

# Homework 1 - Introduction to Probabilistic Graphical Models

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## 1 Bayesian Networks

1. Consider a simple Markov Chain structure  $X \rightarrow Y \rightarrow Z$ , where all variables are binary. You are required to:
  - (a) Write a code (using your preferred programming language) that generates a distribution (not necessarily a valid BN one) over the 3 variables.  
[ in the notebook ]
  - (b) Write a code that verifies whether a distribution is a valid BN distribution.  
[ in the notebook ]
  - (c) Using your code, generate 10000 distributions and compute the fraction of distributions that are valid BN distributions.  
[ in the notebook ]
2. Given the following Bayesian Network

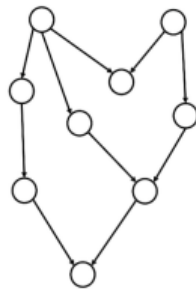


Figure 1: A Bayesian network.

Figure 1: Bayesian Network

- (a) Propose a topological ordering of this graph  
In Figure 2, the topological ordering is:
  - i.  $A \rightarrow B \rightarrow C \rightarrow D \rightarrow E \rightarrow F \rightarrow G \rightarrow H \rightarrow I$
  - ii.  $A \rightarrow B \rightarrow C \rightarrow E \rightarrow D \rightarrow F \rightarrow G \rightarrow I \rightarrow H$
- (b) Let  $\mathbf{X}$  be a random vector that is Markov with respect to the graph. We assume that the random variables  $X_i$  are binary. Write all the local conditional independence

$X_A$  has no parents, so no independence condition applies here.

$X_B$  has no parents, so no independence condition applies here.

$X_C$  is conditionally independent of all other nodes given its parent  $X_A, X_B$  :

$$X_C \perp \{X_D, X_E, X_F, X_G, X_H, X_I\} \mid \{X_A, X_B\}$$

$X_D$  is conditionally independent of all other nodes given its parent  $X_A$  :

$$X_D \perp \{X_B, X_C, X_E, X_F, X_G, X_H, X_I\} \mid X_A$$

$X_E$  is conditionally independent of all other nodes given its parents  $X_C$  and  $X_D$  :

$$X_E \perp \{X_B, X_C, X_D, X_F, X_G, X_H, X_I\} \mid \{X_C, X_D\}$$

$X_F$  is conditionally independent of all other nodes given its parent  $X_B$  :

$$X_F \perp \{X_A, X_C, X_D, X_E, X_G, X_H, X_I\} \mid X_B$$

$X_G$  is conditionally independent of all other nodes given its parents  $X_E$  and  $X_F$  :

$$X_G \perp \{X_A, X_B, X_C, X_D, X_H, X_I\} \mid \{X_E, X_F\}$$

$X_H$  is conditionally independent of all other nodes given its parent  $X_D$  :

$$X_H \perp \{X_A, X_B, X_C, X_E, X_F, X_G, X_I\} \mid X_D$$

$X_I$  is conditionally independent of all other nodes given its parents  $X_G$  and  $X_H$  :

$$X_I \perp \{X_A, X_B, X_C, X_D, X_E, X_F\} \mid \{X_G, X_H\}$$

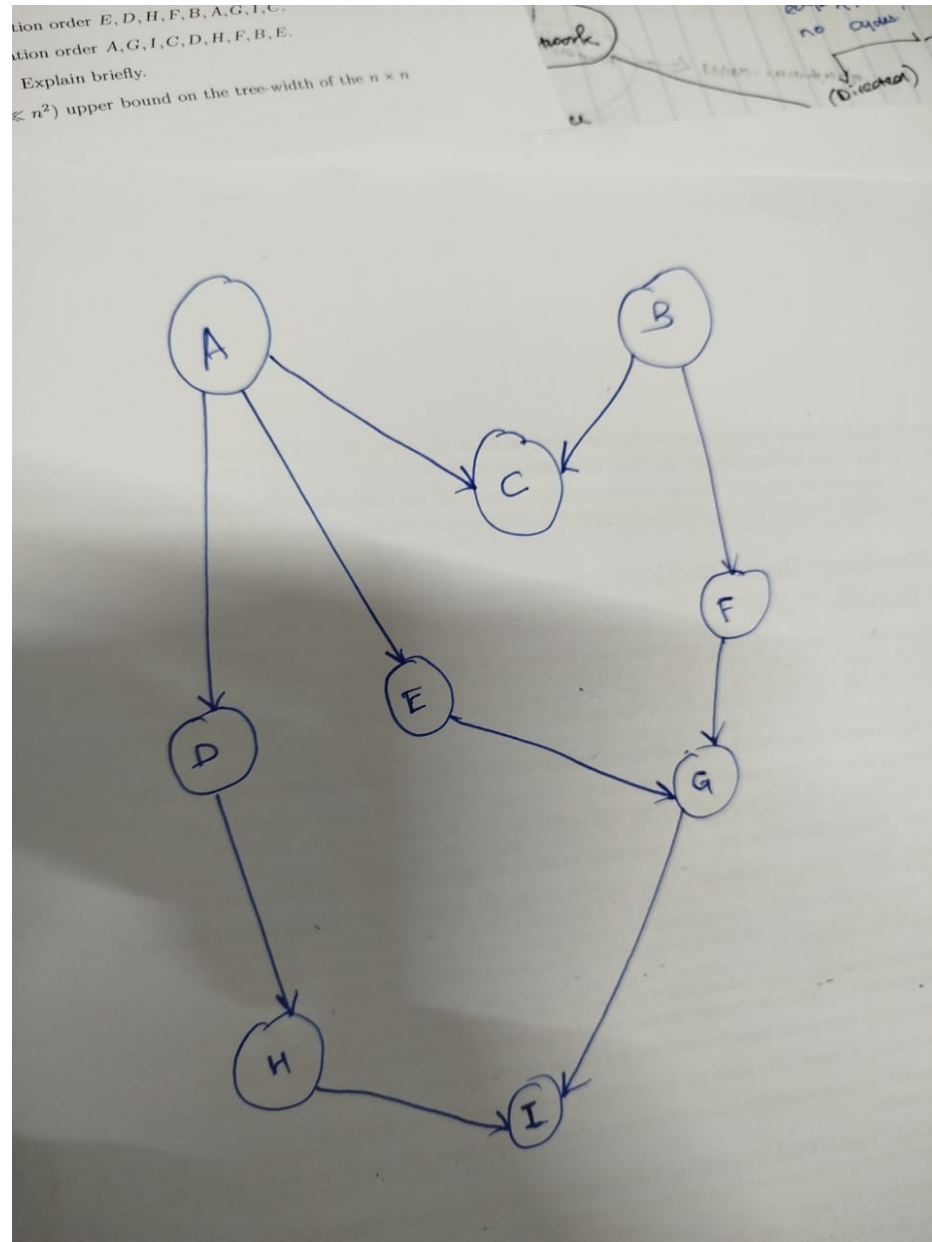


Figure 2: Bayesian Network

3. State True or False, and briefly justify your answer within 3 lines. The statements are either direct consequences of theorems in Koller and Friedman (2009, Ch. 3), or have a short proof. In the follows,  $P$  is a distribution and  $G$  is a BN structure.
- (a) If  $A \perp B \mid C$  and  $A \perp C \mid B$ , then  $A \perp B$  and  $A \perp C$ . (Suppose the joint distribution of  $A, B, C$  is positive.) (This is a general probability question not related to BNs.)
- **False.** Whilst local independence implies global independence, global independence does not imply local independence.

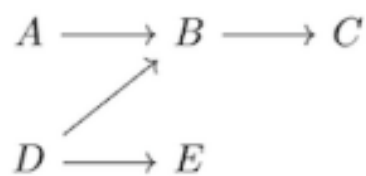


Figure 2: A Bayesian network.

Figure 3: Bayesian Network

- (b) In Figure 2,  $E \perp C \mid B$   
**True.** The path from  $E$  to  $C$  is blocked by the node  $B$ , so  $E$  is conditionally independent of  $C$  given  $B$ .
- (c) in Figure 2,  $A \perp E \mid C$   
**False.**  $A$  is not conditionally independent of  $E$  given  $C$  because there is a path from  $A$  to  $E$  that is not blocked by  $C$ .

## 2 Markov Networks

Let  $\mathbf{X} = (X_1, \dots, X_d)$  be a random vector with mean  $\boldsymbol{\mu}$  and covariance matrix  $\boldsymbol{\Sigma}$ . The partial correlation matrix  $\mathbf{R}$  of  $\mathbf{X}$  is a  $d \times d$  matrix where each entry  $R_{ij} = \rho(X_i, X_j \mid \mathbf{X}_{-ij})$  is the partial correlation between  $X_i$  and  $X_j$  given the  $d - 2$  remaining variables  $\mathbf{X}_{-ij}$ . Let  $\boldsymbol{\Theta} = \boldsymbol{\Sigma}^{-1}$  be the inverse covariance matrix of  $\mathbf{X}$ .

We will prove the relation between  $\mathbf{R}$  and  $\boldsymbol{\Theta}$ , and furthermore how  $\boldsymbol{\Theta}$  characterizes conditional independence in Gaussian graphical models.

$$\begin{pmatrix} \Theta_{ii} & \Theta_{ij} \\ \Theta_{ji} & \Theta_{jj} \end{pmatrix} = \begin{pmatrix} \text{Var}[e_i] & \text{Cov}[e_i, e_j] \\ \text{Cov}[e_i, e_j] & \text{Var}[e_j] \end{pmatrix}^{-1}$$

for any  $i, j \in [d], i \neq j$ . Here  $e_i$  is the residual resulting from the linear regression of  $X_{-ij}$  to  $X_i$ , and similarly  $e_j$  is the residual resulting from the linear regression of  $X_{-ij}$  to  $X_j$ .

1. (10 points) Show that

$$R_{ij} = -\frac{\Theta_{ij}}{\sqrt{\Theta_{ii}\Theta_{jj}}}$$

We start from the given matrix equation:

$$\begin{pmatrix} \Theta_{ii} & \Theta_{ij} \\ \Theta_{ji} & \Theta_{jj} \end{pmatrix} = \begin{pmatrix} \text{Var}[e_i] & \text{Cov}[e_i, e_j] \\ \text{Cov}[e_i, e_j] & \text{Var}[e_j] \end{pmatrix}^{-1}.$$

let us find the determinant of the matrix on the right hand side:

$$\det \begin{pmatrix} \text{Var}[e_i] & \text{Cov}[e_i, e_j] \\ \text{Cov}[e_i, e_j] & \text{Var}[e_j] \end{pmatrix} = \text{Var}[e_i]\text{Var}[e_j] - \text{Cov}[e_i, e_j]^2.$$

Using the determinant, let us find the inverse of the matrix on the right hand side:

$$\frac{1}{\text{Var}[e_i]\text{Var}[e_j] - \text{Cov}[e_i, e_j]^2} \begin{pmatrix} \text{Var}[e_j] & -\text{Cov}[e_i, e_j] \\ -\text{Cov}[e_i, e_j] & \text{Var}[e_i] \end{pmatrix}$$

Let us push the determinant into the matrix:

$$\begin{pmatrix} \frac{\text{Var}[e_j]}{\text{Var}[e_i]\text{Var}[e_j] - \text{Cov}[e_i, e_j]^2} & -\frac{\text{Cov}[e_i, e_j]}{\text{Var}[e_i]\text{Var}[e_j] - \text{Cov}[e_i, e_j]^2} \\ -\frac{\text{Cov}[e_i, e_j]}{\text{Var}[e_i]\text{Var}[e_j] - \text{Cov}[e_i, e_j]^2} & \frac{\text{Var}[e_i]}{\text{Var}[e_i]\text{Var}[e_j] - \text{Cov}[e_i, e_j]^2} \end{pmatrix}$$

We can now equate the two matrices:

$$\begin{pmatrix} \Theta_{ii} & \Theta_{ij} \\ \Theta_{ji} & \Theta_{jj} \end{pmatrix} = \begin{pmatrix} \frac{\text{Var}[e_j]}{\text{Var}[e_i]\text{Var}[e_j] - \text{Cov}[e_i, e_j]^2} & -\frac{\text{Cov}[e_i, e_j]}{\text{Var}[e_i]\text{Var}[e_j] - \text{Cov}[e_i, e_j]^2} \\ -\frac{\text{Cov}[e_i, e_j]}{\text{Var}[e_i]\text{Var}[e_j] - \text{Cov}[e_i, e_j]^2} & \frac{\text{Var}[e_i]}{\text{Var}[e_i]\text{Var}[e_j] - \text{Cov}[e_i, e_j]^2} \end{pmatrix}$$

To show that  $R_{ij} = -\frac{\Theta_{ij}}{\sqrt{\Theta_{ii}\Theta_{jj}}}$ , we need to show that:

$$\begin{aligned} R_{ij} &= -\frac{\Theta_{ij}}{\sqrt{\Theta_{ii}\Theta_{jj}}} \\ R_{ij} &= -\frac{-\frac{\text{Cov}[e_i, e_j]}{\text{Var}[e_i]\text{Var}[e_j] - \text{Cov}[e_i, e_j]^2}}{\sqrt{\frac{\text{Var}[e_i]}{\text{Var}[e_i]\text{Var}[e_j] - \text{Cov}[e_i, e_j]^2} \frac{\text{Var}[e_j]}{\text{Var}[e_i]\text{Var}[e_j] - \text{Cov}[e_i, e_j]^2}}} \\ R_{ij} &= -\frac{-\frac{\text{Cov}[e_i, e_j]}{\text{Var}[e_i]\text{Var}[e_j] - \text{Cov}[e_i, e_j]^2}}{\sqrt{\frac{\text{Var}[e_i]\text{Var}[e_j]}{(\text{Var}[e_i]\text{Var}[e_j] - \text{Cov}[e_i, e_j]^2)^2}}} \\ R_{ij} &= -\frac{-\frac{\text{Cov}[e_i, e_j]}{\text{Var}[e_i]\text{Var}[e_j] - \text{Cov}[e_i, e_j]^2}}{\frac{\sqrt{\text{Var}[e_i]\text{Var}[e_j]}}{\text{Var}[e_i]\text{Var}[e_j] - \text{Cov}[e_i, e_j]^2}} \end{aligned}$$

But we know that:

$$\begin{aligned} \Theta_{ij} &= -\frac{\text{Cov}[e_i, e_j]}{\text{Var}[e_i]\text{Var}[e_j] - \text{Cov}[e_i, e_j]^2} \\ R_{ij} &= -\frac{-\text{Cov}[e_i, e_j]}{\sqrt{\text{Var}[e_i]\text{Var}[e_j]}} \\ R_{ij} &= \frac{\text{Cov}[e_i, e_j]}{\sqrt{\text{Var}[e_i]\text{Var}[e_j]}} \end{aligned}$$

let us rearrange this equation:

$$\text{Cov}[e_i, e_j] = R_{ij} \sqrt{\text{Var}[e_i]\text{Var}[e_j]}$$

Then we know:

$$\begin{aligned} \theta_{ij} &= -\frac{\text{Cov}[e_i, e_j]}{\text{Var}[e_i]\text{Var}[e_j] - \text{Cov}[e_i, e_j]^2} \\ \theta_{ij} &= -\frac{R_{ij} \sqrt{\text{Var}[e_i]\text{Var}[e_j]}}{\text{Var}[e_i]\text{Var}[e_j] - \text{Cov}[e_i, e_j]^2} \end{aligned}$$

let us have  $R_{ij}$  in the LHS:

$$R_{ij} = -\frac{\theta_{ij}\text{Var}[e_i]\text{Var}[e_j] - \text{Cov}[e_i, e_j]^2}{\sqrt{\text{Var}[e_i]\text{Var}[e_j]}}$$

we know  $\theta_{ii}$  and  $\Theta_{jj}$  as:

$$\theta_{ii} = \frac{\text{Var}[e_i]}{\text{Var}[e_i]\text{Var}[e_j] - \text{Cov}[e_i, e_j]^2}$$

$$\theta_{jj} = \frac{\text{Var}[e_j]}{\text{Var}[e_i]\text{Var}[e_j] - \text{Cov}[e_i, e_j]^2}$$

let us multiply the two:

$$\theta_{ii}\theta_{jj} = \frac{\text{Var}[e_i]\text{Var}[e_j]}{(\text{Var}[e_i]\text{Var}[e_j] - \text{Cov}[e_i, e_j]^2)^2}$$

we then proceed to find the square root of the product:

$$\sqrt{\theta_{ii}\theta_{jj}} = \frac{\sqrt{\text{Var}[e_i]\text{Var}[e_j]}}{\text{Var}[e_i]\text{Var}[e_j] - \text{Cov}[e_i, e_j]^2}$$

Let us rearrange the equation:

$$\text{Var}[e_i]\text{Var}[e_j] - \text{Cov}[e_i, e_j]^2 = \frac{\sqrt{\text{Var}[e_i]\text{Var}[e_j]}}{\sqrt{\theta_{ii}\theta_{jj}}}$$

We can then proceed and replace this in the equation for  $R_{ij}$ :

$$R_{ij} = -\frac{\theta_{ij} \frac{\sqrt{\text{Var}[e_i]\text{Var}[e_j]}}{\sqrt{\theta_{ii}\theta_{jj}}}}{\sqrt{\text{Var}[e_i]\text{Var}[e_j]}}$$

This simplifies to:

$$R_{ij} = -\frac{\theta_{ij}}{\sqrt{\theta_{ii}\theta_{jj}}}$$

We have therefore shown that  $R_{ij} = -\frac{\Theta_{ij}}{\sqrt{\Theta_{ii}\Theta_{jj}}}$ .

2. (15 points) From the above result and the relation between independence and correlation, we know

$$\Theta_{ij} = 0 \iff R_{ij} = 0 \implies X_i \perp X_j \mid X_{-ij}$$

Note the last implication only holds in one direction. Now suppose  $X \sim N(\mu, \Sigma)$  is jointly Gaussian. Show that  $R_{ij} = 0 \implies X_i \perp X_j \mid X_{-ij}$ .

$$R_{ij} = -\frac{\Theta_{ij}}{\sqrt{\Theta_{ii}\Theta_{jj}}}.$$

If  $R_{ij} = 0$ , then  $\Theta_{ij} = 0$ .

For a jointly Gaussian distribution, the inverse covariance matrix  $\Theta = \Sigma^{-1}$  encodes conditional independence. Specifically:

$$\Theta_{ij} = 0 \iff X_i \perp X_j \mid X_{-ij}.$$

This is because the off-diagonal elements of  $\Theta$  represent the conditional dependence between variables after accounting for all other variables.

Since  $R_{ij} = 0$  implies  $\Theta_{ij} = 0$ , it follows that:

$$X_i \perp X_j \mid X_{-ij}.$$

### 3 Exact Inference - Variable Elimination

Reference materials for this problem:

- Jordan textbook Ch. 3, available at <https://people.eecs.berkeley.edu/~jordan/prelims/chapter3.pdf>
- Koller and Friedman (2009, Ch. 9 and Ch. 10)

#### 3.1 Variable elimination on a grid [10 points]

Consider the following Markov network:

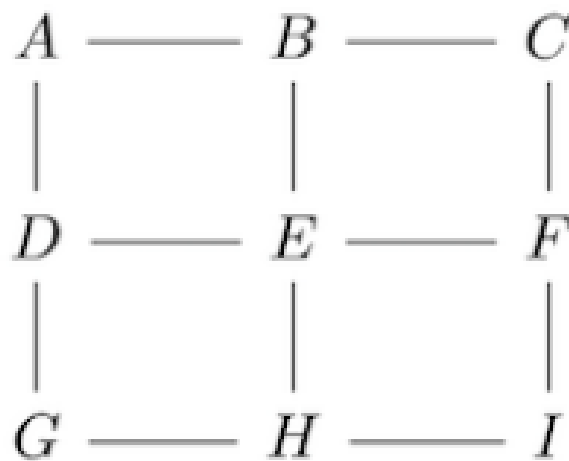


Figure 4: Markov Network

We are going to see how tree-width, a property of the graph, is related to the intrinsic complexity of variable elimination of a distribution

- (5 points) Write down largest clique(s) for the elimination order  $E, D, H, F, B, A, G, I, C$ .  
 We start by eliminating  $E$ , its neighbors are  $D, F, H, B$ , so the clique here is:  $\{D, F, H, B\}$ . size: **4**  
 We then eliminate  $D$ , its neighbors are  $A, B, F, G, H$ , so the clique here is:  $\{A, B, F, G, H\}$ . size: **5**  
 We then eliminate  $H$ , its neighbors are  $G, A, I, B, F$ , so the clique here is:  $\{G, A, I, B, F\}$ . size: **5**  
 We then eliminate  $F$ , its neighbors are  $A, B, C, G, I$ , so the clique here is:  $\{A, B, C, G, I\}$ . size: **5**  
 We then eliminate  $B$ , its neighbors are  $A, C, G, I$ , so the clique here is:  $\{A, C, G, I\}$ . size: **4**  
 We then eliminate  $A$ , its neighbors are  $C, G, I$ , so the clique here is:  $\{C, G, I\}$ . size: **3**  
 We then eliminate  $G$ , its neighbors are  $C, I$ , so the clique here is:  $\{C, I\}$ . size: **2**  
 We then eliminate  $I$ , its neighbors are  $C$ , so the clique here is:  $\{C\}$ . size: **1**  
 The largest clique(s) for the elimination order  $E, D, H, F, B, A, G, I, C$  is **5**
- (5 points) Write down largest clique(s) for the elimination order  $A, G, I, C, D, H, F, B, E$ .  
 We start by eliminating  $A$ , its neighbors are  $D, B$ , so the clique here is:  $\{D, B\}$ . size: **2**  
 We then eliminate  $G$ , its neighbors are  $D, H$ , so the clique here is:  $\{D, H\}$ . size: **2**  
 We then eliminate  $I$ , its neighbors are  $H, F$ , so the clique here is:  $\{H, F\}$ . size: **2**  
 We then eliminate  $C$ , its neighbors are  $F, B$ , so the clique here is:  $\{F, B\}$ . size: **2**  
 We then eliminate  $D$ , its neighbors are  $H, B, E$ , so the clique here is:  $\{H, B, E\}$ . size: **3**  
 We then eliminate  $H$ , its neighbors are  $F, B, E$ , so the clique here is:  $\{F, B, E\}$ . size: **3**  
 We then eliminate  $F$ , its neighbors are  $B, E$ , so the clique here is:  $\{B, E\}$ . size: **2**  
 We then eliminate  $B$ , its neighbors are  $E$ , so the clique here is:  $\{E\}$ . size: **1**  
 The largest clique(s) for the elimination order  $A, G, I, C, D, H, F, B, E$  is **3**
- (5 points) Which of the above ordering is preferable? Explain briefly.  
 The second ordering is preferable because it results in a smaller largest clique size (3) compared to the first ordering (5). A smaller clique size generally leads to more efficient computations in variable elimination algorithms.
- (10 points) Using this intuition, give a reasonable ( $\ll n^2$ ) upper bound on the tree-width of the  $n \times n$  grid.  
 The tree-width of an  $n \times n$  grid is at most  $2n - 2$ . This is because the grid can be decomposed into a tree structure where each node has at most  $2n - 2$  neighbors. This upper bound is reasonable and much smaller than  $n^2$ , which would be the case for a fully connected graph.