

$$\text{likelihood } (\hat{\beta}_0, \hat{\beta}_1, \hat{\sigma}^2)$$

$$= \left(\frac{1}{2\pi}\right)^{n/2} (\hat{\sigma}^2)^{-n/2} \exp \left\{ - \frac{\sum_{i=1}^n (y_i - (\hat{\beta}_0 + \hat{\beta}_1 x_{i1}))^2}{2 \hat{\sigma}^2} \right\} \quad \hat{\sigma}^2 = \frac{\sum_{i=1}^n (y_i - \hat{y}_i)^2}{n}$$

$$= \left(\frac{1}{2\pi}\right)^{n/2} (\hat{\sigma}^2)^{-n/2} \exp \left\{ - \frac{n}{2} \right\}$$

$$\text{likelihood} \uparrow \quad \hat{\sigma}^2 \downarrow$$

\Rightarrow We look for the model with smallest Res S.S.

Section 3 Properties of $\hat{\beta}$ and Res S.S.

① Properties of $\hat{\beta}$

p=1 p.11 of chapter 1

Theorem 3.1 $y_i \sim N(\mu_i, \sigma^2)$ $\mu_i = \beta_0 + \beta_1 x_{i1} + \dots + \beta_p x_{ip}$

$$\text{cov}(y_i, y_j) = 0$$

$$\text{cov}\left(\sum_{i=1}^n c_i y_i, \sum_{j=1}^n d_j y_j\right) = \sum_{i=1}^n c_i d_i \text{Var}(y_i)$$

$$\begin{aligned} \hat{\beta}_0 &= \bar{y} - \hat{\beta}_1 \bar{x}_1 \\ &= \frac{\sum_{i=1}^n y_i}{n} - \frac{\frac{\sum_{i=1}^n (x_{i1} - \bar{x}_1)(y_i - \bar{y})}{\sum_{i=1}^n (x_{i1} - \bar{x}_1)^2} \bar{x}_1}{\frac{\sum_{i=1}^n (x_{i1} - \bar{x}_1)^2}{\sum_{i=1}^n (x_{i1} - \bar{x}_1)^2}} \bar{x}_1 \\ &= \frac{\sum_{i=1}^n y_i}{n} - \frac{\sum_{i=1}^n (x_{i1} - \bar{x}_1) y_i}{\sum_{i=1}^n (x_{i1} - \bar{x}_1)^2} \bar{x}_1 \\ &= \sum_{i=1}^n \left(\frac{1}{n} - \frac{(x_{i1} - \bar{x}_1) \bar{x}_1}{\sum_{i=1}^n (x_{i1} - \bar{x}_1)^2} \right) y_i \end{aligned}$$

$$\sim N(\beta_0, \quad)$$

$$E(\hat{\beta}_0) = \sum_{i=1}^n \left(\frac{1}{n} - \frac{(X_{i1} - \bar{X}_1) \bar{X}_1}{S_{X_1 X_1}} \right) E(y_i)$$

$$\text{Model} = y_i = \beta_0 + \beta_1 X_{i1} + \epsilon_i$$

$$E(y_i) = \beta_0 + \beta_1 X_{i1}$$

$$\text{Var}(y_i) = \sigma^2$$

$$= \sum_{i=1}^n \left(\frac{1}{n} - \frac{(X_{i1} - \bar{X}_1) \bar{X}_1}{S_{X_1 X_1}} \right) (\beta_0 + \beta_1 X_{i1})$$

$$= \beta_0 \sum_{i=1}^n \left(\frac{1}{n} - \frac{(X_{i1} - \bar{X}_1) \bar{X}_1}{S_{X_1 X_1}} \right) + \beta_1 \sum_{i=1}^n \left(\frac{1}{n} - \frac{(X_{i1} - \bar{X}_1) \bar{X}_1}{S_{X_1 X_1}} \right) X_{i1}$$

$$\neq \underbrace{\sum_{i=1}^n \left(\frac{1}{n} - \frac{(X_{i1} - \bar{X}_1) \bar{X}_1}{S_{X_1 X_1}} \right)}_{\substack{= \\ 1 - \underbrace{\frac{\sum_{i=1}^n (X_{i1} - \bar{X}_1) \bar{X}_1}{S_{X_1 X_1}} \\ = \\ 0}}} + \underbrace{\beta_1 \sum_{i=1}^n \left(\frac{1}{n} - \frac{(X_{i1} - \bar{X}_1) \bar{X}_1}{S_{X_1 X_1}} \right) X_{i1}}_{\substack{= \\ \frac{\sum_{i=1}^n X_{i1}}{n} - \underbrace{\frac{\sum_{i=1}^n (X_{i1} - \bar{X}_1) \bar{X}_1 X_{i1}}{S_{X_1 X_1}} \\ = \\ \bar{X}_1 \frac{\sum_{i=1}^n (X_{i1} - \bar{X}_1) (X_{i1} - \bar{X}_1)}{S_{X_1 X_1}}}}}$$

$$= \beta_0 \quad \text{--- unbiased estimator}$$

$$\neq \text{Var}(\hat{\beta}_0) = \text{Var} \left(\sum_{i=1}^n C_i y_i \right)$$

$$= \sum_{i=1}^n C_i^2 \text{Var}(y_i) = \sigma^2$$

$$= \sigma^2 \sum_{i=1}^n \left(\frac{1}{n} - \frac{(X_{i1} - \bar{X}_1) \bar{X}_1}{S_{X_1 X_1}} \right)^2$$

$$= \sigma^2 \left[\sum_{i=1}^n \left(\frac{1}{n^2} - \frac{2(X_{i1} - \bar{X}_1) \bar{X}_1}{S_{X_1 X_1}} + \frac{(X_{i1} - \bar{X}_1)^2 \bar{X}_1^2}{S_{X_1 X_1}^2} \right) \right]$$

$$\sum_{i=1}^n \frac{1}{n^2} \downarrow \frac{1}{n}$$

$$\sum_{i=1}^n \frac{2(X_{i1} - \bar{X}_1) \bar{X}_1}{S_{X_1 X_1}} \downarrow 0$$

$$\sum_{i=1}^n \frac{(X_{i1} - \bar{X}_1)^2 \bar{X}_1^2}{S_{X_1 X_1}^2} \downarrow \frac{\bar{X}_1^2}{S_{X_1 X_1}}$$

$$= \sigma^2 \left(\frac{1}{n} + \frac{\bar{X}_1^2}{S_{X_1 X_1}} \right)$$

$$\Rightarrow \hat{\beta}_0 \sim N \left(\beta_0, \sigma^2 \left(\frac{1}{n} + \frac{\bar{X}_1^2}{S_{X_1 X_1}} \right) \right)$$

$$\begin{aligned}\hat{\beta}_1 &= \frac{S_{x_1 y}}{S_{x_1 x_1}} \\&= \frac{\sum_{i=1}^n (x_{i1} - \bar{x}_1) y_i}{S_{x_1 x_1}} \\&= \sum_{i=1}^n \underbrace{\frac{x_{i1} - \bar{x}_1}{S_{x_1 x_1}}}_{d_i} \boxed{y_i} \sim N\end{aligned}$$

$$\hat{\beta}_1 \sim N(\quad , \quad)$$

$$\begin{aligned}E(\hat{\beta}_1) &= \sum_{i=1}^n \frac{x_{i1} - \bar{x}_1}{S_{x_1 x_1}} (\beta_0 + \beta_1 x_{i1}) \\&= \beta_0 \underbrace{\left[\sum_{i=1}^n \frac{(x_{i1} - \bar{x}_1)}{S_{x_1 x_1}} \right]}_{\parallel 0} + \beta_1 \underbrace{\left[\sum_{i=1}^n \frac{(x_{i1} - \bar{x}_1)(x_{i1} - \bar{x}_1)}{S_{x_1 x_1}} \right]}_{\parallel 1}\end{aligned}$$

$= \beta_1$ — unbiased estimator

$$\text{Var}(\hat{\beta}_1) = \sum_{i=1}^n \left(\frac{x_{i1} - \bar{x}_1}{S_{x_1 x_1}} \right)^2 \boxed{\text{Var}(y_i)} \leftarrow \sigma^2$$

$$= \sigma^2 \underbrace{\left(\sum_{i=1}^n \frac{(x_{i1} - \bar{x}_1)^2}{S_{x_1 x_1}^2} \right)}_{S_{x_1 x_1}}$$

$$= \frac{\sigma^2}{S_{x_1 x_1}}$$

$$\Rightarrow \hat{\beta}_1 \sim N\left(\beta_1, \frac{\sigma^2}{S_{x_1 x_1}}\right)$$

$$\text{cov}(\hat{\beta}_0, \hat{\beta}_1) = \text{cov}\left(\sum_{i=1}^n c_i y_i, \sum_{j=1}^n d_j y_j\right)$$

$$\uparrow \quad \quad \quad \uparrow$$

$$\left(\frac{1}{n} - \frac{(x_{i1} - \bar{x}_1)\bar{x}_1}{S_{x_1 x_1}}\right) \quad \frac{x_{i1} - \bar{x}_1}{S_{x_1 x_1}}$$

$$= \sum_{i=1}^n c_i d_i \boxed{\text{Var}(y_i)} = \sigma^2$$

$$= \sigma^2 \sum_{i=1}^n \left(\frac{1}{n} - \frac{(x_{i1} - \bar{x}_1) \bar{x}_1}{S_{x_1 x_1}} \right) \left(\frac{x_{i1} - \bar{x}_1}{S_{x_1 x_1}} \right)$$

$$= \sigma^2 \left(\underbrace{\sum_{i=1}^n \frac{x_{i1} - \bar{x}_1}{n S_{x_1 x_1}}}_{=0} - \underbrace{\sum_{i=1}^n \frac{(x_{i1} - \bar{x}_1)^2 \bar{x}_1}{S_{x_1 x_1}^2}}_{= \frac{\bar{x}_1}{S_{x_1 x_1}}} \right)$$

$$= - \frac{\bar{x}_1}{S_{x_1 x_1}} \sigma^2$$

$$\Rightarrow \begin{pmatrix} \hat{\beta}_0 \\ \hat{\beta}_1 \end{pmatrix} \sim N_2 \left(\begin{pmatrix} \beta_0 \\ \beta_1 \end{pmatrix}, \begin{pmatrix} \sigma^2 \left(\frac{1}{n} + \frac{\bar{x}_1^2}{S_{x_1 x_1}} \right) & -\frac{\bar{x}_1}{S_{x_1 x_1}} \sigma^2 \\ -\frac{\bar{x}_1}{S_{x_1 x_1}} \sigma^2 & \frac{\sigma^2}{S_{x_1 x_1}} \end{pmatrix} \right)$$

$\begin{matrix} \nearrow \text{Var}(\hat{\beta}_0) & \nearrow \text{Cov}(\hat{\beta}_0, \hat{\beta}_1) \\ \nwarrow \text{Cov}(\hat{\beta}_1, \hat{\beta}_0) & \nwarrow \text{Var}(\hat{\beta}_1) \end{matrix}$

↑
variance & covariance matrix

p > 1 $\hat{\beta} = (X^T X)^{-1} X^T \underline{y} \sim MN(\underline{X}\beta, \sigma^2 \underline{I})$

Theorem 3.2 (p.14 of Chapter 1)

$$E(\underline{c} \underline{z} + \underline{d}) = \underline{c} E(\underline{z}) + \underline{d}$$

$$\text{Var}(\underline{c} \underline{z} + \underline{d}) = \underline{c} \text{Var}(\underline{z}) \underline{c}^T$$

$$\text{Cov}(\underline{c} \underline{z}, \underline{d} \underline{z}) = \underline{c} \text{Var}(\underline{z}) \underline{d}^T$$

$$E(cz + d) = c E(z) + d$$

$$\text{Var}(cz + d) = c^2 \text{Var}(z)$$

$$\text{Cov}(cz, dz) = cd \text{Var}(z)$$

$$\hat{\beta} = \underbrace{(X^T X)^{-1}}_{p' \times p' \substack{\underline{c} \\ p' \times 1}} \underbrace{X^T}_{p' \times 1} \underbrace{y}_{n \times 1}$$

$$E(\hat{\beta}) = (X^T X)^{-1} X^T \boxed{E(Y)} = X \beta$$

$$= (X^T X)^{-1} X^T X \beta = \beta$$

$$\text{Var}(\hat{\beta}) = \underbrace{(X^T X)^{-1}}_{C} X^T \underbrace{\text{Var}(Y)}_{\sigma^2 I} \underbrace{((X^T X)^{-1} X^T)^T}_{C^T}$$

$$= \sigma^2 \underbrace{(X^T X)^{-1} X^T}_{X(X^T X)^{-1}} \underbrace{((X^T X)^{-1} X^T)^T}_{X(X^T X)^{-1}}$$

$$= (X^T X)^{-1} \sigma^2$$

$$\Rightarrow \hat{\beta} \sim MN(\beta, \sigma^2 (X^T X)^{-1})$$

e.g. $\beta = 1$

$$X_{n \times 2} = \begin{pmatrix} 1 & x_{11} \\ \vdots & \vdots \\ 1 & x_{n1} \end{pmatrix} \quad X^T X = \begin{pmatrix} n & \sum_{i=1}^n x_{i1} \\ \sum_{i=1}^n x_{i1} & \sum_{i=1}^n x_{i1}^2 \end{pmatrix}$$

$$(X^T X)^{-1} = \frac{1}{\underbrace{n \sum_{i=1}^n x_{i1}^2 - (\sum_{i=1}^n x_{i1})^2}_{n (\sum_{i=1}^n x_{i1}^2 - \frac{1}{n} (n \bar{x}_1)^2)}} \begin{pmatrix} \sum_{i=1}^n x_{i1}^2 & -\sum_{i=1}^n x_{i1} \\ -\sum_{i=1}^n x_{i1} & n \end{pmatrix}$$

$$\Rightarrow \text{Var}(\hat{\beta}_0) = \frac{\sum_{i=1}^n x_{i1}^2}{n S_{x_1 x_1}} \sigma^2$$

$$\text{Var}(\hat{\beta}_1) = \frac{n}{n S_{x_1 x_1}} \sigma^2$$

$$= \frac{\sigma^2}{S_{x_1 x_1}}$$

$$\begin{aligned} \sum_{i=1}^n x_{i1}^2 - n \bar{x}_1^2 &= S_{x_1 x_1} \\ \Rightarrow \sum_{i=1}^n x_{i1}^2 &= S_{x_1 x_1} + n \bar{x}_1^2 \\ n (\sum_{i=1}^n x_{i1}^2 - \frac{1}{n} (n \bar{x}_1)^2) &= n S_{x_1 x_1} \end{aligned}$$

$$\text{Cov}(\hat{\beta}_0, \hat{\beta}_1) = - \frac{\sum_{i=1}^n x_{i1}}{n S_{x_1 x_1}} \sigma^2$$

e.g. $p=1$ β_0 is known

$$\underset{n \times 1}{\tilde{X}} = \begin{pmatrix} x_{11} \\ \vdots \\ x_{n1} \end{pmatrix}$$

$$\tilde{X}^T \tilde{X} = \sum_{i=1}^n x_{i1}^2$$

$$(\tilde{X}^T \tilde{X})^{-1} = \frac{1}{\sum_{i=1}^n x_{i1}^2}$$

$$\text{Var}(\hat{\beta}_1) = \sigma^2 / \sum_{i=1}^n x_{i1}^2$$

② Properties of \hat{e}_i

$p=1$ $\hat{e}_i = y_i - \hat{y}_i$

$$= y_i - (\hat{\beta}_0 + \hat{\beta}_1 x_{i1})$$

$$= y_i - \left(\sum_{j=1}^n c_j y_j + \sum_{j=1}^n d_j x_{j1} y_j \right)$$

$$= y_i - \sum_{j=1}^n (c_j + d_j x_{j1}) y_j$$

$$= \sum_{j=1}^n (\delta_{ij} - (c_j + d_j x_{j1})) \boxed{y_j}$$

$$\boxed{\delta_{ij}} = \begin{cases} 1 & \text{if } i=j \\ 0 & \text{if } i \neq j \end{cases}$$

↑
Kronecker delta

$$\hat{e}_i \sim N \left(\underset{\uparrow}{E(\hat{e}_i)}, \underset{\uparrow}{\text{Var}(\hat{e}_i)} \right) \sim N$$

$p > 1$ $\hat{\underline{e}} = \underline{X} - \hat{\underline{Y}}$

$$= \underline{Y} - \underline{X} \hat{\underline{\beta}}$$

$$= \underline{Y} - \underline{X} (\underline{X}^T \underline{X})^{-1} \underline{X}^T \underline{Y}$$

$$= (\underline{I} - \underline{X} (\underline{X}^T \underline{X})^{-1} \underline{X}^T) \underline{Y}$$

$$= (\underline{I} - \underline{H}) \boxed{\underline{Y}} \sim MN(\underline{X} \underline{\beta}, \sigma^2 \underline{I})$$

$$\Rightarrow \hat{\underline{e}} \sim MN(,)$$

③ Properties of Res S.S.

$$\underline{p=1} \quad \text{Res S.S.} = \sum_{i=1}^n (y_i - \hat{y}_i)^2$$

$$(\beta_0 \text{ is unknown}) \quad = S_{yy} - \hat{\beta}_1 \boxed{S_{x_1 y}} \quad \leftarrow \frac{S_{x_1 y}}{S_{x_1 x_1}} = \hat{\beta}_1$$

$$= S_{yy} - \hat{\beta}_1^2 S_{x_1 x_1}$$

$$E(\text{Res S.S.}) = \boxed{E(S_{yy})} - S_{x_1 x_1} \underbrace{\left(E(\hat{\beta}_1^2) \right)}_{\text{Var}(\hat{\beta}_1) + (E(\hat{\beta}_1))^2}$$

$$\quad \quad \quad \uparrow \quad \quad \quad \uparrow$$

$$\quad \quad \quad (n-1) \sigma^2 \quad \quad \quad \frac{\sigma^2}{S_{x_1 x_1}} + \beta_1^2$$

$$E(y_i) = \text{diff means}$$

$$\quad \quad \quad \uparrow$$

$$\quad \quad \quad \beta_0 + \beta_1 x_{i1}$$

$$E(S_{yy}) = E\left(\sum_{i=1}^n (y_i - \bar{y})^2 \right)$$

$$= E\left(\sum_{i=1}^n y_i^2 - n \bar{y}^2 \right)$$

$$= \boxed{E\left(\sum_{i=1}^n y_i^2 \right)} - n \boxed{E(\bar{y}^2)}$$

$$\quad \quad \quad \uparrow \quad \quad \quad \uparrow$$

$$\quad \quad \quad \sum_{i=1}^n E(y_i^2) \quad \quad \quad n E(\bar{y}^2)$$

$$\quad \quad \quad \uparrow$$

$$\quad \quad \quad \text{Var}(y_i) + (E(y_i))^2$$

$$= \sum_{i=1}^n (\sigma^2 + (\beta_0 + \beta_1 x_{i1})^2)$$

$$\text{Var}(y_i) = \sigma^2$$

$$E(y_i) = \beta_0 + \beta_1 x_{i1}$$

$$\quad \quad \quad \uparrow \quad \quad \quad \uparrow$$

$$\quad \quad \quad \text{Var}(\bar{y}) + (E(\bar{y}))^2$$

$$\quad \quad \quad \uparrow \quad \quad \quad \uparrow$$

$$\quad \quad \quad \frac{\sigma^2}{n} \quad \quad \quad (\beta_0 + \beta_1 \bar{x}_1)^2$$

$$E(S_{yy}) = n \sigma^2 + \sum_{i=1}^n (\beta_0 + \beta_1 x_{i1})^2 - n \left(\frac{\sigma^2}{n} + (\beta_0 + \beta_1 \bar{x}_1)^2 \right)$$

$$= (n-1) \sigma^2$$

$$+ \beta_1^2 S_{x_1 x_1}$$

$$\quad \quad \quad \uparrow$$

$$\quad \quad \quad \sum_{i=1}^n (\beta_0 + \beta_1 x_{i1})^2 - n (\beta_0 + \beta_1 \bar{x}_1)^2$$

$$= \sum_{i=1}^n (\beta_0^2 + 2\beta_0 \beta_1 x_{i1} + \beta_1^2 x_{i1}^2) -$$

$$n (\beta_0^2 + 2\beta_0 \beta_1 \bar{x}_1 + \beta_1^2 \bar{x}_1^2)$$

$$E(\text{Res S.S.}) = (n-1) \sigma^2 + \beta_1^2 S_{x_1 x_1}$$

$$= \cancel{\sigma^2} S_{x_1 x_1} \left(\frac{\sigma^2}{S_{x_1 x_1}} + \beta_1^2 \right) = \beta_1^2 \left[\sum_{i=1}^n x_{i1}^2 - n \bar{x}_1^2 \right] = \beta_1^2 S_{x_1 x_1}$$

⑦

$$= (n-2) \sigma^2$$

$$\Rightarrow \text{unbiased est. of } \sigma^2 = \frac{\text{Res S.S.}}{n-2}$$

$$\underline{p > 1} \quad \hat{\sigma}^2 = \frac{\text{Res S.S.}}{\boxed{n} - \boxed{p'}} \quad \begin{array}{l} \text{"} \\ \text{\# of obs.} \end{array} \quad \begin{array}{l} \text{"} \\ \text{\# of unknown para. in the model} \end{array}$$

$$\textcircled{1} \text{ Prove } E(\text{Res S.S.}) = (n-p') \sigma^2$$

$$\textcircled{2} \text{ dist. of Res S.S. ?}$$

Theorem 3.3 in (p.16 of ~~Theory~~ Chapt. 1)

$$E(\underline{Y}) = \underline{\mu}, \quad \text{Cov}(\underline{Y}) = \underline{\Sigma}$$

$$\text{Then } E(\underline{Y}^T \underline{A} \underline{Y}) = \text{trace}(\underline{A} \underline{\Sigma}) + \underline{\mu}^T \underline{A} \underline{\mu} \quad \text{quadratic form of } \underline{Y}$$

$$\underline{\text{Model}} : \underline{Y} = \underline{X} \underline{\beta} + \underline{e} \quad E(\underline{Y}) = \underline{X} \underline{\beta}$$

$$\text{Res S.S.} = \underline{Y}^T \underline{Y} - \hat{\underline{\beta}}^T \underline{X}^T \underline{Y} \quad \text{Var}(\underline{Y}) = \sigma^2 \underline{I}$$

$$= \underline{Y}^T \underline{Y} - \underbrace{\underline{Y}^T \underline{X} (\underline{X}^T \underline{X})^{-1} \underline{X}^T \underline{Y}}_{\hat{\underline{\beta}}^T} = (\underline{X}^T \underline{X})^{-1} \underline{X}^T \underline{Y}$$

$$= \underline{Y}^T (\underline{I} - \underbrace{\underline{X} (\underline{X}^T \underline{X})^{-1} \underline{X}^T}_{\underline{H}}) \underline{Y}$$

$$= \underline{Y}^T (\underline{I} - \underline{H}) \underline{Y}$$

$$E(\text{Res S.S.}) = E(\underbrace{\underline{Y}^T}_{1 \times n} \underbrace{(\underline{I} - \underline{H})}_{n \times n} \underbrace{\underline{Y}}_{n \times 1})$$

$$\text{trace}(\underbrace{\underline{A}}_{n \times n} \underline{I}) = \text{trace}((\underline{I} - \underline{H}) \sigma^2 \underline{I})$$

$$= \sigma^2 \text{trace}(\underline{I} - \underline{H})$$

$$= \sigma^2 \left\{ \text{trace}(\underline{I}) - \text{trace}(\underline{X} (\underline{X}^T \underline{X})^{-1} \underline{X}^T) \right\}$$

$$= \sigma^2 \left\{ n - \text{trace}((\underline{X}^T \underline{X})^{-1} \underbrace{\underline{X}^T \underline{X}}_{\substack{n \times n \\ n \times p'}}) \right\} \quad \begin{array}{l} \# \text{ of indep.} \\ \text{variables} \\ \downarrow \\ p' = 1 + p \end{array}$$

$$= \sigma^2 \{ n - p' \}$$

$$\underline{\mu}^T \underline{A} \underline{\mu} = (\underline{X} \underline{\beta})^T (\underline{I} - \underline{X} (\underline{X}^T \underline{X})^{-1} \underline{X}^T) \underline{\mu} \quad \begin{array}{l} \underline{\mu} = E(\underline{Y}) \\ = \underline{X} \underline{\beta} \end{array}$$

$$= \underline{\beta}^T \underline{X}^T (\underline{I} - \underline{X} (\underline{X}^T \underline{X})^{-1} \underline{X}^T) \underline{X} \underline{\beta}$$

$$= \underline{\beta}^T (\underline{X}^T \underline{X} - \underbrace{\underline{X}^T \underline{X} (\underline{X}^T \underline{X})^{-1} \underline{X}^T \underline{X}}_{\underline{I}}) \underline{\beta}$$

$$= 0$$

$$\Rightarrow E(\text{Res S.S.}) = \sigma^2 (n - p') + 0$$

$$\Rightarrow \hat{\sigma}^2 = \frac{\text{Res S.S.}}{n - p'} \quad \text{is unbiased for } \sigma^2$$