## STAT234: Lecture 4 - Sums of random Variables

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# Binomial Distribution and Normal approx.

More generally if  $X_1, X_2, \dots, X_n$  be independent identically distributed (iid) Ber(p) random variables, then

$$Y_n = \sum_{i=1}^n X_i \sim Bin(n, p)$$

So,

$$E(Y_n) = np$$
  $var(Y_n) = np(1-p)$ 

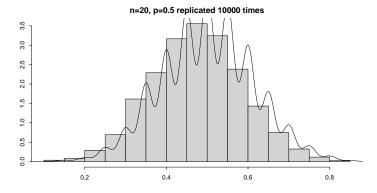
As  $n \to \infty$ 

$$Y_n \approx N(np, np(1-p))$$

Lets look at variables generated at Bin(20, 0.5) and repeat the process 10,000 times.

```
Y1 <- rbinom(10000, 20, p=0.5)/ 20; summary(Y1)
```

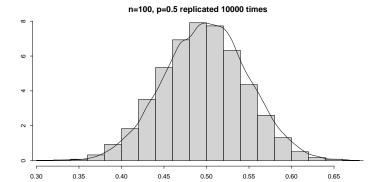
```
Min. 1st Qu. Median Mean 3rd Qu. Max. 0.100 0.400 0.500 0.501 0.600 0.850
```



Now lets look at variables generated at Bin(100, 0.5) and repeat the process 10,000 times.

```
Y2 <- rbinom(10000, 100, p=0.5)/ 100; summary(Y2)
```

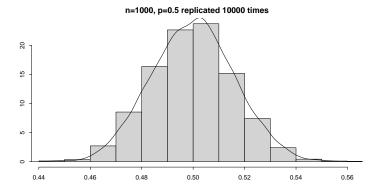
```
Min. 1st Qu. Median Mean 3rd Qu. Max. 0.31 0.47 0.50 0.50 0.53 0.67
```



Now lets look at variables generated at Bin(1000, 0.5) and repeat the process 10,000 times.

```
Y3 <- rbinom(10000, 1000, p=0.5)/ 1000; summary(Y3)
```

```
Min. 1st Qu. Median Mean 3rd Qu. Max. 0.442 0.489 0.500 0.500 0.511 0.561
```



## Normal approx. of Binomial

http://digitalfirst.bfwpub.com/stats\_applet/stats\_applet\_2\_cltbinom.html We define  $\hat{p} = \frac{1}{n} \sum_{i=1}^{n} X_i$ .

$$\hat{p} pprox N\left(p, rac{p(1-p)}{n}
ight)$$

This approximation is better when p is not to close to 0 or 1.

Discrete	Continuous
X = 3	2.5 < X < 3.5
X > 3	X > 3.5
<i>X</i> ≥ 3	X > 2.5
X < 3	X < 2.5
<i>X</i> ≤ 3	X < 3.5

lacktriangleq X is approximately Normal with mean np and sd  $\sqrt{np(1-p)}$ 

- ▶ X is approximately Normal with mean np and sd  $\sqrt{np(1-p)}$
- ▶ But *X* is a discrete random variable and Normal is a continuous one.

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$$P(X \le 3) = P(\frac{X - 9.6}{1.959} \le \frac{3 - 9.6}{1.959})$$

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$$P(X < 3.5) = P(\frac{X - 9.6}{1.959} \le \frac{3.5 - 9.6}{1.959}) = P(z \le -3.114) = 0.00092$$

- ▶ X is approximately Normal with mean np and sd  $\sqrt{np(1-p)}$ 
  - ▶ But X is a discrete random variable and Normal is a
  - continuous one. ►  $X \sim Bin(16, 0.6)$ . How to find  $P(X \le 3)$ ?
  - P(X < 3) = P(X < 3.5)? ▶ By R: pbinom(3,16,0.6) # 0.000938

  - $P(X \le 3) = P(\frac{X 9.6}{1.050} \le \frac{3 9.6}{1.050}) = P(z \le -3.369) = 0.00037$
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  - ► To find P(X > 10) Show the Ree Hise 30.5 or 9.5?

### Continuous distribution

We would want to define Pr(X = x) for a continuous random variable X, but unfortunately it is 0. We will soon see why

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So how do we proceed to build something like a probability table for discrete variable?

We define cumulative density function (cdf)

$$F(x) = Pr(X \le x)$$

If we differentiate this function, we get **probability density function** (pdf)

$$rac{d}{dx}F(x)=f(x)$$
 $f(x)\geq 0 \int\limits_{ ext{Week }\#5: ext{ Page 12 of }33}f(x)dx=1$ 

### Normal Distribution

For normal distribution with mean  $\mu$  and variance  $\sigma^2$ ,

$$\phi(x) := \int \frac{1}{\sqrt{2\pi}} exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right)$$

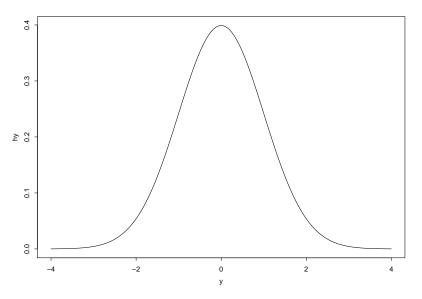
Cumulative distribution

$$\Phi(x) := \int_{-\infty}^{x} \phi(x) dx$$

The mean and variance

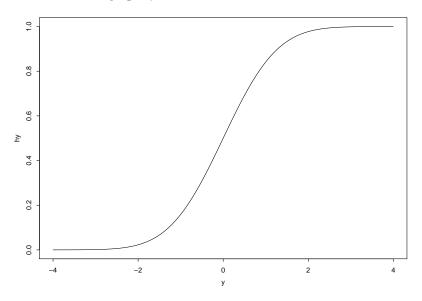
$$E(X) = \mu$$
  $var(X) = \sigma^2$ 

# probability density graph



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# cumulative density graph



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Suppose we consider two random variables X and Y which follow normal distribution.

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  $Y \sim N(\mu, \sigma^2)$ 

Assume now that X and Y are independent.

What is the distribution of X + Y

$$E(X + Y) = E(X) + E(Y) = \mu + \mu = 2\mu$$
  
 $var(X + Y) = var(X) + var(Y) = \sigma^2 + \sigma^2 = 2\sigma^2$ 

To find the distribution, we perform a large number of repetitions (close to infinity) of the experiment of drawing random variables X and Y.

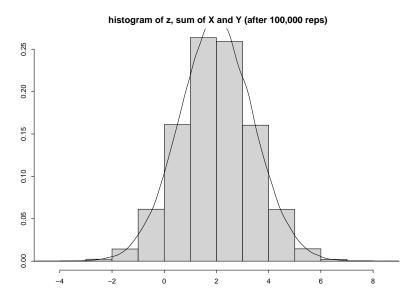
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```
x <- rnorm(100000, 1, 1);
y <- rnorm(100000, 1, 1);
z <- x+y;
length(z)</pre>
[1] 100000
```



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Assume now three random variables X, Y and W and consider their sum

$$Z = X + Y + W$$

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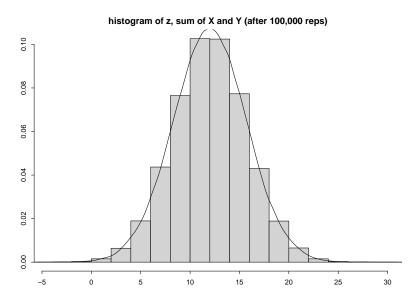
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Suppose we repeat it 100,000 times.

```
x <- rnorm(100000, 1, 1);
y <- rnorm(100000, 1, 3);
w <- rnorm(100000, 10, 2);
z <- x+y+w;
length(z)</pre>
```

[1] 100000



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#### Result

If  $X_1, X_2, \dots, X_n$  are independent normal random variables such that

$$X_i \sim N\left(\mu_i, \sigma_i^2\right)$$

Then if we define

$$Z = X_1 + X_2 + \cdots + X_n$$

then

$$Z \sim N\left(\sum_{i=1}^{n} \mu_i, \sum_{i=1}^{n} \sigma_i^2\right)$$

# Scaling of normal variables

Let X be a random variable that follows a distribution

$$X \sim N(\mu, \sigma^2)$$

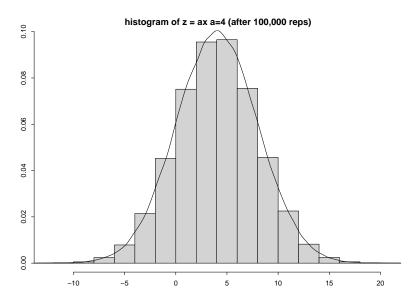
Let a be a constant.

What is the distribution of aX.

```
x <- rnorm(100000, 1, 1);
a <- 4;
z <- a*x
length(z)</pre>
```

[1] 100000

### Sum of normal variables



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#### Result

If X be a normal random variables such that

$$X \sim N(\mu, \sigma^2)$$

Let a be a constant,

$$Z = aX$$

$$E(Z) = aE(X) = a\mu$$

$$var(Z) = var(aX) = a^2 var(X) = a^2 \sigma^2$$

and

$$Z \sim N(a\mu, a^2\sigma^2)$$

#### Result

If  $X_1, X_2, \dots, X_n$  are *independent* normal random variables such that

$$X_i \sim N(\mu_i, \sigma_i^2)$$

Then if we define

$$Z = c_1 X_1 + c_2 X_2 + \cdots + c_n X_n$$

Check

$$E(Z) = \sum_{i=1}^{n} c_i \mu_i$$
  $var(Z) = \sum_{i=1}^{n} c_i^2 \sigma_i^2$ 

then

$$Z \sim N \left( \sum_{\text{Week} = 15: \text{ Page 25 } j \in \$3}^{n} c_i^2 \sigma_i^2 \right)$$

# Corollary of Previous Result

If  $X_1, X_2, \dots, X_n$  are independent normal random variables such that

$$X_i \sim N(\mu, \sigma^2)$$

then if we define

$$Z = \frac{1}{n} \sum_{i=1}^{n} X_i$$

and

$$Z \sim N\left(\mu, \frac{\sigma^2}{n}\right)$$

We showed that sum of independent normal random variables is normal.

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Linear transformation of normal random variables is normal.

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Its Binomial.

#### Central Limit Theorem

All the results we discussed today are true for sum of any n independent variables, where n can be small or large. What about large n?

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All the results we discussed today are true for sum of any n independent variables, where n can be small or large. What about large n?

if  $X_1, X_2, \cdots, X_n$  be independent identically distributed (iid) random variables coming from a distribution with mean  $\mu$  and variance  $\sigma^2$ , then

$$\sum_{i=1}^{n} X_{i} pprox N\left(n\mu, n\sigma^{2}
ight)$$
 n large

and

$$\frac{1}{n}\sum_{i=1}^{n}X_{i}\approx N\left(\mu,\frac{\sigma^{2}}{n}\right)$$
 n large

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As a special case, Sum of a large number of Bernoulli or Binomial random variables is approximately normal.

When the underlying distribution is discrete, remember *continuity correction*.

We define a moment generating function (mgf) as a function of t

$$mgf(t) := E\left(e^{tX}\right) = \int e^{tx} f(x) dx$$

where f(x) is the probability density function (pdf) observed at point x.

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**Very important note**: Moment generating functions characterize distributions for most cases.

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What does that mean?

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What does that mean?

# Normal Moment generating function

For a normal random variable X, it can be shown that the moment generating function has the following form

$$\mathit{mgf}(t) := \exp\left(\mu t + rac{1}{2}t^2\sigma^2
ight)$$

Using the characterizing property of mgf, if suppose we have

$$mgf(t): exp\left(2t+rac{1}{2}6t^2
ight)$$

then this is the mgf of

$$X \sim N(2, 6)$$

# Moment generating function for sums

If  $X_1, X_2, \dots, X_n$  are independent random variables following a distribution with pdf f(x), then the moment generating function of X

$$mgf_X(t): E(e^{tX})$$

If we define

$$Y = X_1 + X_2 + \cdots + X_n$$

$$mgf_Y(t): E(e^{tY}) = E(e^{tX_1+tX_2+\cdots+tX_n})$$

We can show that (check Canvas for proof)

$$mgf_Y(t) = [mgf_X(t)]^n$$

We use the mgf properties to show that sum of normal random variables is normal, or sum of linear transformation of normal random variables is normal.

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