

Homework 10

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Use rats data in the faraway package.

- Description: An experiment was conducted as part of an investigation to combat the effects of certain toxic agents.
- Format: A data frame with 48 observations on the following 3 variables:
- time: survival time in tens of hours
- poison: the poison type - a factor with levels I, II, and III randomly assigned to subjects
- treat: the treatment - a factor with levels A, B, C, and D randomly assigned to subjects

```
data(rats)
summary(rats)
```

	time	poison	treat
Min.	:0.1800	I :16	A:12
1st Qu.	:0.3000	II :16	B:12
Median	:0.4000	III:16	C:12
Mean	:0.4794		D:12
3rd Qu.	:0.6225		
Max.	:1.2400		

```
glimpse(rats)
```

Rows: 48

Columns: 3

```
$ time    <dbl> 0.31, 0.82, 0.43, 0.45, 0.45, 1.10, 0.45, 0.71, 0.46, 0.88, 0.6~
$ poison  <fct> I, I, I, I, I, I, I, I, I, I, I, I, I, I, I, I, II, II, II, II, ~
$ treat   <fct> A, B, C, D, A, B, C, D, A, B, C, D, A, B, C, D, A, B, C, D, A, ~
```

```
rats %>%
  tabyl(poison,treat) # janitor::tabyl()
```

```
poison A B C D
I 4 4 4 4
II 4 4 4 4
III 4 4 4 4
```

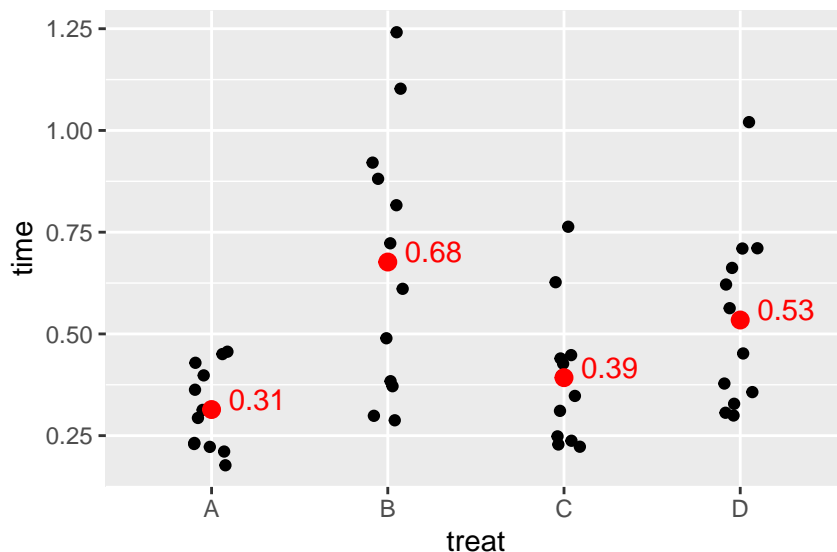
1. Focus on time as the response and treat as the predictor. Plot time as a function of treat using a jittered scatter plot. Add a mean of time for each level of treat to the plot. What pattern do you see?

- Survival time is longest for treatment B followed by treatment D. Treatment A has the shortest survival time and the smallest variability.

```
(means <- aggregate(time~treat,rats, mean))
```

```
  treat      time
1     A 0.3141667
2     B 0.6766667
3     C 0.3925000
4     D 0.5341667
```

```
ggplot(rats,aes(x=treat, y=time))+
  geom_point(position=position_jitter(width=0.1)) +
  stat_summary(fun="mean", colour = "red")+
  geom_text(data = means,
            aes(label = round(time, 2)),
            vjust = 0, hjust = -0.3, size = 4, color="red")
```



2. Perform a one-way ANOVA with time as the response and treat as the predictor. Use `aov()` as well as `lm()`. For each, describe what you observe about what can be concluded.

- They have the same F-statistic (6.48) and p-value (0.000992) indicating that there is significant effect of treatment on time.
- `aov()` neither included the intercept nor the levels of treatments. Its p-value (0.000992) supports rejecting the null that there is no difference among the group means - in favor of the alternative hypothesis (at least one group mean is different).
- `lm()` included the intercept (reference group) and all the other three levels (B,C,D) of treatment. The p-values of the treatment coefficients suggests that the means for treatments B and D are different from the reference group (treatment A) while treatment C is not significantly different from treatment A.

```
summary(aov(time~treat,rats))
```

```

              Df Sum Sq Mean Sq F value    Pr(>F)
treat          3  0.9212  0.30707     6.484 0.000992 ***
Residuals     44  2.0839  0.04736
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
anova(lm(time~treat,rats))
```

Analysis of Variance Table

Response: time

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
treat	3	0.92121	0.307069	6.4836	0.0009921 ***
Residuals	44	2.08387	0.047361		

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```
summary(lm(time~treat,rats))
```

Call:

```
lm(formula = time ~ treat, data = rats)
```

Residuals:

	Min	1Q	Median	3Q	Max
	-0.38667	-0.15292	-0.01417	0.12833	0.56333

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	0.31417	0.06282	5.001	9.62e-06 ***
treatB	0.36250	0.08885	4.080	0.000186 ***
treatC	0.07833	0.08885	0.882	0.382739
treatD	0.22000	0.08885	2.476	0.017196 *

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.2176 on 44 degrees of freedom

Multiple R-squared: 0.3065, Adjusted R-squared: 0.2593

F-statistic: 6.484 on 3 and 44 DF, p-value: 0.0009921

3. Perform pairwise comparisons using Tukey's difference test. Which pairs of treatment are different? How does this compare to your results from #2?

- The following pairs are different: B-A and C-B as they have p-values < 0.05
- The p-values in the Tukey's difference test were larger than in #2 as Tukey's test accounts for the number of comparisons making it more robust while controlling for Type I error.

```
TukeyHSD(aov(time~treat,rats))
```

Tukey multiple comparisons of means
95% family-wise confidence level

```
Fit: aov(formula = time ~ treat, data = rats)
```

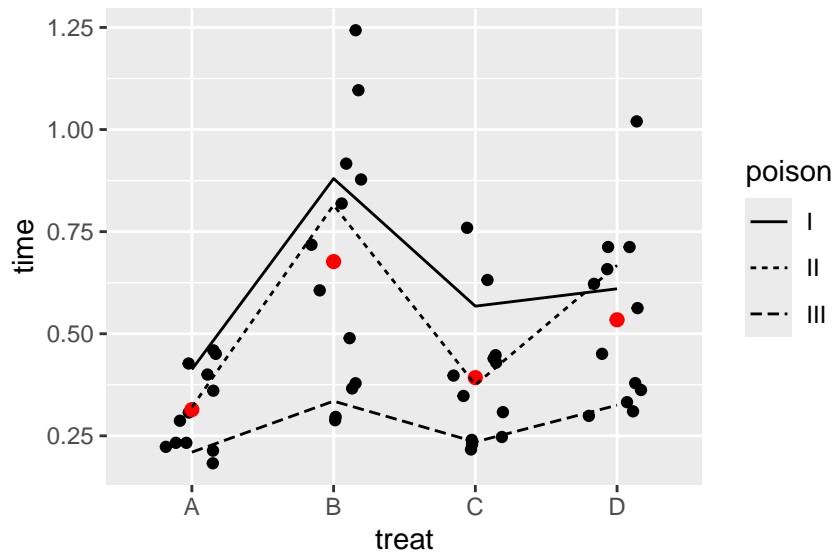
```
$treat
```

	diff	lwr	upr	p adj
B-A	0.36250000	0.12528287	0.59971713	0.0010358
C-A	0.07833333	-0.15888380	0.31555047	0.8143113
D-A	0.22000000	-0.01721713	0.45721713	0.0778376
C-B	-0.28416667	-0.52138380	-0.04694953	0.0131752
D-B	-0.14250000	-0.37971713	0.09471713	0.3869986
D-C	0.14166667	-0.09555047	0.37888380	0.3921830

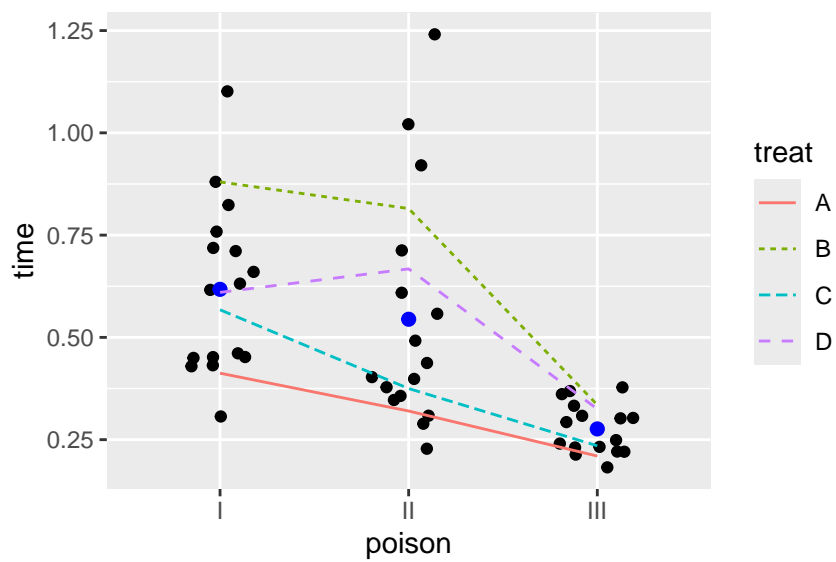
4. Go back to your plot from #1 and add lines showing time as at the interaction of poison and treat. Does the pattern from #1 hold across poison types?

- The pattern from #1 is consistent with the pattern observed here - treatment B has the longest survival time, next is treatment D, then treatment C, while treatment A has the shortest survival time.
- In the first plot we see that treatment B and poison type I and II have higher survival time, followed by the same poison types and treatment D. Flipping the graphic around, we can easily see how treatment B declines across poison types with the largest surviving time for poison type I, followed by II, and lastly III. We see the same trend for treatment types A and C. However, for treatment D, the highest survival time is in combination with poison II, followed by I, and lastly III.

```
ggplot(rats,aes(x=treat,y=time))+  
  geom_point(position=position_jitter(width=0.2))+  
  stat_summary(fun=mean, geom="point", color="red",size=2)+  
  stat_summary(fun=mean, geom="line",  
              aes(group=poison, linetype=poison))
```



```
ggplot(rats,aes(x=poison,y=time))+
  geom_point(position=position_jitter(width=0.2))+
  stat_summary(fun=mean, geom="point", color="blue",size=2)+
  stat_summary(fun=mean, geom="line",
    aes(group=treat, linetype=treat, color=treat))
```



5. Model time as a function of treat, poison, and their interaction in a linear regression. Check whether OLS assumptions hold, and if not, what (if any) transformation of the outcome might present a solution.

- The OLS assumptions do not appear to be met. Data on the qq-plot are positively skewed and the constant variance appears to be violated.
- Since lambda is closer to -0.5, an inverse square root transformation is a potential solution. Other possible recommendations based on choice of lambda are:
 - lambda = 0 => log transformation
 - lambda = -1 => reciprocal transformation

```
summary(rat_ols <- lm(time~treat*poison,rats))
```

Call:

```
lm(formula = time ~ treat * poison, data = rats)
```

Residuals:

	Min	1Q	Median	3Q	Max
	-0.32500	-0.04875	0.00500	0.04312	0.42500

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	0.41250	0.07457	5.532	2.94e-06 ***
treatB	0.46750	0.10546	4.433	8.37e-05 ***
treatC	0.15500	0.10546	1.470	0.1503
treatD	0.19750	0.10546	1.873	0.0692 .
poisonII	-0.09250	0.10546	-0.877	0.3862
poisonIII	-0.20250	0.10546	-1.920	0.0628 .
treatB:poisonII	0.02750	0.14914	0.184	0.8547
treatC:poisonII	-0.10000	0.14914	-0.671	0.5068
treatD:poisonII	0.15000	0.14914	1.006	0.3212
treatB:poisonIII	-0.34250	0.14914	-2.297	0.0276 *
treatC:poisonIII	-0.13000	0.14914	-0.872	0.3892
treatD:poisonIII	-0.08250	0.14914	-0.553	0.5836

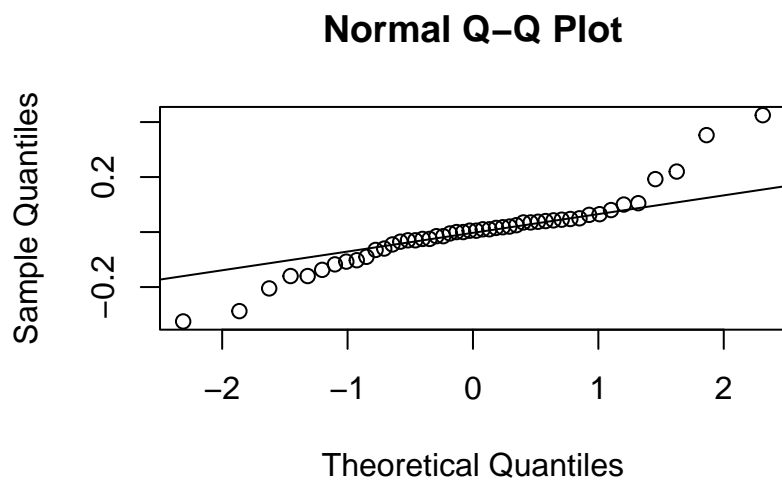
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.1491 on 36 degrees of freedom

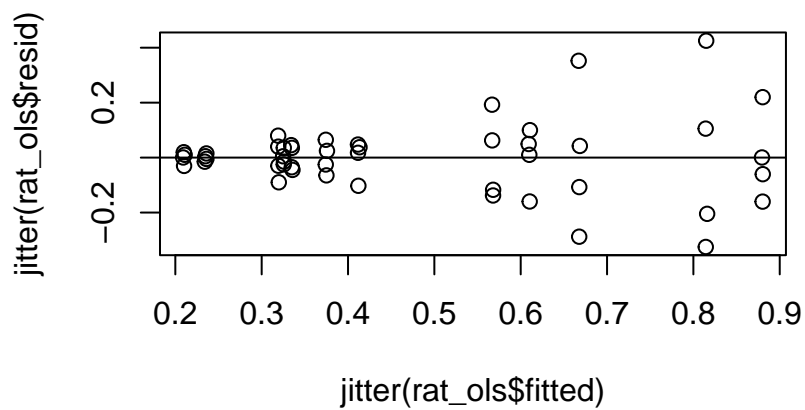
Multiple R-squared: 0.7335, Adjusted R-squared: 0.6521

F-statistic: 9.01 on 11 and 36 DF, p-value: 1.986e-07

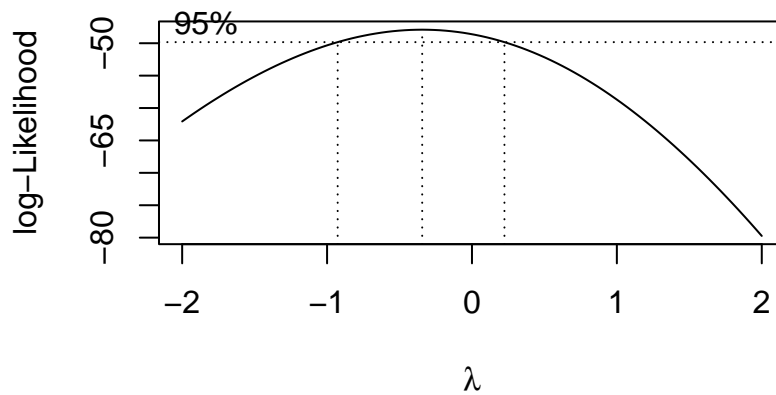
```
qqnorm(rat_ols$resid);qqline(rat_ols$resid)
```



```
plot(jitter(rat_ols$fitted),jitter(rat_ols$resid));abline(h=0)
```



```
boxcox(lm(time~treat,rats))
```

6. Now fit a model of time as a function of treat, but with treat as a random effect. Interpret the effect of treat on time.

- The fixed intercept of .48, survival time was significant ($p < .05$) with a 95% confidence interval between .32 and .64 and does not include zero => reject null hypothesis: variance from treatment does contribute to the overall variance therefore, there may be meaningful differences between treatments. Furthermore, we calculate the intraclass correlation (ICC) to be .31 which is the proportion of variance in time values explained by the grouping structure of treatment.

```
# (1|treat) means we are fitting as random effect
summary(randmod<-lmer(time~(1|treat), rats))
```

```
Linear mixed model fit by REML ['lmerMod']
Formula: time ~ (1 | treat)
Data: rats
```

```
REML criterion at convergence: -0.5
```

```
Scaled residuals:
```

Min	1Q	Median	3Q	Max
-1.6369	-0.7218	-0.1514	0.5186	2.7284

```
Random effects:
```

Groups	Name	Variance	Std.Dev.
treat	(Intercept)	0.02164	0.1471
Residual		0.04736	0.2176

Number of obs: 48, groups: treat, 4

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	0.47937	0.07998	5.993

```
confint(randmod)
```

		2.5 %	97.5 %
.sig01	0.04474072	0.3338817	
.sigma	0.17900826	0.2724161	
(Intercept)	0.30345222	0.6552976	

```
performance::icc(randmod)
```

Intraclass Correlation Coefficient

Adjusted ICC: 0.314
Unadjusted ICC: 0.314