Example 5: Example in Multiple Linear Regression

Xo=(1,3,8,9)

X1=3, X2=8, X3=9

The percent survival of a certain type of animal semen after storage was measured at various combinations of concentrations of three materials used to increase chance of survival. The data are as follows:

(04 : 1)	( 1, (4)	( . 1 . 07)	/ . 1 . 64	- Mylxo = Bo + 3 Bi + 8 Bz+
y (%  survival)	$x_1$ (weight %)	$x_2$ (weight %)	$x_3$ (weight %)	/ 0
25.5	1.74	5.30	10.80	9 B.
31.2	6.32	5.42	9.40	9 73
25.9	6.22	8.41	7.20	=(1389)(80)
38.4	10.52	4.63	8.50	
18.4	1.19	11.60	9.40	\ B <sub>1</sub>
26.7	1.22	5.85	9.90	\ \hat{2}
26.4	4.10	6.62	8.00	/ Pr
25.9	6.32	8.72	9.10	(B)
32.0	4.08	4.42	8.70	= 24.2232
25.2	4.15	7.60	9.20	1.17 (2
39.7	10.15	4.83	9.40	bit 2-4290
35.7	1.72	3.12	7.60	til 6 - 1. 10
26.5	1.70	5.30	8.20	_

## Summary statistics:

$$\sum_{i=1}^{13} y_i = 377.5 \qquad \sum_{i=1}^{13} y_i^2 = 11,400.15 \qquad \sum_{i=1}^{13} x_{i1} = 59.43$$

$$\sum_{i=1}^{13} x_{i2} = 81.82 \qquad \sum_{i=1}^{13} x_{i3} = 115.40 \qquad \sum_{i=1}^{13} x_{i1}^2 = 394.7255$$

$$\sum_{i=1}^{13} x_{i2}^2 = 576.7264 \qquad \sum_{i=1}^{13} x_{i3}^2 = 1035.96 \qquad \sum_{i=1}^{13} x_{i1}y_i = 1877.567$$

$$\sum_{i=1}^{13} x_{i2}y_i = 2246.661 \qquad \sum_{i=1}^{13} x_{i3}y_i = 3337.78 \qquad \sum_{i=1}^{13} x_{i1}x_{i2} = 360.6621$$

$$\sum_{i=1}^{13} x_{i1}x_{i3} = 522.078 \qquad \sum_{i=1}^{13} x_{i2}x_{i3} = 728.31 \qquad n = 13$$

 $\chi_{0}^{T} = (1, 3 - X_{1}, 8 - X_{2}, 9 - X_{3})$   $1 + \chi_{0}^{T} (\chi_{0}^{T} \chi_{0})^{-1} \chi_{0} = (1, 3 - X_{1}, 8 - X_{2}, 9 - X_{3})$ 

 $\begin{pmatrix} 1 \\ 3 - \overline{X}_1 \\ 8 - \overline{X}_2 \\ q - \overline{X}_1 \end{pmatrix} = 0.1267$ 

Coefficient of determination Total S.S. = Reg S.S. + Ros S.S.  $0 \le R^2 = \frac{\text{Reg S.S.}}{\text{total S.S.}} \le 1$ = Total S.S. - Ress.S. = 1 - Res S.S. total SS. - linear st relations hip  $R^{2} = \frac{\text{Rog S.S.}}{\text{totAls S.}} = \frac{\hat{\beta_{1}} S_{X_{1}} y}{\sum_{i=1}^{n} (y_{i} - \overline{y})^{2}} \hat{\beta_{1}} = \frac{S_{X_{1}} y}{S_{X_{1}} X_{1}}$ Ress.s. = XT I - RTXT X  $= \frac{Sxiy}{Sxixi} Sxiy} = \frac{Sxiy}{Sxiy}$   $= \frac{Sxiy}{Syy} = \frac{Sxiy}{Sxiy}$ = Syy - (\$1 Sxiy + βφSxpy).  $= \left(\frac{S_{xy}}{\sqrt{S_{xx}}\sqrt{S_{yy}}}\right)^2$ Reg S.S. r - simple correlation coeff.  $-1 \le r \le 1$ \(\bar{R}^2\) \(\phi > 1\) metiple correlation coeff. Exaple in p. 7 (while linear regression)  $R^2 = 1 - \frac{38.68}{43817} = 0.9117$ 

91.17% of varation in y can be explained by the linear regression model

population simple correlation coefficiet Sinfle linear regression Br = regression weff. of XI  $P = \frac{\text{Cov}(x_i, Y)}{\sqrt{\text{Va}(x_i)}\sqrt{\text{Va}(Y)}} \times X, Y \cdot Y. V.$ or slope of linear regression f(y,x,) = f(y(x,))[f(x,)] $\hat{\gamma} = \gamma = \frac{S_{x_i} y}{\sqrt{S_{x_i} x_i} \sqrt{S_{yy}}}$  $-\hat{\beta}_{1} = \frac{S_{x_{1}}y}{S_{x_{1}}x_{1}}$ Assume  $\begin{bmatrix} x_1 \\ y \end{bmatrix} \sim MN \begin{pmatrix} \mu x_1 \\ \mu y \end{pmatrix}$ ,  $\begin{bmatrix} 6x_1 \\ \beta_1 6x_1 \end{bmatrix} = \begin{bmatrix} 6x_1 \\ 6y_1 \end{bmatrix} = \begin{bmatrix} 6x_1 \\ 6y_1 \end{bmatrix}$ Model = Y = Bo + Bixi + e  $cov(X_1, Y) = cov(X_1, po + p_1 X_1 + e)$ = BI Var(XI) - assume XIde are indep. = B, 6x1 => \$16x1 = 16x1 64 I) \$1=0 => P=0  $= V P = \frac{\beta_1 6 x_1}{64}$ weasures the linear Ho = B1 = 0 relation ship between X, I Y From livear regression model Var ( B) = 62 Sx1x1  $=\frac{\hat{\beta}_1}{6/\sqrt{S_{x_1x_1}}}$  $\hat{G}^2 = \frac{\text{Res S. S.}}{N-2}$ 

= 1 Syy - (B) Sxiy]

 $=\frac{1}{h-2}\left(S_{yy}-\frac{S_{xiy}}{C_{xy}}\right)$ 

(3)

$$t = \frac{\frac{S_{x_1}y}{S_{x_1x_1}}}{\frac{S_{y_1}y}{S_{x_1x_1}}} = \frac{\frac{S_{x_1}y}{S_{x_1x_1}}}{\frac{S_{x_1}y}{S_{x_1x_1}}} = \frac{S_{y_1}y}{\frac{S_{x_1}y}{S_{x_1x_1}}} = \frac{S_{y_1}y}{\frac{S_{x_1}y}{S_{x_1}y}} = \frac{S_{y_1}y}{\frac{S_{x_1}y}{S_{x_1}y}} = \frac{S_{x_1}y}{\frac{S_{x_1}y}{S_{x_1}y}} = \frac{S_{x_1}y}{\frac{S_{x_1}y}{S_{x_1}y}}} = \frac{S_{x_1}y}{\frac{S_{x_1}y}} = \frac{S_{x_1}y}{\frac{S_{x_1}y}{S_{x_1}y}} = \frac{S_{x_1}y}{\frac{S$$

Serting T Added variable plot (partial regression plot) (p.39)

Model =  $y_i = \beta_0 + |\beta_1 \times c_1 + \beta_2 \times c_2 + \beta_3 \times c_4 + \beta_4 \times c_4 + \beta_5 \times c_4 + \beta_6 \times c_4 + \beta_6 \times c_4 + \beta_6 \times c_4 + \beta_6 \times c_4 \times c_4$ 

(4)

= Y - X &

 $= \chi - \chi (\chi^T \chi)^{-1} \chi^T \chi$ 

 $= \left( \frac{1}{\lambda} - \frac{1}{\lambda} \left( \frac{1}{\lambda} \right)^{-1} \frac{1}{\lambda} \right) = \left( \frac{1}{\lambda} - \frac{1}{\lambda} \left( \frac{1}{\lambda} \right)^{-1} \frac{1}{\lambda} \right) = \left( \frac{1}{\lambda} - \frac{1}{\lambda} \left( \frac{1}{\lambda} \right) \right) = \left( \frac{1}{\lambda} - \frac{1}{\lambda} \left( \frac{1}{\lambda} \right) \right) = \left( \frac{1}{\lambda} - \frac{1}{\lambda} \left( \frac{1}{\lambda} \right) \right) = \left( \frac{1}{\lambda} - \frac{1}{\lambda} \left( \frac{1}{\lambda} \right) \right) = \left( \frac{1}{\lambda} - \frac{1}{\lambda} \left( \frac{1}{\lambda} \right) \right) = \left( \frac{1}{\lambda} - \frac{1}{\lambda} \left( \frac{1}{\lambda} \right) \right) = \left( \frac{1}{\lambda} - \frac{1}{\lambda} \left( \frac{1}{\lambda} \right) \right) = \left( \frac{1}{\lambda} - \frac{1}{\lambda} \left( \frac{1}{\lambda} \right) \right) = \left( \frac{1}{\lambda} - \frac{1}{\lambda} \left( \frac{1}{\lambda} \right) \right) = \left( \frac{1}{\lambda} - \frac{1}{\lambda} \left( \frac{1}{\lambda} \right) \right) = \left( \frac{1}{\lambda} - \frac{1}{\lambda} \left( \frac{1}{\lambda} \right) \right) = \left( \frac{1}{\lambda} - \frac{1}{\lambda} \left( \frac{1}{\lambda} \right) \right) = \left( \frac{1}{\lambda} - \frac{1}{\lambda} \left( \frac{1}{\lambda} \right) \right) = \left( \frac{1}{\lambda} - \frac{1}{\lambda} \left( \frac{1}{\lambda} \right) \right) = \left( \frac{1}{\lambda} - \frac{1}{\lambda} \right)$ 

y'= residual of y on the other indep. varables except X, \$ \( \) = (\) - \( \) (\( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) \( \) (I - Xo (X) Xo)- X1) X0  $=(\cancel{x_0}-\cancel{x_0}(\cancel{x_1}\cancel{x_0})^{-1}\cancel{x_1}\cancel{x_0})=\cancel{x_0}$ X' = residual of XI on the other indep. varables, XI & X2 Model = X1 = To + T1 X2 + 82 X3 + P1 ê, = (I - Xo (Xo Xo) -1 XJ) X, question = B1 the regression coefficient of ês on ê, ? ê = ( I - X ( X X ) ( ) ( ) X + & ) + e)  $=\beta_{1}(\overline{\chi}-\chi_{0}(\chi_{0}^{T}\chi_{0})-\chi_{0}^{T})\chi_{1})+(\overline{\chi}-\chi_{0}(\chi_{0}^{T}\chi_{0})-\chi_{0}^{T})\chi_{0}\beta_{1}+$ = BIE, + (I - X (XIX) - XT) & Juestin 2 : [slope in the plot of êo on êi] = [Bi, me.e]?  $X_{1}' = \hat{e}_{1} = (\underline{I} - \chi_{0} (\chi_{0}^{T} \chi_{0}) - \chi_{0}^{T}) \chi_{1}$   $(X_{1}') = \hat{e}_{1} = (\underline{I} - \chi_{0} (\chi_{0}^{T} \chi_{0}) - \chi_{0}^{T}) \chi_{1}$   $(X_{1}') = \hat{e}_{1} = (\underline{I} - \chi_{0} (\chi_{0}^{T} \chi_{0}) - \chi_{0}^{T}) \chi_{1}$   $(X_{1}') = \hat{e}_{1} = (\underline{I} - \chi_{0} (\chi_{0}^{T} \chi_{0}) - \chi_{0}^{T}) \chi_{1}$ y= eo= ( I - Xo ( Xo Xo) + Xo) X  $\left( \begin{array}{c} \chi \overline{\zeta} \left( \overline{\chi} - \chi_0 \left( \chi_0^T \chi_0 \right)^{-1} \chi_0^T \right)^T \left( \overline{\chi} - \chi_0 \left( \chi_0^T \chi_0 \right)^{-1} \chi_0^T \right) \\ \chi_1 \end{array} \right)$ (xī (I - Xo (xō xo)-1xō) ))

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Simple lined begressin between X, dy simple linear regressin between y' and X! Single Correlation coeff. between Partial  $S_{x',y'}$  y' and  $X_i'$   $Y' = \frac{S_{x',y'}}{\sqrt{S_{x',x'}}} \frac{1}{\sqrt{S_{y',y'}}}$ Simple correlation coefficient r = Sxiy VSxixi JSyy TX, y - TX, X = TX x y Simple latter  $\sqrt{1-x_{2}^{2}}\sqrt{1-x_{2}^{2}}$ Simple latter  $\sqrt{1-x_{2}^{2}}\sqrt{1-x_{2}^{2}}$ Fringle correlation welfaint  $\sqrt{1-x_{2}^{2}}\sqrt{1-x_{2}^{2}}$ Called as partial correlation Coeffaint between X, dy X1 ( )/// X2 multiple correlation coeff. measures the linear veg relationship between y and  $\chi = (\chi_1, --, \chi_p)$ partial correlation weff. measure the linear relationship tretween y and X, after adjusted by the other indep varables Simple wordation weff measure the linear relationship between y & X1