

# The Hong Kong University of Science & Technology

## MATH243 – Statistical Inference

### Final Examination – Fall 02/03

Answer ALL questions

All Equal Marks

Date: 12 December 2002 (Thu)

Time allowed: 3 Hours

1. (a) Find the mean and variance of the rectangular (or uniform) distribution on  $(0, \theta)$ , where  $\theta > 0$ .
  - (b) If  $Y_1, \dots, Y_n$  are independently rectangularly distributed on  $(0, \theta)$ , obtain the probability density function of  $Z = \max(Y_1, \dots, Y_n)$ .
  - (c) Find the expectation and variance of  $Z$ .
  - (d) If  $Y_1, \dots, Y_n$  are as in (b) and  $\bar{Y} = \sum Y_j / n$ , find the expectation and variance of  $\bar{Y}$ .
  - (e) If  $\theta$  is an unknown parameter, construct two unbiased estimators of  $\theta$ , one based on  $Z$  and one based on  $\bar{Y}$ .
  - (f) Compare the variances of the two unbiased estimates in (e) and comment briefly.
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2. Suppose that  $X_1, \dots, X_n$  is a random sample from Poisson distribution with mean  $\theta$ .
  - (a) Find the maximum likelihood estimate of  $\theta^2$ .
  - (b) Obtain an unbiased estimate of  $\theta^2$  of the form  $a\bar{X} + b\bar{X}^2$  where  $a$  and  $b$  are constants and  $\bar{X} = \sum X_i / n$ .
  - (c) Find  $\text{Var}(a\bar{X} + b\bar{X}^2)$ .

Hints: (1)  $\left. \frac{d^r m(t)}{dt^r} \right|_{t=0} = E(x^r)$  where  $m(t)$  is moment generating function of  $x$ .

$$(2) \quad \left( \sum_{i=1}^n x_i \right)^3 = \sum_{i=1}^n x_i^3 + 3 \sum_{i=1}^n \sum_{j \neq i} x_i^2 x_j + \sum_{i=1}^n \sum_{j \neq i} \sum_{k \neq i, j} x_i x_j x_k$$

$$(3) \quad \left( \sum_{i=1}^n x_i \right)^4 = \sum_{i=1}^n x_i^4 + 3 \sum_{i=1}^n \sum_{j \neq i} x_i^3 x_j + 4 \sum_{i=1}^n \sum_{j \neq i} x_i^2 x_j^2 + 6 \sum_{i=1}^n \sum_{j \neq i} \sum_{k \neq i, j} x_i^2 x_j x_k \\ + \sum_{i=1}^n \sum_{j \neq i} \sum_{k \neq i, j} \sum_{l \neq i, j, k} x_i x_j x_l x_k$$

- (d) What is the Crámer Rao lower bound for the variance of an unbiased estimator of  $\theta^2$ ? Is the Crámer Rao lower bound attained? Why?

3. (a) The observations  $X_1, \dots, X_m$  are independent and each has a Binomial distribution with index  $n$  and parameter  $\theta$ . Find the minimum variance unbiased estimator of  $\theta^k$ , where  $k$  is an integer ( $1 \leq k \leq mn$ ).
- (b) A random sample,  $X_1, \dots, X_n$ , is taken from the exponential distribution with parameter  $\lambda$ . Find the minimum variance unbiased estimator of  $\lambda^r$  ( $0 < r < n$ ) where  $r$  is an integer.

4. The observations  $X_1, \dots, X_n$  are independent and have an exponential distribution with unknown parameter  $\lambda$ . Find a uniformly most powerful test of the null hypothesis  $H_0: \lambda = \lambda_0$  against the alternative  $H_A: \lambda > \lambda_0$ .

In a particular situation measurement of each individual observation is costly although comparing observations is not. For this reason a different test is suggested in which  $H_0$  is rejected if the observed value of  $T = \max_i(X_i)$  is less than a constant  $c$  which is chosen at significance level  $\alpha$  ( $0 < \alpha < 1$ ). Find the power function of this test. Hence, determine the minimum sample size  $n$  required to obtain power at least 0.8 against the alternative  $\lambda_1 = 0.002$  when  $\lambda_0 = 0.001$  and  $\alpha = 0.05$ .

Is the sample size larger than that required in the previous test? Why? NO calculation is needed.

5. (a)  $X_1, \dots, X_n$  are independent Poisson random variables with  $E(X_i) = \mu_i$  ( $i = 1, \dots, n$ ). Derive the approximate large sample likelihood ratio test for the null hypothesis that  $\mu_1 = \dots = \mu_n$  against the alternative that the  $\mu_i$ s are arbitrary.

By writing  $X_i = \bar{X} + U_i$ , where  $\bar{X} = \sum X_i / n$ , and assuming that the  $U_i$ s are small compared with  $\bar{X}$ , show that the test in (a) approximately compares  $s^2 = \sum (X_i - \bar{X})^2 / (n-1)$  with  $\bar{X}$ . Comment.

*Hint:* Take  $\text{Log}(1 + a) \approx a$  if  $a$  is small.

- (b) Let  $X_1, \dots, X_n$  be i.i.d. random variables, each with the Poisson distribution of parameter  $\mu$ . The best test of  $H_0: \mu = \mu_0$  against  $H_1: \mu = \mu_1 (> \mu_0)$  is  $C_1 = \left\{ x: \sum_{i=1}^n x_i \geq k \right\}$ . By using the central limit theorem to approximate the distribution of  $\sum_{i=1}^n x_i$ , find the smallest value of  $n$  required to obtain power at least 0.9 against the alternative  $\mu_1 = 1.21$  when  $\mu_0 = 1$  and  $\alpha = 0.05$ .

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