

MAS3301 Bayesian Statistics

Problems 2 and Solutions

Semester 2

2008-9

Problems 2

Useful integrals: In solving these problems you might find the following useful.

- Gamma functions: Let a and b be positive. Then

$$\int_0^{\infty} x^{a-1} e^{-bx} dx = \frac{\Gamma(a)}{b^a}$$

where

$$\Gamma(a) = \int_0^{\infty} x^{a-1} e^{-x} dx = (a-1)\Gamma(a-1).$$

If a is a positive integer then $\Gamma(a) = (a-1)!$.

- Beta functions: Let a and b be positive. Then

$$\int_0^1 x^{a-1} (1-x)^{b-1} dx = \frac{\Gamma(a)\Gamma(b)}{\Gamma(a+b)}.$$

1. We are interested in the mean, λ , of a Poisson distribution. We have a prior distribution for λ with density

$$f^{(0)}(\lambda) = \begin{cases} 0 & (\lambda \leq 0) \\ k_0(1+\lambda)e^{-\lambda} & (\lambda > 0) \end{cases}.$$

- (a)
 - i. Find the value of k_0 .
 - ii. Find the prior mean of λ .
 - iii. Find the prior standard deviation of λ .
 - (b) We observe data x_1, \dots, x_n where, given λ , these are independent observations from the Poisson(λ) distribution.
 - i. Find the likelihood.
 - ii. Find the posterior density of λ .
 - iii. Find the posterior mean of λ .
2. We are interested in the parameter, θ , of a binomial(n, θ) distribution. We have a prior distribution for θ with density

$$f^{(0)}(\theta) = \begin{cases} k_0\{\theta^2(1-\theta) + \theta(1-\theta)^2\} & (0 < \theta < 1) \\ 0 & (\text{otherwise}) \end{cases}.$$

- (a)
 - i. Find the value of k_0 .
 - ii. Find the prior mean of θ .
 - iii. Find the prior standard deviation of θ .
- (b) We observe x , an observation from the binomial(n, θ) distribution.

- i. Find the likelihood.
 - ii. Find the posterior density of θ .
 - iii. Find the posterior mean of θ .
3. We are interested in the parameter θ , of a binomial(n, θ) distribution. We have a prior distribution for θ with density

$$f^{(0)}(\theta) = \begin{cases} k_0 \theta^2 (1 - \theta)^3 & (0 < \theta < 1) \\ 0 & (\text{otherwise}) \end{cases}.$$

- (a)
 - i. Find the value of k_0 .
 - ii. Find the prior mean of θ .
 - iii. Find the prior standard deviation of θ .
 - (b) We observe x , an observation from the binomial(n, θ) distribution.
 - i. Find the likelihood.
 - ii. Find the posterior density of θ .
 - iii. Find the posterior mean of θ .
4. In a manufacturing process packages are made to a nominal weight of 1kg. All underweight packages are rejected but the remaining packages may be slightly overweight. It is believed that the excess weight X , in g, has a continuous uniform distribution on $(0, \theta)$ but the value of θ is unknown. Our prior density for θ is

$$f^{(0)}(\theta) = \begin{cases} 0 & (\theta < 0) \\ k_0/100 & (0 \leq \theta < 10) \\ k_0 \theta^{-2} & (10 \leq \theta < \infty) \end{cases}.$$

- (a)
 - i. Find the value of k_0 .
 - ii. Find the prior median of θ .
 - (b) We observe 10 packages and their excess weights, in g, are as follows.

3.8	2.1	4.9	1.8	1.7	2.1	1.4	3.6	4.1	0.8
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Assume that these are independent observations, given θ .

 - i. Find the likelihood.
 - ii. Find a function $h(\theta)$ such that the posterior density of θ is $f^{(1)}(\theta) = k_1 h(\theta)$, where k_1 is a constant.
 - iii. Evaluate the constant k_1 . (*Note that it is a very large number but you should be able to do the evaluation using a calculator*).
5. Repeat the analysis of the Chester Road example in section 6.3, using the same likelihood but with the following prior density.

$$f^{(0)}(\lambda) = \begin{cases} k_0 [1 + (8\lambda)^2]^{-1} & (0 < \lambda < \infty) \\ 0 & (\text{otherwise}) \end{cases}$$

- (a) Find the value of k_0 .
- (b) Use numerical methods in R to do the following.
 - i. Find the posterior density and plot a graph showing both the prior and posterior densities.
 - ii. Find the posterior mean and standard deviation.

Note: For the numerical calculations and the plot in part (b) I suggest that you use a range $0.0 \leq \lambda \leq 0.2$. When plotting the graph, it is easiest to plot the posterior first as this will determine the length of the vertical axis. The value of k_0 can be found analytically. If you do use numerical integration to find it, you will need a much wider range of values of λ .

6. We are interested in the parameter λ of a $\text{Poisson}(\lambda)$ distribution. We have a prior distribution for λ with density

$$f^{(0)}(\lambda) = \begin{cases} 0 & (\lambda < 0) \\ k_0 \lambda^3 e^{-\lambda} & (\lambda \geq 0) \end{cases}.$$

- (a) i. Find the value of k_0 .
 ii. Find the prior mean of λ .
 iii. Find the prior standard deviation of λ .
- (b) We observe x_1, \dots, x_n which are independent observations from the $\text{Poisson}(\lambda)$ distribution.
 i. Find the likelihood function.
 ii. Find the posterior density of λ .
 iii. Find the posterior mean of λ .
7. In a fruit packaging factory apples are examined to see whether they are blemished. A sample of n apples is examined and, given the value of a parameter θ , representing the proportion of apples which are blemished, we regard x , the number of blemished apples in the sample, as an observation from the $\text{binomial}(n, \theta)$ distribution. The value of θ is unknown.

Our prior density for θ is

$$f^{(0)}(\theta) = \begin{cases} k_0(20\theta(1-\theta)^3 + 1) & (0 \leq \theta \leq 1) \\ 0 & (\text{otherwise}) \end{cases}.$$

- (a) i. Show that, for $0 \leq \theta \leq 1$, the prior density can be written as

$$f^{(0)}(\theta) = \frac{1}{2} \left\{ \frac{\Gamma(6)}{\Gamma(2)\Gamma(4)} \theta^{2-1} (1-\theta)^{4-1} + \frac{\Gamma(2)}{\Gamma(1)\Gamma(1)} \theta^{1-1} (1-\theta)^{1-1} \right\}.$$

- ii. Find the prior mean of θ .
 iii. Find the prior standard deviation of θ .
- (b) We observe $n = 10$ apples and $x = 4$.
 i. Find the likelihood function.
 ii. Find the posterior density of θ .
 iii. Find the posterior mean of θ .
 iv. Use R to plot a graph showing both the prior and posterior densities of θ . (*Hint: It is easier to get the vertical axis right if you plot the posterior density and then superimpose the prior density, rather than the other way round.*)

Homework 2

Solutions to Questions 6, 7 of Problems 2 are to be submitted in the Homework Letterbox no later than 4.00pm on Monday February 23rd.

Solutions

1. (a) i.

$$\begin{aligned}\int_0^\infty f^{(0)}(\lambda) d\lambda &= k_0 \left\{ \int_0^\infty e^{-\lambda} d\lambda + \int_0^\infty \lambda e^{-\lambda} d\lambda \right\} \\ &= k_0 \{1 + 1\} = 2k_0\end{aligned}$$

Hence $k_0 = 1/2$.

ii.

$$\begin{aligned}E_0(\lambda) = \int_0^\infty \lambda f^{(0)}(\lambda) d\lambda &= \frac{1}{2} \left\{ \int_0^\infty \lambda e^{-\lambda} d\lambda + \int_0^\infty \lambda^2 e^{-\lambda} d\lambda \right\} \\ &= \frac{1}{2} \{1 + 2\} = \frac{3}{2} = 1.5\end{aligned}$$

iii.

$$\begin{aligned}E_0(\lambda^2) = \int_0^\infty \lambda^2 f^{(0)}(\lambda) d\lambda &= \frac{1}{2} \left\{ \int_0^\infty \lambda^2 e^{-\lambda} d\lambda + \int_0^\infty \lambda^3 e^{-\lambda} d\lambda \right\} \\ &= \frac{1}{2} \{2 + 6\} = \frac{8}{2}\end{aligned}$$

So

$$\text{var}_0(\lambda) = \frac{8}{2} - \left(\frac{3}{2}\right)^2 = \frac{16 - 9}{4} = \frac{7}{4}$$

and

$$\text{std.dev}_0(\lambda) = \sqrt{\frac{7}{4}} = \frac{\sqrt{7}}{2} = 1.323.$$

(b) i. Likelihood

$$L = \prod_{i=1}^n \frac{e^{-\lambda} \lambda^{x_i}}{x_i!} = \frac{e^{-n\lambda} \lambda^{\sum x_i}}{\prod x_i!} = \frac{e^{-n\lambda} \lambda^S}{\prod x_i!}$$

where

$$S = \sum_{i=1}^n x_i.$$

ii. Posterior density proportional to

$$\begin{aligned}f^{(0)}(\lambda)L &\propto (1 + \lambda)e^{-\lambda}e^{-n\lambda}\lambda^S \\ &\propto e^{-(n+1)\lambda}\lambda^S + e^{-(n+1)\lambda}\lambda^{S+1}\end{aligned}$$

The posterior density is

$$f^{(1)}(\lambda) = k_1 \left\{ e^{-(n+1)\lambda}\lambda^S + e^{-(n+1)\lambda}\lambda^{S+1} \right\}$$

where

$$\int_0^\infty f^{(1)}(\lambda) d\lambda = 1 = k_1 \left\{ \frac{\Gamma(S+1)}{(n+1)^{S+1}} + \frac{\Gamma(S+2)}{(n+1)^{S+2}} \right\}.$$

Hence

$$k_1 = \left\{ \frac{\Gamma(S+1)}{(n+1)^{S+1}} + \frac{\Gamma(S+2)}{(n+1)^{S+2}} \right\}^{-1}$$

and

$$f^{(1)}(\lambda) = \left\{ \frac{\Gamma(S+1)}{(n+1)^{S+1}} + \frac{\Gamma(S+2)}{(n+1)^{S+2}} \right\}^{-1} \left\{ e^{-(n+1)\lambda}\lambda^S + e^{-(n+1)\lambda}\lambda^{S+1} \right\}.$$

iii. Posterior mean

$$\begin{aligned}
E_1(\lambda) &= \int_0^\infty \lambda f^{(1)}(\lambda) d\lambda \\
&= k_1 \left\{ \int_0^\infty \lambda^{S+1} e^{-(n+1)\lambda} d\lambda + \int_0^\infty \lambda^{S+2} e^{-(n+1)\lambda} d\lambda \right\} \\
&= k_1 \left\{ \frac{\Gamma(S+2)}{(n+1)^{S+2}} + \frac{\Gamma(S+3)}{(n+1)^{S+3}} \right\} \\
&= \frac{\left\{ \frac{(S+1)}{(n+1)} + \frac{(S+1)(S+2)}{(n+1)^2} \right\}}{\left\{ 1 + \frac{(S+1)}{(n+1)} \right\}}
\end{aligned}$$

2. (a) i.

$$\begin{aligned}
\int_0^1 f^{(0)}(\theta) d\theta &= k_0 \left\{ \int_0^1 \theta^2(1-\theta) d\theta + \int_0^1 \theta(1-\theta)^2 d\theta \right\} \\
&= k_0 \left\{ \frac{\Gamma(3)\Gamma(2)}{\Gamma(5)} + \frac{\Gamma(2)\Gamma(3)}{\Gamma(5)} \right\} \\
&= 2k_0 \frac{\Gamma(3)\Gamma(2)}{\Gamma(5)} = 2k_0 \frac{2!1!}{4!} = \frac{k_0}{6}
\end{aligned}$$

Hence $k_0 = 6$.

ii. Prior mean

$$\begin{aligned}
E_0(\theta) &= \int_0^1 \theta f^{(0)}(\theta) d\theta = k_0 \left\{ \int_0^1 \theta^3(1-\theta) d\theta + \int_0^1 \theta^2(1-\theta)^2 d\theta \right\} \\
&= k_0 \left\{ \frac{\Gamma(4)\Gamma(2)}{\Gamma(6)} + \frac{\Gamma(3)\Gamma(3)}{\Gamma(6)} \right\} \\
&= k_0 \left\{ \frac{3!1! + 2!2!}{5!} \right\} = k_0 \left\{ \frac{6+4}{120} \right\} = \frac{1}{2}
\end{aligned}$$

iii.

$$\begin{aligned}
E_0(\theta^2) &= \int_0^1 \theta^2 f^{(0)}(\theta) d\theta \\
&= k_0 \left\{ \int_0^1 \theta^4(1-\theta) d\theta + \int_0^1 \theta^3(1-\theta)^2 d\theta \right\} \\
&= k_0 \left\{ \frac{\Gamma(5)\Gamma(2)}{\Gamma(7)} + \frac{\Gamma(4)\Gamma(3)}{\Gamma(7)} \right\} \\
&= k_0 \left\{ \frac{4!1! + 3!2!}{6!} \right\} = \left\{ \frac{24+12}{5 \times 24} \right\} = \frac{3}{10}
\end{aligned}$$

Hence

$$\text{var}_0(\theta) = \frac{3}{10} - \left(\frac{1}{2} \right)^2 = \frac{6-5}{20} = \frac{1}{20}$$

and

$$\text{std.dev}_0(\theta) = \frac{1}{\sqrt{20}} = 0.2236.$$

(b) i. Likelihood

$$L = \binom{n}{x} \theta^x (1-\theta)^{n-x}.$$

ii. Posterior density $f^{(1)}(\theta)$ proportional to $f^{(0)}(\theta)L$. Hence

$$\begin{aligned}
f^{(1)}(\theta) &= k_1 \{ \theta^2(1-\theta) + \theta(1-\theta)^2 \} \theta^x (1-\theta)^{n-x} \\
&= k_1 \{ \theta^{x+2}(1-\theta)^{n-x+1} + \theta^{x+1}(1-\theta)^{n-x+2} \}
\end{aligned}$$

Now

$$\begin{aligned}\int_0^1 f^{(1)}(\theta) d\theta = 1 &= k_1 \left\{ \int_0^1 \theta^{x+2}(1-\theta)^{n-x+1} d\theta + \int_0^1 \theta^{x+1}(1-\theta)^{n-x+2} d\theta \right\} \\ &= k_1 \left\{ \frac{\Gamma(x+3)\Gamma(n-x+2)}{\Gamma(n+5)} + \frac{\Gamma(x+2)\Gamma(n-x+3)}{\Gamma(n+5)} \right\}.\end{aligned}$$

Hence

$$k_1 = \left\{ \frac{\Gamma(x+3)\Gamma(n-x+2)}{\Gamma(n+5)} + \frac{\Gamma(x+2)\Gamma(n-x+3)}{\Gamma(n+5)} \right\}^{-1}$$

and

$$\begin{aligned}f^{(1)}(\theta) &= \left\{ \frac{\Gamma(x+3)\Gamma(n-x+2)}{\Gamma(n+5)} + \frac{\Gamma(x+2)\Gamma(n-x+3)}{\Gamma(n+5)} \right\}^{-1} \\ &\quad \times \left\{ \theta^{x+2}(1-\theta)^{n-x+1} + \theta^{x+1}(1-\theta)^{n-x+2} \right\}.\end{aligned}$$

iii. Posterior mean

$$\begin{aligned}E_0(\theta) &= \int_0^1 \theta f^{(1)}(\theta) d\theta \\ &= k_1 \left\{ \int_0^1 \theta^{x+3}(1-\theta)^{n-x+1} d\theta + \int_0^1 \theta^{x+2}(1-\theta)^{n-x+2} d\theta \right\} \\ &= k_1 \left\{ \frac{\Gamma(x+4)\Gamma(n-x+2)}{\Gamma(n+6)} + \frac{\Gamma(x+3)\Gamma(n-x+3)}{\Gamma(n+6)} \right\} \\ &= \frac{\left\{ \frac{\Gamma(x+4)\Gamma(n-x+2)}{\Gamma(n+6)} + \frac{\Gamma(x+3)\Gamma(n-x+3)}{\Gamma(n+6)} \right\}}{\left\{ \frac{\Gamma(x+3)\Gamma(n-x+2)}{\Gamma(n+5)} + \frac{\Gamma(x+2)\Gamma(n-x+3)}{\Gamma(n+5)} \right\}} \\ &= \frac{1}{n+5} \left\{ \frac{\Gamma(x+4)\Gamma(n-x+2) + \Gamma(x+3)\Gamma(n-x+3)}{\Gamma(x+3)\Gamma(n-x+2) + \Gamma(x+2)\Gamma(n-x+3)} \right\} \\ &= \frac{1}{n+5} \left\{ \frac{\Gamma(x+4) + \Gamma(x+3)(n-x+2)}{\Gamma(x+3) + \Gamma(x+2)(n-x+2)} \right\} \\ &= \frac{1}{n+5} \left\{ \frac{(x+3)(x+2) + (x+2)(n-x+2)}{(x+2) + (n-x+2)} \right\} \\ &= \left(\frac{x+2}{n+5} \right) \left\{ \frac{(x+3) + (n-x+2)}{(x+2) + (n-x+2)} \right\} \\ &= \frac{x+2}{n+4}\end{aligned}$$

3. (a) i.

$$\int_0^1 f^{(0)}(\theta) d\theta = k_0 \int_0^1 \theta^2 (1-\theta)^3 d\theta = k_0 \frac{\Gamma(3)\Gamma(4)}{\Gamma(7)}.$$

Hence

$$k_0 = \frac{\Gamma(7)}{\Gamma(3)\Gamma(4)} = \frac{6!}{2!3!} = \frac{6 \times 5 \times 4}{2} = \underline{60}.$$

ii. Prior mean

$$\begin{aligned} E_0(\theta) &= \int_0^1 \theta f^{(0)}(\theta) d\theta = k_0 \int_0^1 \theta^3 (1-\theta)^3 d\theta \\ &= \frac{\Gamma(4)\Gamma(4)}{\Gamma(8)} \frac{\Gamma(7)}{\Gamma(3)\Gamma(4)} = \frac{3}{7} = \underline{0.4286}. \end{aligned}$$

iii.

$$\begin{aligned} E_0(\theta^2) &= \int_0^1 \theta^2 f^{(0)}(\theta) d\theta = k_0 \int_0^1 \theta^4 (1-\theta)^3 d\theta \\ &= \frac{\Gamma(5)\Gamma(4)}{\Gamma(9)} \frac{\Gamma(7)}{\Gamma(3)\Gamma(4)} = \frac{4 \times 3}{8 \times 7} = \frac{3}{14} \end{aligned}$$

Hence

$$\text{var}_0(\theta) = \frac{3}{14} - \left(\frac{3}{7}\right)^2 = \frac{21-18}{98} = \frac{3}{98}$$

and

$$\text{std.dev}_0(\theta) = \sqrt{\frac{3}{98}} = \underline{0.1750}.$$

(b) i. Likelihood

$$L(\theta) = \binom{n}{x} \theta^x (1-\theta)^{n-x}.$$

ii. Posterior density

$$f^{(1)}(\theta) \propto f^{(0)}(\theta)L(\theta) = k_1 \theta^{x+2} (1-\theta)^{n-x+3}.$$

Now

$$\int_0^1 f^{(1)}(\theta) d\theta = 1 = k_1 \int_0^1 \theta^{x+2} (1-\theta)^{n-x+3} d\theta = k_1 \frac{\Gamma(x+3)\Gamma(n-x+4)}{\Gamma(n+7)}.$$

Hence

$$k_1 = \frac{\Gamma(n+7)}{\Gamma(x+3)\Gamma(n-x+4)}$$

and

$$f^{(1)}(\theta) = \frac{\Gamma(n+7)}{\Gamma(x+3)\Gamma(n-x+4)} \theta^{x+2} (1-\theta)^{n-x+3} \quad (0 < \theta < 1).$$

iii. Posterior mean

$$\begin{aligned} E_1(\theta) &= \int_0^1 \theta f^{(1)}(\theta) d\theta = k_1 \int_0^1 \theta^{x+3} (1-\theta)^{n-x+3} d\theta \\ &= \frac{\Gamma(n+7)}{\Gamma(x+3)\Gamma(n-x+4)} \frac{\Gamma(x+4)\Gamma(n-x+4)}{\Gamma(n+8)} = \frac{x+3}{n+7} \end{aligned}$$

4.

5.

6. (a) i. Value of k_0 :

$$\int_0^{\infty} \lambda^3 e^{-\lambda} d\lambda = \int_0^{\infty} \lambda^{4-1} e^{-\lambda} d\lambda = \Gamma(4) = 3! = 6$$

Hence

$$k_0 = \frac{1}{6}.$$

(1 mark)

- ii. Prior mean:

$$\begin{aligned} E_0(\lambda) &= \int_0^{\infty} \lambda k_0 \lambda^3 e^{-\lambda} d\lambda \\ &= k_0 \int_0^{\infty} \lambda^{5-1} e^{-\lambda} d\lambda \\ &= k_0 \Gamma(5) = \frac{4!}{3!} = 4 \end{aligned}$$

(1 mark)

- iii. Prior std.dev.:

$$\begin{aligned} E_0(\lambda^2) &= \int_0^{\infty} \lambda^2 k_0 \lambda^3 e^{-\lambda} d\lambda \\ &= k_0 \int_0^{\infty} \lambda^{6-1} e^{-\lambda} d\lambda \\ &= k_0 \Gamma(6) = \frac{5!}{3!} = 20 \end{aligned}$$

Hence $\text{var}_0(\lambda) = E_0(\lambda^2) - [E_0(\lambda)]^2 = 20 - 16 = 4$ and prior sd is

$$\sqrt{4} = 2.$$

(2 marks)

- (b) i. Likelihood:

$$L = \prod_{i=1}^n \frac{e^{-\lambda} \lambda^{x_i}}{x_i!} = \frac{e^{-n\lambda} \lambda^{\sum x_i}}{\prod x_i!}$$

(1 mark)

- ii. Posterior density proportional to

$$\lambda^3 e^{-\lambda} \times e^{-n\lambda} \lambda^{\sum x_i}$$

That is

$$\lambda^{\sum x_i + 4 - 1} e^{-(n+1)\lambda}$$

Now

$$\int_0^{\infty} \lambda^{a-1} e^{-b\lambda} d\lambda = \frac{\Gamma(a)}{b^a}$$

Hence the posterior density is

$$\frac{(n+1)^{\sum x_i + 4}}{\Gamma(\sum x_i + 4)} \lambda^{\sum x_i + 4 - 1} e^{-(n+1)\lambda}$$

(2 marks)

- iii. To find the posterior mean, increase the power of λ by 1 and integrate. Posterior mean is

$$\frac{(n+1)^{\sum x_i + 4}}{\Gamma(\sum x_i + 4)} \frac{\Gamma(\sum x_i + 5)}{(n+1)^{\sum x_i + 5}} = \frac{\sum x_i + 4}{n+1}$$

(1 mark)

7. (a) i. The expression given is proportional to the prior density since

$$\frac{\Gamma(6)}{\Gamma(2)\Gamma(4)} = \frac{5!}{3!} = 30$$

and $\frac{\Gamma(2)}{\Gamma(1)\Gamma(1)} = 1$

Now we only need to show that

$$\int_0^1 \frac{1}{2} \left\{ \frac{\Gamma(6)}{\Gamma(2)\Gamma(4)} \theta^{2-1} (1-\theta)^{4-1} + \frac{\Gamma(2)}{\Gamma(1)\Gamma(1)} \theta^{1-1} (1-\theta)^{1-1} \right\} d\theta = 1$$

and this follows since

$$\int_0^1 \theta^{2-1} (1-\theta)^{4-1} d\theta = \frac{\Gamma(2)\Gamma(4)}{\Gamma(6)}$$

and $\int_0^1 \theta^{1-1} (1-\theta)^{1-1} d\theta = \frac{\Gamma(1)\Gamma(1)}{\Gamma(2)}.$

(1 mark)

- ii. Prior mean:

$$\begin{aligned} E_0(\theta) &= \int_0^1 \theta f^{(0)}(\theta) d\theta \\ &= \frac{1}{2} \int_0^1 \left\{ \frac{\Gamma(6)}{\Gamma(2)\Gamma(4)} \theta^{3-1} (1-\theta)^{4-1} + \frac{\Gamma(2)}{\Gamma(1)\Gamma(1)} \theta^{2-1} (1-\theta)^{1-1} \right\} d\theta \\ &= \frac{1}{2} \left\{ \frac{\Gamma(6)}{\Gamma(2)\Gamma(4)} \frac{\Gamma(3)\Gamma(4)}{\Gamma(7)} + \frac{\Gamma(2)}{\Gamma(1)\Gamma(1)} \frac{\Gamma(2)\Gamma(1)}{\Gamma(3)} \right\} \\ &= \frac{1}{2} \left\{ \frac{2}{6} + \frac{1}{2} \right\} = \frac{5}{12} = \underline{0.4167} \end{aligned}$$

(2 marks)

- iii. Prior std. dev.:

$$\begin{aligned} E_0(\theta^2) &= \int_0^1 \theta^2 f^{(0)}(\theta) d\theta \\ &= \frac{1}{2} \int_0^1 \left\{ \frac{\Gamma(6)}{\Gamma(2)\Gamma(4)} \theta^{4-1} (1-\theta)^{4-1} + \frac{\Gamma(2)}{\Gamma(1)\Gamma(1)} \theta^{3-1} (1-\theta)^{1-1} \right\} d\theta \\ &= \frac{1}{2} \left\{ \frac{\Gamma(6)}{\Gamma(2)\Gamma(4)} \frac{\Gamma(4)\Gamma(4)}{\Gamma(8)} + \frac{\Gamma(2)}{\Gamma(1)\Gamma(1)} \frac{\Gamma(3)\Gamma(1)}{\Gamma(4)} \right\} \\ &= \frac{1}{2} \left\{ \frac{3 \times 2}{7 \times 6} + \frac{1}{3} \right\} = \frac{5}{21} \end{aligned}$$

Hence

$$\text{var}_0(\theta) = \frac{5}{21} - \left(\frac{5}{12} \right)^2 = 0.06448$$

so the prior std. dev. is

$$\sqrt{0.06448} = \underline{0.2539}.$$

(2 marks)

(b) i. Likelihood:

$$L = \binom{10}{4} \theta^4 (1 - \theta)^6$$

(1 mark)

ii. Posterior density:

Posterior *propto* Prior \times Likelihood

$$f^{(1)}(\theta) = k_1 \left\{ \frac{\Gamma(6)}{\Gamma(2)\Gamma(4)} \theta^{6-1} (1 - \theta)^{10-1} + \frac{\Gamma(2)}{\Gamma(1)\Gamma(1)} \theta^{5-1} (1 - \theta)^{7-1} \right\}$$

$$\begin{aligned} \int_0^1 f^{(1)}(\theta) d\theta &= k_1 \left\{ \frac{\Gamma(6)}{\Gamma(2)\Gamma(4)} \frac{\Gamma(6)\Gamma(10)}{\Gamma(16)} + \frac{\Gamma(2)}{\Gamma(1)\Gamma(1)} \frac{\Gamma(5)\Gamma(7)}{\Gamma(12)} \right\} \\ &= k_1 \left\{ \frac{5 \times 4 \times 3 \times 2 \times 1 \times 9 \times 8 \times 7 \times 6 \times 5 \times 4 \times 3 \times 2 \times 1}{15 \times 14 \times 13 \times 12 \times 11 \times 10 \times 9 \times 8 \times 7 \times 6 \times 5 \times 4 \times 3 \times 2 \times 1} + \frac{4 \times 3 \times 2 \times 1 \times 6 \times 5 \times 4 \times 3 \times 2 \times 1}{11 \times 10 \times 9 \times 8 \times 7 \times 6 \times 5 \times 4 \times 3 \times 2 \times 1} \right\} \\ &= k_1 \left\{ \frac{2}{7 \times 13 \times 3 \times 11} + \frac{1}{11 \times 10 \times 3 \times 7} \right\} \\ &= \frac{k_1}{3 \times 11 \times 7} \left\{ \frac{2}{13} + \frac{1}{10} \right\} \\ &= \frac{k_1}{3 \times 11 \times 7} \left\{ \frac{33}{130} \right\} = \frac{k_1}{7 \times 130} \end{aligned}$$

Hence $k_1 = 7 \times 130 = 910$.

Posterior density:

$$f^{(1)}(\theta) = 910 \{ 20\theta^{6-1}(1 - \theta)^{10-1} + \theta^{5-1}(1 - \theta)^{7-1} \}$$

(2 marks)

iii. Posterior mean:

$$\begin{aligned} E_1(\theta) &= \int_0^1 \theta f^{(1)}(\theta) d\theta \\ &= 910 \int_0^1 \{ 20\theta^{7-1}(1 - \theta)^{10-1} + \theta^{6-1}(1 - \theta)^{7-1} \} d\theta \\ &= 910 \left\{ 20 \frac{\Gamma(7)\Gamma(10)}{\Gamma(17)} + \frac{\Gamma(6)\Gamma(7)}{\Gamma(13)} \right\} \\ &= 910 \left\{ \frac{20 \times 6 \times 5 \times 4 \times 3 \times 2 \times 1 \times 9 \times 8 \times 7 \times 6 \times 5 \times 4 \times 3 \times 2 \times 1}{16 \times 15 \times 14 \times 13 \times 12 \times 11 \times 10 \times 9 \times 8 \times 7 \times 6 \times 5 \times 4 \times 3 \times 2 \times 1} + \frac{5 \times 4 \times 3 \times 2 \times 1 \times 6 \times 5 \times 4 \times 3 \times 2 \times 1}{12 \times 11 \times 10 \times 9 \times 8 \times 7 \times 6 \times 5 \times 4 \times 3 \times 2 \times 1} \right\} \\ &= 910 \left\{ \frac{1}{14 \times 13 \times 11 \times 2} + \frac{1}{11 \times 9 \times 8 \times 7} \right\} = \underline{0.3914} \end{aligned}$$

(2 marks)

iv. Plot: suitable R commands:

```
> theta<-seq(0,1,0.01)
> prior<-0.5*(20*theta*((1-theta)^3)+1)
> post<-910*(20*(theta^5)*((1-theta)^9)+(theta^4)*((1-theta)^6))
> plot(theta,post,type="l",xlab=expression(theta),ylab="Density")
> lines(theta,prior,lty=2)
```

The plot is shown in Figure 1.

(2 marks)

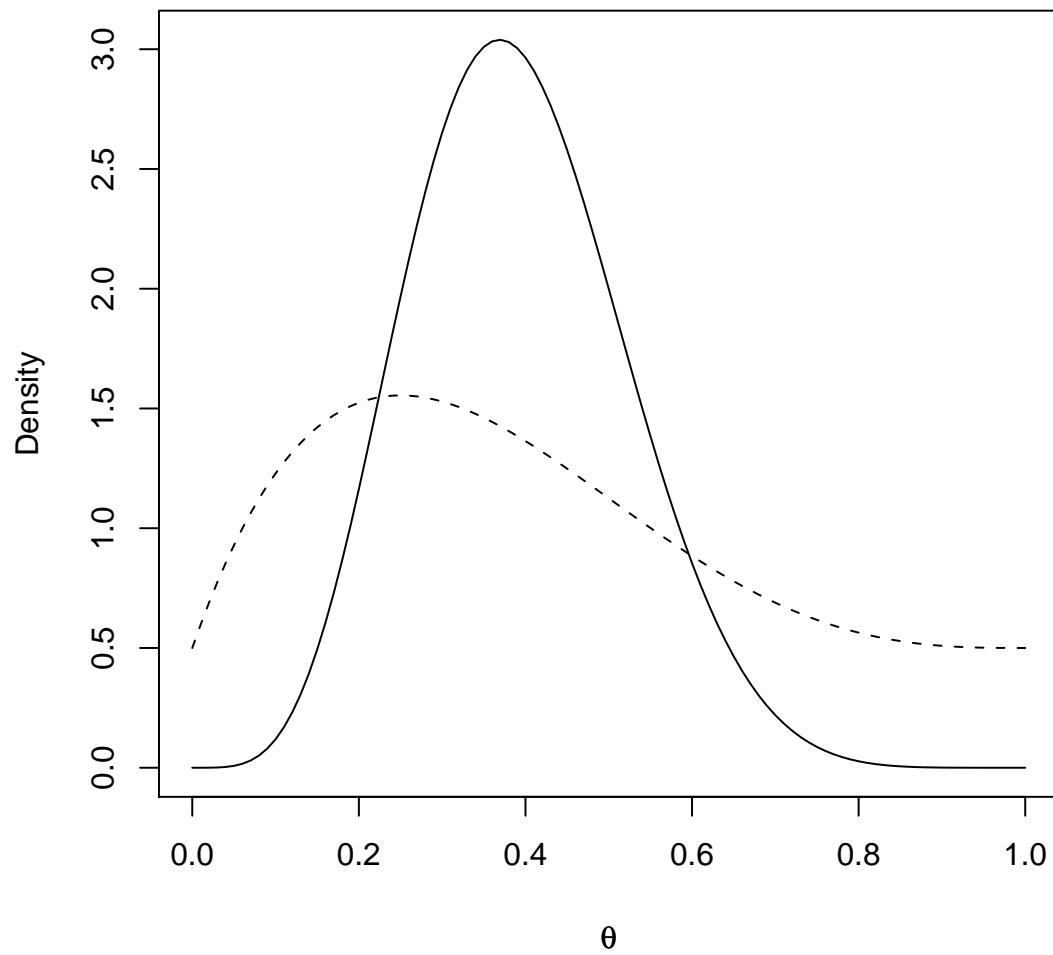


Figure 1: Prior (dashes) and posterior (solid line) density functions for θ .