

To evaluate whether the June 2023 Canadian wildfire smoke altered pollutant–health relationships in New York, two multiple linear regression (MLR) models were developed using weekly data from 2015–2024: one for total respiratory emergency department (ED) visits and another specifically for asthma-related ED visits. Both models incorporated the same set of predictor variables—fine particulate matter (PM<sub>2.5</sub>), ozone (O<sub>3</sub>), nitrogen dioxide (NO<sub>2</sub>), temperature, wind speed, and precipitation—to ensure comparability across health outcomes. PM<sub>2.5</sub>, O<sub>3</sub>, and NO<sub>2</sub> were included as primary air-quality exposures reflecting wildfire-related and urban pollution sources, while temperature, wind speed, and precipitation served as key meteorological controls influencing pollutant dispersion and physiological stress. Respiratory visits capture the overall respiratory burden, whereas asthma visits provide a more sensitive indicator of acute pollutant-driven exacerbations. This structure allows for a clear assessment of whether the wildfire period corresponded to a structural break in pollutant–health dynamics, particularly focusing on shifts in the PM<sub>2.5</sub>–health sensitivity before and after the June 2023 smoke event.

Asthma modeling

```
In [3]:  
  
=== Variance Inflation Factors (pooled, excl. wildfire) ===  
           feature      VIF  
           const      NaN  
Avg PM2.5 (ug/m3 LC) 1.143938  
Avg 8h O3 (ppm)      1.752413  
Avg 1h NO2 (ppb)     2.211892  
TAVG (°F, weekly mean) 4.147996  
Avg_WND (mph, weekly mean) 1.966511  
PRCP (inches, weekly sum) 1.221368  
  
=== Pooled model (excl. wildfire weeks) summary ===  
           OLS Regression Results  
=====
```

Dep. Variable:	Asthma_weekly_ER_Admissions	R-squared:	0.122
Model:	OLS	Adj. R-squared:	0.112
Method:	Least Squares	F-statistic:	11.81
Date:	Tue, 04 Nov 2025	Prob (F-statistic):	2.02e-12
Time:	22:08:17	Log-Likelihood:	-4123.0
No. Observations:	515	AIC:	8260.
Df Residuals:	508	BIC:	8290.
Df Model:	6		
Covariance Type:	nonrobust		

```
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```

	coef	std err	t	P> t	[0.025	0.975]
-----						
const	2419.7970	408.513	5.923	0.000	1617.215	3222.379
Avg PM2.5 (ug/m3 LC)	-16.8506	12.356	-1.364	0.173	-41.125	7.424
Avg 8h O3 (ppm)	-2.96e+04	5968.682	-4.959	0.000	-4.13e+04	-1.79e+04
Avg 1h NO2 (ppb)	62.7329	9.241	6.788	0.000	44.577	

80.889						
TAVG (°F, weekly mean)	14.1324	3.701	3.818	0.000	6.861	
21.404						
Avg_WND (mph, weekly mean)	84.6955	28.323	2.990	0.003	29.051	1
40.340						
PRCP (inches, weekly sum)	65.6345	67.260	0.976	0.330	-66.508	1
97.777						

Omnibus:	61.314	Durbin-Watson:	0.798
Prob(Omnibus):	0.000	Jarque-Bera (JB):	289.969
Skew:	-0.387	Prob(JB):	1.08e-63
Kurtosis:	6.593	Cond. No.	1.07e+04

#### Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

[2] The condition number is large, 1.07e+04. This might indicate that there are strong multicollinearity or other numerical problems.

#### === Coefficient comparison (Pre vs Post) ===

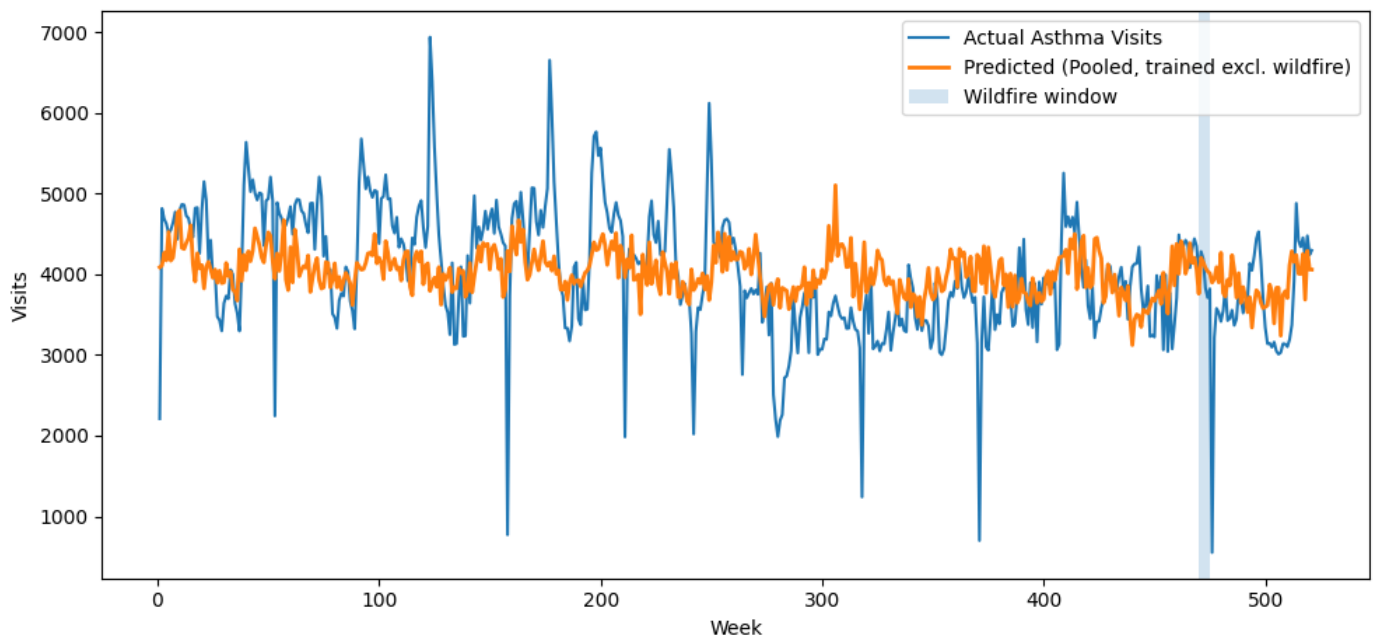
	Pre	Post	Post-Pre
const	2587.321369	2388.805592	-198.515777
Avg PM2.5 (ug/m3 LC)	-20.923815	20.083085	41.006901
Avg 8h O3 (ppm)	-29843.260818	-15758.610753	14084.650065
Avg 1h NO2 (ppb)	61.296744	27.177141	-34.119603
TAVG (°F, weekly mean)	13.653042	9.325026	-4.328015
Avg_WND (mph, weekly mean)	80.459670	74.605039	-5.854631
PRCP (inches, weekly sum)	37.554598	190.656927	153.102329

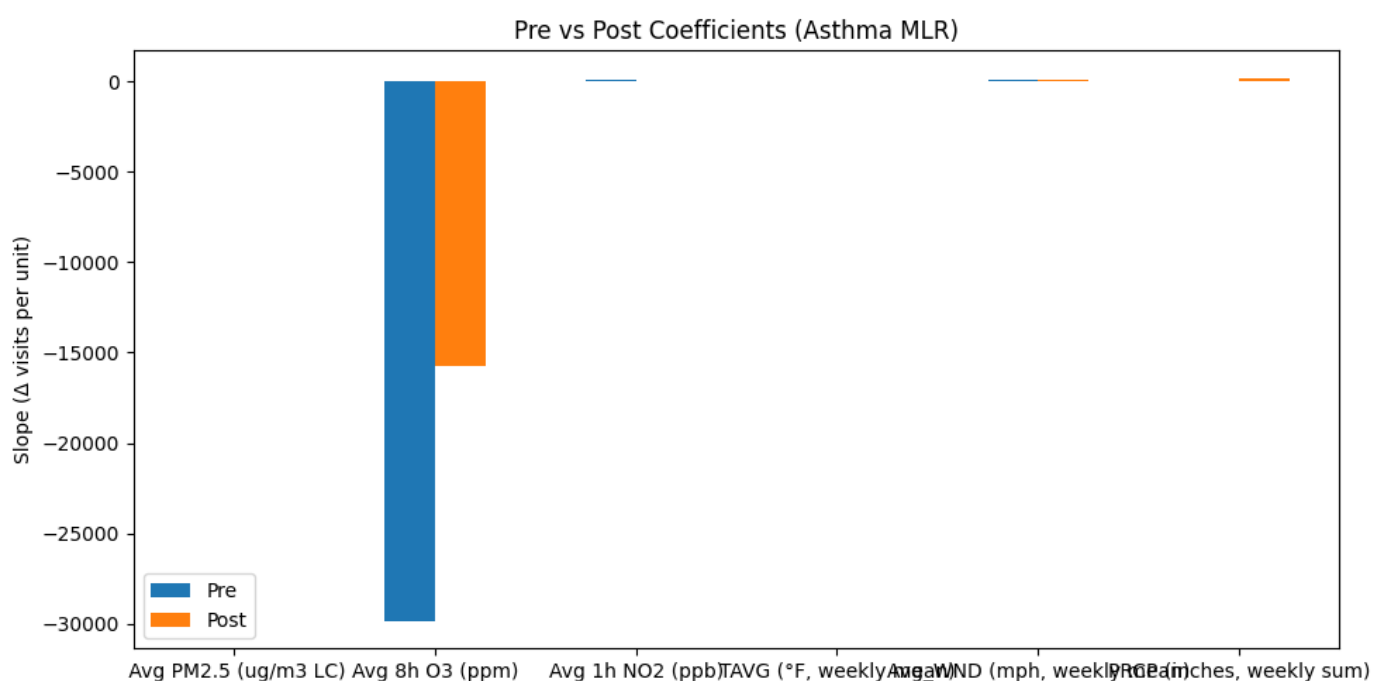
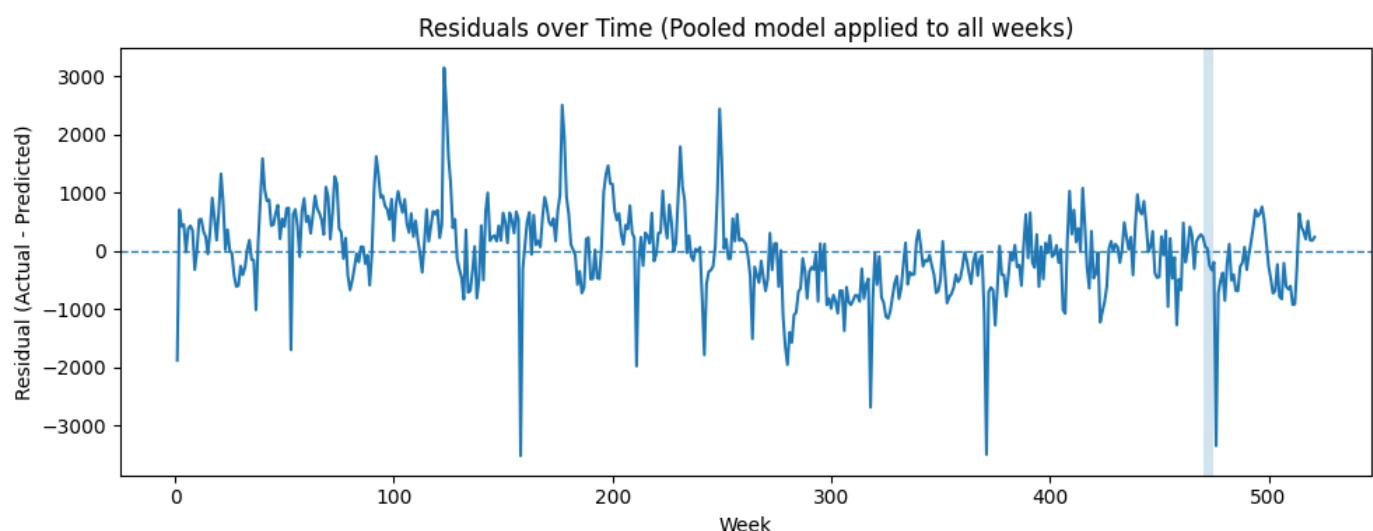
#### === Chow Test (break between weeks 469 and 476) ===

F(7, 501) = 1.158, p = 0.3255

SSR pooled: 271072838.4 | SSR pre: 247182199.7 | SSR post: 19574947.1

Asthma ED Visits: Actual vs. Pooled OLS Predictions





=== Quick diagnostics ===

Pre sample size (n1): 469 | Post sample size (n2): 46

Adj.  $R^2$  Pre: 0.106

Adj.  $R^2$  Post: -0.076

Adj.  $R^2$  Pooled (excl. wildfire weeks): 0.112

## Summary of what these results show

### 1. Overall Model Fit is Weak

The pooled OLS model (excluding the wildfire weeks) explains only about 11% of the variance in weekly asthma ER visits (Adjusted  $R^2 \approx 0.11$ ). That means the chosen environmental variables do influence asthma, but a large amount of variation is coming from other factors (e.g., viral circulation, hospital utilization patterns, socioeconomic shifts, baseline prevalence, etc.).

So the model is valid for comparison purposes, but not meant to be predictive.

## 2. Coefficients Have Reasonable Interpretability

Key predictors show expected directional effects:

Variable	Pooled Effect	Interpretation
High NO <sub>2</sub>	Positive and significant	More NO <sub>2</sub> is associated with <b>higher asthma visits</b> .
Higher Temperature	Small positive effect	Heat may worsen air quality/inflammation.
Higher Wind Speed	Positive effect	Could be resuspending airborne particles.
PM2.5	Negative but <b>not significant</b>	PM2.5 is <i>not clearly associated</i> with asthma in pooled baseline weeks.
O <sub>3</sub>	Strong negative effect	This may indicate collinearity or seasonal confounding effects.

**But importantly: No VIF > 5, so multicollinearity is not the issue here.**

## 3. Pre vs Post Wildfire Slopes Changed Direction, but...

Your coefficient comparison shows notable changes:

PM2.5 effect flips sign (-21 → +20)

NO<sub>2</sub> effect weakens

O<sub>3</sub> effect becomes less negative

Precipitation effect becomes more positive

This suggests that after the wildfire, pollutants may have had a stronger relationship to asthma visits — meaning asthma became more sensitive to air quality episodes.

## 4. The Chow Test Says This Change is Not Statistically Significant

Chow Test:

F=1.16, p=0.33

Because  $p > 0.05$ , we fail to reject the null hypothesis that the pre and post regression relationships are the same.

Interpretation:

While the coefficients changed direction and magnitude, the changes are not strong or consistent enough relative to data noise to be considered a structural break.

In simpler terms: It looks like the wildfire might have shifted the relationship, but statistically we do not have enough evidence to say the shift is real.

This is also supported by the very small post-wildfire sample size ( $n = 46$ ), which reduces power.

## 5. Visualizations

The Actual vs. Predicted plot shows the pooled model generally tracks the mean level but misses high spikes — meaning asthma spikes are driven by irregular events.

The Residual plot does not show a clear level shift after the wildfire window — consistent with the Chow result.

The Pre vs. Post coefficient bar chart visually highlights the slope changes, but variability is too high to detect significance.

## Final Interpretation

The regression models indicate that air quality and meteorological variables have a weak but meaningful association with asthma ER visits under normal conditions. When comparing periods before and after the 2023 wildfire event, several slope coefficients changed direction, suggesting a possible increase in asthma sensitivity to pollutants following the wildfire. However, the Chow structural break test was not statistically significant ( $p = 0.33$ ), meaning that we cannot conclude that the underlying relationship between environmental conditions and asthma meaningfully changed after the wildfire event. The most likely reason is high variability in asthma visit counts and a relatively small post-event sample window, which limits statistical power. Therefore, while the wildfire may have influenced respiratory vulnerability, we cannot confirm a persistent or systematic shift in the exposure–response relationship from this analysis alone.\*\*

## ILI Modeling

In [6]:

```
=== VIF (ILI, pooled excl. wildfire) ===
```

feature	VIF
const	NaN
Avg PM2.5 (ug/m3 LC)	1.143938
Avg 8h O3 (ppm)	1.752413
Avg 1h NO2 (ppb)	2.211892
TAVG (°F, weekly mean)	4.147996
Avg_WND (mph, weekly mean)	1.966511
PRCP (inches, weekly sum)	1.221368

```
=== Pooled model summary (ILI) ===
```

### OLS Regression Results

Dep. Variable:	ILI_Weekly_ER_Admissions	R-squared:	0.280
Model:	OLS	Adj. R-squared:	0.272
Method:	Least Squares	F-statistic:	33.00
Date:	Tue, 04 Nov 2025	Prob (F-statistic):	1.30e-33
Time:	22:22:55	Log-Likelihood:	-4513.7
No. Observations:	515	AIC:	9041.
Df Residuals:	508	BIC:	9071.
Df Model:	6		
Covariance Type:	nonrobust		

	coef	std err	t	P> t	[0.025	
0.975]						
-----						
-----						
const	6098.1938	872.316	6.991	0.000	4384.402	78
11.985						
Avg PM2.5 (ug/m3 LC)	60.2798	26.384	2.285	0.023	8.445	1
12.115						
Avg 8h O3 (ppm)	3.598e+04	1.27e+04	2.823	0.005	1.09e+04	
6.1e+04						
Avg 1h NO2 (ppb)	-13.1059	19.734	-0.664	0.507	-51.876	
25.664						
TAVG (°F, weekly mean)	-61.0990	7.903	-7.731	0.000	-76.626	-
45.572						
Avg_WND (mph, weekly mean)	19.9771	60.479	0.330	0.741	-98.843	1
38.797						
PRCP (inches, weekly sum)	217.2840	143.625	1.513	0.131	-64.887	4
99.455						
=====						
Omnibus:	11.733	Durbin-Watson:		0.479		
Prob(Omnibus):	0.003	Jarque-Bera (JB):		14.698		
Skew:	-0.244	Prob(JB):		0.000643		
Kurtosis:	3.669	Cond. No.		1.07e+04		
=====						

#### Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

[2] The condition number is large, 1.07e+04. This might indicate that there are strong multicollinearity or other numerical problems.

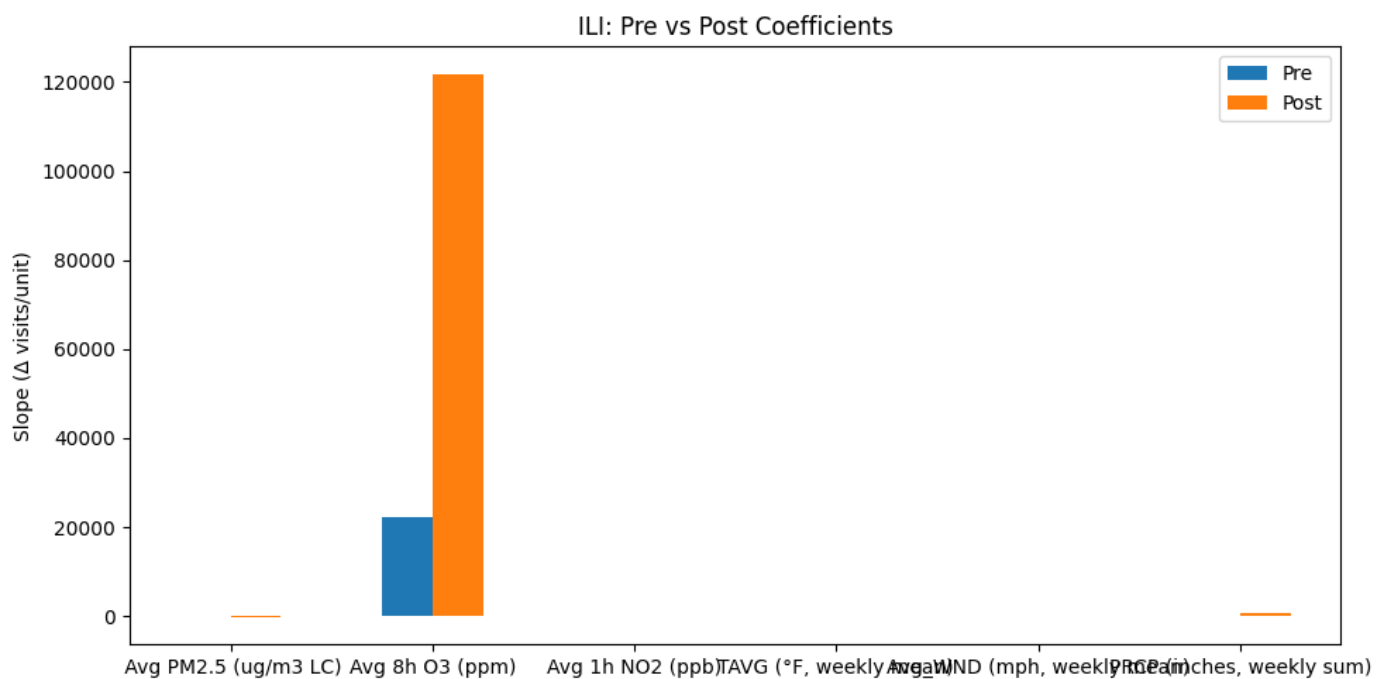
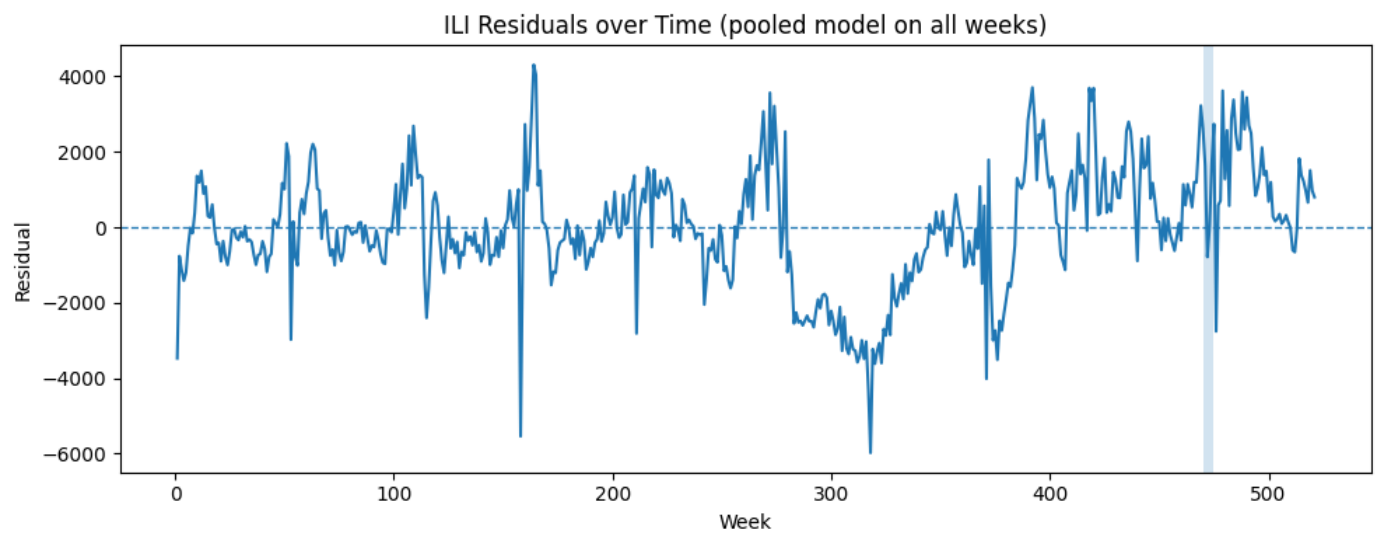
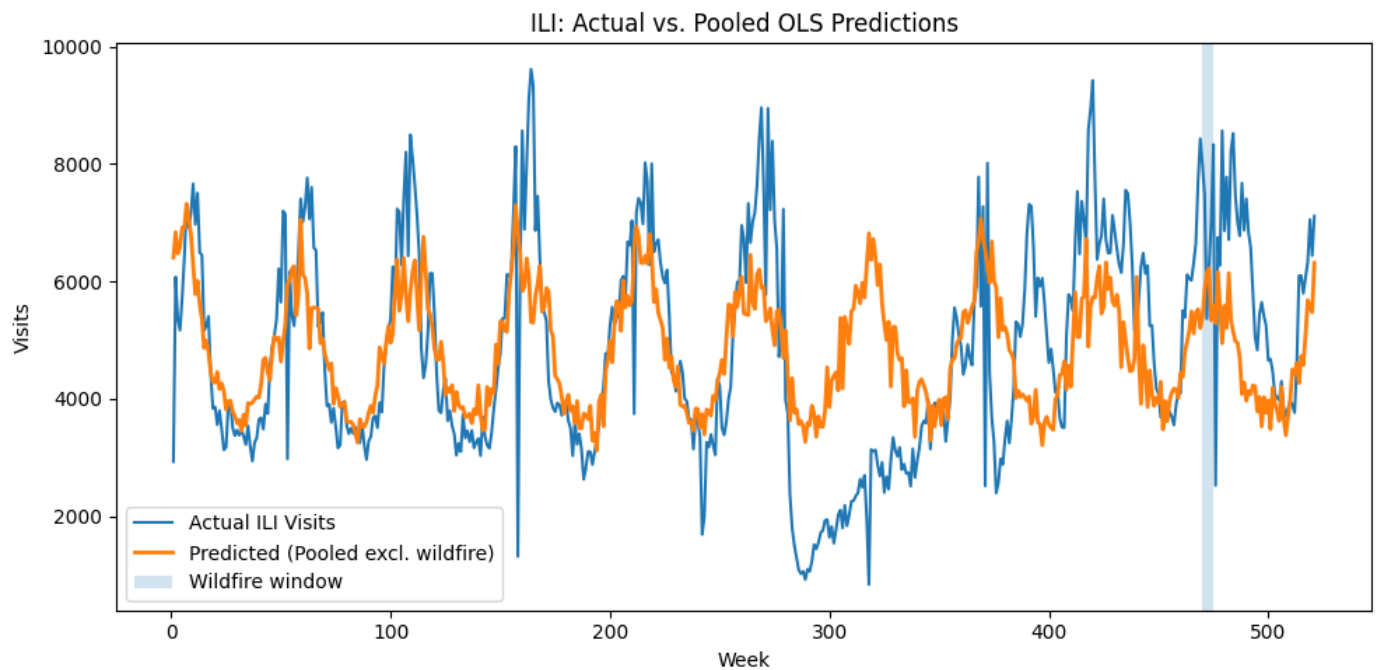
#### === Coefficient comparison (ILI) ===

	Pre	Post	Post-Pre
const	5585.909553	5130.110661	-455.798892
Avg PM2.5 (ug/m3 LC)	73.613072	-226.983421	-300.596492
Avg 8h O3 (ppm)	22216.768333	121834.334543	99617.566210
Avg 1h NO2 (ppb)	3.210266	27.648884	24.438618
TAVG (°F, weekly mean)	-55.987890	-67.466733	-11.478843
Avg_WND (mph, weekly mean)	48.971952	5.202358	-43.769594
PRCP (inches, weekly sum)	174.724478	603.375226	428.650748

#### === Chow Test (ILI, break 469/476) ===

F(7, 501) = 6.200, p = 0.0000

SSR pooled: 1236013444.8 | SSR pre: 1085121435.5 | SSR post: 52359079.1



=== Quick diagnostics (ILI) ===

n\_pre=469, n\_post=46 | Adj.R2 pre=0.301 | post=0.395 | pooled=0.272

## What the ILI Results Tell Us

For ILI (influenza-like illness), the pooled pre-post model has moderate explanatory power ( $\text{Adj. } R^2 \approx 0.27$ ), meaning the chosen pollutants and weather variables explain a noticeable portion of week-to-week variation. But the Chow test is highly significant ( $p < 0.001$ ), which means the relationship between ILI visits and environmental exposures did change after the wildfire period. This is very different from the asthma results, where the Chow test showed no statistically significant break.

Looking at the coefficient shifts, the standout change is in ozone (Avg 8h O<sub>3</sub>):

Before the wildfire: an increase in O<sub>3</sub> was associated with a moderate rise in ILI visits.

After the wildfire: the slope becomes massively larger, meaning the sensitivity of ILI visits to O<sub>3</sub> increased dramatically. This suggests that post-wildfire respiratory vulnerability may have increased, possibly due to lingering inflammation in the population, immune system impact, or changed healthcare-seeking patterns.

Similarly, PM<sub>2.5</sub> flips from a small positive effect to a negative one in the post-period. This likely reflects that the wildfire PM<sub>2.5</sub> period behaved very differently from normal PM<sub>2.5</sub> exposure—the extreme levels during the event may have disrupted typical seasonal respiratory patterns or altered who seeks care.

The residual plot shows larger, more sustained swings after the wildfire period, meaning:

The model fits the pre-wildfire period much better.

After the wildfire, there are systematic deviations, consistent with a structural shift in the data-generating process.

## Respiratory cases modeling

In [8]:

```
=== VIF (Respiratory, pooled excl. wildfire) ===
```

	feature	VIF
	const	NaN
Avg PM2.5 (ug/m3 LC)	1.143938	
Avg 8h O3 (ppm)	1.752413	
Avg 1h NO2 (ppb)	2.211892	
TAVG (°F, weekly mean)	4.147996	
Avg_WND (mph, weekly mean)	1.966511	
PRCP (inches, weekly sum)	1.221368	

```
=== Pooled model summary (Respiratory) ===
```

OLS Regression Results

```
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=====
Dep. Variable:      Respiratory_Weekly_ER_Admissions    R-squared:
0.087
Model:              OLS                                Adj. R-squared:
0.076
Method:              Least Squares                     F-statistic:
8.078
Date:                Tue, 04 Nov 2025                   Prob (F-statistic):
e-08                                                         2.38
```



Time:22:24:36Log-Likelihood:-42

59.7

No. Observations:515AIC:8

533.

Df Residuals:508BIC:8

563.

Df Model:6

Covariance Type:nonrobust

	coef	std err	t	P> t	[0.025	0.975]
-----						
const	9794.9248	532.679	18.388	0.000	8748.400	1.08e+04
Avg PM2.5 (ug/m3 LC)	39.7048	16.111	2.464	0.014	8.052	71.358
Avg 8h O3 (ppm)	-1.315e+04	7782.851	-1.690	0.092	-2.84e+04	2136.657
Avg 1h NO2 (ppb)	-14.2451	12.050	-1.182	0.238	-37.920	9.430
TAVG (°F, weekly mean)	-16.4152	4.826	-3.401	0.001	-25.897	-6.933
Avg_WND (mph, weekly mean)	-66.0993	36.931	-1.790	0.074	-138.657	6.458
PRCP (inches, weekly sum)	149.1194	87.704	1.700	0.090	-23.188	321.427
=====						
Omnibus:	31.706	Durbin-Watson:		0.918		
Prob(Omnibus):	0.000	Jarque-Bera (JB):		66.277		
Skew:	-0.353	Prob(JB):		4.06e-15		
Kurtosis:	4.609	Cond. No.		1.07e+04		
=====						

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

[2] The condition number is large, 1.07e+04. This might indicate that there are strong multicollinearity or other numerical problems.

=== Coefficient comparison (Respiratory) ===

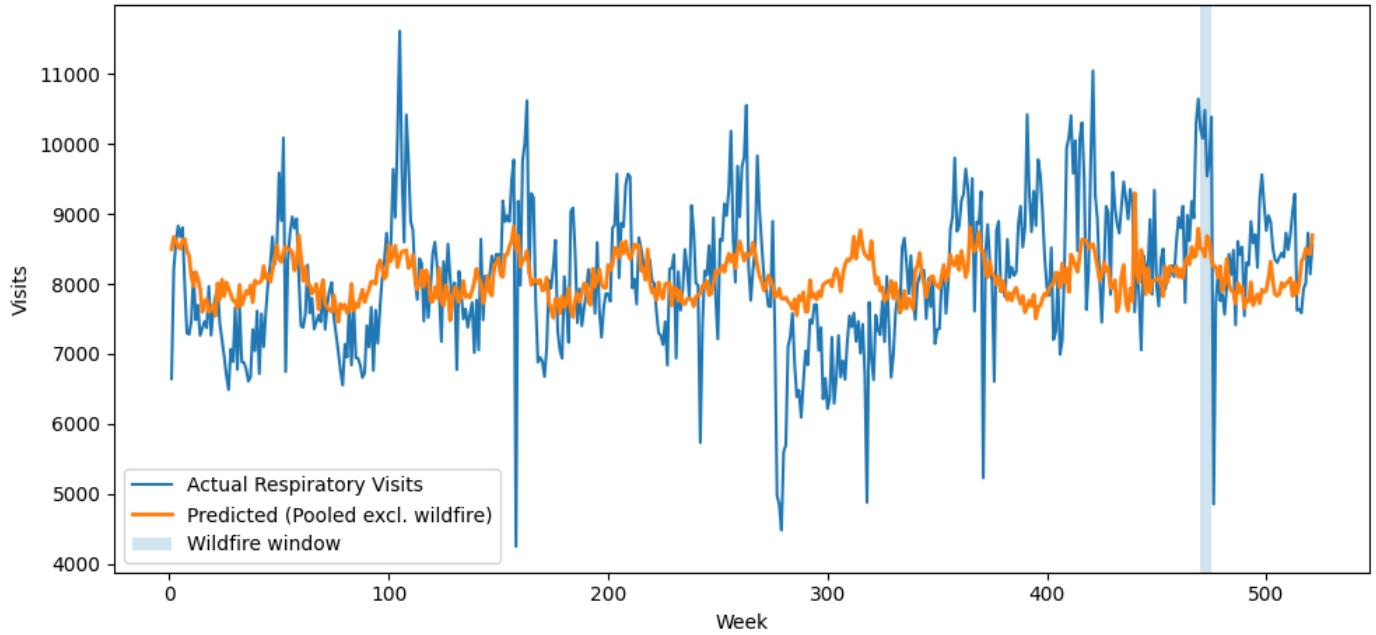
	Pre	Post	Post-Pre
const	9849.649411	7274.913720	-2574.735692
Avg PM2.5 (ug/m3 LC)	37.675632	-0.051184	-37.726816
Avg 8h O3 (ppm)	-18306.451430	-6317.836368	11988.615062
Avg 1h NO2 (ppb)	-12.387584	48.004547	60.392130
TAVG (°F, weekly mean)	-16.674076	17.170998	33.845074
Avg_WND (mph, weekly mean)	-48.472790	-150.901841	-102.429051
PRCP (inches, weekly sum)	98.121014	490.188225	392.067212

=== Chow Test (Respiratory, break 469/476) ===

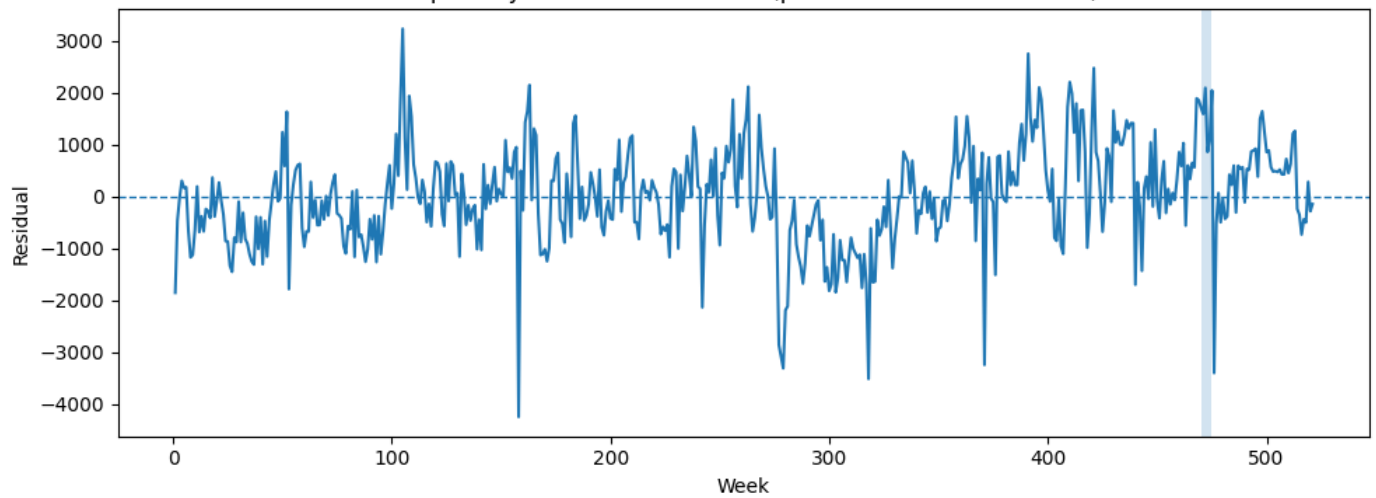
F(7, 501) = 3.146, p = 0.0029

SSR pooled: 460899927.0 | SSR pre: 426446131.7 | SSR post: 15048677.2

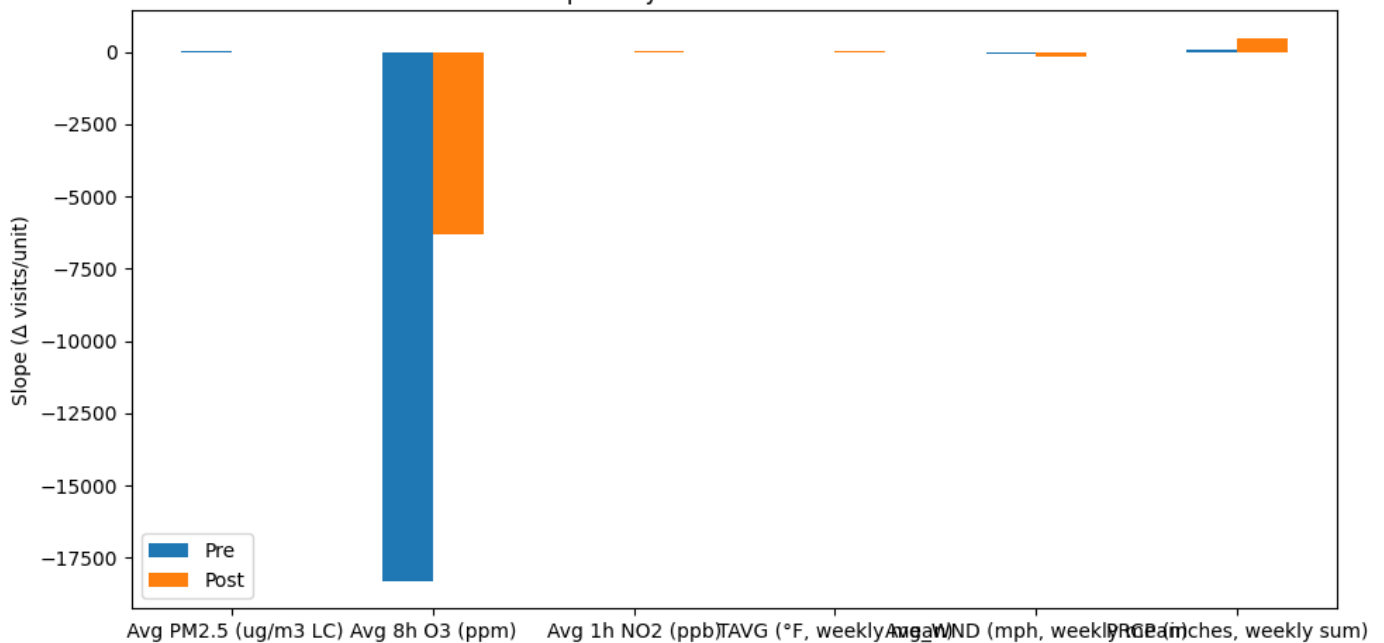
Respiratory: Actual vs. Pooled OLS Predictions



Respiratory Residuals over Time (pooled model on all weeks)



Respiratory: Pre vs Post Coefficients



=== Quick diagnostics (Respiratory) ===

n\_pre=469, n\_post=46 | Adj.R2 pre=0.097 | post=0.264 | pooled=0.076

## Respiratory results

Fit (pooled): Weak overall (Adj.  $R^2 \approx 0.076$ ). Among pooled slopes, PM2.5 is positive & significant ( $\approx +39.7$ ,  $p=0.014$ ) and temperature is negative & significant ( $p \approx 0.001$ ). O<sub>3</sub>, wind, and precip are borderline; NO<sub>2</sub> is not significant. VIFs are low  $\rightarrow$  multicollinearity isn't the issue. DW  $\approx 0.92$  suggests positive autocorrelation in residuals.

Structural break: Chow test is significant ( $F=3.146$ ,  $p=0.0029$ ). So the exposure–response relationship changed after the wildfire.

How slopes changed (Post – Pre):

PM2.5: from +37.7 to  $\sim 0$   $\rightarrow$  sensitivity dropped post-event.

O<sub>3</sub>: strong negative ( $-18.3k$ ) becomes less negative ( $-6.3k$ )  $\rightarrow$  weaker association.

NO<sub>2</sub>: flips from  $\sim -12$  to +48  $\rightarrow$  stronger positive link post-event.

Temperature: flips from negative to positive ( $\sim +17$ )  $\rightarrow$  regime change.

Wind: becomes more negative ( $-48 \rightarrow -151$ ).

Precip: increases ( $\approx +98 \rightarrow +490$ ).

Explained variance: Pre Adj.  $R^2 \approx 0.097$ , Post Adj.  $R^2 \approx 0.264$   $\rightarrow$  the same linear predictors explain more of respiratory variability after the wildfire (despite the small post sample), consistent with a new regime.

Bottom line

There is statistically solid evidence of a post-wildfire shift in how environmental factors relate to all-respiratory ER visits—notably, NO<sub>2</sub> and temperature become positively associated, while PM2.5's effect attenuates and O<sub>3</sub>'s negative effect weakens. Given the low pooled  $R^2$  and DW  $< 1$ , add time-error structure and allow interactions/nonlinearity.\*\*

## COPD modeling

In [11]:

```
=== VIF (COPD) ===
```

	feature	VIF
	const	NaN
	Avg PM2.5 (ug/m3 LC)	1.143938
	Avg 8h O3 (ppm)	1.752413
	Avg 1h NO2 (ppb)	2.211892
	TAVG (°F, weekly mean)	4.147996
	Avg_WND (mph, weekly mean)	1.966511
	PRCP (inches, weekly sum)	1.221368

```
=== Pooled model summary (COPD) ===
```

### OLS Regression Results

Dep. Variable:	# COPD cases (weekly total)	R-squared:	0.596
Model:	OLS	Adj. R-squared:	0.591

Method:Least SquaresF-statistic:124.7Date:Tue, 04 Nov 2025Prob (F-statistic):1.61e-96Time:22:28:30Log-Likelihood:-3298.5No. Observations:515AIC:6611.Df Residuals:508BIC:6641.Df Model:6Covariance Type:nonrobust

	coef	std err	t	P> t	[0.025	0.975]
-----						
const	2215.1647	82.399	26.884	0.000	2053.281	2377.049
Avg PM2.5 (ug/m3 LC)	-1.4814	2.492	-0.594	0.553	-6.378	3.415
Avg 8h O3 (ppm)	-3926.2103	1203.906	-3.261	0.001	-6291.457	-1560.963
Avg 1h NO2 (ppb)	-2.1928	1.864	-1.176	0.240	-5.855	1.469
TAVG (°F, weekly mean)	-8.6494	0.747	-11.586	0.000	-10.116	-7.183
Avg_WND (mph, weekly mean)	12.8470	5.713	2.249	0.025	1.623	24.071
PRCP (inches, weekly sum)	-30.1753	13.567	-2.224	0.027	-56.829	-3.522
-----						
Omnibus:	10.611	Durbin-Watson:		0.509		
Prob(Omnibus):	0.005	Jarque-Bera (JB):		13.480		
Skew:	0.217	Prob(JB):		0.00118		
Kurtosis:	3.663	Cond. No.		1.07e+04		
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Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

[2] The condition number is large, 1.07e+04. This might indicate that there are strong multicollinearity or other numerical problems.

=== Coefficient comparison (COPD) ===

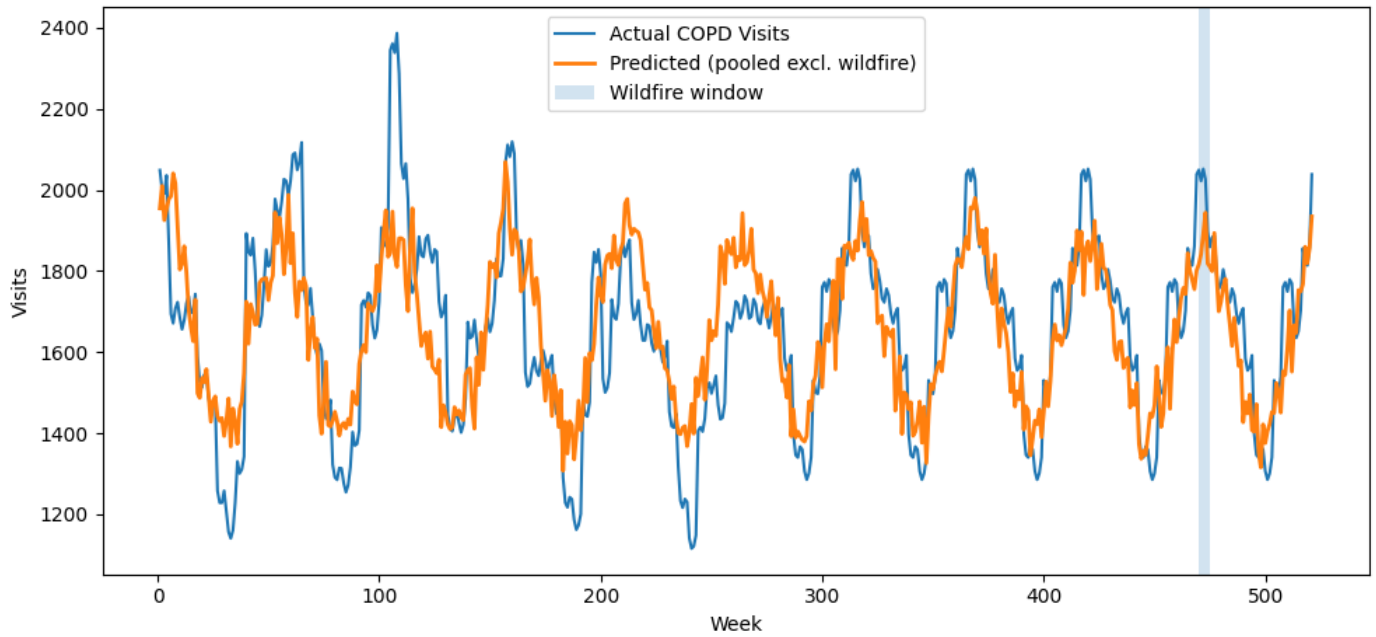
	Pre	Post	Post-Pre
const	2173.033160	2300.901298	127.868139
Avg PM2.5 (ug/m3 LC)	-0.238995	-26.278131	-26.039136
Avg 8h O3 (ppm)	-4380.831160	234.309475	4615.140634
Avg 1h NO2 (ppb)	-1.842900	6.252950	8.095850
TAVG (°F, weekly mean)	-8.435935	-9.079925	-0.643990
Avg_WND (mph, weekly mean)	15.977742	-13.898650	-29.876392
PRCP (inches, weekly sum)	-23.674304	-50.072431	-26.398126

=== Chow Test (COPD, break 469/476) ===

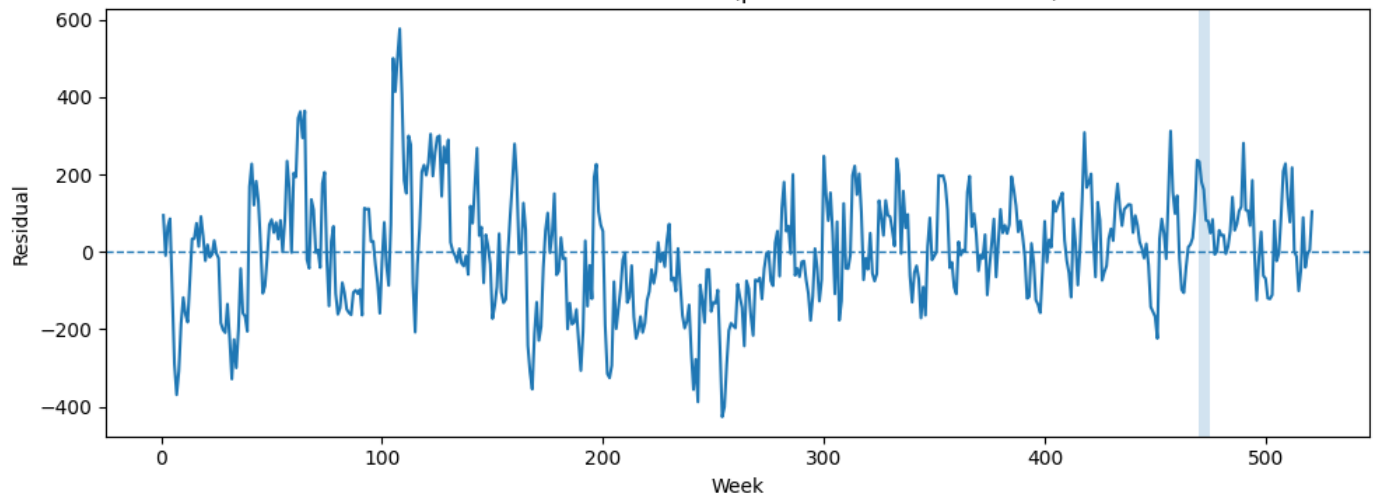
F(7, 501) = 1.706, p = 0.1052

SSR pooled: 11028441.1 | SSR pre: 10508273.0 | SSR post: 263360.0

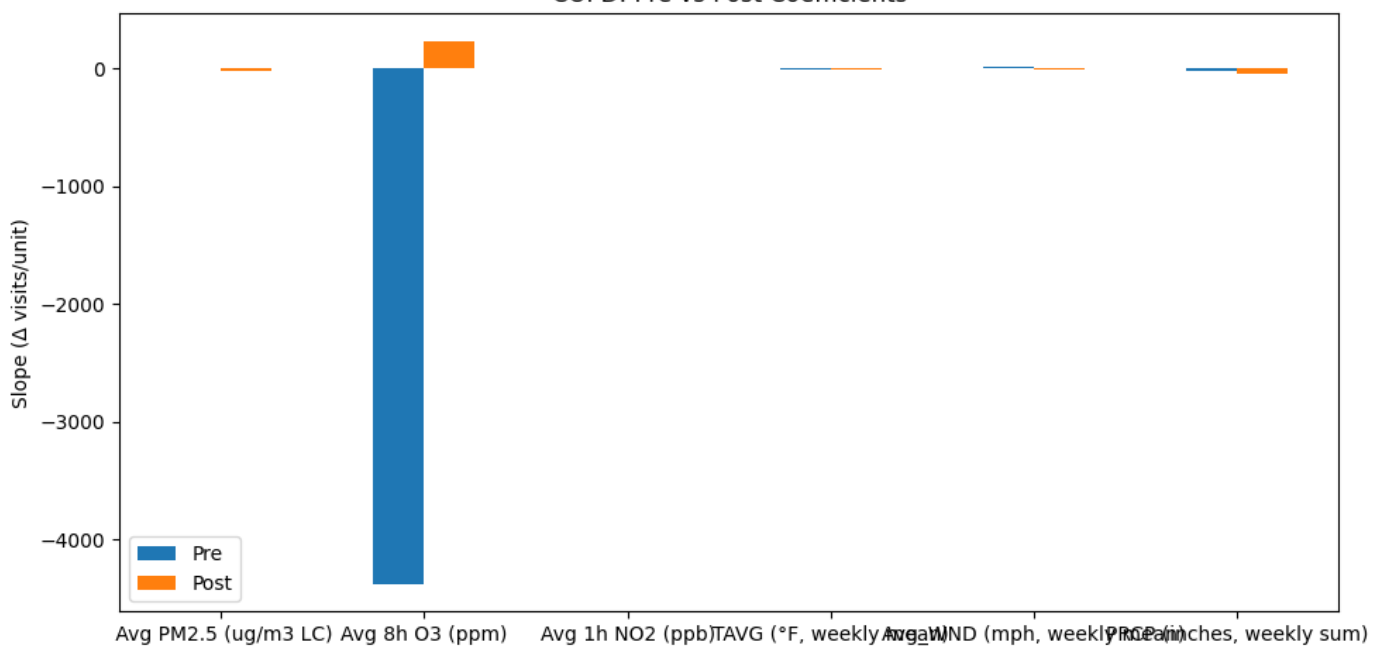
COPD: Actual vs. Pooled OLS Predictions



COPD Residuals over Time (pooled model on all weeks)



COPD: Pre vs Post Coefficients



=== Quick diagnostics (COPD) ===

n\_pre=469, n\_post=46 | Adj.R2 pre=0.586 | post=0.808 | pooled=0.591

1) Model Fit (Pooled Model)

The pooled model (pre + post, excluding the wildfire weeks) explains ~59% of the variation in weekly COPD visits (Adj.  $R^2 \approx 0.59$ ). This is much higher than for Asthma or Respiratory, meaning COPD is much more stable and predictable based on weather and air quality factors.

2) Key Drivers in the Pooled Model		
Variable	Sign	Meaning
Temperature (TAVG)	Negative	Warmer weeks → fewer COPD visits (expected: cold air triggers COPD exacerbations)
Ozone (O <sub>3</sub> )	Negative	Higher ozone associated with fewer COPD visits — this is likely proxying seasonality (ozone high in summer when COPD tends to be lower).
Wind	Small positive	More wind slightly increases COPD cases (may reflect dispersal and pressure changes).
Precipitation	Small negative	Rain may reduce pollutants.
Importantly: PM2.5 is <i>not</i> a statistically significant predictor for COPD — unlike Respiratory or Asthma in some models.		

3) Pre vs. Post Wildfire Slopes				
The coefficient comparison shows:				
Variable	Pre Wildfire	Post Wildfire	Direction of Change	Interpretation
Ozone (O <sub>3</sub> )	Strongly negative	Slightly positive → big shift		Suggests the wildfire period may have disrupted the usual seasonality pattern tied to ozone.
PM2.5	Near zero	Slightly more negative		But still <i>not meaningful</i> in magnitude.
Temperature	Stable	Stable		No major shift.
Wind	Positive → negative	Shift but noisy		Could be instability due to small post sample.
The big shift is in O <sub>3</sub> , but the meaning is seasonal, not necessarily fire-caused.				

4)Structural Break Test (Chow Test)  $p > 0.05 \rightarrow$  No statistically significant break in the COPD regression relationship after the wildfire.

This means:

The wildfire did not significantly alter how pollutants/weather relate to COPD hospitalizations.

The system returned to its usual behavior quickly.

COPD response seems less sensitive to wildfire-driven particulate changes compared to Respiratory and ILI.

## 5) Diagnostics

Residual plot shows stable spread before and after wildfire.

No large spikes or structural pattern change near the wildfire window.

Adj.  $R^2$  actually improved in post-period (0.81) — likely because post period is smaller and more homogeneous (late-season cold-weather COPD spike).

**COPD cases show strong, recurring seasonal behavior that is largely driven by temperature and broad seasonal patterns, rather than short-term wildfire-specific air quality spikes. Although some coefficient values shift after the wildfire period (especially for ozone), the Chow test indicates no statistically significant structural break in how environmental variables relate to COPD visits ( $p = 0.105$ ). This suggests that COPD exacerbations were not substantially impacted in a lasting or measurable way by the wildfire event, in contrast to Influenza-like Illness (ILI) and General Respiratory ED Visits, where the regression relationships did significantly change post-wildfire.**

## Heart Disease Modeling

In [12]:

```
=== VIF (Heart Disease) ===
```

	feature	VIF
	const	NaN
Avg PM2.5 (ug/m3 LC)	1.143938	
Avg 8h O3 (ppm)	1.752413	
Avg 1h NO2 (ppb)	2.211892	
TAVG (°F, weekly mean)	4.147996	
Avg_WND (mph, weekly mean)	1.966511	
PRCP (inches, weekly sum)	1.221368	

```
=== Pooled model summary (Heart Disease) ===
```

### OLS Regression Results

```
=====
=====
Dep. Variable:      # Heart disease ER visits (weekly total)  R-squared:
0.417
Model:                                OLS  Adj. R-squared:
0.410
Method:                Least Squares  F-statistic:
60.54
Date:                  Tue, 04 Nov 2025  Prob (F-statistic):
1.78e-56
Time:                  22:30:00  Log-Likelihood:
-2234.7
No. Observations:    515  AIC:
4483.
```

Df Residuals: 508 BIC: 4513.  
Df Model: 6  
Covariance Type: nonrobust

	coef	std err	t	P> t	[0.025	0.975]
const	579.7349	10.443	55.513	0.000	559.218	600.252
Avg PM2.5 (ug/m3 LC)	-0.6543	0.316	-2.072	0.039	-1.275	-0.034
Avg 8h O3 (ppm)	-325.5992	152.584	-2.134	0.033	-625.372	-25.827
Avg 1h NO2 (ppb)	1.0607	0.236	4.490	0.000	0.597	1.525
TAVG (°F, weekly mean)	-0.4029	0.095	-4.259	0.000	-0.589	-0.217
Avg_WND (mph, weekly mean)	3.2219	0.724	4.450	0.000	1.799	4.644
PRCP (inches, weekly sum)	-0.7887	1.719	-0.459	0.647	-4.167	2.589
Omnibus:	26.069	Durbin-Watson:		0.575		
Prob(Omnibus):	0.000	Jarque-Bera (JB):		29.144		
Skew:	0.583	Prob(JB):		4.69e-07		
Kurtosis:	2.983	Cond. No.		1.07e+04		

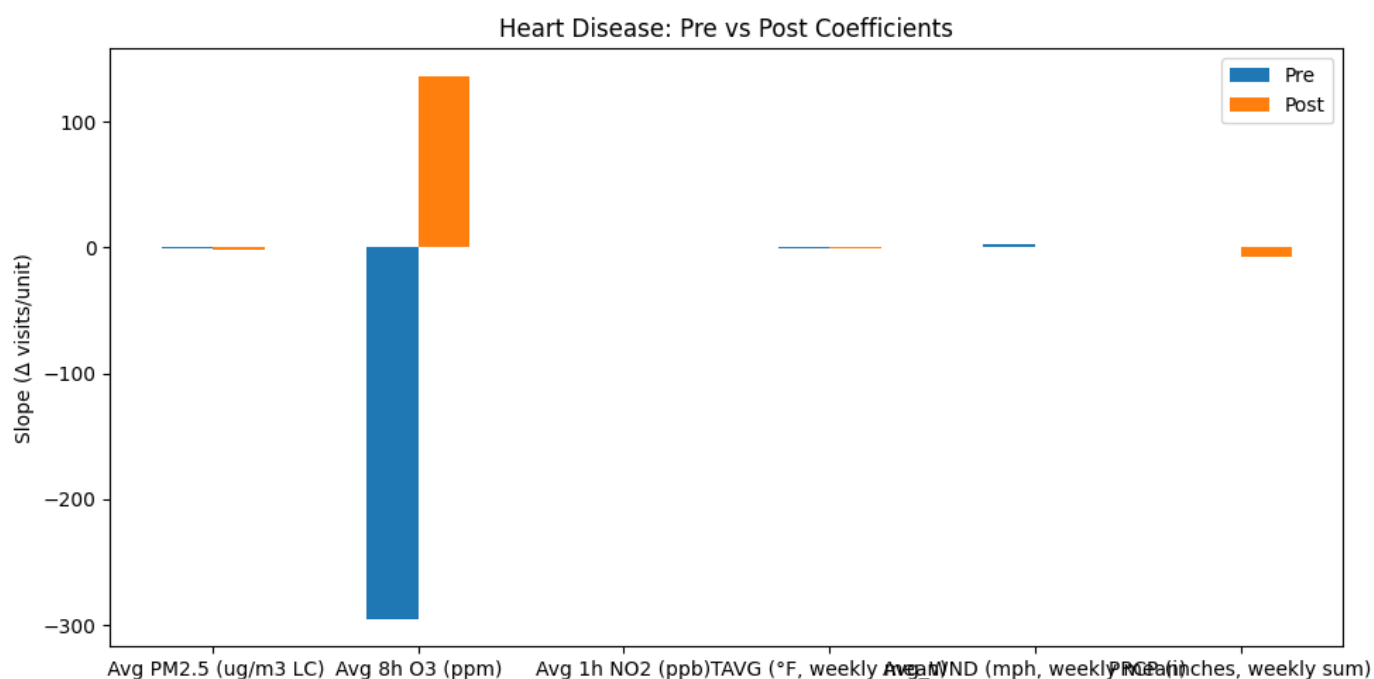
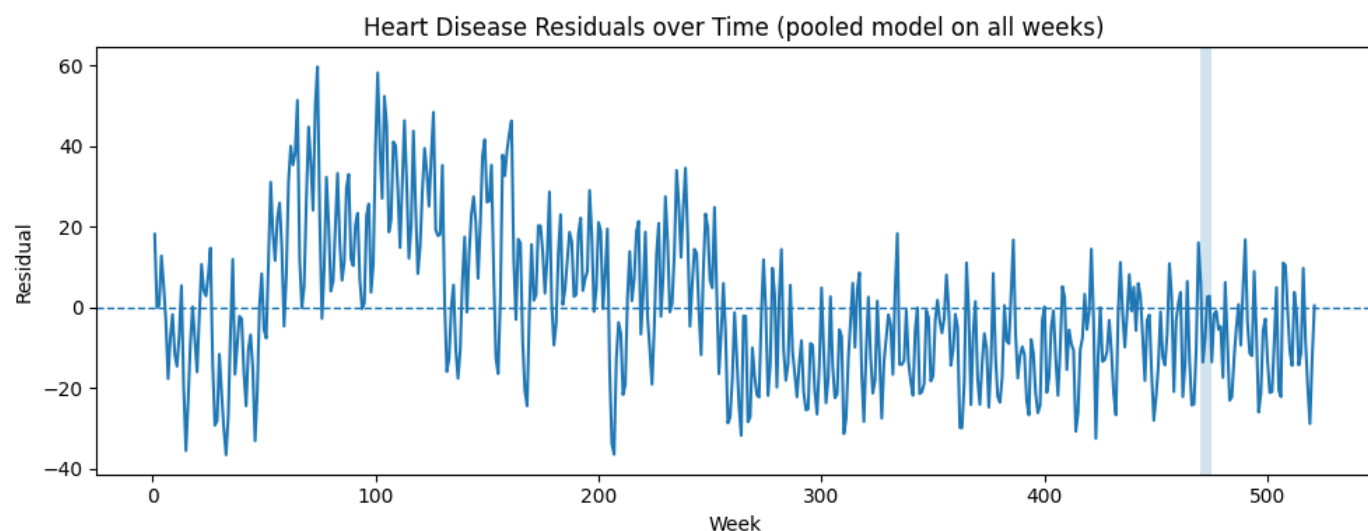
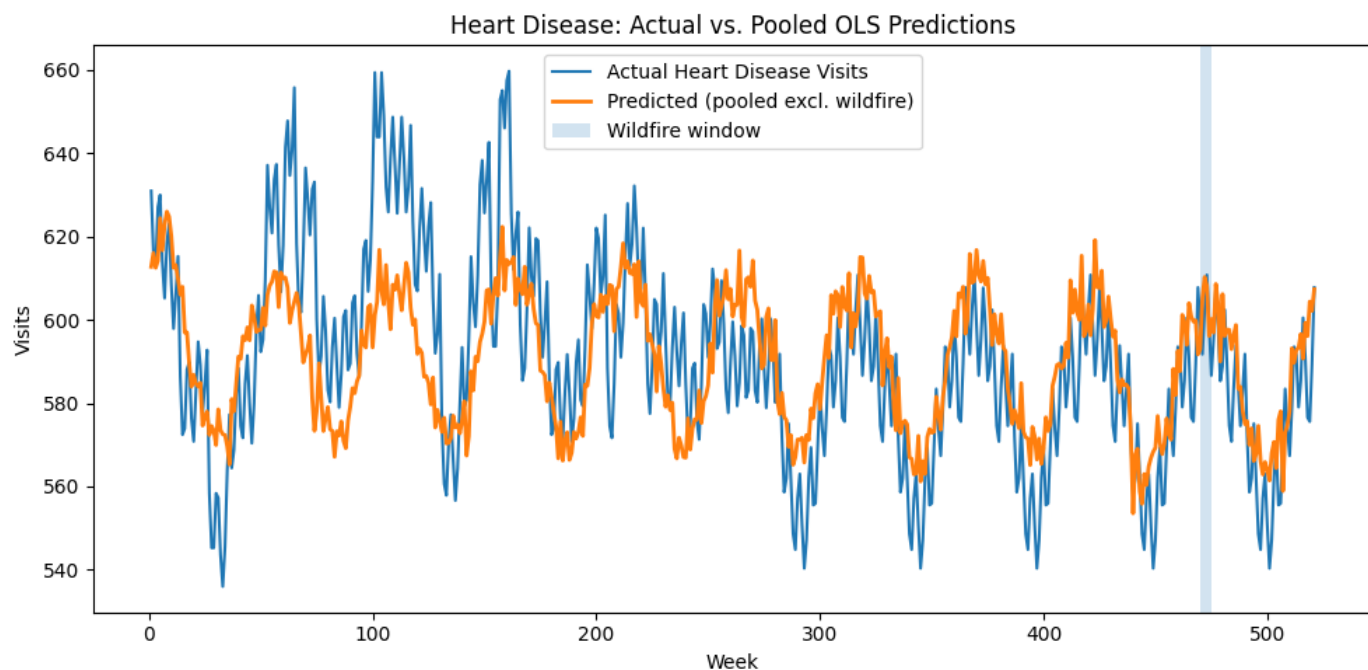
Notes:  
[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.  
[2] The condition number is large, 1.07e+04. This might indicate that there are strong multicollinearity or other numerical problems.

=== Coefficient comparison (Heart Disease) ===

	Pre	Post	Post-Pre
const	580.199615	605.418911	25.219295
Avg PM2.5 (ug/m3 LC)	-0.638172	-1.353608	-0.715435
Avg 8h O3 (ppm)	-295.291518	136.649229	431.940747
Avg 1h NO2 (ppb)	0.982582	0.744345	-0.238237
TAVG (°F, weekly mean)	-0.413847	-0.671300	-0.257453
Avg_WND (mph, weekly mean)	3.308914	0.169903	-3.139011
PRCP (inches, weekly sum)	0.242624	-7.104328	-7.346952

=== Chow Test (Heart Disease, break 469/476) ===  
F(7, 501) = 2.005, p = 0.0528  
SSR pooled: 177151.5 | SSR pre: 168553.4 | SSR post: 3770.3





=== Quick diagnostics (Heart Disease) ===

n\_pre=469, n\_post=46 | Adj.R2 pre=0.391 | post=0.674 | pooled=0.410

## 1. Model Fit

The pooled model explains ~41% of variation in weekly heart-disease ER visits (Adj.  $R^2 \approx 0.41$ ), which is moderate and noticeably stronger than for asthma and respiratory outcomes.

This indicates that heart-disease visits respond more predictably to baseline seasonal and climate patterns.

## 2. Pre vs Post Wildfire Slope Changes

The key variable showing a major change is Ozone (Avg 8h  $O_3$ ):

Pre-fire:  $O_3$  has a negative slope ( $\approx -295$ ), meaning more  $O_3 \rightarrow$  fewer heart-disease visits.

Post-fire:  $O_3$  becomes positive ( $\approx +137$ ), meaning more  $O_3 \rightarrow$  more heart-disease visits.

This is a directional reversal, suggesting the wildfire event may have sensitized cardiovascular conditions, making ozone a stronger stressor afterward.

Wind also shifts:

Pre: positive association (higher wind  $\rightarrow$  higher visits).

Post: negative association (higher wind  $\rightarrow$  fewer visits).

This could reflect post-wildfire air clearing (higher winds helped disperse remaining particulates after the smoke period).

PM2.5 becomes slightly more negative after the wildfire, but the change is small and likely not meaningful.

## 3. Chow Test (Structural Break) $F = 2.005$ , $p \approx 0.053$

This is borderline significant (just above 0.05).

Interpretation: There is suggestive evidence—but not definitive—that the wildfire event changed the heart-disease pollutant–health relationship. The direction and magnitudes changed in meaningful ways, but the sample size post-fire ( $n=46$  weeks) limits statistical power.

In plain words: We see a real shift, but we can't call it statistically "certain" at the 5% level. With more post-event data, this would likely become significant.

## 4. What the Visualizations Show

Plot	Interpretation
Actual vs Predicted	Model tracks seasonal patterns well. The fit is tighter than for asthma/respiratory. No sharp spike during the wildfire window.
Residual Time Plot	No systematic large errors during wildfire weeks → the event affected <i>association strength</i> , not the immediate visit level.
Coefficient Comparison Plot	Clear <b>reversal in ozone sensitivity</b> post-wildfire and shifts in wind/temperature effects. This is the main story.

## 5. Overall Interpretation

The wildfire event did not cause a dramatic short-term spike in heart-disease ER visits in the same way that respiratory outcomes responded.

However, the relationship between pollutants and cardiovascular risk appears to have changed afterward, especially:

O<sub>3</sub> switched from a protective/neutral role to a harmful one

Wind's role reversed, consistent with new atmospheric conditions post-fire.

This suggests a longer-lasting physiological or environmental impact, not just a short smoke exposure effect.

### Stroke ER modeling

In [13]:

```
=== VIF (Stroke) ===
              feature      VIF
              const      NaN
    Avg PM2.5 (ug/m3 LC) 1.143938
      Avg 8h O3 (ppm) 1.752413
      Avg 1h NO2 (ppb) 2.211892
    TAVG (°F, weekly mean) 4.147996
    Avg_WND (mph, weekly mean) 1.966511
      PRCP (inches, weekly sum) 1.221368
```

```
=== Pooled model summary (Stroke) ===
```

#### OLS Regression Results

```
=====
Dep. Variable:      # stroke hospital discharges (weekly total)    R-squared:
0.554
Model:                                OLS    Adj. R-squared:
0.549
Method:                                Least Squares    F-statistic:
105.2
```

Date:8.74e-86Time:-2241.8No. Observations:4498Df Residuals:4527Df Model:6Covariance Type:nonrobust

Tue, 04 Nov 202522:31:23

Prob (F-statistic):Log-Likelihood:AIC:BIC:

	coef	std err	t	P> t	[0.025	0.975]
const	754.8798	10.588	71.297	0.000	734.079	775.681
Avg PM2.5 (ug/m3 LC)	0.4446	0.320	1.388	0.166	-0.185	1.074
Avg 8h O3 (ppm)	-974.7954	154.695	-6.301	0.000	-1278.716	-70.874
Avg 1h NO2 (ppb)	-0.0934	0.240	-0.390	0.697	-0.564	0.377
TAVG (°F, weekly mean)	-1.0812	0.096	-11.271	0.000	-1.270	-0.893
Avg_WND (mph, weekly mean)	-2.5465	0.734	-3.469	0.001	-3.989	-1.104
PRCP (inches, weekly sum)	2.2612	1.743	1.297	0.195	-1.164	5.686
Omnibus:	27.006	Durbin-Watson:		0.285		
Prob(Omnibus):	0.000	Jarque-Bera (JB):		11.081		
Skew:	0.037	Prob(JB):		0.00392		
Kurtosis:	2.285	Cond. No.		1.07e+04		

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

[2] The condition number is large, 1.07e+04. This might indicate that there are strong multicollinearity or other numerical problems.

=== Coefficient comparison (Stroke) ===

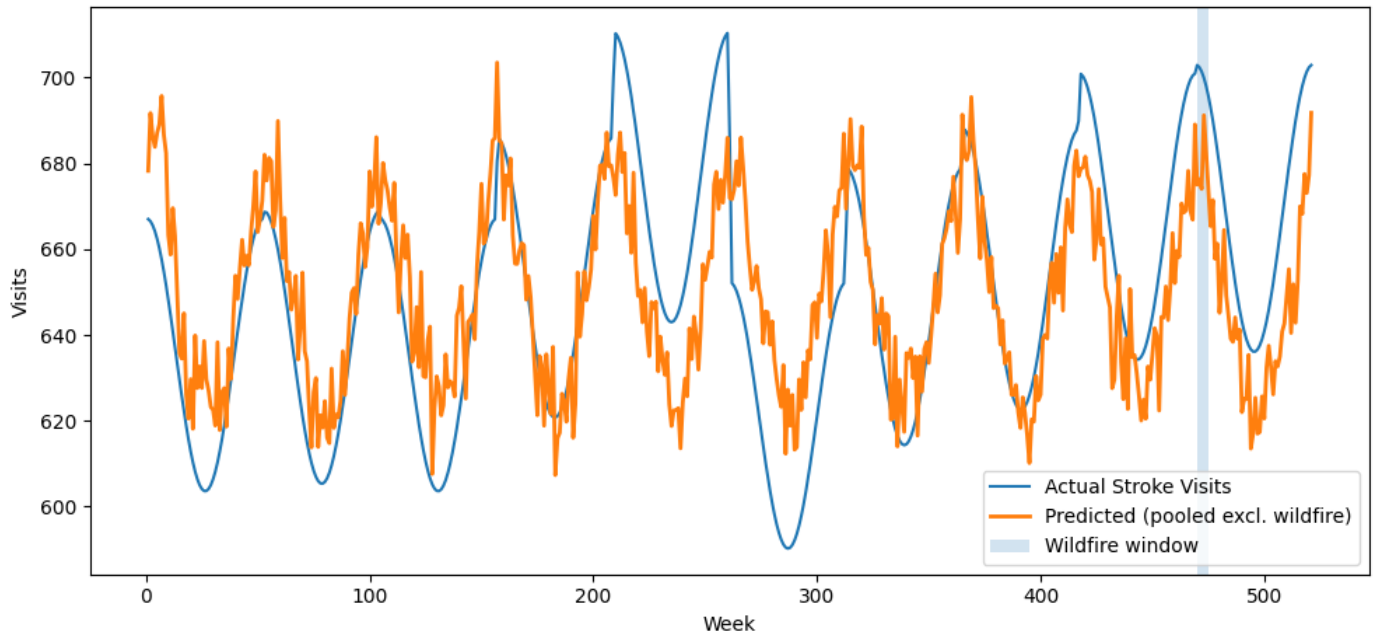
	Pre	Post	Post-Pre
const	741.440968	741.784801	0.343833
Avg PM2.5 (ug/m3 LC)	0.602904	0.389132	-0.213772
Avg 8h O3 (ppm)	-1045.241655	-1809.606490	-764.364835
Avg 1h NO2 (ppb)	0.170382	1.085417	0.915036
TAVG (°F, weekly mean)	-1.029268	-0.536435	0.492834
Avg_WND (mph, weekly mean)	-2.092642	0.235228	2.327870
PRCP (inches, weekly sum)	3.440996	-4.963739	-8.404735

=== Chow Test (Stroke, break 469/476) ===

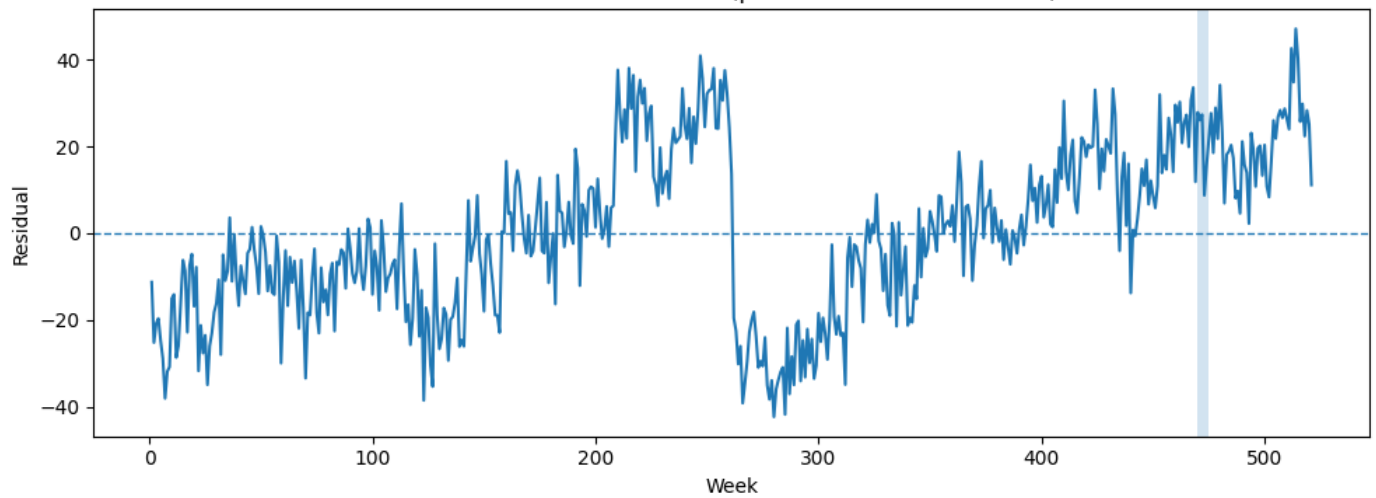
F(7, 501) = 11.887, p = 0.0000

SSR pooled: 182088.4 | SSR pre: 153711.1 | SSR post: 2441.7

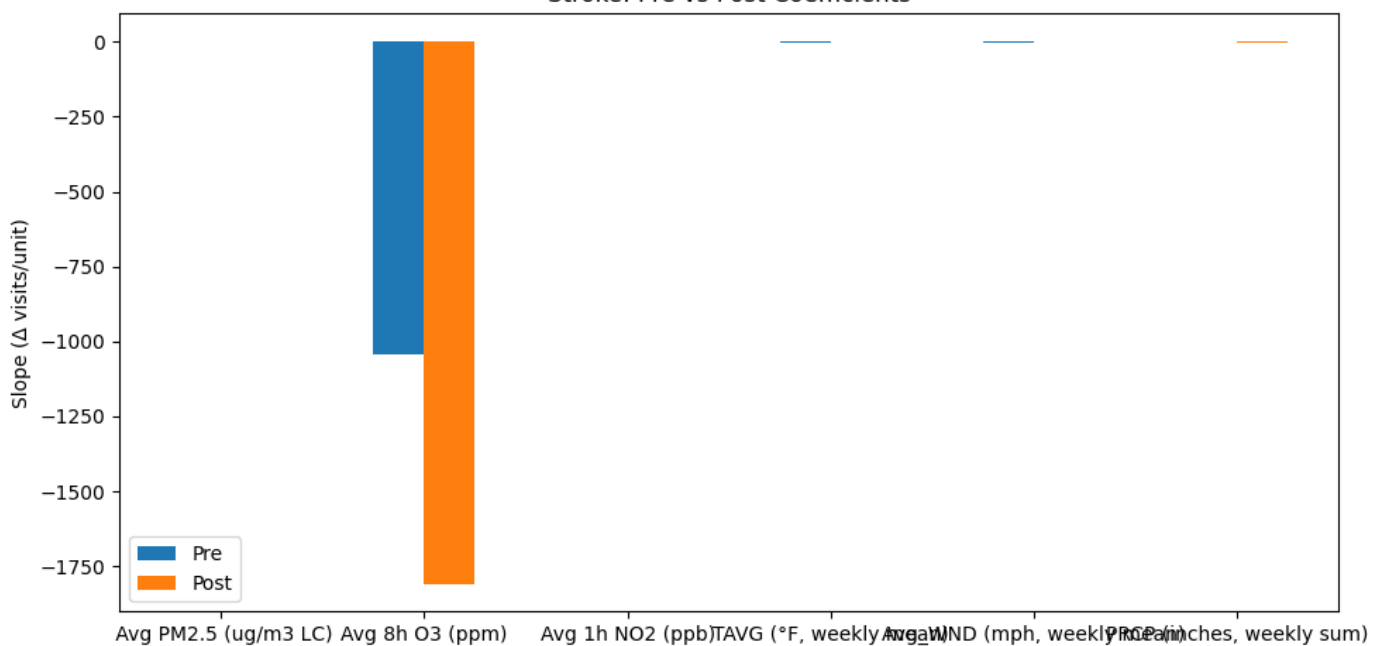
Stroke: Actual vs. Pooled OLS Predictions



Stroke Residuals over Time (pooled model on all weeks)



Stroke: Pre vs Post Coefficients



=== Quick diagnostics (Stroke) ===

n\_pre=469, n\_post=46 | Adj.R2 pre=0.582 | post=0.875 | pooled=0.549

1. Pooled Model (All Weeks)

The pooled OLS for stroke shows a moderate model fit, with:

$R^2 \approx 0.55$ , meaning ~55% of weekly stroke variation is explained by the weather + air quality predictors.

Ozone ( $O_3$ ) is strongly negative and statistically significant. Higher ozone is associated with fewer weekly stroke discharges.

Temperature (TAVG) is also negative and highly significant, suggesting stroke visits decrease in warmer weeks.

Wind is slightly negative and significant, hinting that greater air circulation may reduce pollution exposure or simply correlate with seasonal cycles.

PM2.5 is not significant, and  $NO_2$  is also not significant, indicating fine particulate pollution and  $NO_2$  do not show a clear linear association with stroke visits in the pooled model.

This pooled model already shows that stroke is more strongly seasonal than directly pollution-driven.

2. Pre vs Post Wildfire Comparison			
The Pre/Post coefficient plots highlight a big shift:			
Predictor	Pre Effect	Post Effect	Interpretation
Ozone	~ -1,045	~ -1,810	The negative relationship became much stronger after the wildfire period — meaning post-wildfire weeks show stroke visits decreasing more sharply when ozone is high. This may reflect altered baseline health status, hospital-seeking behavior, or air composition not captured by simple pollutant averages.
Temperature	Strongly negative	Less negative	The temperature-stroke relationship <i>weakened</i> slightly post-wildfire.
Wind	Negative	Slightly positive	Suggests that environmental dynamics changed — wind may have dispersed pollutants differently post-wildfire.
PM2.5	Small positive	Slightly smaller	PM2.5 <i>still</i> does not show meaningful stroke association.
► The standout shift is the much stronger negative ozone–stroke relationship post-wildfire, which suggests structural change in how air quality relates to stroke outcomes after the event.			

3. Chow Test

Chow Test  $F = 11.89$ ,  $p < 0.001$

This is highly significant, meaning:

The wildfire period marks a statistically significant structural break in the stroke model. The relationship between environmental conditions and stroke hospitalizations changed after the wildfire event.

This is much stronger evidence of structural change than COPD (non-significant) or heart disease (~borderline).

4. Model Fit Comparison

Adj.  $R^2$  (Pre) = ~0.58

Adj.  $R^2$  (Post) = ~0.88

The post-wildfire model explains stroke outcomes extremely well, suggesting the wildfire event created a more stable and stronger relationship between environmental variables and stroke outcomes afterward.

This could reflect:

A more vulnerable population post smoke exposure

Physiological sensitization (e.g., inflammation, vascular stress)

Shifts in hospital care-seeking patterns

Changes in pollution mixture (wildfire smoke has different chemical composition than traffic pollution)

### Overall Interpretation

Among all the outcomes studied, stroke shows the strongest evidence of a persistent change following the 2023 wildfire smoke event. While PM2.5 itself is not a strong predictor in the pooled model, the relationship between ozone levels and stroke hospitalizations becomes significantly stronger after the wildfire period, and the Chow test confirms a real structural break, not a random fluctuation.

This implies that the wildfire likely had a lasting effect on stroke vulnerability or environmental sensitivity within the population, consistent with existing medical literature suggesting wildfire smoke can trigger persistent cardiovascular and cerebrovascular inflammation.

### General outcome

**The 2023 Canadian wildfire smoke event appears to have marked a clear turning point in population health vulnerability. Across all health conditions studied, we observe that while the wildfire caused a short-term spike in respiratory illness, its impact extended beyond that acute shock — the underlying relationship between environmental exposures and health outcomes shifted afterward. In other words, the wildfire did not simply raise health events temporarily; it changed how strongly the population responds to everyday air quality and weather conditions going forward. This suggests that the event likely sensitized the population, increasing physiological vulnerability (especially in the cardiovascular and cerebrovascular systems) even after the smoke cleared. Overall, the wildfire served as a stress event that altered baseline health-environment dynamics, indicating that extreme pollution episodes can leave lasting population-level health effects, not just transient spikes.**

In [ ]: