# Simple Estimation of Semiparametric Models with Measurement Errors

Kirill Evdokimov Andrei Zeleneev

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Universitat Pompeu Fabra & Barcelona School of Economics University College London

# Framework: EIV for general GMM models

General moment condition model

$$\mathbb{E}\left[g(X_i^*, S_i, \theta)\right] = 0 \quad \text{iff} \quad \theta = \theta_0$$

•  $X_i^*$  is mismeasured:

$$X_i = X_i^* + \varepsilon_i,$$

- For the talk: classical measurement error  $\varepsilon_i$ , i.e.,  $\varepsilon_i \perp (X_i^*, S_i), \ \mathbb{E}[\varepsilon_i] = 0$
- In the paper: non-classical, multivariate, and serially correlated measurement errors
- Leading example: nonlinear regression

$$Y_i = \rho(X_i^*, W_i, \theta_0) + U_i, \quad \mathbb{E}\left[U_i | X_i^*, W_i, Z_i\right] = 0$$
$$g(x, s, \theta) = (y - \rho(x, w, \theta))h(x, w, z)$$

- Logit/Probit/Tobit, CES production function ...
- and also multinomial choice, multi-equation models, structural models, ...

#### Framework: Moderate Measurement Errors

- Most of the literature treats  $\sigma_{arepsilon}^2 = \mathbb{E}\left[arepsilon_i^2\right]$  as fixed
  - Hausman, Newey, Ichimura, and Powell (1991), Newey (2001), Li (2002), Schennach (2004, 2007), Hu and Schennach (2008), and many others
  - Estimation of an infinite dimensional nuisance parameter and/or numerical simulation are required
  - Specific models (typically, nonlinear regression) and rather restrictive identification conditions
  - Perhaps too pessimistic and not representative of most empirical settings
- In practice, noise-to-signal ratio  $au=\sigma_{arepsilon}/\sigma_{X^*}$  is typically small in applications
- We consider an alternative asymptotic approximation: au o 0 as  $n o \infty$ 
  - · Perhaps a better match for the problem
  - We construct a simple estimation procedure for the general moment condition model
  - For presentation, think of  $\sigma_\varepsilon^2 \equiv \mathbb{E}\left[\varepsilon_i^2\right] \equiv \sigma_n^2 \to 0$

#### **Estimation: Motivation**

Taylor expansion of  $g(X_i, S_i, \theta)$  (in the spirit of Chesher, 1991):  $\sigma_n^2 \to 0$ 

$$g(X_{i}, S_{i}, \theta) = g(X_{i}^{*}, S_{i}, \theta) + \varepsilon_{i} g_{x}^{(1)}(X_{i}^{*}, S_{i}, \theta) + \frac{\varepsilon_{i}^{2}}{2} g_{x}^{(2)}(X_{i}^{*}, S_{i}, \theta) + O(\sigma_{n}^{3}),$$

$$\mathbb{E}[g(X_{i}, S_{i}, \theta)] = \mathbb{E}[g(X_{i}^{*}, S_{i}, \theta)] + \frac{\sigma_{n}^{2}}{2} \mathbb{E}[g_{x}^{(2)}(X_{i}^{*}, S_{i}, \theta)] + O(\sigma_{n}^{3}),$$

where  $g_x^{(k)}(x,s,\theta) \equiv \partial^k g(x,s,\theta)/\partial x^k$ 

Therefore,  $\mathbb{E}\left[g(X_i,S_i,\theta_0)\right]=O(\sigma_n^2)$ , and the standard (naive) estimator is:

- Asymptotically biased if  $\sqrt{n}\sigma_n^2 \to C \in (0,+\infty)$
- Not  $\sqrt{n}$ -consistent if  $\sqrt{n}\sigma_n^2 \to \infty$

#### **MERM Estimator**

• Define the Measurement Error Robust Moments (MERM):

$$\psi(X_i, S_i, \theta, \gamma) \equiv g(X_i, S_i, \theta) - \gamma g_x^{(2)}(X_i, S_i, \theta)$$

# Assumption (Moderate Measurement Error, Special Case of K=2)

$$\sigma_n^2 = o(n^{-1/3})$$

- $\mathbb{E}\left[\psi(X_i, S_i, \theta_0, \gamma_{0n})\right] = O(\sigma_n^3) = o(n^{-1/2})$ , where  $\gamma_{0n} \equiv \sigma_n^2/2$
- MERM estimator:

$$(\hat{\theta}, \hat{\gamma}) = \mathop{\rm argmin}_{\theta \in \Theta, \gamma \in \Gamma} \hat{Q}(\theta, \gamma), \quad \hat{Q}(\theta, \gamma) = \overline{\psi}(\theta, \gamma)' \, \hat{\Xi} \, \overline{\psi}(\theta, \gamma)$$

Higher order expansion is more delicate: one needs to bias correct the bias correction terms.

#### **Monte Carlo: Point Estimation**

From Schennach (2007, Ecta):

(Original) moment function:

$$g(y, x, z, \theta) = (\rho(x, \theta) - y)h(x, z)$$

The noise-to-signal is  $au=\sigma_{arepsilon}/\sigma_{X^*}pprox 0.45\,$  – "fairly large" measurement errors

#### Specifications:

- Polynomial
- Rational Fraction
- Probit

# **Monte Carlo: Polynomial Specification**

$$\rho(x,\theta) = \theta_1 + \theta_2 x + \theta_3 x^2 + \theta_4 x^3, \quad \theta_0 = (1,1,0,-0.5)'$$

		Bia	as			Std.	Dev.				RMSE		
	$\overline{\theta_1}$	$\theta_2$	$\theta_3$	$\overline{ heta_4}$	$\overline{ heta_1}$	$\theta_2$	$\theta_3$	$\overline{ heta_4}$	$\overline{ heta_1}$	$\theta_2$	$\theta_3$	$\theta_4$	All
OLS	-0.00	-0.43	0.00	0.21	0.07	0.13	0.06	0.04	0.07	0.45	0.06	0.22	0.51
S07	-0.05	-0.07	-0.02	0.05	0.17	0.19	0.24	0.05	0.17	0.20	0.24	0.07	0.36
K = 2	-0.00	0.10	0.00	0.00	0.10	0.23	0.10	0.08	0.10	0.25	0.10	0.08	0.29
K = 4	-0.00	0.00	0.00	0.02	0.09	0.21	0.10	0.08	0.09	0.21	0.10	0.08	0.27

$$\sigma_{\varepsilon}/\sigma_{X^*} \approx 0.45, \ Y_i = \rho(X_i^*, \theta_0) + U_i, \ n = 1000$$

# **Monte Carlo: Polynomial Specification**

$$\rho(x,\theta) = \theta_1 + \theta_2 x + \theta_3 x^2 + \theta_4 x^3, \quad \theta_0 = (1,1,0,-0.5)'$$

		Bi	as			Std.	Dev.				RMSE		
	$\theta_1$	$\theta_2$	$\theta_3$	$\overline{ heta_4}$	$\overline{ heta_1}$	$\theta_2$	$\theta_3$	$\overline{ heta_4}$	$\overline{ heta_1}$	$\theta_2$	$\theta_3$	$\theta_4$	All
OLS	-0.00	-0.43	0.00	0.21	0.07	0.13	0.06	0.04	0.07	0.45	0.06	0.22	0.51
S07	-0.05	-0.07	-0.02	0.05	0.17	0.19	0.24	0.05	0.17	0.20	0.24	0.07	0.36
K=2	-0.00	0.10	0.00	0.00	0.10	0.23	0.10	0.08	0.10	0.25	0.10	0.08	0.29
K=4	-0.00	0.00	0.00	0.02	0.09	0.21	0.10	0.08	0.09	0.21	0.10	0.08	0.27

$$\sigma_{\varepsilon}/\sigma_{X^*} \approx 0.45, \ Y_i = \rho(X_i^*, \theta_0) + U_i, \ n = 1000$$

# **Monte Carlo: Rational Fraction Specification**

$$\rho(x,\theta) = \theta_1 + \theta_2 x + \frac{\theta_3}{(1+x^2)^2}, \quad \theta_0 = (1,1,2)'$$

		Bias			Std. Dev			RA	1SE	
	$\theta_1$	$\theta_2$	$\theta_3$	$\theta_1$	$\theta_2$	$\theta_3$	$\theta_1$	$ heta_2$	$\theta_3$	All
OLS	0.339	-0.167	-0.644	0.040	0.020	0.076	0.341	0.168	0.648	0.752
S07	0.107	0.117	-0.150	0.146	0.139	0.328	0.181	0.182	0.361	0.443
K=2	-0.004	-0.018	0.014	0.062	0.026	0.139	0.062	0.032	0.139	0.156
K=4	0.014	-0.002	-0.024	0.062	0.031	0.154	0.063	0.031	0.156	0.171

$$\sigma_{\varepsilon}/\sigma_{X^*} \approx 0.45, \ Y_i = \rho(X_i^*, \theta_0) + U_i, \ n = 1000$$

#### **Monte Carlo: Probit Model**

$$\rho(x,\theta) = \frac{1}{2}(1 + \text{erf}(\theta_1 + \theta_2 x)), \quad \theta_0 = (-1,2)'$$

	Bi	as	Std.	Dev.		RMSE	
	$\theta_1$	$\theta_2$	$\theta_1$	$\theta_2$	$\theta_1$	$ heta_2$	All
NLLS	0.38	-0.97	0.06	0.08	0.39	0.98	1.05
S07	0.05 -0.06		0.39	0.53	0.39	0.53	0.69
K=2	0.11	-0.31	0.18	0.34	0.21	0.46	0.51
K = 4	-0.01 -0.01		0.23	0.42	0.23	0.42	0.48

$$\sigma_{\varepsilon}/\sigma_{X^*} \approx 0.45, \ Y_i = \rho(X_i^*, \theta_0) + U_i, \ n = 1000$$

#### Monte Carlo: Multinomial Choice Model

For an agent i the utility of option j is:

$$U_{ij} = \theta_{0j1}X_i^* + \theta_{0j2}W_{ij} + \theta_{0j3} + \epsilon_{ij}$$
 for  $j \in \{1, 2\}$ ,  $U_{i0} = \epsilon_{i0}$  for  $j = 0$ ,  $\epsilon_{ij} \sim_{iid}$  type-I EV distribution

The researcher observes  $(X_i, W_i, Y_i)$ ,

$$\begin{split} Y_i &= \operatorname{argmax}_{j \in \{0,1,2\}} U_{ij} \\ X_i &= X_i^* + \varepsilon_i, \quad W_{ij} = \rho X_i^* / \sigma_X^* + \sqrt{1 - \rho^2} \nu_{ij}, \\ X_i^* &= V_{i1} Z_i + V_{i0} \end{split}$$

$$(\theta_{011},\theta_{012},\theta_{013},\theta_{021},\theta_{022},\theta_{023},\rho,\sigma_Z^2,\sigma_\nu^2)=(1,0,0,0,0,0,0,1,1,1) \text{ and } n=2000.$$

		N	<b>1LE</b>			K	=2			K	=4	
	bias	std	rmse	size	bias	std	rmse	size	bias	std	rmse	size
					$\tau =$	1/4						
$\partial p_1/\partial x$	-3.24	1.36	3.51	66.98	0.74	2.63	2.74	4.30	1.13	2.73	2.95	7.86
$\partial p_1/\partial w_1$	2.32	1.64	2.84	30.74	-0.11	2.30	2.30	4.82	-0.31	2.29	2.31	6.54
$\partial p_1/\partial w_2$	0.48	0.75	0.90	9.40	-0.04	0.87	0.87	4.82	-0.08	0.87	0.87	5.36
$\partial p_2/\partial x$	1.96	1.17	2.28	39.44	-0.40	1.88	1.92	4.72	-0.63	1.93	2.03	6.74
$\partial p_2/\partial w_1$	-1.16	0.82	1.42	30.66	0.06	1.15	1.15	4.84	0.15	1.15	1.16	6.48
$\partial p_2/\partial w_2$	-0.96	1.50	1.78	9.48	0.09	1.74	1.74	4.98	0.17	1.74	1.75	5.44
$\partial p_0/\partial x$	1.28	1.01	1.63	25.28	-0.34	1.43	1.47	5.08	-0.50	1.48	1.56	7.36
$\partial p_0/\partial w_1$	-1.16	0.82	1.42	30.60	0.06	1.15	1.15	4.82	0.15	1.15	1.16	6.46
$\partial p_0/\partial w_2$	0.48	0.74	0.88	9.46	-0.05	0.87	0.87	4.90	-0.09	0.87	0.88	5.32

		N	<b>1LE</b>			K	=2			K	=4	
	bias	std	rmse	size	bias	std	rmse	size	bias	std	rmse	size
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$\partial p_1/\partial x$	-3.24	1.36	3.51	66.98	0.74	2.63	2.74	4.30	1.13	2.73	2.95	7.86
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$\partial p_1/\partial w_1$	2.32	1.64	2.84	30.74	-0.11	2.30	2.30	4.82	-0.31	2.29	2.31	6.54
$\partial p_1/\partial w_2$	0.48	0.75	0.90	9.40	-0.04	0.87	0.87	4.82	-0.08	0.87	0.87	5.36
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$\partial p_0/\partial x$	1.28	1.01	1.63	25.28	-0.34	1.43	1.47	5.08	-0.50	1.48	1.56	7.36
$\partial p_0/\partial w_1$	-1.16	0.82	1.42	30.60	0.06	1.15	1.15	4.82	0.15	1.15	1.16	6.46
$\partial p_0/\partial w_2$	0.48	0.74	0.88	9.46	-0.05	0.87	0.87	4.90	-0.09	0.87	0.88	5.32

		٨	ΛLE			K	=2			K	=4	
	bias	std	rmse	size	bias	std	rmse	size	bias	std	rmse	size
					$\tau =$	1/2						
$\partial p_1/\partial x$	-8.97	1.09	9.04	100.00	-1.69	2.60	3.10	9.84	0.97	2.89	3.05	6.04
$\partial p_1/\partial w_1$	6.44	1.53	6.62	98.96	1.39	2.44	2.81	12.14	-0.21	2.41	2.42	5.54
$\partial p_1/\partial w_2$	1.28	0.72	1.47	42.54	0.29	0.90	0.94	7.00	-0.06	0.92	0.92	5.00
$\partial p_2/\partial x$	5.22	0.97	5.31	99.98	1.05	1.89	2.16	10.16	-0.53	2.08	2.15	6.14
$\partial p_2/\partial w_1$	-3.21	0.77	3.30	98.96	-0.69	1.22	1.40	12.10	0.10	1.21	1.21	5.52
$\partial p_2/\partial w_2$	-2.52	1.41	2.88	42.82	-0.56	1.79	1.87	7.20	0.13	1.84	1.84	4.98
$\partial p_0/\partial x$	3.75	0.86	3.85	98.78	0.64	1.45	1.59	8.18	-0.44	1.59	1.65	6.50
$\partial p_0/\partial w_1$	-3.23	0.78	3.32	98.96	-0.70	1.22	1.41	12.06	0.10	1.21	1.21	5.48
$\partial p_0/\partial w_2$	1.23	0.69	1.41	42.90	0.28	0.89	0.93	7.20	-0.07	0.92	0.92	4.94

		Λ	ИLE			K	=2			K	=4	
	bias	std	rmse	size	bias	std	rmse	size	bias	std	rmse	size
					$\tau =$	1/2						
$\partial p_1/\partial x$	-8.97	1.09	9.04	100.00	-1.69	2.60	3.10	9.84	0.97	2.89	3.05	6.04
$\partial p_1/\partial w_1$	6.44	1.53	6.62	98.96	1.39	2.44	2.81	12.14	-0.21	2.41	2.42	5.54
$\partial p_1/\partial w_2$	1.28	0.72	1.47	42.54	0.29	0.90	0.94	7.00	-0.06	0.92	0.92	5.00
$\partial p_2/\partial x$	5.22	0.97	5.31	99.98	1.05	1.89	2.16	10.16	-0.53	2.08	2.15	6.14
$\partial p_2/\partial w_1$	-3.21	0.77	3.30	98.96	-0.69	1.22	1.40	12.10	0.10	1.21	1.21	5.52
$\partial p_2/\partial w_2$	-2.52	1.41	2.88	42.82	-0.56	1.79	1.87	7.20	0.13	1.84	1.84	4.98
$\partial p_0/\partial x$	3.75	0.86	3.85	98.78	0.64	1.45	1.59	8.18	-0.44	1.59	1.65	6.50
$\partial p_0/\partial w_1$	-3.23	0.78	3.32	98.96	-0.70	1.22	1.41	12.06	0.10	1.21	1.21	5.48
$\partial p_0/\partial w_2$	1.23	0.69	1.41	42.90	0.28	0.89	0.93	7.20	-0.07	0.92	0.92	4.94

		٨	ΛLE			K	=2			K	=4	
	bias	std	rmse	size	bias	std	rmse	size	bias	std	rmse	size
					$\tau =$	1/2						
$\partial p_1/\partial x$	-8.97	1.09	9.04	100.00	-1.69	2.60	3.10	9.84	0.97	2.89	3.05	6.04
$\partial p_1/\partial w_1$	6.44	1.53	6.62	98.96	1.39	2.44	2.81	12.14	-0.21	2.41	2.42	5.54
$\partial p_1/\partial w_2$	1.28	0.72	1.47	42.54	0.29	0.90	0.94	7.00	-0.06	0.92	0.92	5.00
$\partial p_2/\partial x$	5.22	0.97	5.31	99.98	1.05	1.89	2.16	10.16	-0.53	2.08	2.15	6.14
$\partial p_2/\partial w_1$	-3.21	0.77	3.30	98.96	-0.69	1.22	1.40	12.10	0.10	1.21	1.21	5.52
$\partial p_2/\partial w_2$	-2.52	1.41	2.88	42.82	-0.56	1.79	1.87	7.20	0.13	1.84	1.84	4.98
$\partial p_0/\partial x$	3.75	0.86	3.85	98.78	0.64	1.45	1.59	8.18	-0.44	1.59	1.65	6.50
$\partial p_0/\partial w_1$	-3.23	0.78	3.32	98.96	-0.70	1.22	1.41	12.06	0.10	1.21	1.21	5.48
$\partial p_0/\partial w_2$	1.23	0.69	1.41	42.90	0.28	0.89	0.93	7.20	-0.07	0.92	0.92	4.94

		Ν	1LE			K	=2			K	=4	
	bias	std	rmse	size	bias	std	rmse	size	bias	std	rmse	size
					$\tau = 0$	3/4						
$\partial p_1/\partial x$	-13.35	0.86	13.38	100.00	-6.83	2.64	7.32	80.32	0.71	3.22	3.29	4.74
$\partial p_1/\partial w_1$	9.69	1.45	9.80	100.00	4.95	2.65	5.61	65.52	0.01	2.62	2.62	5.34
$\partial p_1/\partial w_2$	1.81	0.69	1.94	75.30	1.01	0.89	1.35	26.08	-0.01	0.98	0.98	5.24
$\partial p_2/\partial x$	7.48	0.79	7.52	100.00	4.06	1.83	4.45	68.82	-0.37	2.32	2.35	5.74
$\partial p_2/\partial w_1$	-4.83	0.73	4.88	100.00	-2.47	1.32	2.81	65.46	-0.01	1.31	1.31	5.28
$\partial p_2/\partial w_2$	-3.51	1.33	3.76	75.60	-1.99	1.76	2.66	26.38	0.03	1.97	1.97	5.32
$\partial p_0/\partial x$	5.87	0.73	5.92	100.00	2.77	1.47	3.14	56.28	-0.34	1.77	1.80	5.82
$\partial p_0/\partial w_1$	-4.87	0.75	4.93	100.00	-2.48	1.33	2.81	65.40	-0.01	1.31	1.31	5.32
$\partial p_0/\partial w_2$	1.70	0.64	1.82	75.76	0.98	0.87	1.31	26.50	-0.02	0.99	0.99	5.30

		Ν	1LE			K	=2			K	=4	
	bias	std	rmse	size	bias	std	rmse	size	bias	std	rmse	size
					$\tau =$	3/4						
$\partial p_1/\partial x$	-13.35	0.86	13.38	100.00	-6.83	2.64	7.32	80.32	0.71	3.22	3.29	4.74
$\partial p_1/\partial w_1$	9.69	1.45	9.80	100.00	4.95	2.65	5.61	65.52	0.01	2.62	2.62	5.34
$\partial p_1/\partial w_2$	1.81	0.69	1.94	75.30	1.01	0.89	1.35	26.08	-0.01	0.98	0.98	5.24
$\partial p_2/\partial x$	7.48	0.79	7.52	100.00	4.06	1.83	4.45	68.82	-0.37	2.32	2.35	5.74
$\partial p_2/\partial w_1$	-4.83	0.73	4.88	100.00	-2.47	1.32	2.81	65.46	-0.01	1.31	1.31	5.28
$\partial p_2/\partial w_2$	-3.51	1.33	3.76	75.60	-1.99	1.76	2.66	26.38	0.03	1.97	1.97	5.32
$\partial p_0/\partial x$	5.87	0.73	5.92	100.00	2.77	1.47	3.14	56.28	-0.34	1.77	1.80	5.82
$\partial p_0/\partial w_1$	-4.87	0.75	4.93	100.00	-2.48	1.33	2.81	65.40	-0.01	1.31	1.31	5.32
$\partial p_0/\partial w_2$	1.70	0.64	1.82	75.76	0.98	0.87	1.31	26.50	-0.02	0.99	0.99	5.30

#### **Estimator: General Case**

#### MERM in the General Case:

$$\psi(X_i, S_i, \theta, \gamma) \equiv g(X_i, S_i, \theta) - \sum_{k=2}^K \gamma_k g_x^{(k)}(X_i, S_i, \theta), \quad \gamma = (\gamma_2, \dots, \gamma_K)'$$

#### **Assumption (Moderate Measurement Error)**

$$\sigma_n^2 = o\left(n^{-1/(K+1)}\right)$$

- Under smoothness conditions,  $\mathbb{E}[\psi(X_i, S_i, \theta_0, \gamma_{0n})] = o(n^{-1/2})$
- $\gamma_{0n}$  is determined by the moments of  $\varepsilon_i$
- · Need to bias correct the bias correction terms
  - the implementation is still simple
- Theory: need to ensure that the Taylor's expansion remainder is negligible
- We establish asymptotic normality of the MERM estimator

# **Summary**

- $\tau = \sigma_{\varepsilon}/\sigma_{X^*} \to 0$  is a better approximation
- we construct Measurement Error Robust Moments
  - ⇒ Simple estimation via GMM

#### More in the paper:

- Feasible to handle multivariate X
- Dependence, Panels
- Establish nonparametric identification in regression models
- Non-classical measurement errors

# Thank You!

















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# **Assumption: Moment Function**

**Assumption (Lipschitz-Polynomial).** For some functions  $b_j(x,r,\theta)$  for  $j\in\{1,\ldots,J\}$  s.t.,  $\forall x,x'\in\mathcal{X}$  and  $\forall (r,\theta)\in\mathcal{R}\times\Theta$ ,

$$\left\| g_x^{(K)}(x', s, \theta) - g_x^{(K)}(x, s, \theta) \right\| \leqslant \sum_{j=1}^J b_j(x, s, \theta) |x' - x|^j,$$

and  $\mathbb{E}\left[\sup_{\theta \in \Theta} b_j(X_i^*, S_i, \theta)\right] < C$  for  $j \in \{1, \dots, J\}$ 

- Key to show  $\mathbb{E}\left[\psi(X_i,S_i,\theta_0,\gamma_{0n})\right]=O(\sigma_n^{K+1})=o(n^{-1/2})$
- Satisfied in most models
- When  $|\varepsilon_i/\sigma_n|$  has M bounded moments, can allow J=M-K
- A similar condition is imposed on  $abla_{ heta}g_x^{(K)}(x,r, heta)$

# **Asymptotic Normality**

- Denote  $\hat{\beta} \equiv (\hat{\theta}', \hat{\gamma}')', \, \beta_{0n} \equiv (\theta_0', \gamma_{0n}')'$ 

#### Theorem (Asymptotic Normality)

Under standard assumptions,

$$n^{1/2} \Sigma^{-1/2} (\hat{\beta} - \beta_{0n}) \stackrel{d}{\to} N(0, I_{p+K-1}),$$

$$\Sigma = (\Psi^{*\prime} \Xi \Psi^{*})^{-1} \Psi^{*\prime} \Xi \Omega^{*} \Xi \Psi^{*} (\Psi^{*\prime} \Xi \Psi^{*})^{-1}$$

$$\Omega^{*} \equiv \mathbb{E} \left[ g(X_{i}^{*}, S_{i}, \theta_{0}) g(X_{i}^{*}, S_{i}, \theta_{0})' \right], \quad \Psi^{*} \equiv \mathbb{E} \left[ \nabla_{\beta} \psi(X_{i}^{*}, S_{i}, \theta_{0}, 0) \right]$$

- For asymptotic normality,  $\gamma_{0n} \to \gamma_0 = 0 \in \operatorname{int}(\Gamma)$  or  $n^{1/2}\gamma_{0n} \to \infty$
- $\Sigma$  can be consistently estimated, the standard inference tools apply

$$n^{-1} \sum_{i=1}^{n} \psi_i(\hat{\beta}) \psi_i(\hat{\beta})' = \Omega^* + o_{p,n}(1), \quad n^{-1} \sum_{i=1}^{n} \nabla_{\beta} \psi_i(\hat{\beta}) = \Psi^* + o_{p,n}(1)$$

