

Tutorial for 2020-10-13 and 2020-10-16

Maximum Likelihood Estimation (MLE)



- How to use MLE → method

$$\left\{ \begin{array}{l} \Theta = \{\theta_1, \theta_2, \dots, \theta_n \\ X_1, X_2, \dots, X_n \stackrel{iid}{\sim} f(x) \end{array} \right.$$

- To find likelihood function $L(\theta)$

$$\underline{L(\theta)} = L(\theta_1, \theta_2, \dots, \theta_n | x) = \begin{cases} \prod_{i=1}^n f(x_i | \theta) & \text{continuous.} \\ \prod_{i=1}^n P(x_i | \theta) & \text{discrete.} \end{cases}$$

- To find log likelihood fun $\ell(\theta)$

$$\ell(\theta) = \log [L(\theta)] = \begin{cases} \sum_{i=1}^n \log f(x_i | \theta) & \text{continuous} \\ \sum_{i=1}^n \log P(x_i | \theta) & \text{discrete.} \end{cases}$$

$a \times b \rightarrow \log(a \times b) \rightarrow \log a + \log b$.

③ Take derivatives.

k parameter. $\Leftarrow k$ times.

$\Rightarrow ? \text{, } 1 \text{ solution.}$

$$\left\{ \begin{array}{l} \frac{\partial l(\theta)}{\partial \theta_1} = 0 \\ \frac{\partial l(\theta)}{\partial \theta_2} = 0 \\ \vdots \\ \frac{\partial l(\theta)}{\partial \theta_k} = 0 \end{array} \right. \quad \begin{array}{l} \text{equations} \\ k \text{ unknown param} \end{array} \quad \Rightarrow \hat{\theta}_{MLE} = \{ \hat{\theta}_1, \hat{\theta}_2, \dots, \hat{\theta}_k \} \quad \begin{array}{l} MLE \text{ candidate} \\ 1 \text{ solution.} \end{array}$$

④ Double check.

$k=1$. 1 parameter $\theta = \{ \hat{\theta}_1 \}$

$$L''(\theta) = \frac{\partial^2 l(\theta)}{\partial \theta^2} < 0.$$

• Invariance property of MLE.

$$\left\{ X \sim f_{\theta}(x) \quad \hat{\theta} \xrightarrow{MLE} \theta \right.$$

$$g(\cdot)$$

$$g(\theta) \xleftarrow{MLE} \underbrace{g(\hat{\theta})}_{\text{---}}$$

• ⑤ MLE may not be unbiased

$$\left\{ E(\hat{\theta}) \neq \theta \rightarrow \text{biased.} \right.$$

$$\left. E(\hat{\theta}) = \theta \rightarrow \text{unbiased} \right.$$

See it later!!!

Song TnT

② MLE may not be unique.

e.g. $X_1, X_2, \dots, X_n \sim U(\theta - \frac{1}{2}, \theta + \frac{1}{2})$

DGR

Q: $\hat{\theta}_{MLE} ?$

$\hat{\theta} \in [x_{(n)} - \frac{1}{2}, x_{(1)} + \frac{1}{2}]$

③ MLE \rightarrow asymptotically normal.

approximate / Limited

$t > 0, t$ is very small.

$$P\left(\frac{n(\hat{\theta} - \theta)}{\theta} \leq t\right) \xrightarrow{d} \text{Normal}$$

e.g. $X_1, X_2, \dots, X_n \sim U(0, \theta)$

Q. MLE for θ

$$\hat{\theta}_{MLE} = X_{(n)}$$

$$P\left(\frac{n(\hat{\theta} - \theta)}{\theta} \leq t\right) \rightarrow 1 - e^{-t}$$

$\rightarrow \hat{\theta}$ is exp(I)

Information Inequality: $\theta \in \mathbb{R}$

$\theta \in \mathbb{R}^k$

$\hat{\theta}_n$ is { Consistent: $\lim_{n \rightarrow \infty} P(\|\hat{\theta}_n - \theta\| > \epsilon) = 0$ }
 asymptotically unbiased $E(\hat{\theta}) = \theta$ + Limitat.
 efficient $(E(\hat{\theta}) - \theta)^T P(E(\hat{\theta}) - \theta) \xrightarrow{P} 0$
 normal distributed.



$$\sqrt{n}(\hat{\theta}_n - \theta) \xrightarrow{d} N_k(0, \frac{1}{I_X(\theta)})$$

Fisher Information Matrix

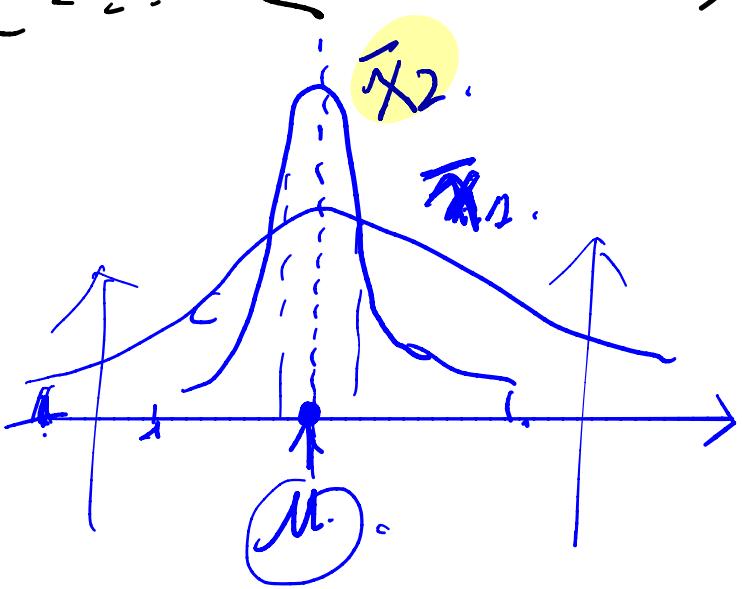
$$I_X(\theta) := E \left[\begin{matrix} \frac{\partial \log f_X(X|\theta)}{\partial \theta_i} \\ \frac{\partial \log f_X(X|\theta)}{\partial \theta_j} \end{matrix} \right]_{k \times k}, \quad i = 1, 2, \dots, k, \quad j = 1, 2, \dots, k.$$

$$= E \left[\underline{\underline{L'(\theta)}}^T \cdot \underline{\underline{L'(\theta)}}_{k \times k} \right]$$

Efficient: (Without asymptotically)

$$\hat{\theta} \rightarrow$$

- { $\text{Var}(\hat{\theta}) = \frac{1}{I_X(\theta)} \rightarrow \hat{\theta}$ is efficient
- - - > $\frac{1}{I_X(\theta)}$ $\rightarrow \hat{\theta}$ is sub-efficient
- - - < - - - $\rightarrow \hat{\theta}$ is super-efficient



$$X_1, X_2, \dots, X_n \sim N(\mu, \frac{\sigma^2}{n})$$

σ^2 known

$$E(\bar{N}) = \mu \quad \frac{\sigma^2}{n}$$

$$\bar{N} = \bar{x} \sim N(\mu, \text{Var})$$

$$\text{Var}(\hat{\theta}) = \text{Var}(\bar{x})$$

$$\frac{1}{I_X(\theta)} = \frac{\sigma^2}{n}$$

$$I_X(\theta) = \frac{n}{\sigma^2}$$

Example:

Let $X_1, X_2, \dots, X_n \sim \text{Bin}(1, P) \quad P \in (0, 1)$

Q MLE for P Pdf $\stackrel{\leftarrow}{P}^{x_1} (1-P)^{1-x_1}$

Solution: $\theta = P$

$$\textcircled{1} \quad J(P) = \prod_{i=1}^n f(x_i) = \prod_{i=1}^n P^{x_i} (1-P)^{1-x_i}$$

$$\textcircled{2} \quad l(P) = \log J(P) = n \left[\bar{x} \log P + (1-\bar{x}) \cdot \log (1-P) \right]$$

$$= \boxed{n \sum_{i=1}^n x_i / n}$$

$$\textcircled{3} \quad l'(P) = \frac{d l(P)}{d P} = n \left[\frac{\bar{x}}{P} - \frac{1-\bar{x}}{1-P} \right] = 0$$

$$\Rightarrow \hat{P} = \bar{x}$$

$$\textcircled{4} \quad l''(P) = 1 - \frac{\bar{x}}{P^2} - \frac{(1-\bar{x})}{(1-P)^2} < 0.$$

\bar{x} is a mLG for P // $\hat{P} \rightarrow g(\hat{P}) = \bar{x}$

Q2: $P^2 - \hat{P}(1+\hat{P})$

invariance. $P^2 =: g(P)$

$$\left\{ \begin{array}{l} \bar{x} \xrightarrow{\text{MLE}} P \\ g(P) = P^2 \end{array} \right. \rightarrow g(\bar{x}) = \boxed{\frac{(\bar{x})^2}{T}} \xrightarrow{\text{MLE}} g(P) = P^2$$

