

$$\begin{aligned}
 1(a). \quad P(X=Y) &= \sum_{k=1}^n P(X=k, Y=k) \\
 &= \sum_{k=1}^n P(X=k)P(Y=k) \quad (\text{independence}) \\
 &= \sum_{k=1}^n \frac{1}{n} \cdot \frac{1}{n} \\
 &= \frac{1}{n}
 \end{aligned}$$

1(b) Method 1:

$$\begin{aligned}
 P(X < Y) &= \sum_{i=1}^{n-1} \sum_{j=i+1}^n P(X=i, Y=j) \\
 &= \sum_{i=1}^{n-1} \sum_{j=i+1}^n P(X=i)P(Y=j) \quad (\text{independence}) \\
 &= \sum_{i=1}^{n-1} \sum_{j=i+1}^n \frac{1}{n^2} \\
 &= \sum_{i=1}^{n-1} \frac{n-i}{n^2} \\
 &= \frac{1}{n} \cdot (n-1) - \frac{1}{n^2} \cdot \frac{(n-1) \cdot [(n-1)+1]}{2} \\
 &= \frac{n-1}{n} - \frac{n-1}{2n} \\
 &= \frac{n-1}{2n}
 \end{aligned}$$

Method 2: By symmetry (X and Y follows the same distribution),

$$P(X > Y) = P(X < Y).$$

$$\begin{aligned}
 \text{So } 1 &= P(X > Y) + P(X < Y) + P(X=Y) \\
 &= 2P(X < Y) + P(X=Y)
 \end{aligned}$$

$$\begin{aligned}
 \text{And } P(X < Y) &= \frac{1}{2}[1 - P(X=Y)] \\
 &= \frac{1}{2}(1 - \frac{1}{n}) \quad (\text{from 1(a)}) \\
 &= \frac{n-1}{2n}
 \end{aligned}$$

$$1(c) \quad P(\max(X, Y) = k)$$

$$= P(X = k, Y < k) + P(X < k, Y = k) + P(X = k, Y = k)$$

$$= P(X = k)P(Y < k) + P(X < k)P(Y = k) + P(X = k)P(Y = k)$$

$$= \frac{1}{n} \cdot \frac{k-1}{n} + \frac{k-1}{n} \cdot \frac{1}{n} + \frac{1}{n^2}$$

$$= \frac{2(k-1)+1}{n^2}$$

$$= \frac{2k-1}{n^2}$$

□

2(a) For $z > 0$,

$$P(Z > z) = P(\min(X, Y) > z)$$

$$= P(X > z, Y > z)$$

$$= P(X > z)P(Y > z)$$

$$= e^{-\lambda z} \cdot e^{-\mu z}$$

$$= e^{-(\lambda+\mu)z}$$

Hence the cdf of Z is, for $z > 0$,

$$F_Z(z) = 1 - P(Z > z) = 1 - e^{-(\lambda+\mu)z}$$

and the pdf is

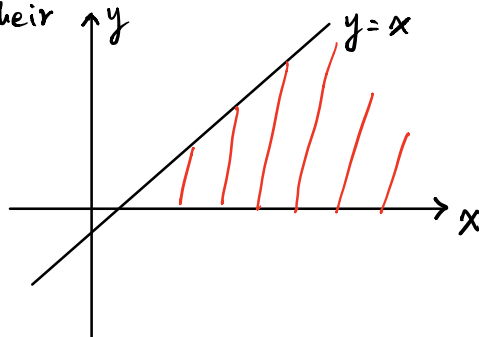
$$f_Z(z) = F'_Z(z) = (\lambda+\mu)e^{-(\lambda+\mu)z}, \text{ for } z > 0.$$

So $Z \sim \text{Exp}(\lambda+\mu)$.

2(b). Since X, Y are independent, their joint pdf is

$$f(x, y) = \lambda \mu e^{-\lambda x - \mu y}$$

The shaded area in red is the region of $\{X \geq Y\}$.



And hence

$$P(X \geq Y) = \int_0^{+\infty} \int_0^x \lambda \mu e^{-\lambda x - \mu y} dy dx$$

$$= - \int_0^{+\infty} \lambda e^{-\lambda x} \left[\int_0^x e^{-\mu y} d(-\mu y) \right] dx$$

$$= - \int_0^{+\infty} \lambda e^{-\lambda x} \left(e^{-\mu y} \Big|_{y=0}^{y=x} \right) dx$$

$$= \int_0^{+\infty} \lambda e^{-\lambda x} (1 - e^{-\mu x}) dx$$

$$= \lambda \int_0^{+\infty} e^{-\lambda x} dx - \lambda \int_0^{+\infty} e^{-(\lambda+\mu)x} dx$$

$$= -e^{-\lambda x} \Big|_{x=0}^{x=+\infty} + \frac{\lambda}{\lambda+\mu} e^{-(\lambda+\mu)x} \Big|_{x=0}^{x=+\infty}$$

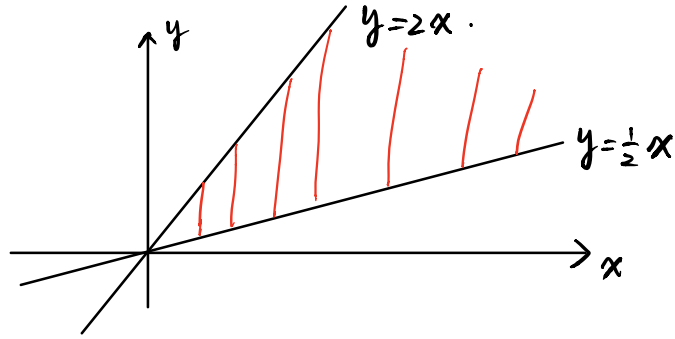
$$= 1 - \frac{\lambda}{\lambda+\mu}$$

$$= \frac{\mu}{\lambda+\mu}$$

2(c). When $\lambda = \mu$, the joint pdf of (X, Y) becomes

$$f(x, y) = \lambda^2 e^{-\lambda(x+y)}$$

Then the region of $\left\{ \frac{1}{2} < \frac{X}{Y} < 2 \right\} = \left\{ \frac{1}{2}X < Y < 2X \right\}$ is



$$\begin{aligned}
 \text{Therefore, } P\left(\frac{1}{2} < \frac{X}{Y} < 2\right) &= \int_0^{+\infty} \int_{\frac{1}{2}x}^{2x} \lambda^2 e^{-\lambda(x+y)} dy dx \\
 &= -\int_0^{+\infty} \lambda e^{-\lambda x} \int_{\frac{1}{2}x}^{2x} e^{-\lambda y} d(-\lambda y) dx \\
 &= -\int_0^{+\infty} \lambda e^{-\lambda x} \left(e^{-\lambda y} \Big|_{y=\frac{1}{2}x}^{y=2x} \right) dx \\
 &= -\int_0^{+\infty} \lambda e^{-\lambda x} \left(e^{-2\lambda x} - e^{-\frac{1}{2}\lambda x} \right) dx \\
 &= \int_0^{+\infty} \lambda e^{-\frac{3}{2}\lambda x} dx - \int_0^{+\infty} \lambda e^{-3\lambda x} dx \\
 &= -\frac{2}{3} e^{-\frac{3}{2}\lambda x} \Big|_{x=0}^{x=+\infty} + \frac{1}{3} e^{-3\lambda x} \Big|_{x=0}^{x=+\infty} \\
 &= -\frac{2}{3} (0-1) + \frac{1}{3} (0-1) \\
 &= \frac{2}{3} - \frac{1}{3} \\
 &= \frac{1}{3}
 \end{aligned}$$

□