

Statistical Inference for Nanopore Sequencing with a Biased Random Walk Model

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ABSTRACT Nanopore sequencing promises long read-lengths and single-molecule resolution, but the stochastic motion of the DNA molecule inside the pore is a current barrier to high accuracy reads. We develop a method of statistical inference that explicitly accounts for this error and demonstrate that high accuracy (>99.9%) sequence inference is feasible even under highly diffusive motion by using a hidden Markov model to jointly analyze multiple stochastic reads. Using this model, we place bounds on achievable inference under a range of experimental parameters.

Received for publication 1 June 2013 and in final form 1 June 2013.

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Rapid advances in DNA sequencing technologies have led to an explosion in available nucleotide sequence data, greatly enhancing our understanding of the genomic basis of many biological processes. However, the short length of the raw reads means high coverage is required for reliable sequence assembly. Nanopore sequencing has emerged as a candidate to supersede current generation sequencing and allow for theoretically unlimited read length (1). A number of strategies have been proposed, with the common basis of detecting individual nucleotides as they pass through a nanometer-scale aperture in a thin membrane separating two electrolytes (2). To date, a significant obstacle of **nanopore approaches** has been overcoming the fast stochastic motion of the individual molecules as they are driven through the pore (3,4). Ideally the passage of DNA through the pore would be unidirectional and each base would have a well-resolved signal. Recent methods have demonstrated an ability to controllably ‘ratchet’ DNA molecules through a nanopore one base at a time, although motion of the molecule can still occur in both forward and backward directions within a single read (5–7). Unidirectional motion remains difficult to reliably achieve, leading to a source of error in the read sequence not previously addressed by existing algorithms.

In this letter, we consider a simple physical model of DNA translocation as a one-dimensional biased random walk and analyze how diffusive molecular motion affects the achievable read accuracy of nanopore sequencing. To do so, we simulate noisy data where the true sequence is unknown, and develop a statistical technique to estimate the most likely DNA sequence associated with the simulated read signal. Our method uses hidden Markov models (HMMs), which have recently been used to study multi-base resolution in a nanopore sequencer (8), but have not yet been applied to the problem of diffusive motion inside the pore. This work shows that combining multiple reads from a given input DNA sequence allows accurate estimates of the sequence, both in the presence of highly diffusive molecular motion

and at high base-call error rates.

We model motion through the nanopore as one-dimensional diffusion with driving force F and unit base step size a , which results in a random walk with fixed forward bias (9)

$$\beta = \frac{1}{2} + \frac{Fa}{4k_B T}. \quad (1)$$

Given an input DNA sequence of length L , we generate a single output read by stepping through the sequence with a bias β , at each step making a base-call with error probability ϵ , which is independent of error introduced due to backward motion. Finally, we assume an appropriate method of making a base-call from the raw signal, which has been computationally studied in (10). From this simple physical model, we generate a set of N simulated reads X_n , which each have a unique length $T_n \geq L$. N is a notion of sequence coverage in terms of identical molecules sequenced. Coverage could result from either amplification or resequencing of the initial molecule. A schematic representation of a nanopore sequencing device is shown in Fig. 1 A. In Fig. 1 B a series of random walks at several different forward biases is shown, along with the resulting read length distribution. The expected length of a read is

$$\langle T \rangle = \frac{1}{2(\beta - 0.5)} L. \quad (2)$$

{ok there actually is something wrong here. In the limit $\beta \rightarrow 1$, the read length converges to $T \rightarrow L$. Conversely, in the limit $\beta \rightarrow 1/2$ however, the length tends to $T \rightarrow \pi^2 L^2 / 4$. Does this relationship follow from a linear expansion around $\beta = 1/2$?}

Given this set of N read sequences, the statistical task is to infer the sequence most likely to have generated the observed data. An experiment similar to this model was demonstrated on very short sequences with tunneling current data in (11). Here we extend the approach to the longer sequences expected from a nanopore device.

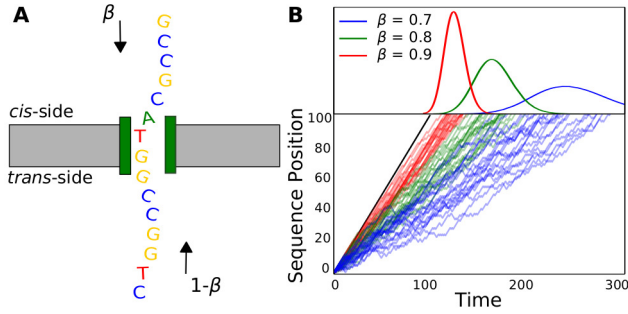


FIGURE 1 Random walk model of nanopore sequencing. (A) Schematic representation of an idealized nanopore sequencing device. (B) Transit time distribution of a set of random walks for different forward biases (sequence length $L = 100$).

In our HMM formulation, each output read is modeled as a discrete set of observed states, $\mathbf{x} = \{x_1 \dots x_T\}$, $x_i \in (A, C, G, T)$, a vector of observed bases, and a discrete set of hidden states, $\mathbf{z} = \{z_1 \dots z_T\}$, $z_i \in (1 \dots L)$, the unknown position along the sequence. An HMM is described by three model parameters, the initial state distribution $\pi = p(z_1)$, the hidden state transition matrix $A = p(z_t | z_{t-1})$, and an emission distribution $S = p(\mathbf{x}_n | \mathbf{z}_n)$. π and A are fixed by the experimental conditions. The inference problem in this model is to estimate the emission distribution, S , which acts as an implicit representation of our sequence,

$$S_{dl} = (1 - \epsilon)p(\mathbf{x}_n = d | \mathbf{z}_n = l) + \epsilon/4 \quad (3)$$

In practice, S is a $4 \times L$ matrix with a multinomial distribution over the possible nucleotides at each position (see Fig. 2). We maximize the likelihood, $p(\mathbf{X} | \theta)$, with respect to the model parameters using an implementation of the expectation-maximization algorithm (13). The joint probability of data and states can be written as a product over the independent output reads, $p(\mathbf{X}, \mathbf{Z} | \theta) = \prod_n p(\mathbf{X}_n, \mathbf{Z}_n | \theta)$, from which follows that we can perform expectation updates on each read individually before averaging results in the maximization step (see Supporting Material for full model derivation). The resulting shared parameter estimation scheme incorporates all reads while allowing an efficient, parallel calculation.

After a convergence criterion on the likelihood is satisfied, we recover an estimated emission distribution S , which can be converted to an estimate for the DNA sequence by taking $\max_d S_{dl}$. The final inference accuracy is measured as the Levenshtein distance between the input sequence and inferred sequence, normalized by L . The algorithm has complexity $O(NLT)$, where T is $O(L)$ in the limit of $\beta = 1$, and $O(L^2)$ in the limit of $\beta = 0.5$. *{note: this $O(L)$ vs $O(L^2)$ argument needs to be consistent with equation 2}* An example of the output of this algorithm showing the relationship between the true sequence and the inferred sequence distribution is shown in Fig. 2 A. Additionally, the

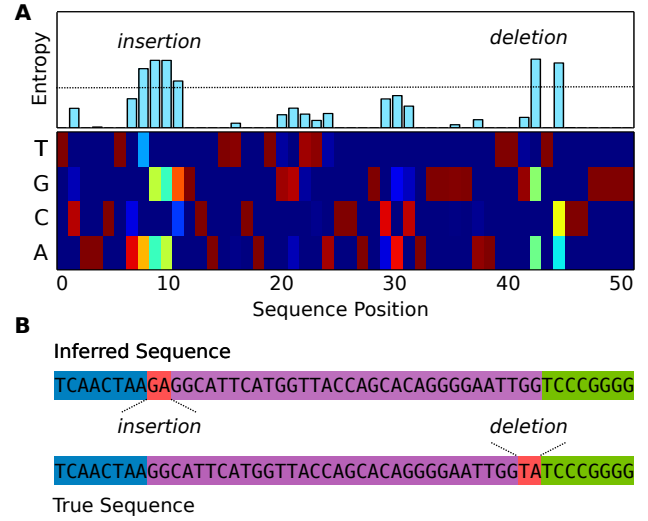


FIGURE 2 The output of the statistical model and its relationship to the true sequence. (A) The output sequence inference distribution, S . The inference entropy is plotted above S . (B) Taking argmax of each columns yields the sequence inference, which maps onto the true sequence in chunks, with errors in the form of insertion and deletion artifacts. The entropy signature heuristically identifies these positions.

column-wise entropy of the sequence inference, defined as $H_l = -\sum_d S_{dl} \log S_{dl}$, is shown in Fig. 2 C. Inference errors in this model are generally associated with increases in the inference entropy around the error position l , suggesting a possible approach that further analysis might systematically identify and remove the inference errors due to bidirectional motion.

We examined the performance of the algorithm over a range of parameter values in order to identify a minimal experimental configuration capable of sequence inference at a given accuracy. First, we consider how to select realistic values for our parameters. $\beta = 0.5$ corresponds to completely symmetric diffusion and is unlikely to reach the end of the sequence without first exiting from the *cis* side of the nanopore (these reads are discarded in the data generation). $\beta = 1.0$ corresponds to nondiffusive motion and is a trivial case in this model. Lu et al. report an experimental bias term of $Fa/4k_B T = 0.2$ ($\beta = 0.7$), for which they claim accurate sequence recovery is near impossible (4). We take this as a starting point for investigation, exploring the range $\beta = 0.7$ to $\beta = 1.0$. An appropriate base-call error rate will be device specific and is presently unknown, so we examine the range $\epsilon = 0.0$ to $\epsilon = 0.5$. The number of reads, N , is potentially unlimited, and we examine until convergence.

A sweep across the parameter space is shown in Fig. 3. For $N=1$, the average inference accuracy goes roughly as $\sim \beta$, a baseline measure of performance (Fig. 3 A). In Fig. 3 B we plot the relationship between forward bias and the number of reads. As expected, the strongest determinant of inference accuracy is the number of reads, N . This is expected

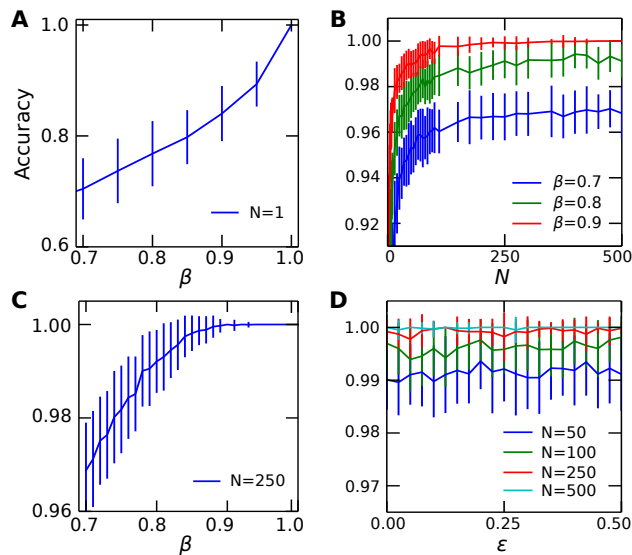


FIGURE 3 Parameter Sweeps. Inference accuracy (measured as $1 - \text{Levenshtein distance}/\text{length}$) is plotted as a function of experimental parameters. (A) Baseline performance for a single read, $N = 1$. (B) Parameter sweep across number of reads, N . (C) Asymptotic performance (large N) performance vs forward bias β . (D) error rate ϵ vs N . Accuracy is independent of error rate.

because each read contributes an independent observation of the input DNA sequence. We observe an inference error rate that falls rapidly with increasing N , even at highly diffusive motion. Generally, the achievable accuracy converges at a fixed β even with the addition of more reads. This asymptotic performance is plotted in Fig. 3 C. A threshold accuracy of 99% is achievable for β as low as 0.8; an accuracy exceeding 99.9% is achievable with $\beta > 0.88$. Finally, inference accuracy is independent of the base-call error rate, provided each nucleotide in the sequence is visited sufficient times (Fig. 3 D).

Precise control over the motion of the DNA sequence inside the nanopore has previously been assumed a necessary prerequisite for accurate nanopore sequencing. Here we have demonstrated that high accuracy inference is possible, even in the presence of diffusive motion, with the integration of data under an appropriate statistical model. The strength of the method lies in the way in which multiple reads of the input sequence are independently modeled and then iteratively combined to yield a joint estimate of the true sequence. Under this model, reasonably accurate sequence inference is achievable even within the experimental constraints of today's nanopore sequencers ($\beta \approx 0.7$). Designers of nanopore sequencers may find that there is a tradeoff between sequencing speed and maintaining unidirectional molecular motion. In this scenario, techniques such as the random walk model may allow increased throughput without a loss of accuracy.

SUPPORTING MATERIAL

The full inference model is available at BPJ Online. The source code used to perform the analysis in this paper will shortly be available via <http://stochseq.github.io>

ACKNOWLEDGEMENTS

The authors gratefully acknowledge helpful discussions with David Blei and David Pfau. KJE and CHW were supported by NIH grant U54-CA121852 (National Center for Multi-scale Analysis of Genomic and Cellular Networks). JWM was supported through the NWO Rubicon Fellowship 680-50-1016.

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