n) \dots M_{Y_n}(t/n)$$

$$= \prod_{i=1}^n M_{Y_i}(t/n). \qquad (*)$$

We are given that $Y_i \sim N(\mu, \sigma^2)$ which has MGF

$$M_{Y_i}(t) = \exp\left\{\mu t + \frac{t^2 \sigma^2}{2}\right\}$$

Using (*) we have

$$M_{\bar{Y}}(t) = \left[\exp\left\{ \frac{\mu t}{n} + \frac{t^2 \sigma^2}{2n^2} \right\} \right]^n$$

$$= \exp\left\{ \sum_{i=1}^n \frac{\mu t}{n} + \sum_{i=1}^n \frac{t^2 \sigma^2}{2n^2} \right\}$$

$$= \exp\left\{ \mu t + \frac{t^2 \sigma^2}{2n} \right\}.$$

Which by the uniqueness of MGFs proves that \bar{Y} is normally distributed with mean μ and variance σ^2/n .

2. Let Z_1, Z_2, \dots, Z_p be i.i.d. N(0,1) random variables.

a. Show that the moment generating function $M_{Z_i^2}(t)$ of Z_i^2 is given by $(1-2t)^{-\frac{1}{2}}$ for each $i=1,2,\ldots,p$.

Solutions:

$$\begin{split} M_{Z_i^2}(t) &= E[e^{tZ_i^2}] \\ &= \int_{-\infty}^{\infty} e^{tz^2} \frac{1}{\sqrt{2\pi}} e^{-z^2/2} dz \\ &= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{z^2(t-1/2)} dz \\ &= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-z^2(1-2t)/2} dz \\ &= \frac{1}{(1-2t)^{1/2}} \int_{-\infty}^{\infty} \frac{(1-2t)^{1/2}}{\sqrt{2\pi}} e^{-z^2(1-2t)/2} dz \end{split}$$

but the integral is one since it is the pdf of $N(0, \frac{1}{1-2t})$. So we have

$$M_{Z^2}(t) = \frac{1}{(1-2t)^{1/2}}.$$

b. Hence, show that the moment generating function of $X = \sum_{i=1}^{p} Z_i^2$ is

$$M_X(t) = \frac{1}{(1-2t)^{p/2}}.$$

Solutions:

Using independence of the Z_i we can conclude that the Z_i^2 are independent. Then using the property of independent MGFs we get that:

$$M_X(t) = M_{\sum_{i=1}^p Z_i^2}(t)$$

$$= \prod_{i=1}^p M_{Z_i^2}(t)$$

$$= \prod_{i=1}^p (1 - 2t)^{-\frac{1}{2}}$$

$$= \frac{1}{(1 - 2t)^{p/2}}.$$

3. Let $Y \sim Bin(n, p)$ and consider two estimators for p; namely

$$\hat{p}_1 = \frac{Y}{n} \text{ and } \hat{p}_2 = \frac{Y+1}{n+2}.$$

a. Show that \hat{p}_1 is unbiased for p.

Solutions:

$$E[\hat{p}_1] = E[Y/n]$$

$$= \frac{1}{n}E[Y]$$

$$= \frac{np}{n}$$

$$= p.$$

b. Derive the bias of \hat{p}_2 .

Solutions:

$$b_{\hat{p}_2}(p) = E[\hat{p}_2] - p$$

$$= E\left[\frac{Y+1}{n+2}\right] - p$$

$$= \frac{np+1}{n+2} - p$$

$$= \frac{np+1 - p(n+2)}{n+2}$$

$$= \frac{1-2p}{n+2}.$$

c. Find $MSE_{\hat{p}_1}(p)$ and $MSE_{\hat{p}_2}(p)$.

Solutions:

$$MSE_{\hat{p}_{1}}(p) = Var(\hat{p}_{1}) + b_{\hat{p}_{1}}(p)^{2}$$

$$= Var(Y/n) + 0$$

$$= \frac{1}{n^{2}}Var(Y)$$

$$= \frac{np(1-p)}{n^{2}}$$

$$= \frac{p(1-p)}{n}.$$

$$MSE_{\hat{p}_{2}}(p) = Var(\hat{p}_{2}) + b_{\hat{p}_{2}}(p)^{2}$$

$$= Var\left(\frac{Y+1}{n+2}\right) + \left(\frac{1-2p}{n+2}\right)^{2}$$

$$= \frac{Var(Y) + (1-2p)^{2}}{(n+2)^{2}}$$

$$= \frac{np(1-p) + (1-2p)^{2}}{(n+2)^{2}}.$$

d. If p = 0.5, which estimator has the largest MSE?

Solutions:

$$MSE_{\hat{p}_1}(p) = \frac{1}{4n}$$

$$MSE_{\hat{p}_2}(p) = \frac{n}{4(n+2)^2}$$

$$= \frac{1}{4(n+4+4/n)}$$

So $MSE_{\hat{p}_2}(p) < MSE_{\hat{p}_1}(p)$

