Topic 4: Simple Linear Regression-Estimation Part B

- ➤ So how well does the estimated model explain the dependent variable?
- ➤ R² is the fraction of the variation of Y that is explained by the model
- ➤ Here model is linear functional form and one explanatory variable X
- The dependent variable Y varies from observation to observation, and so we just want to know how much of that variation we have captured with our model.

3) Fit

Econ 3334

- Explained Sum of Squares $ESS = \sum (\hat{Y}_i \bar{Y})^2$ Sum of Squared Residuals $SSR = \sum (Y_i \bar{Y})^2$
- ightharpoonup Total Sum of Squares $TSS = \sum (Y_i \bar{Y})^2$
- \triangleright We can show ESS + SSR = TSS
- What does this mean?

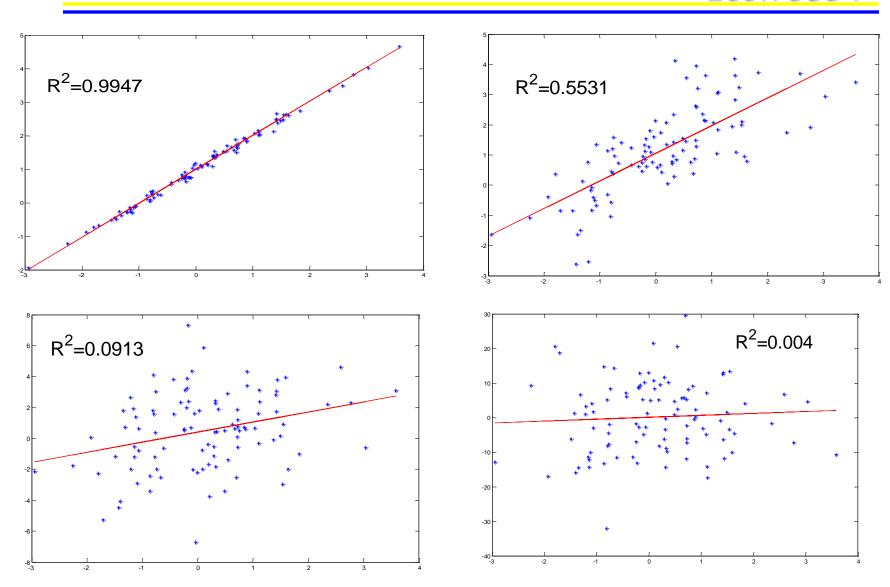
$$Y_i = \hat{Y}_i + \hat{u}_i = \mathsf{OLS} \ \mathsf{prediction} + \mathsf{OLS} \ \mathsf{residual}$$

- \Rightarrow sample variance(\hat{Y}_i) = sample variance(\hat{Y}_i) + sample variance(\hat{u}_i)
- ⇒total sum of squares (TSS: total variation) =
 - "explained" sum of squares (ESS: variation explained by OLS)
 - +"residual" sum of squares (SSR: variation that cannot be explained by OLS)

$$ho R^2 = \frac{ESS}{TSS}$$
 or $R^2 = 1 - \frac{SSR}{TSS}$

- $0 \le R^2 \le 1$
- ➤ R²near zero means model explained virtually none of the variation of Yi

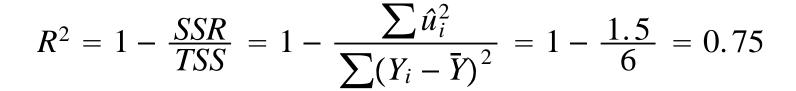
 $ightharpoonup \mathbb{R}^2$ near one means model explained virtually all of the variation of Y_i .



The numerical example: recall:

X_i	Y_i	\Rightarrow	\hat{Y}_i	\hat{u}_i	\hat{u}_i^2	$(Y_i - \overline{Y})^2$
1	2		2.5	-0.5	0.25	4
2	5		4	1	1	1
3	5		5.5	-0.5	0.25	1

Recall:
$$\hat{Y}_i = 1 + 1.5X_i$$



3) Fit

- \triangleright Is R²=0.75 good? bad?
- ➤ Shouldn't think of R² as good or bad.
- \triangleright If it's low, e.g. $R^2 = 0.05$ Why?
 - ➤ A lot of things could be causing Y other than just X. So a low R² could mean we need to include other explanatory variables.

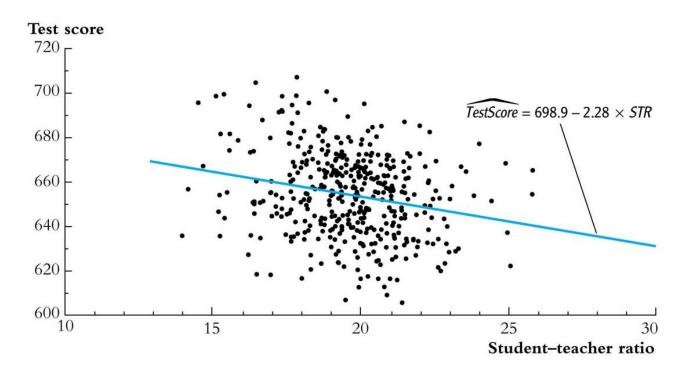
3) Fit

- > Standard error of regression (SER)
- The SER is an estimator of the standard deviation of the error u_i. It is a measure of spread (just like standard deviation).

$$SER = \sqrt{\frac{1}{n-2} \sum_{i} (\hat{u}_i - \bar{\hat{u}})^2} = \sqrt{\frac{1}{n-2} \sum_{i} \hat{u}_i^2} = \sqrt{\frac{1}{n-2} SSR}$$

- Note that $\bar{\hat{u}} = \frac{1}{n} \sum \hat{u}_i = 0$.
- ➤ Here it is divided by n-2, since we "used" up to 2 degree of freedom by estimating beta0 and beta1.
- > SER measures the average "size" of the OLS residual (the average "mistake" made by the OLS regression line)

\triangleright Example of the R^2 and SER



$$\hat{T}estScore = 698.9 - 2.28 \times STR, R^2 = .05, SER = 18.6$$

- ➤ Why do we use the OLS estimator?
- ➤ What are the properties of the OLS estimator?
- ➤ Under what conditions is the OLS estimator "nice"?

 Unbiased, consistent, distributed normally, efficient.
- To answer these questions, we need to make some assumptions about how Y and X are related to each other, and about how they are collected (the sampling scheme)
- These assumptions are known as the Least Squares Assumptions.

$$Y_i = \beta_0 + \beta_1 X_i + u_i, i = 1,..., n$$

Assumption 1: The conditional distribution of u_i given X_i has mean zero, that is, $E(u_i \mid X_i) = 0$.

• This is the key assumption to ensure that β_1 is unbiased. It may not be satisfied in practice.

Assumption 2: (X_i, Y_i) , i = 1, ..., n, are i.i.d.

• This is true if (X_i, Y_i) are collected by simple random sampling

Assumption 3: Large outliers in X_i and Y_i are rare.

- Technically, X_i and Y_i have finite fourth moments
- Outliers can result in meaningless values of $\hat{\beta}_1$

- ightharpoonup Assumption 1: $E(u_i|X_i) = 0$.
- \triangleright Given the explanatory variable X_i , the expected value of the error is zeros.
- \triangleright This implies that X_i and u_i are not correlated.

$$E(u_i|X_i) = 0 \Rightarrow E(u_i) = E[E(u_i|X_i)] = 0$$
 by law of iterated expectation $cov(X_i, u_i) = E(X_iu_i) - E(X_i)E(u_i) = E(X_iu_i)$
= $E[E(X_iu_i|X_i)]$ by law of iterated expectation

$$=E\left[\begin{array}{c}X_{i}\bullet E(u_{i}|X_{i})\\ =0\end{array}\right]=0$$

- \triangleright If X_i and u_i are not correlated, this does NOT imply E(u_i|X_i)=0
- \triangleright If X_i and u_i are correlated, we know that Assumption 1 is false. 12

- > **Assumption 1**: $E(u_i|X_i) = 0$.
- A benchmark for thinking about this assumption is to consider an ideal randomized controlled experiment:
- \triangleright X_i is randomly assigned to people. Randomization is done by computer using no information about the individual.
- \triangleright Because X_i is assigned randomly, all other individual characteristics the things that make up u_i are independently distributed of X_i
- Thus, in an ideal randomized controlled experiment, $E(u_i|X_i) = 0$ (that is, Assumption #1 holds)
- With observational data, we will need to think hard about whether $E(u_i|X_i) = 0$ holds.

ightharpoonup Assumption 1: $E(u_i|X_i) = 0$.

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(i) X_i and u_i are independent and E(u_i) = 0
 \downarrow (Yes) \quad \uparrow (No)
(ii) E(u_i|X_i) = 0
 \downarrow (Yes) \quad \uparrow (No)
(iii) cov(u_i,X_i) = 0
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- Assumption 1 is a strong assumption. It can fail easily. In practice, it may be hard to justify.
- Fortunately, econometricians have developed many other methods for the case where Assumption 1 does not hold.

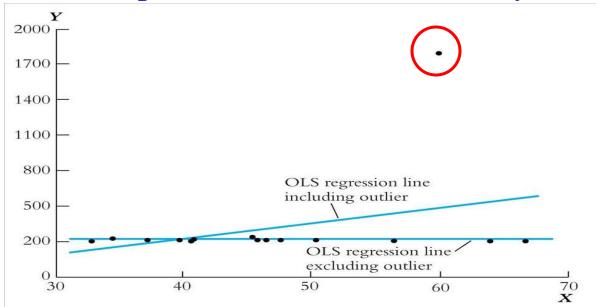
- \triangleright **Assumption 2**: (X_i, Y_i) are i.i.d.
- This arises automatically if the entity (individual, district) is sampled by simple random sampling: the entity is selected then, for that entity, X and Y are observed (recorded).
- ➤ The main place we will encounter non-i.i.d. sampling is when data are recorded over time ("time series data") this will introduce some extra complications.

- > **Assumption 3**: outliers are unlikely.
- > Technically,

$$0 < E(X_i^4) < \infty$$
 and $0 < E(Y_i^4) < \infty$

- \triangleright A large outlier is an extreme value of X or Y
- ➤ On a technical level, if *X* and *Y* are bounded, then they have finite fourth moments. (Standardized test scores automatically satisfy this; *STR*, family income, etc. satisfy this too).
- ➤ However, the substance of this assumption is that a large outlier can strongly influence the results

> Assumption 3: outliers are unlikely.



- ➤ Is the lone point an outlier in X or Y?
- ➤ In practice, outliers often are data glitches (coding/recording problems) so check your data for outliers!

 $X_i \quad Y_i$

10 3000

8 2400

6 2300

↓ oops

 $X_i \quad Y_i$

10 (30000)

8 2400

6 2300

- ➤ Why do we need to make these assumptions?
- > We need them to prove that OLS estimators are
 - Unbiased and Consistent
 - Asymptotically normally distributed (the distributions of OLS estimators are close to normal distribution)
 - Efficient (if we further assume homoskedasticity, $var(u_i|X_i)$ =constant, i.e., conditional variance of u_i given X_i is constant)

THE GAUSS-MARKOV THEOREM FOR $\hat{\boldsymbol{\beta}}_1$

If the three least squares assumptions in Key Concept 4.3 hold and if errors are homoskedastic, then the OLS estimator $\hat{\beta}_1$ is the Best (most efficient) Linear conditionally Unbiased Estimator (is BLUE).

- The data we use comes from a sample taken from a underlying population
- The data differ from sample to sample, and thus so does the OLS estimators $\hat{\beta}_0, \hat{\beta}_1$
- $\geq \hat{\beta}_0, \hat{\beta}_1$ are random variables since they come from random samples (just like $\bar{\gamma}$ in Chapter 3)
- ➤ We need to understand their probability distribution, so we can make inferences (e.g., hypothesis testing, confidence interval)

- \triangleright Like \overline{Y} , $\hat{\beta}_1$ has a sampling distribution.
- What is $E(\hat{\beta}_1)$? (where is it centered?) If $E(\hat{\beta}_1) = \beta_1$, then OLS is unbiased – a good thing!
- \triangleright What is var($\hat{\beta}_1$)? (measure of sampling uncertainty)
- ➤ What is the distribution of in small samples? It can be very complicated in general
- ➤ What is the distribution of in large samples?

 It turns out to be relatively simple in large samples, is normally distributed.

- \triangleright Under the Key OLS Assumption 1: $E(\hat{\beta}_0) = \beta_0$ and $E(\hat{\beta}_1) = \beta_1$
- \triangleright Given the OLS assumptions, the large sample distribution of $\hat{\beta}_0, \hat{\beta}_1$ are:

$$\hat{\beta}_{1} \sim N\left(\beta_{1}, \sigma_{\hat{\beta}_{1}}^{2}\right),$$
where
$$\sigma_{\hat{\beta}_{1}}^{2} = \frac{1}{n} \frac{var[(X_{i} - \mu_{X})u_{i}]}{[var(X_{i})]^{2}}$$

$$\hat{\beta}_{0} \sim N\left(\beta_{0}, \sigma_{\hat{\beta}_{0}}^{2}\right),$$
where
$$\sigma_{\hat{\beta}_{0}}^{2} = \frac{1}{n} \frac{var(H_{i}u_{i})}{[E(H_{i}^{2})]^{2}}, H_{i} = 1 - \left(\frac{\mu_{X}}{E(X_{i}^{2})}\right)X_{i}$$

> We see that

$$var(\hat{\beta}_1) = \frac{1}{n} \frac{var[(X_i - \mu_X) \cdot u_i]}{[var(X_i)]^2}$$
$$var(\hat{\beta}_0) = \frac{1}{n} \frac{var(H_i u_i)}{[E(H_i^2)]^2}$$

- Variance of $\hat{\beta}_0$ and $\hat{\beta}_1$ are proportionate to n. Thus as n goes to infinite, the variance becomes smaller and smaller. This means that $\hat{\beta}_0$ and $\hat{\beta}_1$ become more and more concentrated around true β_0 and β_1 .
- \triangleright Thus $\hat{\beta}_0$ and $\hat{\beta}_1$ are consistent.