

MATH243 Statistical Inference

Exercise 3 (Parameter Estimation II: Rao-Blackwell Theorem)

Cramer-Rao lower bound:

1. Let \mathcal{X} be a random sample of size n from $\text{Bi}(1, \theta)$. Find the maximum likelihood estimator for θ and verify in this case the Cramer-Rao Inequality. Comment on what you have found.
2. Let \mathcal{X} be a random sample of size n from an exponential distribution with parameter $\frac{1}{\theta}$. The p.d.f. is

$$f(x; \theta) = \frac{1}{\theta} \exp\left(-\frac{x}{\theta}\right), \quad x \in (0, \infty)$$

Find the Cramer-Rao lower bound and deduce whether or not \bar{X} is a fully efficient estimator for θ .

Exponential family, Sufficient Statistic, Rao-Blackwell Theorem

3. Show that the following p.d.f.'s belong to the exponential family (discrete case)
 - (a) Poisson;
 - (b) Binomial;
 - (c) Negative Binomial.
4. Show that the following p.d.f.'s belong to the exponential family (continuous cases).
 - (a) Gamma $(\theta; k)$, where $k > 0$ is known;
 - (b) $N(\theta; 1)$;
 - (c) $N(0; \theta)$.
5. Let X_1, X_2, \dots, X_n be a random sample of size n from a geometric distribution that has p.d.f. $f(x; \theta) = (1 - \theta)^x \theta$, $x = 0, 1, 2, \dots$, $0 < \theta < 1$, zero elsewhere. Show that $\sum_{i=1}^n X_i$ is a sufficient statistic for θ .
6. Show that the sum of the items of a random sample of size n from a gamma distribution which has p.d.f. $f(x; \theta) = (1/\theta)e^{-x/\theta}$, $0 < x < \infty$, $0 < \theta < \infty$, zero elsewhere, is a sufficient statistic for θ .
7. Suppose we have a random sample \mathcal{X} of size n from a normal distribution $N(0, \theta)$, $\theta \in R^+$. Show that
$$T = t(\mathcal{X}) = \sum_{i=1}^n X_i^2$$
is a sufficient statistic for θ .
8. Let $\mathcal{X} = (X_1, \dots, X_n)$ be a random sample from a uniform distribution over the interval $[0, \theta]$. Show that $Y = \max(X_1, X_2, \dots, X_n)$ is sufficient for θ .
9. Prove that the sum of the items of a random sample of size n from a Poisson distribution having parameter θ , $0 < \theta < \infty$, is a sufficient statistic for θ .

10. Consider a random sample of size n drawn from a Bernoulli distribution. Show that the single statistic, $X_k, k \in \{1, 2, \dots, n\}$, is an unbiased estimator for θ , the probability of success, but not sufficient for θ . Show that $Y = \sum_{i=1}^n X_i$ is sufficient for θ . Find an improved unbiased estimator provided by the Rao-Blackwell Theorem.
11. Let the random variable X denote the time to failure of television tubes and assume that X is exponentially distributed with unknown mean θ . Let $\underline{X} = (X_1, X_2)$ be a random sample from this distribution
- Show that X_1, X_2 are unbiased estimator for θ but not sufficient for θ .
 - Show that $Y = X_1 + X_2$ is sufficient for θ .
 - Hence obtain an unbiased estimator for θ which has variance which is not greater than that of X_1 or X_2 for all $\theta (0 < \theta < \infty)$.

12. If X_1, X_2 denote a random sample of size 2 from a distribution having p.d.f.

$$f(x; \theta) = (1/\theta)e^{-x/\theta}, \quad 0 < x < \infty, \quad 0 < \theta < \infty.$$

- Find the joint p.d.f. of $Y_1 = X_1 + X_2$, and $Y_2 = X_2$.
 - Show that Y_2 is an unbiased estimator for θ having variance θ^2 , and apply Rao-Blackwell Theorem to find an improved estimator. Does the variance of the improved estimator attain the Cramer-Rao lower bound?
13. Write the p.d.f.

$$f(x; \theta) = \frac{1}{6\theta^4} x^3 e^{-x/\theta}, \quad 0 < x < \infty, \quad 0 < \theta < \infty,$$

zero elsewhere, in the exponential form. If X_1, X_2, \dots, X_n is a random sample from this distribution, find a complete sufficient statistic Y_1 for θ and the unique continuous function $\varphi(Y_1)$ of this statistic which is the best statistic for θ . Is $\varphi(Y_1)$ itself a complete sufficient statistic?

14. Let X_1, X_2, \dots, X_n denote a random sample of size $n > 2$ from a distribution with p.d.f. $f(x; \theta) = \theta e^{-\theta x}$, $0 < x < \infty$, zero elsewhere, and $\theta > 0$. Then $Y = \sum_{i=1}^n X_i$ is a sufficient statistic for θ . Prove that $(n-1)/Y$ is the best statistic for θ .
15. Let X_1, X_2, \dots, X_n denote a random sample from a distribution which is $b(1, \theta)$. Find the best statistic for the variance $n\theta(1-\theta)$ of $Y = \sum X_i$.
16. Let X_1, X_2, \dots, X_n denote a random sample from a distribution which is $n(0, \theta)$. Then $Y = \sum X_i^2$ is a sufficient statistic for θ . Find the best statistic for θ^2 .
17. Let X_1, X_2, \dots, X_n denote a random sample of size $n > 2$ from a distribution with p.d.f. $f(x; \theta) = \theta e^{-\theta x}$, $0 < x < \infty$, zero elsewhere, and $\theta > 0$. Find the best estimator for θ .
18. Let (X_1, \dots, X_n) be a sample from $P_0(\lambda)$, find the UMVUE of $\tau(\lambda) = \frac{\lambda^k e^{-\lambda}}{k!}$ for any $k = 0, 1, 2, \dots$.
19. Let X be a r.v. having the Negative Binomial distribution with known k and parameter $\theta \in \Omega = (0, 1)$. Find the UMVU estimator of $g(\theta) = 1/\theta$ and determine its variance.
20. Let X_1, \dots, X_n be i.i.d. r.v.'s from the Gamma distribution with α known and $\theta \in \Omega = (0, \infty)$ unknown, i.e.

$$f(x, \theta) = \frac{x^{\alpha-1} e^{-x/\theta}}{\theta^\alpha \Gamma(\alpha)}.$$

Then show that the UMVU estimator of θ is

$$U(X_1, \dots, X_n) = \frac{1}{n\alpha} \sum_{j=1}^n X_j$$

and its variance attains the Cramer-Rao bound.

21. Let X be a r.v. denoting the life length of an equipment. Then the reliability of the equipment at time x , $R(x)$, is defined as the probability that $X > x$. If X has the exponential distribution with parameter $\theta \in \Omega = (0, \infty)$, find the UMVU estimator of the reliability $R(x, \theta)$ on the basis of n observations on X .

Difficult questions:

22. Let $Y_1 < Y_2 < Y_3 < Y_4 < Y_5$ be the order statistics of a random sample of size 5 from the uniform distribution having p.d.f. $f(x; \theta) = 1/\theta$, $0 < x < \theta$, $0 < \theta < \infty$, zero elsewhere. Show that $2Y_3$ is an unbiased statistic for θ . Determine the joint p.d.f. of Y_3 and the sufficient statistic Y_5 for θ . Find the conditional expectation $E(2Y_3|y_5) = \varphi(y_5)$. Compare the variances of $2Y_3$ and $\varphi(Y_5)$.
23. Let the random variables X and Y have the joint p.d.f. $f(x, y) = (2/\theta^2)e^{-(x+y)/\theta}$, $0 < x < y < \infty$, zero elsewhere.
- Show that the mean and the variance of Y are, respectively, $3\theta/2$ and $5\theta^2/4$.
 - Show that $E(Y|x) = x + \theta$. In accordance with the Rao-Blackwell theorem, the expected value of $X + \theta$ is that of Y , namely, $3\theta/2$, and the variance of $X + \theta$ is less than that of Y . Show that the variance of $X + \theta$ is in fact $\theta^2/4$.
24. Let a random sample of size n be taken from a distribution of the discrete type with p.d.f. $f(x; \theta) = 1/\theta$, $x = 1, 2, \dots, \theta$, zero elsewhere, where θ is an unknown positive integer.
- Show that the largest item, say Y , of the sample is a complete sufficient statistic for θ .
 - Prove that
$$[Y^{n+1} - (Y-1)^{n+1}]/[Y^n - (Y-1)^n]$$
is the unique best statistic for θ .
25. Suppose X_1, \dots, X_n are independent random variable with distribution $B(1, p)$.
- Find the maximum likelihood estimator of $\theta = (1-p)^2$.
 - Show that $\hat{\theta}$ is an unbiased estimator of θ , where $\hat{\theta} = 1$ if $X_1 + X_2 = 0$ and $\hat{\theta} = 0$ otherwise.
 - Find the best estimator for θ .

26. Let X_1, \dots, X_n be i.i.d. r.v.'s from the $U(\theta, 2\theta)$, $\theta \in \Omega = (0, \infty)$ distribution and set

$$U_1 = \frac{n+1}{2n+1}Y_n \quad \text{and} \quad U_2 = \frac{n+1}{5n+4}[2Y_n + Y_1].$$

Then show that both U_1 and U_2 are unbiased estimators of θ and the U_2 is better than U_1 .

27. Let X_1, \dots, X_n be independent r.v.'s distributed as $N(\theta, 1)$. Show that $\bar{X}^2 - (1/n)$ is the UMVU estimator of $g(\theta) = \theta^2$. Also show that the Cramer-Rao bound is not attained.
28. Let X_1, \dots, X_n be independent r.v.'s distributed as $N(\mu, \sigma^2)$, where both μ and σ^2 are unknown. Find the UMVU estimator of μ/σ .
29. Let X_1, \dots, X_m and Y_1, \dots, Y_n be two independent random samples with the same mean θ and known variances σ_1^2 and σ_2^2 , respectively. Then show that for every $c \in [0, 1]$, $U = c\bar{X} + (1-c)\bar{Y}$ is an unbiased estimator of θ . Also find the value of c for which the variance of U is minimum.