

Math 3423 Tutorial

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About me 😊

- CHENG Yu
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- 2 tutorial sessions:
 - 18:00 – 18:50 on Tuesday
 - 19:00 – 19:50 on Friday

Tutorial (For 2020-09-15 and 2020-09-18)

1. What is statistical inference = {① To known unknown things by Sampling
② parametric method}

- Target: To find good estimators to study population

- Population

Population

(Distribution,
cdf. pdf)

know → USE IT.

don't know

SAMPLING



Introduces Uncertainty.

Random Variable (r.v.) \Leftarrow (Don't have a certain solution/answer)

- Estimator: $T(\cdot)$

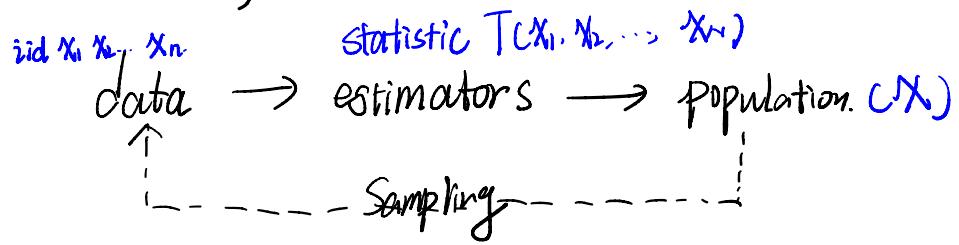
To estimate the uncertainty

{ Parametric method $= \theta \Rightarrow F_\theta \Rightarrow F$ (where θ is a parameter)

{ non-parametric $\cdots =$ don't assume a distribution

- No distribution assumption
- Kendall T (Tau): Compare two variables to get their correlation.
 - Statistical bootstrap method: estimate the accuracy distribution of a statistic
 - Anderson - Darling Test: To test whether a R.V. follow a certain distribution.

Hence,



e.g. If population $X \sim N(\mu, \sigma^2)$

$\mu, \sigma^2 \rightarrow$ Normal

∴ $T(X_1, X_2, \dots, X_n)$ estimate μ and σ^2 .

2. 3 important issues:

① Point estimation:

Using 1 value to make an estimation.

e.g. $X \sim N(\mu, \sigma^2)$ and σ^2 is known.

X_1, X_2, \dots, X_n (from X) → \bar{X} mean.

② Interval estimation:

Using an interval, instead of 1 value to contain the actual value

e.g. $\bar{X} \pm \frac{\sigma}{\sqrt{n}}$ for μ (σ is known)

100(1- α)% chance that this interval contains μ .

Question:

There is 100(1- α)% chance that μ falls into the interval $[\bar{X} - \frac{\sigma}{\sqrt{n}}, \bar{X} + \frac{\sigma}{\sqrt{n}}]$

③ Hypothesis test:

e.g. $\begin{cases} H_0: \mu = \mu_0 \\ H_1: \mu > \mu_0 \end{cases}$ one-way HT

$\begin{cases} H_0: \mu = \mu_0 \\ H_1: \mu \neq \mu_0 \end{cases}$ Two-way HT

3. Summary.

If we have a population $X \sim N(\mu, \sigma^2)$. And we have n copies of X : x_1, x_2, \dots, x_n (iid)

Target	μ	σ		
μ		✓	<ul style="list-style-type: none"> $\bar{X} = \frac{\sum_i^n x_i}{N}$ C.I. $[\bar{X}_n - t_{\frac{\alpha}{2}} \cdot \frac{\sigma}{\sqrt{n}}, \bar{X}_n + t_{\frac{\alpha}{2}} \cdot \frac{\sigma}{\sqrt{n}}]$ 	Z-distribut
μ		✗	<ul style="list-style-type: none"> $\bar{X} = \frac{\sum_i^n x_i}{N}$ C.I. $[\bar{X}_n - t_{\frac{\alpha}{2}} \cdot \frac{(n-1)S_n}{\sqrt{n}}, \bar{X}_n + t_{\frac{\alpha}{2}} \cdot \frac{(n-1)S_n}{\sqrt{n}}]$ 	t-distribut
σ	✓		<ul style="list-style-type: none"> $S_{n-1}^2 \rightarrow \sigma^2$ C.I. $[\frac{(n-1)S_n^2}{\chi_{\frac{\alpha}{2}}^2(n-1)}, \frac{(n-1)S_n^2}{\chi_{1-\frac{\alpha}{2}}^2(n-1)}]$ 	Chi-dist
σ	✗			

• Some useful distribution:

① Multivariate normal distribution.

$$f_X(\mathbf{x}) = (2\pi)^{-\frac{p}{2}} |\Sigma|^{-\frac{1}{2}} e^{-\frac{1}{2} \underline{\mathbf{x}-\mu}^T \Sigma^{-1} (\mathbf{x}-\mu)},$$

↓
quadratic form

denoted by $\mathbf{X} \sim N_p(\boldsymbol{\mu}, \Sigma)$.

$$\begin{aligned} \star A^2 &= -\frac{1}{2} (\mathbf{x}-\boldsymbol{\mu})^T \Sigma^{-1} (\mathbf{x}-\boldsymbol{\mu}) \\ &= -\frac{1}{2} \mathbf{x}^T \Sigma^{-1} \mathbf{x} + \mathbf{x}^T \Sigma^{-1} \boldsymbol{\mu} + \text{constant}. \end{aligned}$$

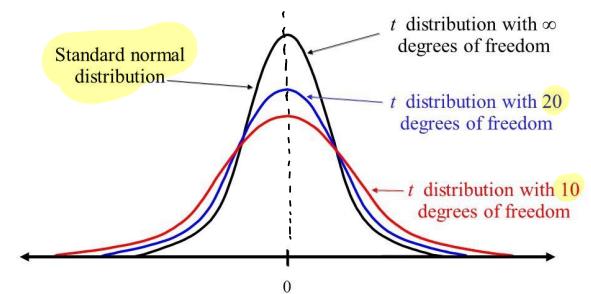
Gaussian distribution \Leftrightarrow Constant

② Chi-square distribution.

- $X_i \sim N(0, 1) \Rightarrow X_i^2 \sim \chi^2(1)$
- $\Rightarrow \sum_{i=1}^N X_i^2 \sim \chi^2(N)$
- $Y_i \sim \chi^2(r_i) \Rightarrow \sum_{i=1}^k Y_i \sim \chi^2(\sum_{i=1}^k r_i)$

t Distribution

The *t*-distribution is used when n is small and σ is unknown.



Definition (*t* distribution): If $Z \sim N(0, 1)$, $U \sim \chi^2(r)$, and they are independent, then the distribution of

$$T := \frac{Z}{\sqrt{U/r}} = \frac{\text{standard Normal}}{\text{Chi-square}}$$

is called a *t* distribution with r degrees of freedom, denoted by $t(r)$.

Note that the pdf of $T \sim t(r)$ can be shown to be

$$f_T(t) = \frac{\Gamma[(r+1)/2]}{\sqrt{\pi r} \Gamma(r/2)} (1 + t^2/r)^{-(r+1)/2},$$

Standard Normal:

$$f_X(x) = \frac{1}{\sqrt{2\pi \sigma^2}} \cdot \exp \left[-\frac{1}{2} \left(\frac{x-\mu}{\sigma} \right)^2 \right]$$

When $r \rightarrow \infty$

$$\begin{aligned} f_T(t) &= \frac{\Gamma(\frac{r+1}{2})}{\sqrt{\pi r} \Gamma(\frac{r}{2})} \cdot \left(1 + \frac{t^2}{r}\right)^{-\frac{(r+1)}{2}} \\ &= \frac{1}{\sqrt{\pi r}} \cdot \left(1 + \frac{t^2}{r}\right)^{-\frac{1}{2}} \cdot \left(1 + \frac{t^2}{r}\right)^{-\frac{r}{2}} \\ &= \frac{1}{\sqrt{\pi r}} \cdot \left(1 + \frac{t^2}{r}\right)^{-\frac{1}{2}} \cdot \left(1 + \frac{t^2}{r}\right)^{\frac{r}{2}} \cdot \left(1 + \frac{t^2}{r}\right)^{-\frac{r}{2}} \\ &\quad \exp(-\frac{t^2}{2r}) \cdot \exp(-\frac{t^2}{2}) \end{aligned}$$

$$\begin{aligned} &\Gamma(\frac{r}{2} + \frac{1}{2}) \quad \Gamma(\frac{r}{2}) \\ &\left(1 + \frac{t^2}{r}\right)^{-\frac{1}{2}} \quad \left(1 + \frac{t^2}{r}\right)^{-\frac{r}{2}} \\ &\left(1 + \frac{t^2}{r}\right)^{\frac{r}{2}} \quad \left(1 + \frac{t^2}{r}\right)^{-\frac{r}{2}} \\ &\exp(-\frac{t^2}{2r}) \quad \exp(-\frac{t^2}{2}) \end{aligned}$$

④ F distribution

Definition (Fdistribution): If $X \sim \chi^2(r_1)$, $Y \sim \chi^2(r_2)$, and they are independent, then the distribution

$$F(r_1, r_2) := \frac{X/r_1}{Y/r_2} = \frac{\text{Chi-square / df}}{\text{Chi-square / df}}$$

is called a F distribution with r_1 and r_2 degrees of freedom, denoted by $F(r_1, r_2)$.

Note that the pdf of $F \sim F(r_1, r_2)$ can be shown to be

$$f_F(w) = \frac{\Gamma[(r_1 + r_2)/2]}{\Gamma(r_1/2) \Gamma(r_2/2)} \left(\frac{r_1}{r_2}\right)^{\frac{r_1}{2}} w^{\frac{r_1}{2}-1} (1 + r_1 w/r_2)^{-(r_1+r_2)/2},$$

Appendix

