17 September 2020

Properties of estimators

1. **Â**

For p=1

By the Theorem 3.1 in page 11,

$$\hat{\beta}_{0} = \sum_{i=1}^{n} \left(\frac{1}{n} - \frac{(x_{i1} - \bar{x}_{1})\bar{x}_{1}}{S_{x_{1}x_{1}}} \right) y_{i}$$

$$\hat{\beta}_{1} = \sum_{i=1}^{n} \left(\frac{x_{i1} - \bar{x}_{1}}{S_{x_{1}x_{1}}} \right) y_{i}$$

Then

$$Var(\hat{\beta}_0) = \left(\frac{1}{n} + \frac{\bar{x}_1^2}{S_{x_1 x_1}}\right) \sigma^2$$

$$= \frac{\sigma^2 \sum_{i=1}^n x_{i1}^2}{n S_{x_1 x_1}}$$

$$Var(\hat{\beta}_1) = \frac{\sigma^2}{S_{x_1 x_1}}$$

$$Cov(\hat{\beta}_0, \hat{\beta}_1) = -\frac{\sigma^2 \sum_{i=1}^n x_{i1}}{n S_{x_1 x_1}}$$

$$\begin{pmatrix} \hat{\beta}_0 \\ \hat{\beta}_1 \end{pmatrix} \sim \mathcal{N} \begin{pmatrix} \beta_0 \\ \beta_1 \end{pmatrix}, \quad \begin{pmatrix} \sigma^2(\frac{1}{n} + \frac{\bar{x}_1^2}{S_{x_1 x_1}}) & -\frac{\bar{x}_1 \sigma^2}{S_{x_1 x_1}} \\ -\frac{\bar{x}_1 \sigma^2}{S_{x_1 x_1}} & \frac{\sigma^2}{S_{x_1 x_1}} \end{pmatrix} \end{pmatrix}$$

For any p

By the Theorem 3.2 in page 14,

$$\hat{\beta} \sim N(\beta, \sigma^2(X^TX)^{-1})$$

 $2. \hat{e}_i$

For p = 1

$$\hat{e}_{i} = \sum_{j=1}^{n} (\delta_{ij} - (c_{j} + d_{j}x_{i1})) y_{j}$$
 where $c_{j} = \frac{1}{n} - \frac{(x_{j1} - \bar{x}_{1})\bar{x}_{1}}{S_{x_{1}x_{1}}}, d_{j} = \frac{x_{j1} - \bar{x}_{1}}{S_{x_{1}x_{1}}} \text{ and } \delta_{ij} = \begin{cases} 1 & \text{if } i = j \\ 0 & \text{if } i \neq j \end{cases}$

For any p

$$\hat{e} = (\underline{I} - \underline{H}) \underline{Y}$$

3. Residual Sum of Squares

For p=1

$$E(S_{yy}) = (n-1)\sigma^2 + \beta_1^2 S x_1 x_1 \text{ and}$$

$$E(\hat{\beta}_1^2) = Var(\hat{\beta}_1) + (E(\hat{\beta}_1))^2$$

$$E(RSS) = (n-2)\sigma^2$$

$$\Rightarrow \hat{\sigma}^2 = \frac{RSS}{n-2}$$

For any p

By the Theorem 3.3 in page 16,

$$E(RSS) = (n - p')\sigma^2.$$

Thus,

$$\hat{\sigma}^2$$
 (unbiased estimator) = $\frac{RSS}{n-p'}$