

Q2

@ We have

(i)

$$X = \begin{bmatrix} 1 & x_1 \\ 1 & x_2 \\ \vdots & \vdots \\ 1 & x_n \end{bmatrix}$$

$$\bar{x} = \frac{1}{n} \sum_{i=1}^n x_i$$

$$\therefore n\bar{x} = \sum_{i=1}^n x_i$$

$$\therefore X^T = \begin{bmatrix} 1 & 1 & \dots & 1 \\ x_1 & x_2 & \dots & x_n \end{bmatrix}$$

$$X^T X = \begin{bmatrix} 1 & 1 & \dots & 1 \\ x_1 & x_2 & \dots & x_n \end{bmatrix} \begin{bmatrix} 1 & x_1 \\ 1 & x_2 \\ \vdots & \vdots \\ 1 & x_n \end{bmatrix}$$

$$= \begin{bmatrix} 1 \times 1 + 1 \times 1 + \dots + 1 \times 1 & x_1 + x_2 + \dots + x_n \\ x_1 + x_2 + \dots + x_n & x_1^2 + x_2^2 + \dots + x_n^2 \end{bmatrix}$$

$$= \begin{bmatrix} n & \sum_{i=1}^n x_i \\ \sum_{i=1}^n x_i & \sum_{i=1}^n x_i^2 \end{bmatrix}$$

$$\text{and } n\bar{x} = \sum_{i=1}^n x_i$$

$$\therefore X^T X = \begin{bmatrix} n & n\bar{x} \\ n\bar{x} & \sum_{i=1}^n x_i^2 \end{bmatrix}$$

(ii) We have:

$$\therefore \det(X^T X) = \begin{vmatrix} n & n\bar{x} \\ n\bar{x} & \sum_{i=1}^n x_i^2 \end{vmatrix}$$

$$= n \sum_{i=1}^n x_i^2 - n^2 (\bar{x})(\bar{x})$$

$$= n \sum_{i=1}^n x_i^2 - n^2 \bar{x}^2$$

$$= n \left(\sum_{i=1}^n x_i^2 - n \bar{x}^2 \right)$$

$$\text{and } S_{xx} = \sum_{i=1}^n x_i^2 - n \bar{x}^2$$

$$\therefore \det(X^T X) = n S_{xx}$$

⑥. ~~Proof~~ According to Lemma 6, if $X (n \times 2)$ is a matrix with linearly independent columns then the symmetric, 2×2 matrix $X^T X$ is invertible.

~~By the way~~ We have:

$$X = \begin{bmatrix} 1 & x_1 \\ 1 & x_2 \\ \vdots & \vdots \\ 1 & x_n \end{bmatrix} \text{ with}$$

and for $X^T X$ to be invertible, $\begin{pmatrix} 1 \\ 1 \\ \vdots \\ 1 \end{pmatrix}$ and $\begin{pmatrix} x_1 \\ x_2 \\ \vdots \\ x_n \end{pmatrix}$ are linearly independent

Hence, we can not express $\begin{pmatrix} x_1 \\ x_2 \\ \vdots \\ x_n \end{pmatrix}$ in term of $\begin{pmatrix} 1 \\ 1 \\ \vdots \\ 1 \end{pmatrix}$

Therefore, there has to be at least one of x_i 's (for $i=1, 2, \dots, n$) that is different from the remaining remaining.

⑦. We have

$$X^T X = \begin{bmatrix} n & n\bar{x} \\ n\bar{x} & \sum_{i=1}^n x_i^2 \end{bmatrix}$$

$$\begin{aligned} X^T y &= \begin{bmatrix} 1 & 1 & \dots & 1 \\ x_1 & x_2 & \dots & x_n \end{bmatrix} \begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_n \end{bmatrix} \\ &= \begin{bmatrix} y_1 + y_2 + \dots + y_n \\ x_1 y_1 + x_2 y_2 + \dots + x_n y_n \end{bmatrix} \\ &= \begin{bmatrix} \sum_{i=1}^n y_i \\ \sum_{i=1}^n x_i y_i \end{bmatrix} \quad \text{and} \end{aligned}$$

$$\text{and } n\bar{y} = \sum_{i=1}^n y_i$$

$$\therefore X^T y = \begin{bmatrix} n\bar{y} \\ \sum_{i=1}^n x_i y_i \end{bmatrix}$$

$X^T X$ is invertible

$$\begin{aligned} \therefore (X^T X)^{-1} &= \frac{1}{n \sum_{i=1}^n x_i^2 - n^2 \bar{x}^2} \begin{bmatrix} \sum_{i=1}^n x_i^2 & -n\bar{x} \\ -n\bar{x} & n \end{bmatrix} \\ &= \frac{1}{n S_{xx}} \begin{bmatrix} \sum_{i=1}^n x_i^2 & -n\bar{x} \\ -n\bar{x} & n \end{bmatrix} \end{aligned}$$

(from ⑥ (i))

$$\bar{y} = \frac{1}{n} \sum_{i=1}^n y_i$$

$$\therefore n\bar{y} = \sum_{i=1}^n y_i$$

$$\begin{aligned}
 \therefore (X^T X)^{-1} X^T y &= \frac{1}{n S_{xx}} \begin{bmatrix} \sum_{i=1}^n x_i^2 & -n\bar{x} \\ -n\bar{x} & n \end{bmatrix} \begin{bmatrix} n\bar{y} \\ \sum_{i=1}^n x_i y_i \end{bmatrix} \\
 &= \frac{1}{n S_{xx}} \begin{bmatrix} n\bar{y} \sum_{i=1}^n x_i^2 - n\bar{x} \sum_{i=1}^n x_i y_i \\ -n^2 \bar{x} \bar{y} + n \sum_{i=1}^n x_i y_i \end{bmatrix} \\
 &= \frac{1}{n S_{xx}} \begin{bmatrix} \bar{y} \sum_{i=1}^n x_i^2 - \bar{x} \sum_{i=1}^n x_i y_i \\ \sum_{i=1}^n x_i y_i - n\bar{x} \bar{y} \end{bmatrix}
 \end{aligned}$$

$$\text{and } S_{xy} = \sum_{i=1}^n x_i y_i - n\bar{x} \bar{y}$$

$$\therefore (X^T X)^{-1} X^T y = \frac{1}{S_{xx}} \begin{bmatrix} \bar{y} \sum_{i=1}^n x_i^2 - \bar{x} \sum_{i=1}^n x_i y_i \\ S_{xy} \end{bmatrix} \quad (1)$$

$$\begin{aligned}
 \bar{y} \sum_{i=1}^n x_i^2 - \bar{x} \sum_{i=1}^n x_i y_i &= \bar{y} \sum_{i=1}^n x_i^2 - n\bar{x}^2 \bar{y} + n\bar{x}^2 \bar{y} - \bar{x} \sum_{i=1}^n x_i y_i \\
 &= \bar{y} \left(\sum_{i=1}^n x_i^2 - n\bar{x}^2 \right) - \bar{x} \left(\sum_{i=1}^n x_i y_i - n\bar{x} \bar{y} \right)
 \end{aligned}$$

$$\text{and } S_{xx} = \sum_{i=1}^n x_i^2 - n\bar{x}^2 ; S_{xy} = \sum_{i=1}^n x_i y_i - n\bar{x} \bar{y}$$

$$\therefore \bar{y} \sum_{i=1}^n x_i^2 - \bar{x} \sum_{i=1}^n x_i y_i = \bar{y} S_{xx} - \bar{x} S_{xy} \quad (2)$$

From (1) and (2).

$$\begin{aligned}
 \therefore (X^T X)^{-1} X^T y &= \frac{1}{S_{xx}} \begin{bmatrix} \bar{y} S_{xx} - \bar{x} S_{xy} \\ S_{xy} \end{bmatrix} \\
 &= \begin{bmatrix} \bar{y} - \frac{S_{xy}}{S_{xx}} \bar{x} \\ \frac{S_{xy}}{S_{xx}} \end{bmatrix}
 \end{aligned}$$

(Q1)

①. We have:

$$H = \sqrt{n-3} (z - \operatorname{arctanh}(p))$$

$$\begin{aligned} E[H] &= E[\sqrt{n-3} (z - \operatorname{arctanh}(p))] \\ &= \sqrt{n-3} E[z - \operatorname{arctanh}(p)] \\ &= \sqrt{n-3} (E[z] - E[\operatorname{arctanh}(p)]) \\ \text{and } z &\sim N\left(\operatorname{arctanh}(p), \frac{1}{n-3}\right) \end{aligned}$$

$$\begin{aligned} \therefore E[H] &= \sqrt{n-3} (\operatorname{arctanh}(p) - \operatorname{arctanh}(p)) \\ &= 0 \end{aligned}$$

$$\begin{aligned} \operatorname{Var}[H] &= \operatorname{Var}(\sqrt{n-3} (z - \operatorname{arctanh}(p))) \\ &= (n-3) \operatorname{Var}(z - \operatorname{arctanh}(p)) \\ &= (n-3) (\operatorname{Var}(z) + \operatorname{Var}(\operatorname{arctanh}(p))) \\ &= (n-3) \left(\frac{1}{n-3} + 0 \right) \\ &= \frac{n-3}{n-3} = 1. \end{aligned}$$

$$\therefore H \sim N(0, 1)$$

• Since ~~it is known that~~ $H \sim N(0, 1)$, which is a known distribution that does not depend on p , and $H = \sqrt{n-3} (z - \operatorname{arctanh}(p))$, ~~we obtain~~

$\therefore H$ is a pivotal quantity for p

②. We have:

The symmetric $(1-\alpha)$ 100% confidence interval for p :

$$\begin{aligned} (1-\alpha) 100\% &= P(L \leq p \leq U) \\ &= P(\operatorname{arctanh}(L) \leq \operatorname{arctanh}(p) \leq \operatorname{arctanh}(U)) \\ &= P(\operatorname{arctanh}(L) - \operatorname{arctanh}(p) \geq 0 \geq \operatorname{arctanh}(U) - \operatorname{arctanh}(p)) \\ &= P(\sqrt{n-3} (z - \operatorname{arctanh}(L)) \geq \sqrt{n-3} (z - \operatorname{arctanh}(p)) \geq \sqrt{n-3} (z - \operatorname{arctanh}(U))) \\ &= P(\sqrt{n-3} (z - \operatorname{arctanh}(L)) \geq H \geq \sqrt{n-3} (z - \operatorname{arctanh}(U))) \\ &\text{where } H \sim N(0, 1) \end{aligned}$$

We want a symmetric confidence interval:

$$P(p \leq L) = \frac{\alpha}{2} \quad \text{and} \quad P(p \geq U) = \frac{\alpha}{2}$$

And we know:

$$P(H \geq Z_{\alpha/2}) = \frac{\alpha}{2} \quad \text{and} \quad P(H \leq -Z_{\alpha/2}) = \frac{\alpha}{2} \quad \text{where } H \sim N(0,1)$$

This implies:

$$Z_{\alpha/2} = \sqrt{n-3} (z - \operatorname{arctanh}(L))$$

$$-Z_{\alpha/2} = \sqrt{n-3} (z - \operatorname{arctanh}(U))$$

$$\therefore \frac{1}{\sqrt{n-3}} Z_{\alpha/2} = z - \operatorname{arctanh}(L)$$

$$\therefore \frac{-1}{\sqrt{n-3}} Z_{\alpha/2} = z - \operatorname{arctanh}(U)$$

$$\therefore \operatorname{arctanh}(L) = z - \frac{Z_{\alpha/2}}{\sqrt{n-3}}$$

$$\therefore \operatorname{arctanh}(U) = z + \frac{Z_{\alpha/2}}{\sqrt{n-3}}$$

$$\therefore L = \tanh\left(z - \frac{Z_{\alpha/2}}{\sqrt{n-3}}\right)$$

$$\therefore U = \tanh\left(z + \frac{Z_{\alpha/2}}{\sqrt{n-3}}\right)$$

\therefore The symmetric $(1-\alpha)$ 100% confidence interval for p :

$$\left(\tanh\left(z - \frac{Z_{\alpha/2}}{\sqrt{n-3}}\right), \tanh\left(z + \frac{Z_{\alpha/2}}{\sqrt{n-3}}\right) \right)$$

©. ~~We need to calculate~~ Based on the given dataset, we can calculate 95% confidence interval for the true correlation p between X and Y (using R code)

~~Given data~~

```
data <- read.csv("bivariate_normal_data.csv")
```

```
sx = sd(data$X) ≈ 0.994
```

```
sy = sd(data$Y) ≈ 1.044
```

```
sxy = cov(data$X, data$Y) ≈ 0.522
```

$$\therefore r = \frac{s_{xy}}{s_x s_y} = \frac{0.522}{(0.994)(1.044)} \approx 0.503$$

$$\therefore z = \operatorname{arctanh}(r) = \frac{1}{2} \log\left(\frac{1+r}{1-r}\right) = \frac{1}{2} \log\left(\frac{1+0.503}{1-0.503}\right) \approx 0.553$$

~~Since it is a 95% confidence interval~~ Since it is a 95% confidence interval, the confidence level is 5%

$$\therefore \alpha = 0.05$$

$$Z_{\alpha/2} = Z_{0.05/2} = Z_{0.025} = \operatorname{qnorm}(0.025, \text{lower.tail} = \text{FALSE}) \approx 1.960$$

$$\text{and } n = 100$$

. Therefore, the 95% confidence interval for ρ :

$$\begin{aligned} & \left(\tanh \left(z - \frac{z_{\alpha/2}}{\sqrt{n-3}} \right), \tanh \left(z + \frac{z_{\alpha/2}}{\sqrt{n-3}} \right) \right) \\ &= \left(\tanh \left(0.553 - \frac{1.960}{\sqrt{100-3}} \right), \tanh \left(0.553 + \frac{1.960}{\sqrt{100-3}} \right) \right) \\ &\approx (0.340, 0.636) \end{aligned}$$

④

④. We have:

$$H_0: \rho = \frac{1}{2}$$

$$\text{and } \alpha = 0.05$$

. Since $\rho = \frac{1}{2}$ is in the 95% confidence interval for ρ (0.340, 0.636), there is insufficient evidence to reject the null hypothesis at the 5% level.

~~In context, this says it is unreasonable to assume~~

. In context, we do not have sufficient evidence to say that the true correlation ρ between X and Y is ~~also~~ not equal to $\frac{1}{2}$.