

# Sparse Models

CMPUT 466/551

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# Outline

- **Introduction to Dimension Reduction**
- **Linear Regression and Least Squares (Review)**
- **Subset Selection**
- **Shrinkage Method**
- **Beyond Lasso**

# Part 1: Introduction to Dimension Reduction

- **Introduction to Dimension Reduction**
  - **General notations**
  - **Motivations**
  - **Feature selection and feature extraction**
  - **Feature Selection**
    - **Wrapper method**
    - **Filter method**
    - **Embedded method**
  - **Feature Extraction**
    - **PCA, ICA...**
- Linear Regression and Least Squares (Review)
- Subset Selection
- Shrinkage Method
- Beyond Lasso

# General Notations

## Dataset

- $\mathbf{X}$ : columnwise centered  $N \times p$  matrix
  - $N$  : # samples,  $p$  : # features
- $\mathbf{y}$ :  $N \times 1$  vector of labels(classification) or continuous values(regression)

## Basic Model

- Linear Regression
  - Assumption: the regression function  $E(Y|X)$  is linear

$$f(X) = X\beta$$

- $\beta$ :  $p \times 1$  vector of coefficients

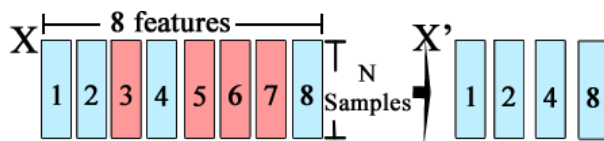
# Motivations

- Dimension reduction is about transforming data with high dimensionality into data of much lower dimensionality
  - **Computational efficiency**: less dimensions require less computations
  - **Accuracy**: lower risk of overfitting

- **Categories**

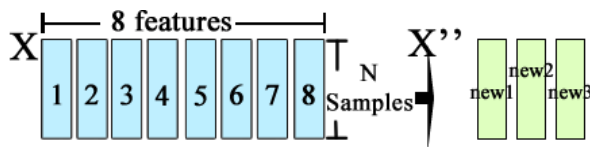
- Feature Selection:

- chooses a subset of features from the original feature set



- Feature Extraction:

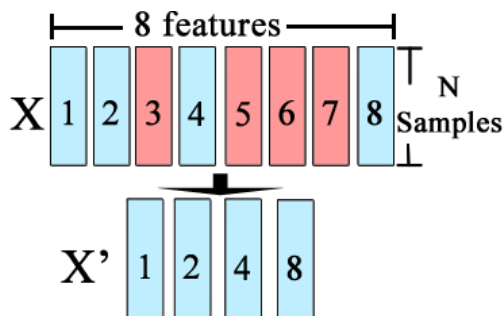
- transforms the original features into new ones, linearly or non-linearly
    - e.g. projects data from high dimensions to low dimensions



# Feature Selection and Feature Extraction

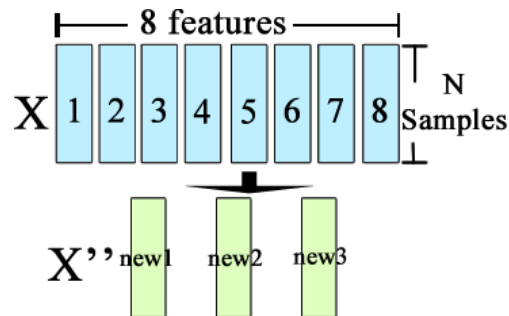
## Feature Selection

- Interpretation
- Cost constraint: computation, cost, etc.



## Feature Extraction

- More flexible. Feature selection is a special case of linear feature extraction



# Feature Selection and Feature Extraction

## Example 1: Prostate Cancer

- **Response:** level of prostate-specific antigen (*lpsa*).
- **Inputs:**

$\{lcavol, lweight, age, lbph, svi, lcp, gleason, pgg45\}.$

- **Task:**
  - predict *lpsa* from measurements of features

Feature selection applies better

- Cost: Measuring features cost money
- Interpretation: Doctors can see which features are important

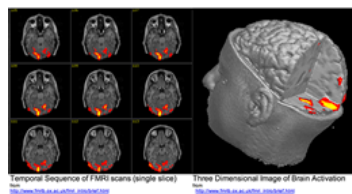
# Feature Selection and Feature Extraction

## Example 2: classification with fMRI data

- fMRI data are 4D images, with one dimension being the time slot.
- Each image is  $\sim 50 \times 50 \times 50$  (spatial)  $\times 200$  (times) =  $25m$  dimensions

Feature extraction applies better,

- Interpretation is not very important in this task
- Cost is not correlated with #features
- Feature extraction offers more flexibility in transforming features, which potentially results in better accuracy

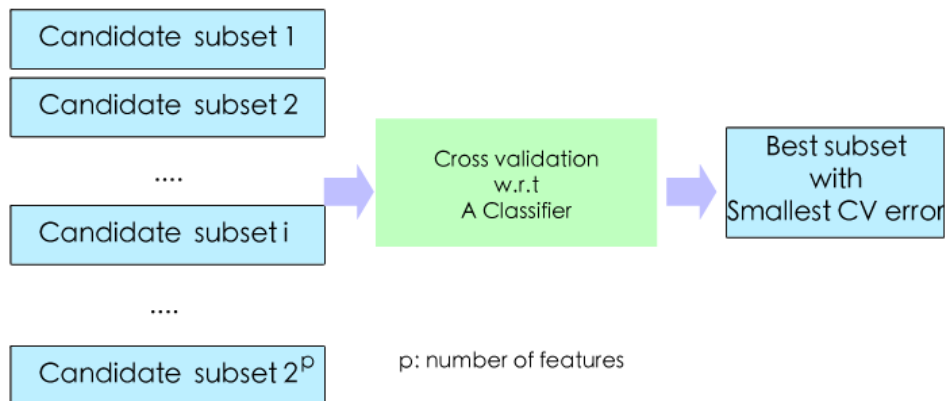




# Feature Selection Methods

## Wrapper Methods

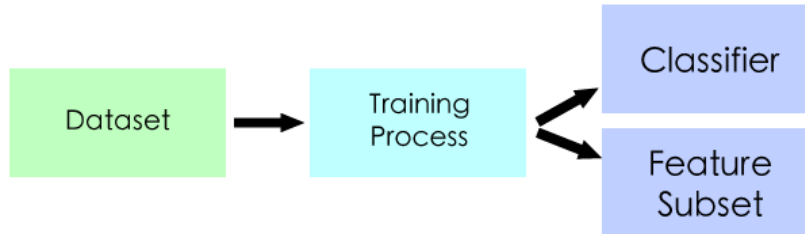
- search the space of feature subsets
- use the training/validation accuracy of a particular classifier as the measure of utility for a candidate subset



# Feature Selection Methods

## Embedded Methods

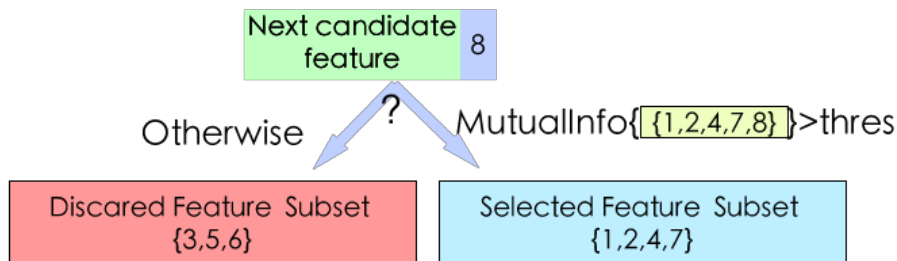
- exploit the structure of specific classes of learning models to guide the feature selection process
- e.g. LASSO. It is embedded as part of the model construction process



# Feature Selection Methods

## Filter Methods

- use some general rules/criteria to measure the feature selection results independent of the classifiers
- e.g. mutual information based method



# Feature Selection

## Comparison

	WRAPPER	FILTER	EMBEDDED
Speed	Low	High	Mid
Chance of Overfitting	High	Low	Mid
Classifier-Independent	No	Yes	No

# Feature Extraction

## Principle Components Analysis

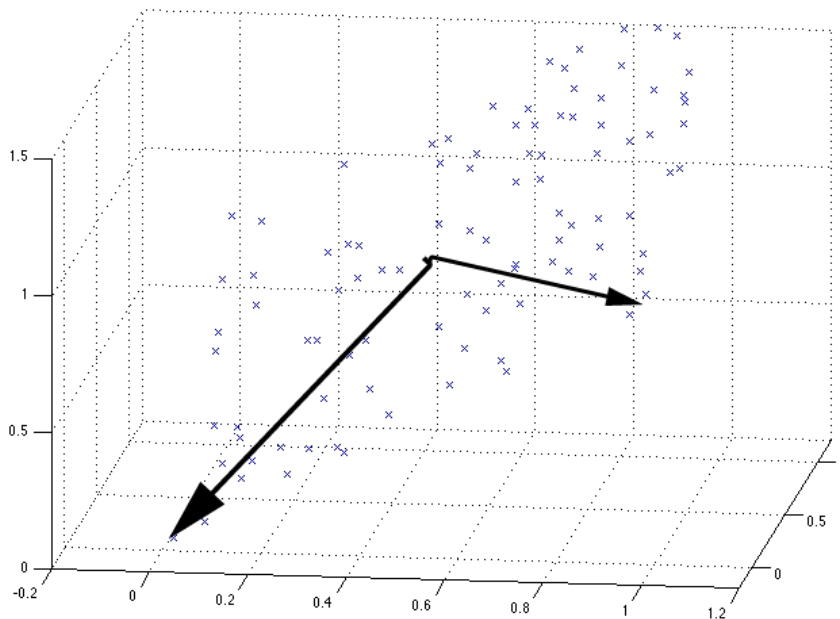
- **A graphical explanation**

- Each data sample has two features
- Often prefer the direction with larger variance
- Original features are transformed into new ones

- **Example**

- For fMRI images, we usually have millions of dimensions. PCA can project the data from millions of dimensions to only thousands of dimensions, or even less

- Other feature extraction methods: ICA, Kernel PCA , etc..



# Part 2: Linear Regression and Least Squares (Review)

- Introduction to Dimension Reduction
- **Linear Regression and Least Squares (Review)**
  - **Least Square Fit**
  - **Gauss Markov**
  - **Bias-Variance tradeoff**
  - **Problems**
- Subset Selection
- Shrinkage Method
- Beyond Lasso

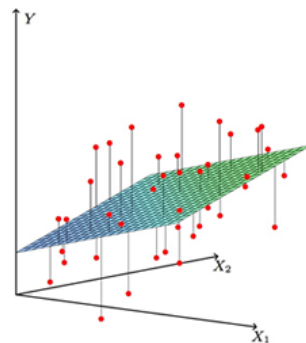
# Linear Regression and Least Squares (Review)

## Least Squares Fit

$$RSS(\beta) = (\mathbf{y} - \mathbf{X}\beta)^T (\mathbf{y} - \mathbf{X}\beta)$$

$$\frac{\partial RSS}{\partial \beta} = -2\mathbf{X}^T (\mathbf{y} - \mathbf{X}\beta) = 0$$

$$\hat{\beta} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$$

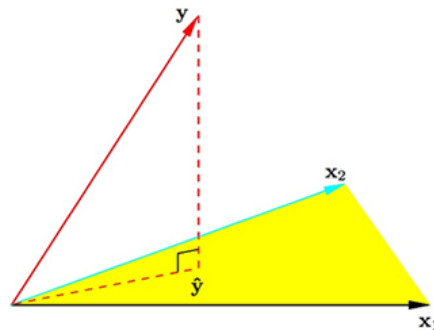


## Gauss Markov Theorem

The least squares estimates of the parameters  $\beta$  have the smallest variance among all linear unbiased estimates.

## Question

Is it good to be unbiased?



# Linear Regression and Least Squares (Review)

## Bias-Variance tradeoff

$$\begin{aligned}MSE(\tilde{\theta}) &= E[(\tilde{\theta} - \theta)^2] \\&= Var(\tilde{\theta}) + [E[\tilde{\theta}] - \theta]^2\end{aligned}$$

where  $\theta = \alpha^T \beta$ . We can trade some bias for much less variance.

## Problems of Least Squares

- **Prediction accuracy:** unbiased, but high variance compared to many biased estimators, overfitting noise and sensitive to outlier
- **Interpretation:**  $\hat{\beta}$  involves all of the features. Better to have SIMPLER linear model, that involves only a few features...
- $(\mathbf{X}^T \mathbf{X})$  may be **not invertible** and thus no closed form solution



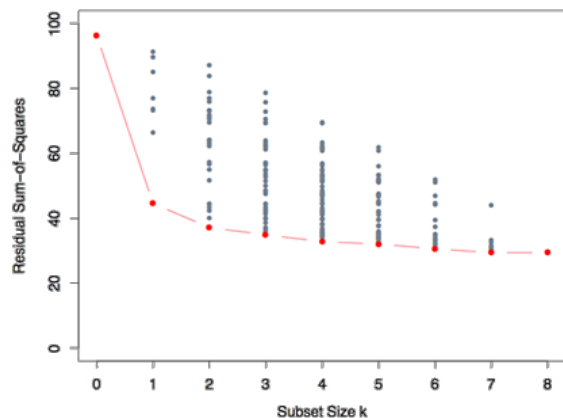
# Part 3: Subset Selection

- Introduction to Dimension Reduction
- Linear Regression and Least Squares (Review)
- **Subset Selection**
  - **Best-subset selection**
  - **Forward stepwise selection**
  - **Forward stagewise selection**
  - **Problems**
- Shrinkage Method
- Beyond Lasso

# Subset Selection

## Best-subset selection

- Best subset regression finds for each  $k \in \{0, 1, 2, \dots, p\}$  the subset of features of size  $k$  that gives smallest residual sum of squares. Then cross validation is utilized to choose the best  $k$
- An efficient algorithm, the leaps and bounds procedure (Furnival and Wilson, 1974), makes this feasible for  $p$  as large as 30 or 40.



**FIGURE 3.5.** All possible subset models for the prostate cancer example. At each subset size is shown the residual sum-of-squares for each model of that size.

# Subset Selection

## Forward-STEPWISE selection

Instead of searching all possible subsets, we can seek a good path through them. This is a **sequential greedy** algorithm.

Forward-Stepwise Selection builds a model sequentially, adding one variable at a time.

- Initialization
  - Active set  $\mathcal{A} = \emptyset$ ,  $\mathbf{r} = \mathbf{y}$ ,  $\beta = 0$
- At each step, it
  - identifies the best variable (with the highest correlation with the residual error)

$$\mathbf{k} = \operatorname{argmax}_j (|\operatorname{correlation}(\mathbf{x}_j, \mathbf{r})|)$$

- $A = A \cup \mathbf{k}$
- then updates the least squares fit  $\beta$ ,  $\mathbf{r}$  to include all the active variables

# Subset Selection

## Forward-STAGewise Regression

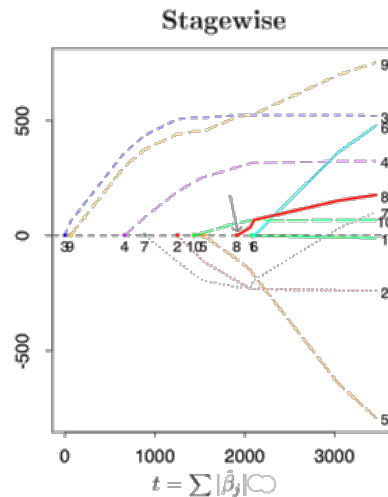
Suppose that  $\mathbf{X}$  is columnwise centered

- Initialize the fit vector  $\mathbf{f} = 0$
- For each time step
  - Compute the correlation vector  $\mathbf{c} = (c_1, \dots, c_p)$ ,  $c_j$  represents the correlation between  $\mathbf{x}_j$  and the residual error
  - $k = \operatorname{argmax}_{j \in \{1, 2, \dots, p\}} |c_j|$
  - Coefficients and fit vector are updated

$$\mathbf{f} \leftarrow \mathbf{f} + \alpha \cdot \operatorname{sign}(\mathbf{c}_k) \mathbf{x}_k$$

$$\beta_k \leftarrow \beta_k + \alpha \cdot \operatorname{sign}(\mathbf{c}_k)$$

where  $\alpha$  is the learning rate



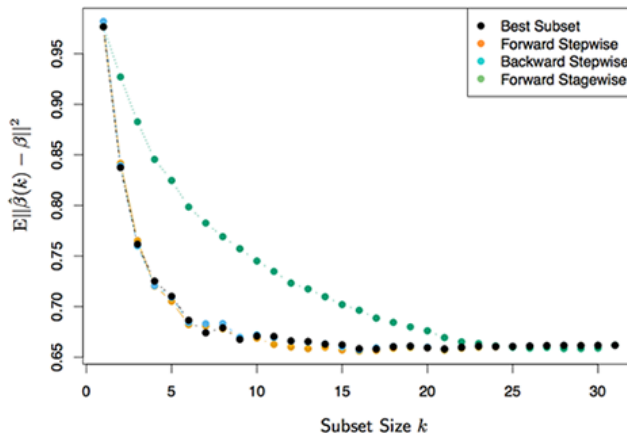
$m = 10$

$n = 442$

# Subset Selection

## Comparison

- Forward-STEPWISE selection:
  - algorithm stops in  $p$  steps
- Forward-STAGEWISE selection:
  - is a slow fitting algorithm, at each time step, only  $\beta_k$  is updated. It can take more than  $p$  steps for the algorithm to stop
  - Forward stagewise is useful in high dimensional problem



- $N = 300$  Observations
- $p = 31$  features
- averaged over 50 simulations

# Subset Selection

## Pros

- More interpretable
- More compact

## Cons

- It is a discrete process, and thus has high variance and sensitivity to the change in dataset.
  - If the dataset changes a little, the feature selection result may be very different
- Thus may not be able to lower prediction error

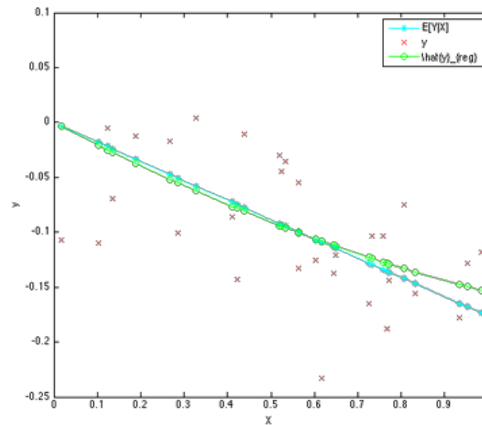
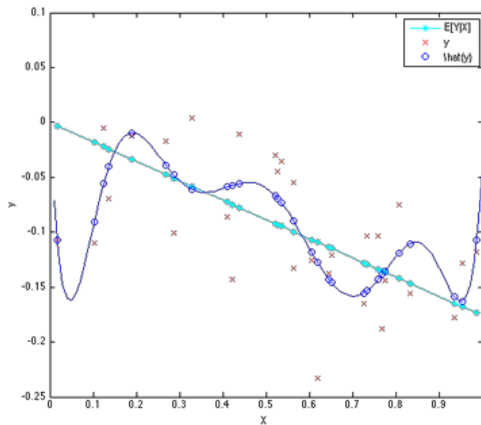
# Part 4 Shrinkage Method

- Introduction to Dimension Reduction
- Linear Regression and Least Squares (Review)
- Subset Selection
- **Shrinkage Method**
  - **Ridge Regression**
    - **Formulations and closed form solution**
    - **Singular value decomposition**
    - **Degree of Freedom**
  - Lasso
- Beyond Lasso

# Ridge Regression

- Least squares with quadratic constraints

$$\hat{\beta}^{\text{ridge}} = \underset{\beta}{\operatorname{argmin}} \sum_{i=1}^N (y_i - \beta_0 - \sum_{j=1}^p \mathbf{x}_{ij} \beta_j)^2, \text{ s. t. } \sum_{j=1}^p \beta_j^2 \leq t$$





# Ridge Regression

- **Least squares with quadratic constraints**

$$\hat{\beta}^{ridge} = \underset{\beta}{\operatorname{argmin}} \sum_{i=1}^N (y_i - \beta_0 - \sum_{j=1}^p \mathbf{x}_{ij} \beta_j)^2, \text{ s. t. } \sum_{j=1}^p \beta_j^2 \leq t$$

- **Its Lagrange form**

$$\hat{\beta}^{ridge} = \underset{\beta}{\operatorname{argmin}} \left\{ \sum_{i=1}^N (y_i - \beta_0 - \sum_{j=1}^p \mathbf{x}_{ij} \beta_j)^2 + \lambda \sum_{j=1}^p \beta_j^2 \right\}$$

- The  $l_2$ -regularization can be viewed as a Gaussian prior on the coefficients, and our estimates are the posterior means
- **Solution**

$$\begin{aligned} RSS(\lambda) &= (\mathbf{y} - \mathbf{X}\beta)^T (\mathbf{y} - \mathbf{X}\beta) + \lambda \beta^T \beta \\ \hat{\beta}^{ridge} &= (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I})^{-1} \mathbf{X}^T \mathbf{y} \end{aligned}$$

# Ridge Regression

## Singular Value Decomposition (SVD)

SVD offers some additional insight into the nature of ridge regression.

- The SVD of  $\mathbf{X}$ :

$$\mathbf{X} = \mathbf{U}\mathbf{D}\mathbf{V}^T$$

- $\mathbf{U}$ :  $N \times p$  **orthogonal** matrix with columns spanning the column space of  $\mathbf{X}$ .  $\mathbf{u}_j$  is the  $j$ th column of  $\mathbf{U}$
- $\mathbf{V}$ :  $p \times p$  **orthogonal** matrix with columns spanning the row space of  $\mathbf{X}$ .  $\mathbf{v}_j$  is the  $j$ th column of  $\mathbf{V}$
- $\mathbf{D}$ :  $p \times p$  **diagonal** matrix with diagonal entries  $d_1 \geq d_2 \geq \dots \geq d_p \geq 0$  being the singular values of  $\mathbf{X}$

# Ridge Regression

## Singular Value Decomposition (SVD)

- For least squares

$$\begin{aligned}\hat{\mathbf{X}}\hat{\boldsymbol{\beta}}^{ls} &= \mathbf{X}(\mathbf{X}^T\mathbf{X})^{-1}\mathbf{X}^T\mathbf{y} \\ &= \mathbf{U}\mathbf{U}^T\mathbf{y}\end{aligned}$$

- For ridge regression

$$\begin{aligned}\hat{\mathbf{X}}\hat{\boldsymbol{\beta}}^{ridge} &= \mathbf{X}(\mathbf{X}^T\mathbf{X} + \lambda\mathbf{I})^{-1}\mathbf{X}^T\mathbf{y} \\ &= \sum_{j=1}^p \mathbf{u}_j \frac{d_j^2}{d_j^2 + \lambda} \mathbf{u}_j^T \mathbf{y}\end{aligned}$$

- Compared with the solution of least squares, we have an additional shrinkage term  $\frac{d_j^2}{d_j^2 + \lambda}$ , the smaller  $d$  is and the larger  $\lambda$  is, the more shrinkage we have.
- The SVD of the centered matrix  $X$  is another way of expressing the principal components of the variables in  $X$ .

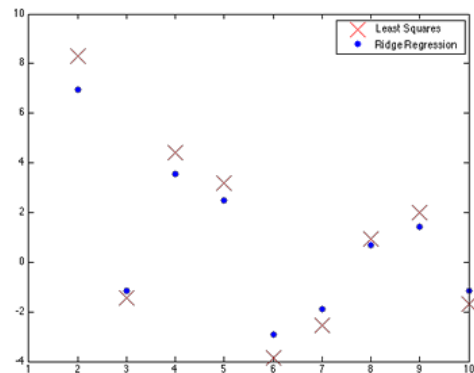
# Ridge Regression

## Singular Value Decomposition (SVD)

•  $N = 100, p = 10$

$$\hat{\beta}^{ls} = u_1 \times (u_1^T y) + u_2 \times (u_2^T y) + \dots + u_p \times (u_p^T y)$$

$$\hat{\beta}^{ridge} = u_1 \times (u_1^T y) \frac{d_1^2}{d_1^2 + \lambda} + u_2 \times (u_2^T y) \frac{d_2^2}{d_2^2 + \lambda} + \dots + u_p \times (u_p^T y) \frac{d_p^2}{d_p^2 + \lambda}$$



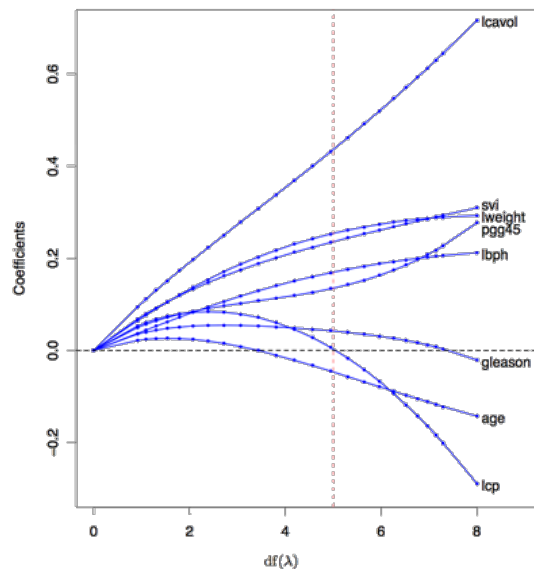
# Ridge Regression

## Degree of Freedom

- In statistics, the number of degrees of freedom is the number of values in the final calculation of a statistic that are free to vary.
- Computation

$$\begin{aligned}d(\lambda) &= \text{tr}[\mathbf{X}(\mathbf{X}^T\mathbf{X} + \lambda\mathbf{I})^{-1}\mathbf{X}^T] \\ &= \sum_{j=1}^p \frac{d_j^2}{d_j^2 + \lambda}\end{aligned}$$

- The larger  $\lambda$  is, the less degree of freedom we have. And then our model will be more constrained.



# Ridge Regression

## Pros

- $(\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I})$  is invertible and thus the closed form solution always exist
- Ridge regression controls the complexity with regularization term via  $\lambda$ , which is less prone to overfitting compared with least squares fit, e.g. sometimes a wildly large coefficient on one variable can be cancelled by another wildly large coefficient of a correlated variable
- Possibly higher prediction accuracy, as the estimates of ridge regression trade a little bias for less variance, compared with least squares

## Cons

- Interpretability and compactness: Though coefficients are shrunk, but not to zero. For high dimensional problem, it may cause efficiency issue.

# Part4 Shrinkage Method - LASSO

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  - **Lasso**
    - **Formulations**
    - **Comparisons with ridge regression and subset selection**
    - **Quadratic Programming**
    - **Least Angle Regression**
    - **Viewed as approximation for  $l_0$ -regularization**
- Beyond Lasso

# LASSO

## Linear regression with $l_1$ -regularization

- Formulations

- Least squares with constraints

$$\hat{\beta}^{lasso} = \underset{\beta}{\operatorname{argmin}} \sum_{i=1}^N (y_i - \beta_0 - \sum_{j=1}^p \mathbf{x}_{ij} \beta_j)^2, \quad \text{s.t.} \sum_{j=1}^p |\beta_j| \leq t$$

- Its lagrange form

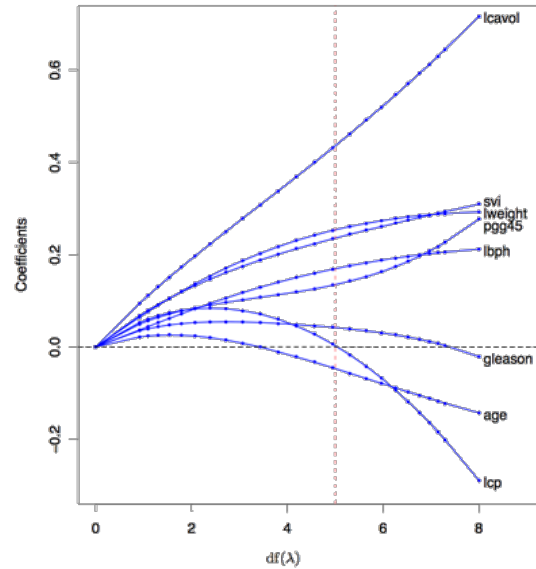
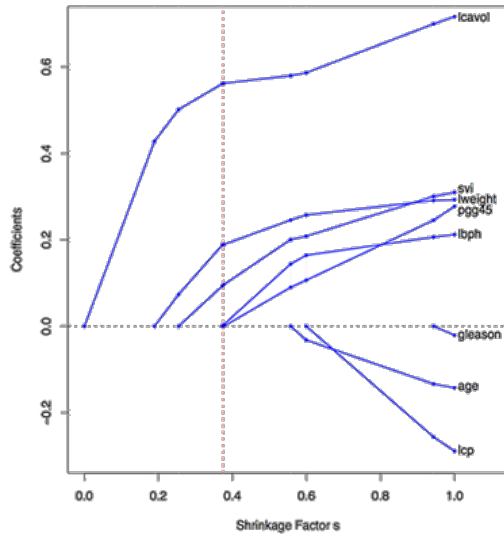
$$\hat{\beta}^{lasso} = \underset{\beta}{\operatorname{argmin}} \left\{ \sum_{i=1}^N (y_i - \beta_0 - \sum_{j=1}^p \mathbf{x}_{ij} \beta_j)^2 + \lambda \sum_{j=1}^p |\beta_j| \right\}$$

- The  $l_1$ -regularization can be viewed as a Laplace prior on the coefficients



# LASSO

- $s = \frac{t}{\sum_{j=1}^p |\hat{\beta}_j|}$ , where  $\hat{\beta}$  is the least square estimates.
- Redlines represent the  $s$  and  $df(\lambda)$  with the best cross validation errors



# LASSO

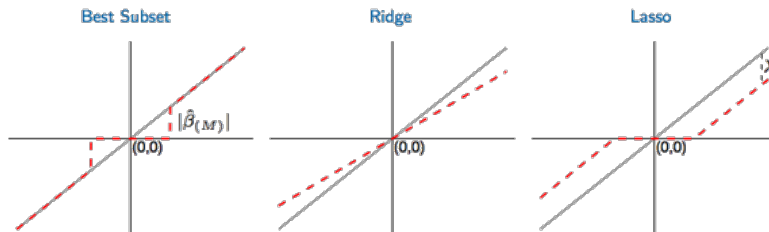
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      - **Orthonormal inputs**
      - **Non-orthonormal inputs**
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# LASSO

## Comparison

- **Orthonormal Input X**

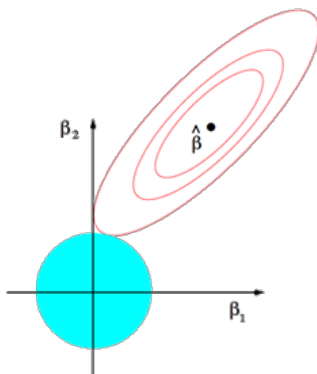
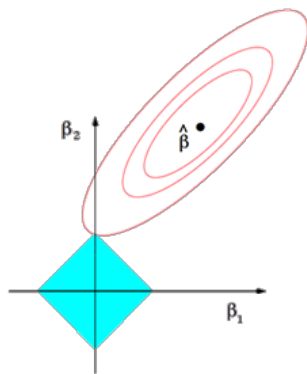
- **Best subset**: [Hard thresholding] Only keep the top  $M$  largest coefficients of  $\hat{\beta}^{ls}$
- **Ridge**: [Pure shrinkage] does proportional shrinkage of  $\hat{\beta}^{ls}$
- **Lasso**: [Soft thresholding] translates each coefficient of  $\hat{\beta}^{ls}$  by  $\lambda$  towards 0, truncating at 0



# LASSO

## Comparison

- Non-orthonormal Input X



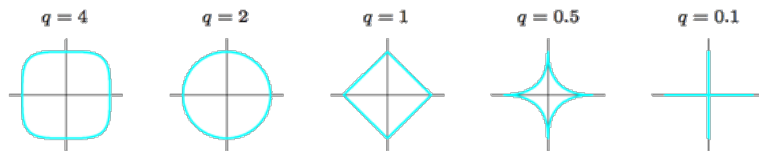
**Solid blue area:** the constraints

- left:  $|\beta_1| + |\beta_2| \leq t$
- right:  $\beta_1^2 + \beta_2^2 \leq t^2$

$\hat{\beta}$ : least squares fit

# LASSO

## Other unit circles for different $p$ -norms



**FIGURE 3.12.** Contours of constant value of  $\sum_j |\beta_j|^q$  for given values of  $q$ .

	CONVEX	SMOOTH	SPARSE
$q < 1$	No	No	Yes
$q > 1$	Yes	Yes	No
$q = 1$	Yes	No	Yes

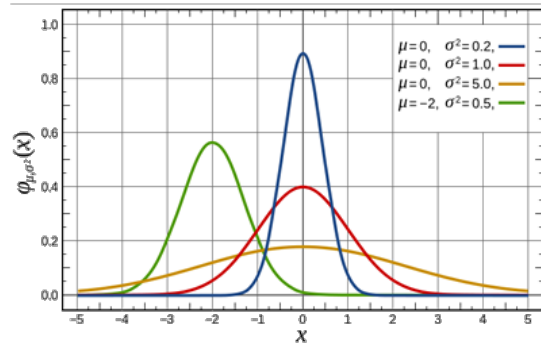
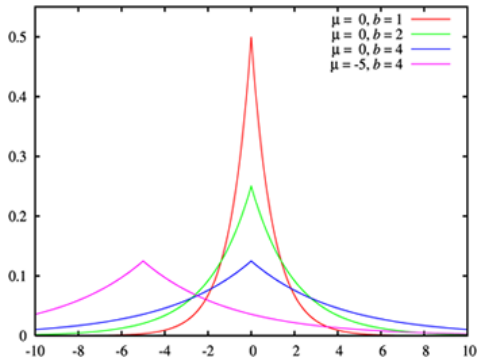
Here  $q = 0$  is the pure variable selection procedure, as it is counting the **number of non-zero coefficients**.

# LASSO

## Regularizations as priors

$|\beta_j|^q$  can be viewed as the log-prior density for  $\beta_j$ , these three methods are bayes estimates with different priors

- **Subset selection:** corresponds to  $q = 0$
- **LASSO:** corresponds to  $q = 1$ , Laplace prior,  $\text{density} = (\frac{1}{\tau})\exp(\frac{-|\beta|}{\tau})$ ,  $\tau = \sigma/\lambda$
- **Ridge regression:** corresponds to  $q = 2$ , Gaussian Prior,  $\beta \sim N(0, \tau\mathbf{I})$ ,  $\lambda = \frac{\sigma^2}{\tau^2}$



# LASSO

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# LASSO

## Quadratic Programming

- Formulation

$$\min_{\beta} \left\{ \frac{1}{2} (\mathbf{X}\beta - \mathbf{y})^T (\mathbf{X}\beta - \mathbf{y}) + \lambda \|\beta\|_1 \right\}$$

is equivalent to

$$\min_{w, \xi} \left\{ \frac{1}{2} (\mathbf{X}\beta - \mathbf{y})^T (\mathbf{X}\beta - \mathbf{y}) + \lambda \mathbf{1}^T \xi \right\}$$

$$\begin{aligned} s. t. \quad & \beta_j \leq \xi_j \\ & \beta_j \geq -\xi_j \end{aligned}$$

- Note that QP can only solve LASSO for a given  $\lambda$ . Later in this slide, a method called least angle regression can solve LASSO for all  $\lambda$



# LASSO

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# Least Angle Regression

## Notations

- $\mathcal{A}_k$ : active set, the set of features we already included in the model at time step  $k$
- $\beta_{\mathcal{A}_k}$ : coefficients vector at the beginning of time step  $k$
- $\beta_{\mathcal{A}_k}(\alpha)$ : coefficients vector in time step  $k$  w.r.t.  $\alpha$ ,
- $\mathbf{f}_k$ : the fit vector at the beginning of time step  $k$ ,  $\mathbf{f}_0 = 0$
- $\mathbf{f}_k(\alpha)$ : the fit vector in time step  $k$  w.r.t.  $\alpha$
- $\mathbf{r}_k$ : residual vector at the beginning of time step  $k$ ,  $\mathbf{r}_0 = \mathbf{y} - \bar{\mathbf{y}}$
- $\mathbf{r}_k(\alpha)$ : residual vector in time step  $k$ , w.r.t.  $\alpha$

# LAR Algorithm

- Initialization:

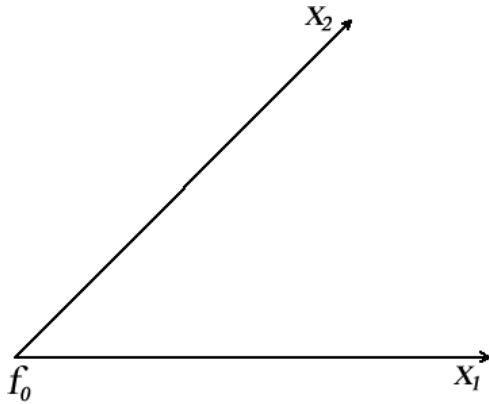
- Standardized all predictors s.t.  $\bar{\mathbf{x}}_j = 0, \mathbf{x}_j^T \mathbf{x}_j = 1; \mathbf{r}_0 = \mathbf{y} - \bar{\mathbf{y}}; \beta = \mathbf{0}; \mathbf{f}_0 = \mathbf{0}; \mathcal{A}_k = \emptyset$
- $k = \operatorname{argmax}_j |\mathbf{x}_j^T \mathbf{r}_0|, \mathcal{A}_1 = \{k\}$

- Main

- for time step  $t$

- $\mathbf{r}_t = \mathbf{y} - \mathbf{X}_{\mathcal{A}_t} \beta_{\mathcal{A}_t}, \mathbf{f}_t = \mathbf{X}_{\mathcal{A}_t} \beta_{\mathcal{A}_t}$
- Search  $\alpha$ 
  - $\beta_{\mathcal{A}_t}(\alpha) = \beta_{\mathcal{A}_t} + \alpha \cdot \delta_t$ , where  $\delta_t = (\mathbf{X}_{\mathcal{A}_t}^T \mathbf{X}_{\mathcal{A}_t})^{-1} \mathbf{X}_{\mathcal{A}_t}^T \mathbf{r}_t$
  - Concurrently,  $\mathbf{f}_t(\alpha) = \mathbf{f}_t + \alpha \cdot \mathbf{u}_t$ , where  $\mathbf{u}_t = \mathbf{X}_{\mathcal{A}_t} \delta_t$
- Until  $|\mathbf{X}_{\mathcal{A}_t} \mathbf{r}_t(\alpha)| = \max_{\mathbf{x}_j \in \tilde{\mathcal{A}}_t} |\mathbf{x}_j^T \mathbf{r}_t(\alpha)|$
- $k = \operatorname{argmax}_{j \in \tilde{\mathcal{A}}_t} |\mathbf{x}_j^T \mathbf{r}_t(\alpha)|$
- $\mathcal{A}_{t+1} = \mathcal{A}_t \cup \{k\}$

# Lar - Example



- Initialization:

- Standardized all predictors

- s.t.  $\bar{\mathbf{x}}_j = 0$ ,  $\mathbf{x}_j^T \mathbf{x}_j = 1$

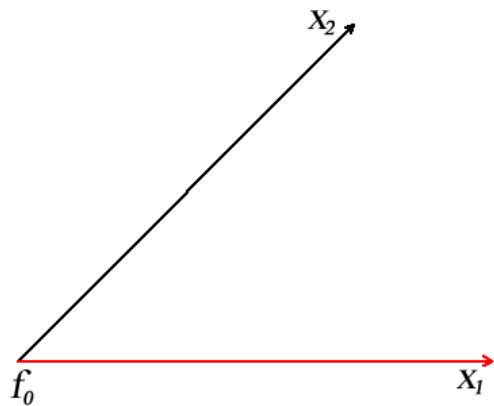
- $\mathbf{r}_0 = \mathbf{y} - \bar{\mathbf{y}}$ ;

- $\beta = (0, 0)^T$ ;

- $\mathcal{A}_0 = \emptyset$

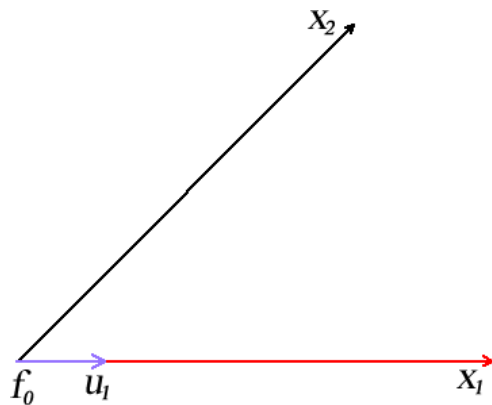
- $\mathbf{f}_0$  is the current fit at time 0 and  $\mathbf{f}_0 = (0, 0)^T$

# Lar - Example



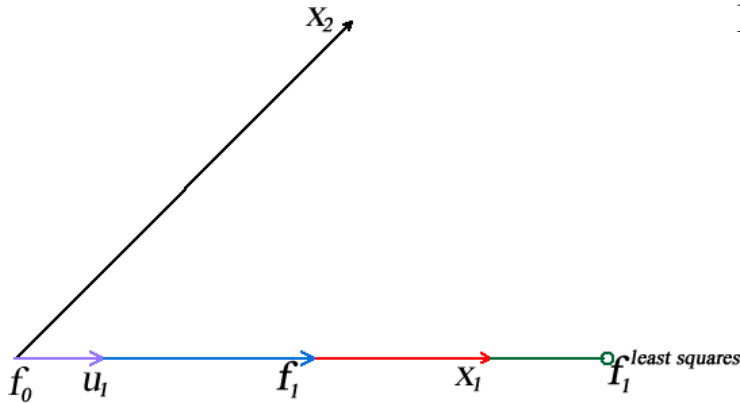
- $k = \operatorname{argmax}_j |\mathbf{x}_j^T \mathbf{r}_0| = 1$
- $\mathcal{A}_1 = \{1\}$

# Lar - Example



- $\mathbf{r}_1 = \mathbf{y} - \mathbf{X}_{\mathcal{A}_1} \beta_{\mathcal{A}_1}$
- $\delta_1 = (\mathbf{X}_{\mathcal{A}_1}^T \mathbf{X}_{\mathcal{A}_1})^{-1} \mathbf{X}_{\mathcal{A}_1}^T \mathbf{r}_1$
- $\mathbf{u}_1 = \mathbf{X}_{\mathcal{A}_1} \delta_1$

# Lar - Example



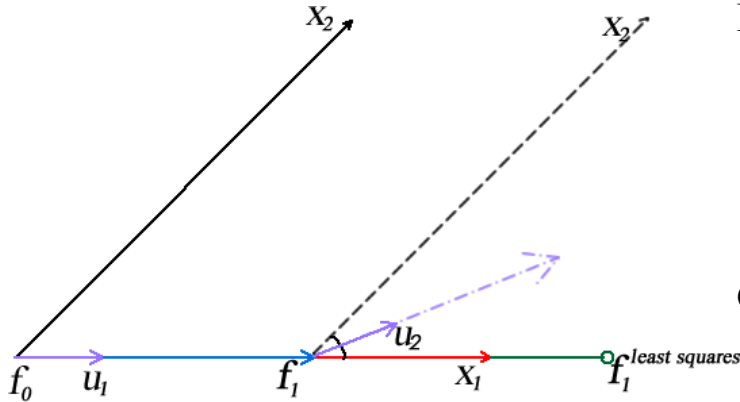
## Explanations

- Search  $\alpha$ 
  - $\beta_{\mathcal{A}_1}(\alpha) = \beta_{\mathcal{A}_1} + \alpha \cdot \delta_1$ ,
  - $\mathbf{f}_1(\alpha) = \mathbf{f}_1 + \alpha \cdot \mathbf{u}_1$
  - $\mathbf{r}_1(\alpha) = \mathbf{y} - \mathbf{X}_{\mathcal{A}_1} \beta_{\mathcal{A}_1}(\alpha)$
- Until  $|\mathbf{X}_{\mathcal{A}_1} \mathbf{r}_1(\alpha)| = \max_{j \in \tilde{\mathcal{A}}_1} |\mathbf{x}_j^T \mathbf{r}_1(\alpha)|$
- $2 = \operatorname{argmax}_{j \in \tilde{\mathcal{A}}_1} |\mathbf{x}_j^T \mathbf{r}_1(\alpha)|$
- $\mathcal{A}_2 = \{1, 2\}$

## Comments

- $f_t$  is approaching  $f_t^{\text{LeastSquares}}$ , but never reaches it, except for the final step

# Lar - Example



## Explanations

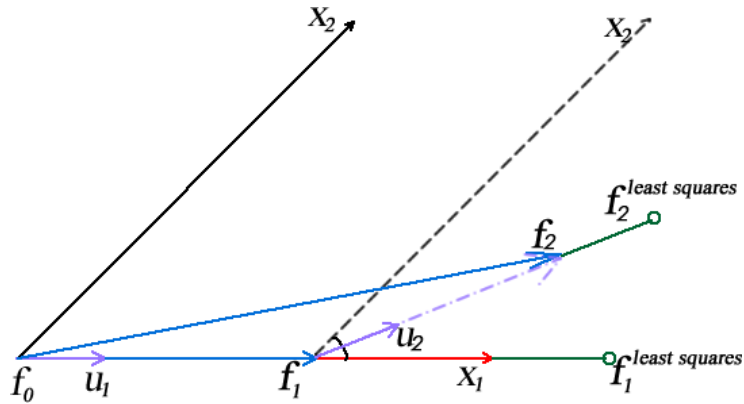
- $\mathbf{r}_2 = \mathbf{y} - \mathbf{X}_{\mathcal{A}_2} \beta_{\mathcal{A}_2}$
- $\delta_2 = (\mathbf{X}_{\mathcal{A}_2}^T \mathbf{X}_{\mathcal{A}_2})^{-1} \mathbf{X}_{\mathcal{A}_2}^T \mathbf{r}_2$
- $\mathbf{u}_2 = \mathbf{X}_{\mathcal{A}_2} \delta_2$

## Comments

- the direction  $\mathbf{u}_k = \mathbf{X}_{\mathcal{A}_k} \delta_k$  that our fit  $\mathbf{f}_k(\alpha)$  increases actually has the same angle with any  $\mathbf{x}_j \in \mathcal{A}_k$ .



# Lar - Example



- If  $p = 2$ 
  - $\mathbf{f}_2 = \mathbf{f}_2^{\text{LeastSquares}}$
- The absolute values of correlations of  $\mathbf{x}_j \in \mathcal{A}_k, \forall j$  with the residual error  $\mathbf{r}_t \alpha$  are tied and decrease at the same rate during searching  $\alpha$ .

# LAR

## More Comments

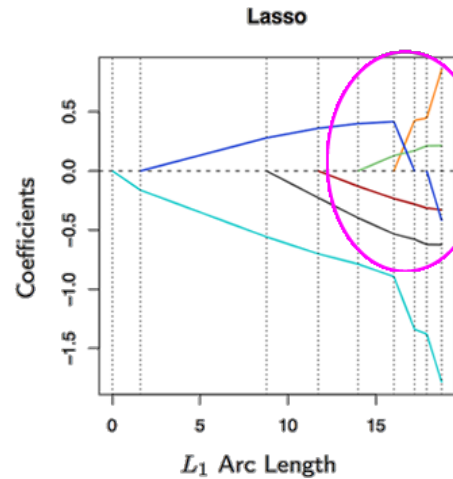
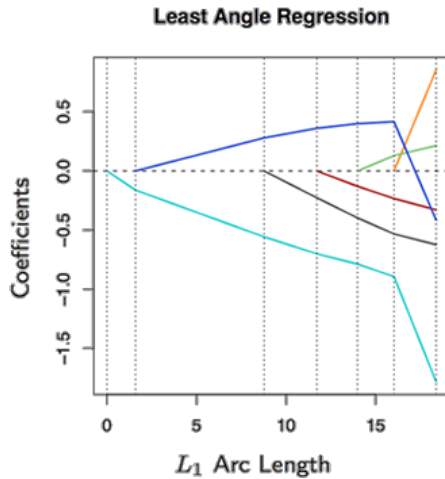
- The procedure of searching is approaching the least-squares coefficients of fitting  $\mathbf{y}$  on  $\mathcal{A}_k$
- LAR solves the subset selection problem for all  $t, s. t. \|\beta\| \leq t$
- Actually,  $\alpha$  can be computed instead of searching
- LAR algorithm ends in  $\min(p, N - 1)$  steps

# LAR

## Result compared with LASSO

### Observations

When the blue line coefficient cross zero, LAR and LASSO become different.

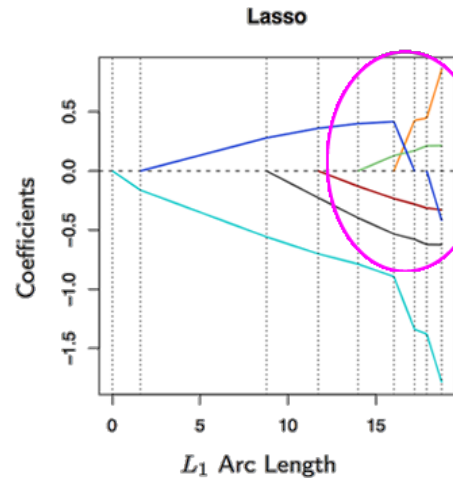
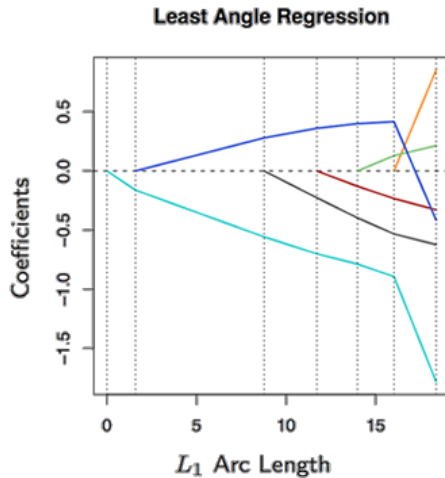


# LAR

## Result compared with LASSO

### Modification for LASSO

During the searching procedure, if a non-zero coefficient hits zero, drop this variable from  $\mathcal{A}_k$ , and recompute the direction  $\delta_k$



# LAR

## Some heuristic analysis

- At a certain time point, we know that all  $\mathbf{x}_j \in \mathcal{A}$  share the same absolute values of correlations with the residual error. That is

$$\mathbf{x}_j^T (\mathbf{y} - \mathbf{X}\beta) = \gamma \cdot s_j, \forall j \in \mathcal{A}$$

where  $s_j \in \{-1, 1\}$  indicates the sign of the left hand inner product and  $\gamma$  is the common value. We also know that  $|\mathbf{x}_j^T (\mathbf{y} - \mathbf{X}\beta)| \leq \gamma, \forall \mathbf{x}_j \notin \mathcal{A}$

- Consider LASSO for a fixed  $\lambda$ . Let  $\mathcal{B}$  be the set of indices of non-zero coefficients, then we differentiate the objective function w.r.t. those coefficients in  $\mathcal{B}$  and set the gradient to zero. We have

$$\mathbf{x}_j^T (\mathbf{y} - \mathbf{X}\beta) = \lambda \cdot \text{sign}(\beta_j), \forall j \in \mathcal{B}$$

- They are identical only if  $\text{sign}(\beta_j)$  matches the sign of the lefthand side. In  $\mathcal{A}$ , we allow for the  $\beta_j$ , where  $\text{sign}(\beta_j) \neq \text{sign}(\mathbf{x}_j^T (\mathbf{y} - \mathbf{X}\beta))$ , while this is forbidden in  $\mathcal{B}$ .

# LAR

## Some heuristic analysis

- For LAR, we have

$$|\mathbf{x}_j^T(\mathbf{y} - \mathbf{X}\beta)| \leq \gamma, \forall \mathbf{x}_j \notin \mathcal{A}$$

- According to the stationary conditions, for LASSO, we have

$$|\mathbf{x}_j^T(\mathbf{y} - \mathbf{X}\beta)| \leq \lambda, \forall \mathbf{x}_j \notin \mathcal{B}$$

- These two algorithms match for variables with zero coefficients too.

# LASSO

- Introduction to Dimension Reduction
- Linear Regression and Least Squares (Review)
- Subset Selection
- **Shrinkage Method**
  - Ridge Regression
  - **Lasso**
    - Formulations
    - Comparisons with ridge regression and subset selection
    - Quadratic Programming
    - Least Angle Regression
    - **Viewed as approximation for  $l_0$ -regularization**
- Beyond Lasso

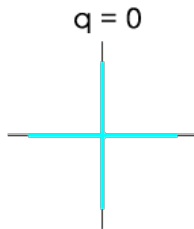
# Viewed as approximation for $l_0$ -regularization

## Pure variable selection

$$\hat{\beta}^{ridge} = \underset{\beta}{\operatorname{argmin}} \sum_{i=1}^N (y_i - \beta_0 - \sum_{j=1}^p \mathbf{x}_{ij} \beta_j)^2, \text{ s. t. } \# \text{nonzero} \beta_j \leq t$$

Actually  $\# \text{nonzero} \beta_j = \|\beta\|_0$ , where

$$\|\beta\|_0 = \lim_{q \rightarrow 0} \left( \sum_{j=1}^p |\beta_j|^q \right)^{\frac{1}{q}} = \operatorname{card}(\{\beta_j | \beta_j \neq 0\})$$





# Viewed as approximation for $l_0$ -regularization

## Problem

$l_0$ -norm is not convex, which makes it very hard to optimize.

## Solutions

- **LASSO**: Approximated objective function ( $l_1$ -norm), with exact optimization
- **Subset selection**: Exact objective function, with approximated optimization (greedy strategy)

# Part5 Beyond LASSO

- Introduction to Dimension Reduction
- Linear Regression and Least Squares (Review)
- Subset Selection
- Shrinkage Method
- **Beyond LASSO**
  - **Elastic-Net**
  - **Fused Lasso**
  - **Group Lasso**
  - $l_1 - lp$  norm
  - **Graph-guided Lasso**

# Beyond LASSO - Elastic Net

## Problems with Lasso

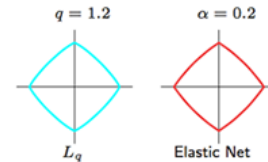
- Lasso tends to rather arbitrarily select one of a group of highly correlated variables (see how LAR works). Sometimes, it is better to select all the relevant variables in a group
- Lasso selects at most  $N$  variables, when  $p > N$ , which may be undesirable when  $p \gg N$
- The performance of Ridge dominates that of Lasso, when  $N > p$  and variables are correlated

## Elastic Net

- **Penalty Term**

$$\lambda \sum_{j=1}^p (\alpha \beta_j^2 + (1 - \alpha) |\beta_j|)$$

which is a compromise between ridge regression and LASSO.



**FIGURE 3.13.** Contours of constant value of  $\sum_j |\beta_j|^q$  for  $q = 1.2$  (left plot), and the elastic-net penalty  $\sum_j (\alpha \beta_j^2 + (1 - \alpha) |\beta_j|)$  for  $\alpha = 0.2$  (right plot). Although visually very similar, the elastic-net has sharp (non-differentiable) corners, while the  $q = 1.2$  penalty does not.

# Beyond LASSO - Elastic Net

## Advantages

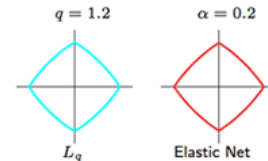
- Solves above problems
- elects variables like lasso, and shrinks together the coefficients of correlated predictors like ridge.
- has considerable computational advantages over the  $l_q$  penalties. See 18.4 [Elements of Statistical Learning]

## Elastic Net

- **Penalty Term**

$$\lambda \sum_{j=1}^p (\alpha \beta_j^2 + (1 - \alpha) |\beta_j|)$$

which is a compromise between ridge regression and LASSO.



**FIGURE 3.13.** Contours of constant value of  $\sum_j |\beta_j|^q$  for  $q = 1.2$  (left plot), and the elastic-net penalty  $\sum_j (\alpha \beta_j^2 + (1 - \alpha) |\beta_j|)$  for  $\alpha = 0.2$  (right plot). Although visually very similar, the elastic-net has sharp (non-differentiable) corners, while the  $q = 1.2$  penalty does not.

# Elastic Net - A simple illustration

- Two independent “hidden” factors  $\mathbf{z}_1$  and  $\mathbf{z}_2$

$$\mathbf{z}_1 \sim U(0, 20), \quad \mathbf{z}_2 \sim U(0, 20),$$

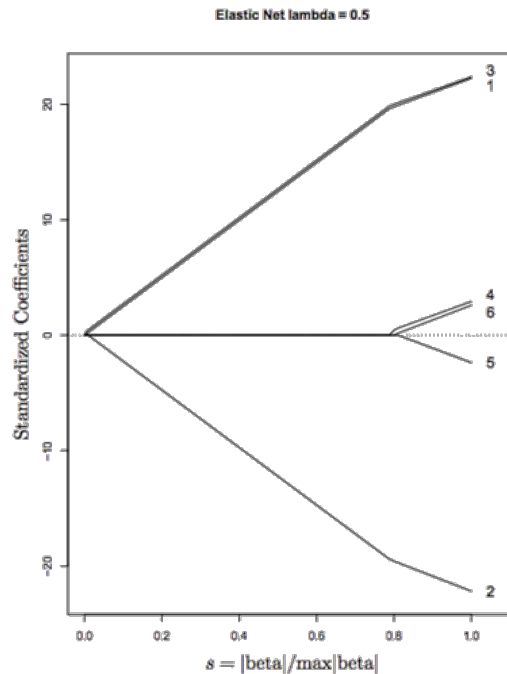
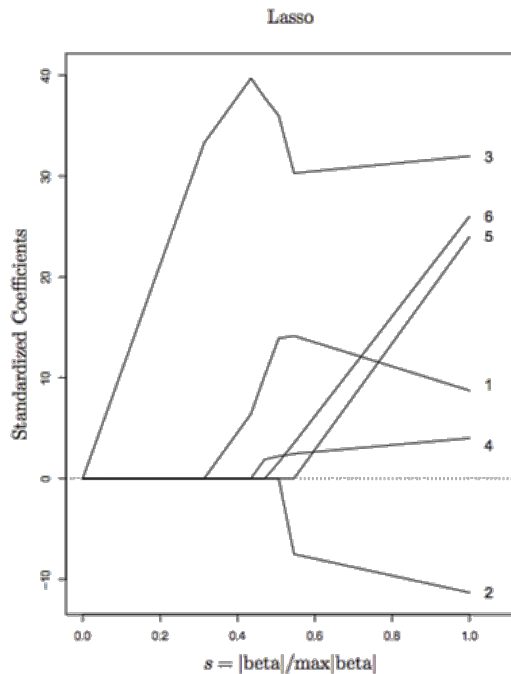
- Generate the response vector  $\mathbf{y} = \mathbf{z}_1 + 0.1\mathbf{z}_2 + N(0, 1)$
- Suppose the observed features are

$$\mathbf{x}_1 = \mathbf{z}_1 + \epsilon_1, \quad \mathbf{x}_2 = -\mathbf{z}_1 + \epsilon_2, \quad \mathbf{x}_3 = \mathbf{z}_1 + \epsilon_3$$

$$\mathbf{x}_4 = \mathbf{z}_2 + \epsilon_4, \quad \mathbf{x}_5 = -\mathbf{z}_2 + \epsilon_5, \quad \mathbf{x}_6 = \mathbf{z}_2 + \epsilon_6$$

- Fit the model on data  $(\mathbf{X}, \mathbf{y})$
- A good model should identify that  $\mathbf{x}_1, \mathbf{x}_2, \mathbf{x}_3$  are important

# Elastic Net - A simple illustration



# Beyond LASSO - Fused Lasso

## Fused Lasso

- **Intuition**

- Fused lasso is a generalization that is designed for problems with features that can be ordered in some meaningful way.
- The fused lasso penalizes the  $L_1$ -norm of both the coefficients and their successive differences.

- **Example**

- Classification with fMRI data: each voxel has about 200 measurements over time. The coefficients for adjacent voxels should be similar

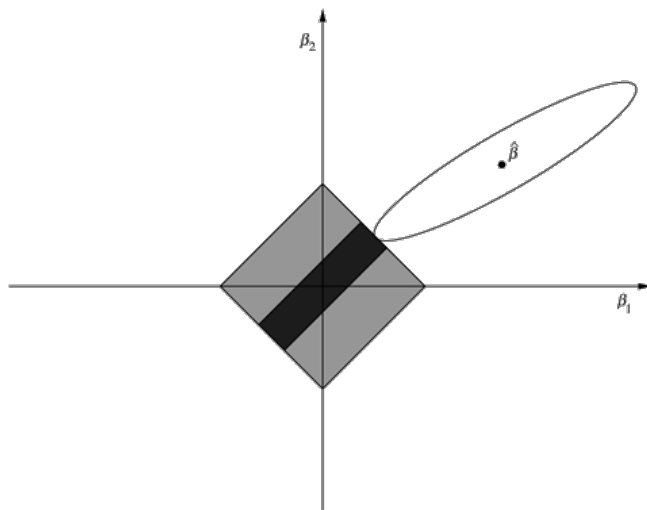
- **Formulation**

$$\hat{\beta} = \operatorname{argmin}_{\beta} \{ \|\mathbf{X}\beta - \mathbf{y}\|_2^2 \}$$

$$s. t. \|\beta\| \leq s_1 \quad \text{and} \quad \sum_{j=2}^p |\beta_j - \beta_{j-1}| \leq s_2$$

# Beyond LASSO - Fused Lasso

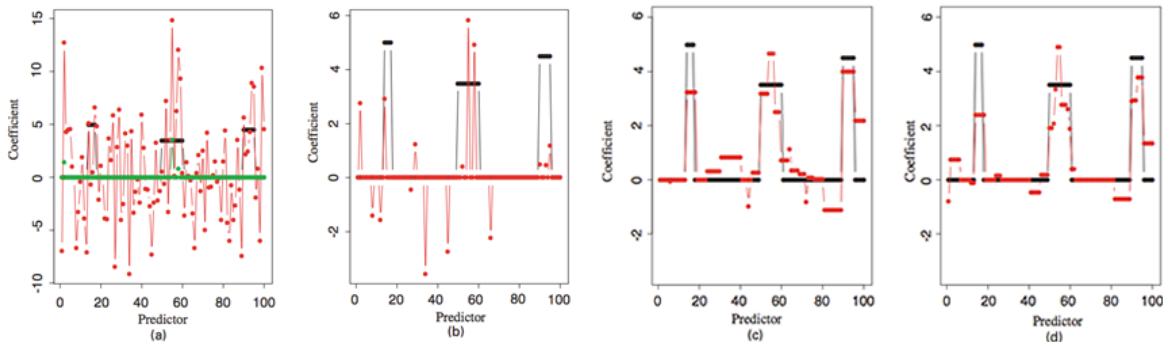
## Fused Lasso



**Fig. 2.** Schematic diagram of the fused lasso, for the case  $N > p = 2$ : we seek the first time that the contours of the sum-of-squares loss function ( $\ominus$ ) satisfy  $\sum_j |\beta_j| = s_1$  ( $\diamond$ ) and  $\sum_j |\beta_j - \beta_{j-1}| = s_2$  ( $\blacklozenge$ )



# Fused Lasso - Simulation results



- $p = 100$ . Black lines are the true coefficients.
- (a) Univariate regression coefficients (red), a soft threshold version of them (green)
- (b) Lasso solution (red),  $s_1 = 35.6, s_2 = \infty$
- (c) Fusion estimate,  $s_1 = \infty, s_2 = 26$
- (d) Fused Lasso,  $s_1 = \sum |\beta_j|, s_2 = \sum |\beta_j - \beta_{j-1}|$

# Beyond LASSO - Group Lasso

## Group Lasso

- **Intuition**

- Features are divided into  $L$  groups
- Features within the same group should share similar coefficients

- **Example**

- Binary dummy variables from one single discrete variable, e.g.  $stage\_cancer \in \{1, 2, 3\}$  can be translated into three binary dummy variables ( $stage1, stage2, stage3$ )

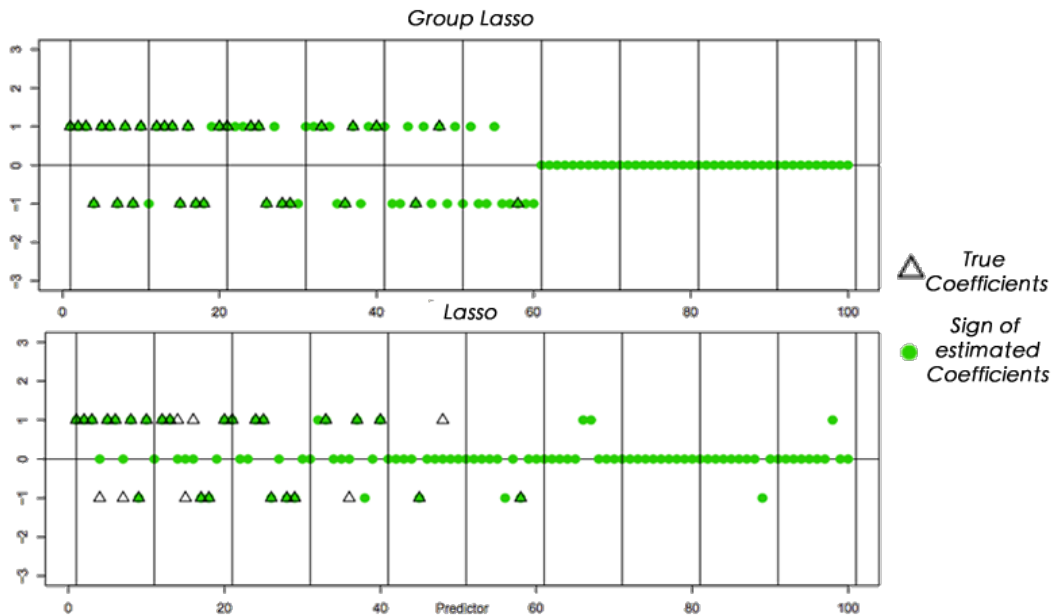
- **Formulations**

$$obj = \|y - \sum_{l=1}^L X_l \beta_l\|_2^2 + \lambda_1 \sum_{l=1}^L \|\beta_l\|_2 + \lambda_2 \|\beta\|_1$$

# Group Lasso - Simulation Results

- Generate  $n = 200$  observations with  $p = 100$ , divided into ten blocks equally
- The last 5 blocks of coefficients  $\beta_j, \forall j \in \{51, 52, \dots, 100\}$  are all 0
- The number of non-zero coefficients in the first five blocks are 10, 8, 6, 4, 2 respectively. The coefficients are either -1 or +1, with the sign being chosen randomly.
- The predictors are standard Gaussian with correlation 0.2 within a group and zero otherwise
- A Gaussian noise with standard deviation 4.0 was added to each observation

# Group Lasso - Simulation Results



# Beyond LASSO - $l_1$ - $l_p$ penalization

## $l_1$ - $l_p$ penalization

- **Applies to multi-task learning**, where the goal is to estimate predictive models for several related tasks.
- **Examples**
  - **Example 1**: recognize speech of different speakers, or handwriting of different writers,
  - **Example 2**: learn to control a robot for grasping different objects or drive in different landscapes, etc.
- **Assumptions about the tasks**
  - sufficiently different that learning a specific model for each task results in improved performance
  - similar enough that they share some common underlying representation that should make simultaneous learning beneficial.
  - focus on the scenario where the different tasks share a subset of relevant features selected from a large common space of features.

# Beyond LASSO - $l_1$ - $l_p$ penalization

## $l_1$ - $l_p$ penalization

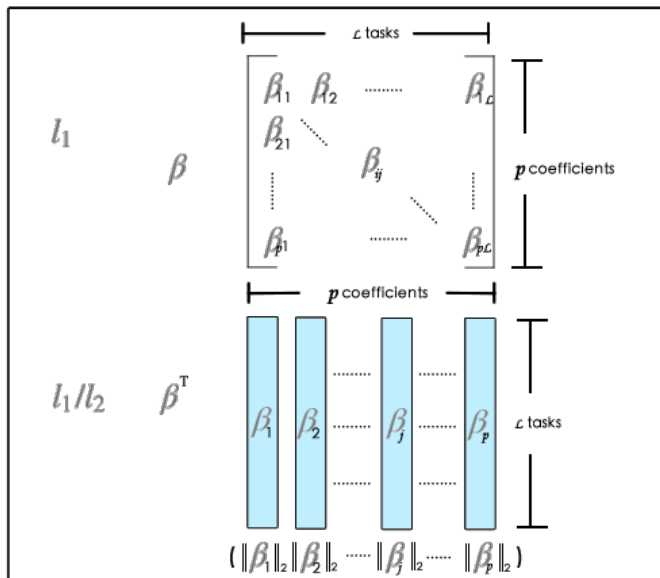
- Formulation
  - $\mathbf{X}_l$ :  $N \times p$  input matrix for task  $l$  and  $L$  is the total number of tasks
  - $\beta$ :  $p \times L$  coefficient matrix
  - $\mathbf{y}$ :  $N \times L$  output matrix
  - objective function

$$obj = \sum_{l=1}^L J(\beta_{:,l}, \mathbf{X}_l, \mathbf{y}_{:,l}) + \lambda \sum_{j=1}^p \|\beta_{j:}\|_2$$

where  $J$  is some loss function and  $\sum_{j=1}^p \|\beta_{j:}\|_2$  is the  $l_1$  norm of vector  $(\|\beta_{:,1}\|_2, \|\beta_{:,2}\|_2, \dots, \|\beta_{:,p}\|_2)$ .

# Beyond LASSO - $l_1 - l_p$ penalization

$l_1 - l_p$  penalization - Coefficient matrix



# Beyond LASSO - $l_1 - l_p$ penalization

$l_1 - l_p$  penalization - Norm ball

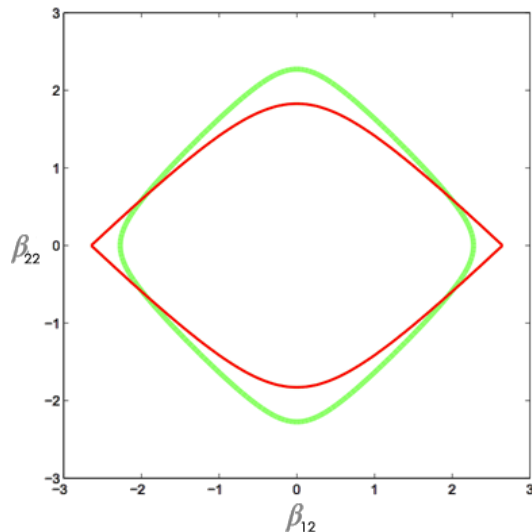


Figure 2: Norm ball induced on the coefficients  $(\beta_{12}, \beta_{22})$  for task 2 as feature coefficients for task 1 vary: thin red contour for  $(\beta_{11}, \beta_{21}) = (0, 1)$  and thick green contour for  $(\beta_{11}, \beta_{21}) = (0.5, 0.5)$ .



# $l_1 - l_p$ penalization - Experiment Result

- Dataset: handwritten words dataset collected by Rob Kassel
  - Contains writings from more than 180 different writers.
  - For each writer, the number of each letter we have is between 4 and 30
  - The letters are originally represented as  $8 \times 16$
- Task: build binary classifiers that discriminate between pairs of letters. Specifically concentrate on the pairs of letters that are the most difficult to distinguish when written by hand.
- Experiment: learned classifications of 9 pairs of letters for 40 different writers

# $l_1 - l_p$ penalization - Experiment Result

- Candidate methods
  - Pooled  $l_1$ : a classifier is trained on all data regardless of writers
  - Independent  $l_1$  regularization: For each writer, a classifier is trained
  - $l_1/l_1$ -regularization:

$$obj = \sum_{l=1}^L J(\beta_{:l}, \mathbf{X}_l, \mathbf{y}_{:l}) + \lambda \sum_{l=1}^L \|\beta_{:l}\|_1$$

- $l_1/l_2$ -regularization:

$$obj = \sum_{l=1}^L J(\beta_{:l}, \mathbf{X}_l, \mathbf{y}_{:l}) + \lambda \sum_{j=1}^p \|\beta_j\|_2$$

---

# $l_1 - l_p$ penalization - Experiment Result

Task	strokes : error(%)				pixels: error (%)			
	$\ell_1/\ell_2$	$\ell_1/\ell_1$	sp. $\ell_1$	pool	$\ell_1/\ell_2$	$\ell_1/\ell_1$	sp. $\ell_1$	pool
<i>c/e</i>	<b>2.5</b>	3.0	3.3	3.0	<b>4.0</b>	8.5	9.0	4.5
	<b>2.0</b>	3.5	3.3	2.5	<b>3.5</b>	7.8	10.3	4.5
<i>g/y</i>	8.4	11.3	<b>8.1</b>	17.8	<b>11.4</b>	16.1	17.2	18.6
	10.3	10.3	<b>9.3</b>	16.9	11.6	<b>9.7</b>	10.9	21.4
<i>g/s</i>	3.3	3.8	<b>3.0</b>	10.7	<b>4.4</b>	10.0	10.3	6.9
	3.8	4.0	<b>2.5</b>	12.0	4.7	6.7	5.0	6.4
<i>m/n</i>	4.4	4.4	<b>3.6</b>	4.7	<b>2.5</b>	6.3	6.9	4.1
	4.1	5.8	<b>3.6</b>	5.3	<b>1.9</b>	2.8	4.1	
<i>a/g</i>	<b>1.4</b>	2.8	2.2	2.8	<b>1.3</b>	3.6	4.1	3.6
	<b>0.8</b>	1.6	1.3	2.5	<b>0.8</b>	1.7	1.4	3.9
<i>i/j</i>	<b>8.9</b>	9.5	9.5	11.5	12.0	14.0	14.0	<b>11.3</b>
	<b>9.2</b>	9.8	11.1	11.3	<b>10.3</b>	12.7	13.5	11.5
<i>a/o</i>	<b>2.0</b>	2.9	2.3	3.8	<b>2.8</b>	4.8	5.2	4.2
	2.7	2.7	<b>1.9</b>	4.3	<b>2.1</b>	3.1	3.5	4.2
<i>f/t</i>	<b>4.0</b>	5.0	6.0	8.1	<b>5.0</b>	6.7	6.1	8.2
	5.8	<b>4.1</b>	5.5	7.5	<b>6.4</b>	11.1	9.6	7.1
<i>h/n</i>	<b>0.9</b>	1.6	1.9	3.4	<b>3.2</b>	14.3	18.6	5.0
	0.9	0.6	<b>0.3</b>	3.7	<b>1.8</b>	3.6	5.0	5.0

- First row contains results for feature selection, the second row uses random projections to obtain a common subspace
- Bold: best of  $l_1/l_2, l_1/l_1, indptl_1$  or pooled  $l_1$ , Boxed : best of cell

# Beyond LASSO - Graph-Guided Fused LASSO

## Graph-Guided Fused LASSO (GFlasso)

- Example

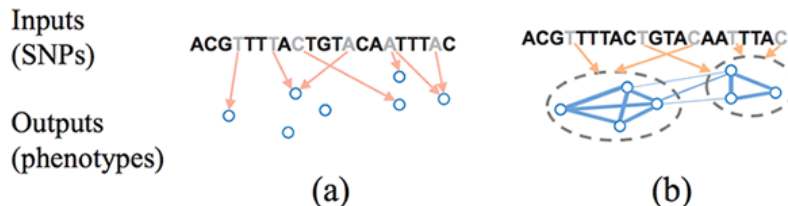


Figure 1: Illustrations of (a) lasso, (b) graph-guided fused lasso.

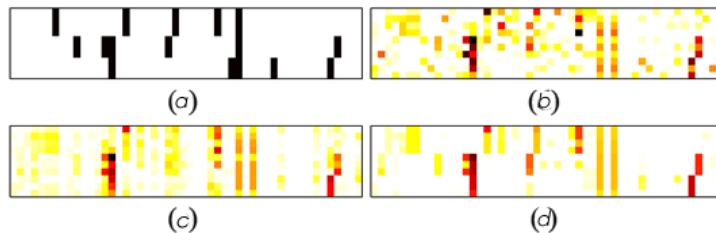
- **Formulation** Graph-Guided Lasso applies to multi-task settings

$$obj = \sum_{l=1}^L loss(\beta_{:,l}, \mathbf{X}_l, \mathbf{y}_{:,l}) + \lambda \|\beta\|_1 + \gamma \sum_{e=(a,b) \in E} \tau(r_{ab}) \sum_{j=1}^p |\beta_{ja} - \text{sign}(r_{a,b})\beta_{jb}|$$

where  $r_{a,b} \in \mathbb{R}$  denotes the weight of the edge and  $\tau(r)$  can be any positive monotonically increasing function of  $|r|$ , e.g.  $\tau(r) = |r|$ .

# Beyond LASSO - Graph-Guided Fused LASSO

## Graph-Guided Fused LASSO



- (a) The true regression coefficients
- (b) lasso
- (c)  $l_1/l_2$ -regularized multi-task regression
- (d) GFLasso

# Summary

## Outline

- **Introduction to Dimension Reduction**
- **Linear Regression and Least Squares (Review)**
- **Subset Selection**
- **Shrinkage Method**
- **Beyond Lasso**

# Summary

## Part 1: Introduction to Dimension Reduction

- **Introduction to Dimension Reduction**
  - **General notations**
  - **Motivations**
  - **Feature selection and feature extraction**
  - **Feature Selection**
    - **Wrapper method**
    - **Filter method**
    - **Embedded method**
  - **Feature Extraction**
    - **PCA, ICA...**
- **Linear Regression and Least Squares (Review)**
- **Subset Selection**
- **Shrinkage Method**

# Summary

## Part 2: Linear Regression and Least Squares (Review)

- Introduction to Dimension Reduction
- **Linear Regression and Least Squares (Review)**
  - **Least Square Fit**
  - **Gauss Markov**
  - **Bias-Variance tradeoff**
  - **Problems**
- Subset Selection
- Shrinkage Method
- Beyond Lasso



# Summary

## Part 3: Subset Selection

- Introduction to Dimension Reduction
- Linear Regression and Least Squares (Review)
- **Subset Selection**
  - **Best-subset selection**
  - **Forward stepwise selection**
  - **Forward stagewise selection**
  - **Problems**
- Shrinkage Method
- Beyond Lasso

# Summary

## Part 4 Shrinkage Method - Ridge Regression

- Introduction to Dimension Reduction
- Linear Regression and Least Squares (Review)
- Subset Selection
- **Shrinkage Method**
  - **Ridge Regression**
    - **Formulations and closed form solution**
    - **Singular value decomposition**
    - **Degree of Freedom**
  - Lasso
- Beyond Lasso

# Summary

## Part4 Shrinkage Method - LASSO

- Introduction to Dimension Reduction
- Linear Regression and Least Squares (Review)
- Subset Selection
- **Shrinkage Method**
  - Ridge Regression
  - **Lasso**
    - **Formulations**
    - **Comparisons with ridge regression and subset selection**
    - **Quadratic Programming**
    - **Least Angle Regression**
    - **Viewed as approximation for  $l_0$ -regularization**
- Beyond Lasso

# Summary

## Part5 Beyond LASSO

- Introduction to Dimension Reduction
- Linear Regression and Least Squares (Review)
- Subset Selection
- Shrinkage Method
- **Beyond LASSO**
  - **Elastic-Net**
  - **Fused Lasso**
  - **Group Lasso**
  - $l_1 - l_p$  norm
  - **Graph-guided Lasso**

# More on the topics skipped here

- More on feature extraction methods:

- [http://www.cs.otago.ac.nz/cosc453/student\\_tutorials/principal\\_components.pdf](http://www.cs.otago.ac.nz/cosc453/student_tutorials/principal_components.pdf)
- Imola K. Fodor, A survey of dimension reduction techniques
- Christopher J. C. Burges, Dimension Reduction: A Guided Tour
- Ali Ghodsi, Dimensionality Reduction A Short Tutorial

- Mutual-info-based feature selection:

- Gavin Brown, Adam Pocock, Ming-Jie Zhao, Mikel Luján; Conditional Likelihood Maximisation: A Unifying Framework for Information Theoretic Feature Selection
- Howard Hua Yang, John Moody. Feature Selection Based on Joint Mutual Information
- Hanchuan Peng, Fuhui Long, and Chris Ding. Feature selection based on mutual information: criteria of max-dependency, max-relevance, and min-redundancy

- Beyond Lasso

- <http://webdocs.cs.ualberta.ca/~mahdavi/ReadingGroup/>

# Sparse Models

**Thank You!**

# Reference

- Trevor Hastie, Robert Tibshirani and Jerome Friedman. Elements of Statistical Learning [p7, p15, p16, p18, p19, p21-22, p26-27, p29-30, p33, p35-37, p42-p43, p50-p54, p56, p59]
- Temporal Sequence of FMRI scans (single slice): from <http://www.midwest-medical.net/mri.sagittal.head.jpg> [p8]
- Three Dimensional Image of Brain Activation from [http://www.fmrib.ox.ac.uk/fmri\\_intro/brief.html](http://www.fmrib.ox.ac.uk/fmri_intro/brief.html) [p8]
- [http://en.wikipedia.org/wiki/Feature\\_selection](http://en.wikipedia.org/wiki/Feature_selection) [p10-12]
- [http://en.wikipedia.org/wiki/Normal\\_distribution](http://en.wikipedia.org/wiki/Normal_distribution) [p38]
- [http://en.wikipedia.org/wiki/Laplacian\\_distribution](http://en.wikipedia.org/wiki/Laplacian_distribution) [p38]
- <http://webdocs.cs.ualberta.ca/~mahdavi/ReadingGroup/Papers/LARS.pdf> [p20]
- Bradley Efron, Trevor Hastie, Iain Johnstone and Robert Tibshirani. Least Angle Regression [p20]

# Reference

- Prof.Schuurmans' notes on Lasso [p40]
- Conditional Likelihood Maximisation: A Unifying Framework for Information Theoretic Feature Selection [p8]
- Hui Zou and Trevor Hastie. Regularization and Variable Selection via the Elastic Net [p59-62]
- [http://www.stanford.edu/~hastie/TALKS/enet\\_talk.pdf](http://www.stanford.edu/~hastie/TALKS/enet_talk.pdf) [p59-62]
- Robert Tibshirani and Michael Saunders, Sparsity and smoothness via the fused lasso [P63-p65]
- Jerome Friedman Trevor Hastie and Robert Tibshirani. A note on the group lasso and a sparse group lasso [p66-68]
- Guillaume Obozinski, Ben Taskar, and Michael Jordan. Multi-task feature selection [p69-70, p72-p75]
- Xi Chen, Seyoung Kim, Qihang Lin, Jaime G. Carbonell, Eric P. Xing. Graph-Structured Multi-task Regression and an Efficient Optimization Method for General Fused Lasso [p76-77]