

# Homework 9

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```
library(jtools)
library(leaps)
library(car)
library(gratia)
library(dplyr)
library(faraway)
library(huxtable)
library(MASS)
library(mgcv)
library(tidyverse)

options(scipen=999)
```

**1. Faraway Chapter 9. Exercise 7. Use the cheddar data for this question.**

```
data("cheddar")
summary(cheddar)
```

taste	Acetic	H2S	Lactic
Min. : 0.70	Min. : 4.477	Min. : 2.996	Min. : 0.860
1st Qu.: 13.55	1st Qu.: 5.237	1st Qu.: 3.978	1st Qu.: 1.250
Median : 20.95	Median : 5.425	Median : 5.329	Median : 1.450
Mean : 24.53	Mean : 5.498	Mean : 5.942	Mean : 1.442
3rd Qu.: 36.70	3rd Qu.: 5.883	3rd Qu.: 7.575	3rd Qu.: 1.667
Max. : 57.20	Max. : 6.458	Max. : 10.199	Max. : 2.010

**1.A. Fit a generalized additive model (GAM) for a response of taste with the other three variables as predictors. Do the predictors appear to have a non-linear relationship with the outcome?**

- Based on our plots for the predictors in the GAM we do not see evidence of non-linear relationship with Taste as the outcome variable.

```
# fit a GAM to the data
summary(mod_taste_gam <- gam(taste ~ s(Acetic) + s(H2S) + s(Lactic), data=cheddar))
```

Family: gaussian

Link function: identity

Formula:

taste ~ s(Acetic) + s(H2S) + s(Lactic)

Parametric coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	24.53	1.85	13.26	0.000000000000441 ***

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Approximate significance of smooth terms:

	edf	Ref.df	F	p-value
s(Acetic)	1	1	0.005	0.94198
s(H2S)	1	1	9.818	0.00425 **
s(Lactic)	1	1	5.196	0.03108 *

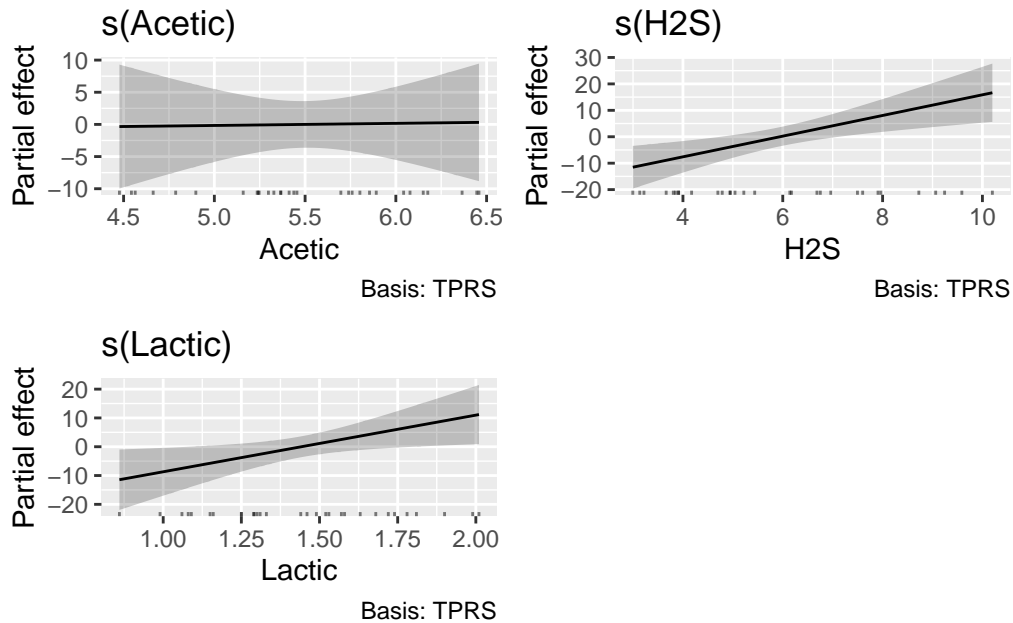
---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

R-sq.(adj) = 0.612 Deviance explained = 65.2%

GCV = 118.42 Scale est. = 102.63 n = 30

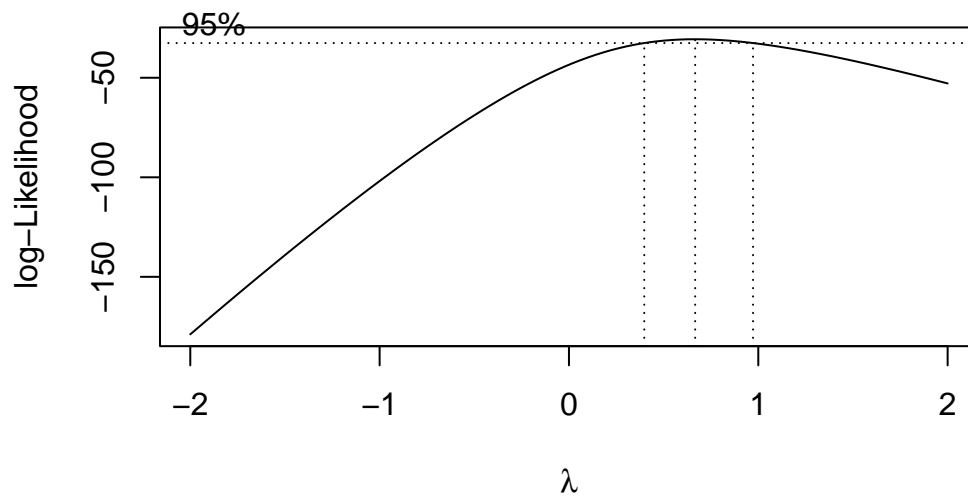
```
gratia::draw(mod_taste_gam)
```



**1.B. Use the Box-Cox method to determine an optimal transformation of the response. Would it be reasonable to leave the response as is (i.e., no transformation)?**

- The Lambda value is between .50 and 1. We can either leave the outcome variable un-transformed or try a squared root transformation.

```
mod_taste_lm <- lm(taste ~., data = cheddar)
boxcox(mod_taste_lm)
```



- If we do decide to squared root transform the outcome response, our residual plot appears to show more evidence on non-constant variance than the residual plot before the transformation.

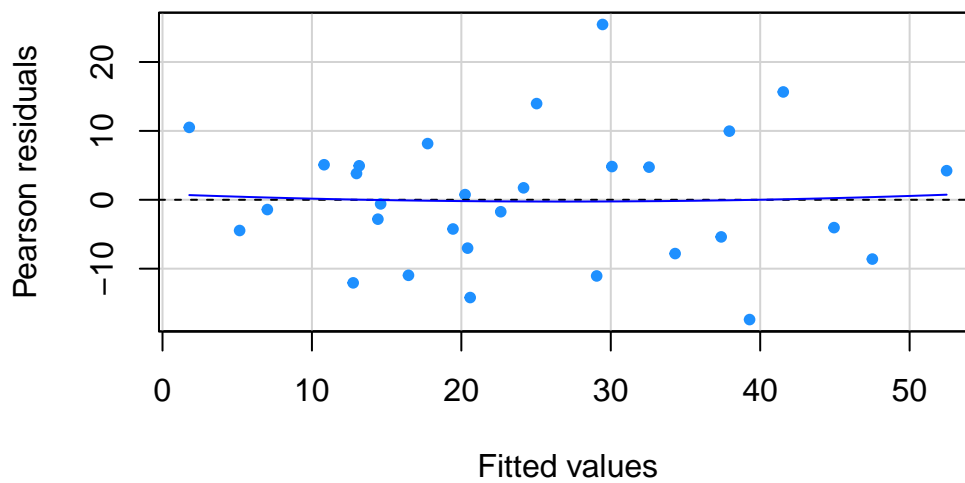
```
mod_taste_lm_t <- lm(I(sqrt(taste)) ~ ., data = cheddar)
export_summs(mod_taste_lm, mod_taste_lm_t)
```

```
residualPlot(mod_taste_lm, pch=20, col= "dodgerblue",
             main="Model w/o transformation")
```

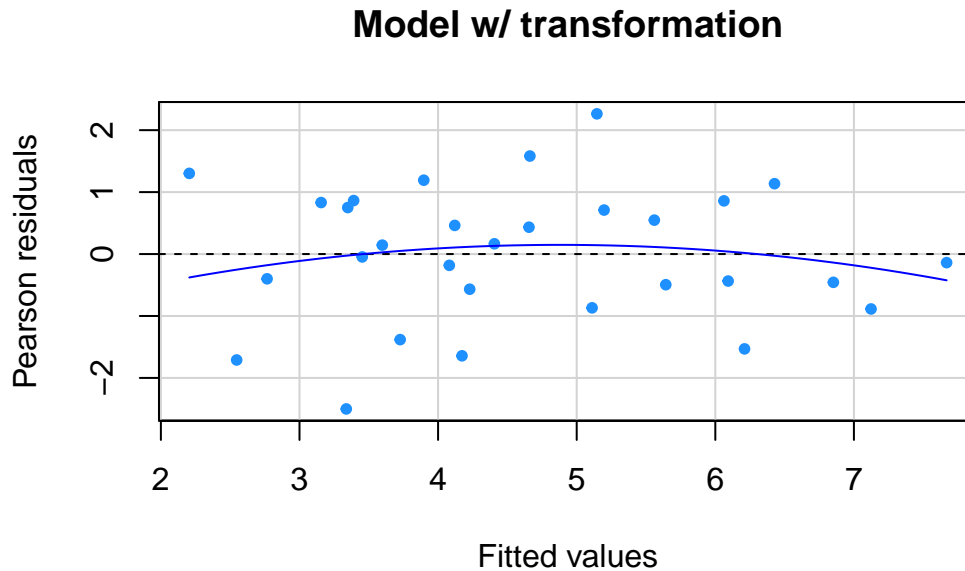
	Model 1	Model 2
(Intercept)	-28.88 (19.74)	-0.92 (2.26)
Acetic	0.33 (4.46)	-0.00 (0.51)
H2S	3.91 ** (1.25)	0.44 ** (0.14)
Lactic	19.67 * (8.63)	2.02 (0.99)
N	30	30
R2	0.65	0.63

\*\*\*  $p < 0.001$ ; \*\*  $p < 0.01$ ; \*  $p < 0.05$ .

### Model w/o transformation



```
residualPlot(mod_taste_lm_t, pch=20, col= "dodgerblue",
             main="Model w/ transformation")
```



**2. Faraway Chapter 10. Exercise 2. Using the teengamb dataset with gamble as the response and the other variables. Implement the following variable selection methods to determine the “best” model.**

```
data("teengamb")
glimpse(teengamb)
```

Rows: 47

Columns: 5

```
$ sex      <int> 1, 1, 1, 1, 1, 1, 1, 1, 1, 1, 1, 1, 1, 1, 1, 1, 1, 1, 1, 0, 0, ~
$ status   <int> 51, 28, 37, 28, 65, 61, 28, 27, 43, 18, 18, 43, 30, 28, 38, 38, ~
$ income   <dbl> 2.00, 2.50, 2.00, 7.00, 2.00, 3.47, 5.50, 6.42, 2.00, 6.00, 3.0~
$ verbal   <int> 8, 8, 6, 4, 8, 6, 7, 5, 6, 7, 6, 6, 4, 6, 6, 8, 8, 5, 8, 9, 8, ~
$ gamble   <dbl> 0.00, 0.00, 0.00, 7.30, 19.60, 0.10, 1.45, 6.60, 1.70, 0.10, 0.~
```

## 2.A. Backward elimination (based on the significance of predictors)

- The full model contains gamble as a function of sex, income, status, and verbal. Given our significance level, we remove status, and finally status and verbal together to refit the models.

```
summ(mod_back <- lm(gamble ~., teengamb))
```

MODEL INFO:

Observations: 47

Dependent Variable: gamble

Type: OLS linear regression

MODEL FIT:

$F(4,42) = 11.69$ ,  $p = 0.00$

$R^2 = 0.53$

Adj.  $R^2 = 0.48$

Standard errors:OLS

	Est.	S.E.	t val.	p
(Intercept)	22.56	17.20	1.31	0.20
sex	-22.12	8.21	-2.69	0.01
status	0.05	0.28	0.19	0.85
income	4.96	1.03	4.84	0.00
verbal	-2.96	2.17	-1.36	0.18

```
# drop status
```

```
summ(mod_back_status <- update(mod_back, ~. -status))
```

MODEL INFO:

Observations: 47

Dependent Variable: gamble

Type: OLS linear regression

MODEL FIT:

$F(3,43) = 15.93$ ,  $p = 0.00$

$R^2 = 0.53$

Adj.  $R^2 = 0.49$

Standard errors:OLS

	Est.	S.E.	t val.	p
(Intercept)	24.14	14.77	1.63	0.11
sex	-22.96	6.77	-3.39	0.00
income	4.90	0.96	5.13	0.00
verbal	-2.75	1.83	-1.50	0.14

```
# drop status and verbal
summ(mod_back_status_verbal <- update(mod_back, .~. -status-verbal))
```

MODEL INFO:

Observations: 47

Dependent Variable: gamble

Type: OLS linear regression

MODEL FIT:

$F(2,44) = 22.12$ ,  $p = 0.00$

$R^2 = 0.50$

Adj.  $R^2 = 0.48$

Standard errors:OLS

	Est.	S.E.	t val.	p
(Intercept)	4.04	6.39	0.63	0.53
sex	-21.63	6.81	-3.18	0.00
income	5.17	0.95	5.44	0.00



## 2.B. Now use AIC. Which is the “best” model?

- The model with sex + income + verbal has the lower AIC, and we determine it is the “best” model for these data based on this criterion.

```
back.mod <- stepAIC(mod_back, direction = "backward")
```

Start: AIC=298.18

```
gamble ~ sex + status + income + verbal
```

	Df	Sum of Sq	RSS	AIC
- status	1	17.8	21642	296.21
<none>			21624	298.18
- verbal	1	955.7	22580	298.21
- sex	1	3735.8	25360	303.67
- income	1	12056.2	33680	317.00

Step: AIC=296.21

```
gamble ~ sex + income + verbal
```

	Df	Sum of Sq	RSS	AIC
<none>			21642	296.21
- verbal	1	1139.8	22781	296.63
- sex	1	5787.9	27429	305.35
- income	1	13236.1	34878	316.64

```
back.mod$aova
```

Step	Df	Deviance	Resid. Df	Resid. Dev	AIC
			42	2.16e+04	298
- status	1	17.8	43	2.16e+04	296

## 2.C. Now use adjusted R2. Which is the “best” model?

- “Best” model is  $\text{gamble} \sim \text{sex} + \text{income} + \text{verbal}$ , the adjusted r-squared is close to the full model while dropping status or one less parameter to estimate.

```
huxtable::huxreg(  
  mod_back, mod_back_status, mod_back_status_verbal,  
  statistics=c( "adj.r.squared"))
```

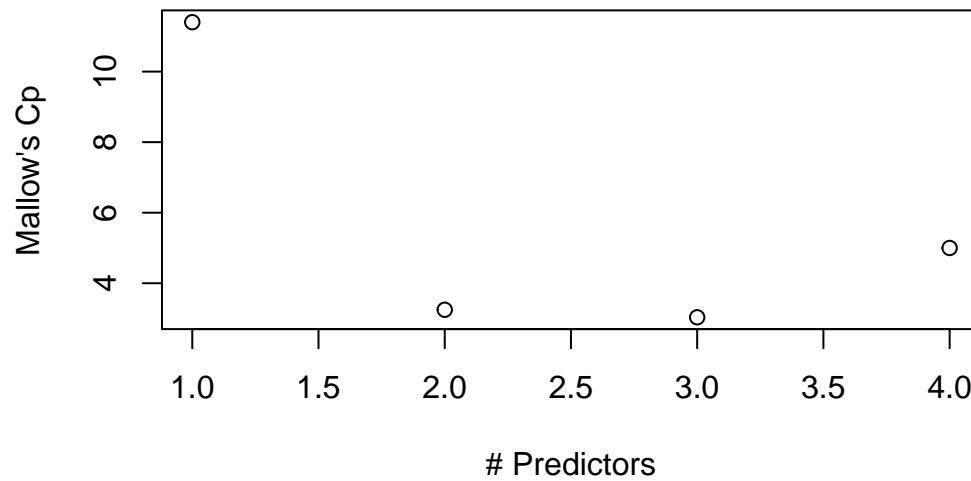
	(1)	(2)	(3)
(Intercept)	22.556 (17.197)	24.139 (14.769)	4.041 (6.394)
sex	-22.118 * (8.211)	-22.960 ** (6.771)	-21.634 ** (6.809)
status	0.052 (0.281)		
income	4.962 *** (1.025)	4.898 *** (0.955)	5.172 *** (0.951)
verbal	-2.959 (2.172)	-2.747 (1.825)	
adj.r.squared	0.482	0.493	0.479

\*\*\* p < 0.001; \*\* p < 0.01; \* p < 0.05.

## 2.D. Now use Mallows Cp. Which is the “best” model?

- “Best” model is  $\text{gamble} \sim \text{sex} + \text{income} + \text{verbal}$  based on having the lowest Mallows  $c_p$  value.

```
sub1<-regsubsets(gamble ~.,teengamb)
rsub1<-summary(sub1)
plot(I(1:4), rsub1$cp, ylab="Mallow's Cp", xlab="# Predictors")
```



```
rsub1$which
```

	(Intercept)	sex	status	income	verbal
1	TRUE	FALSE	FALSE	TRUE	FALSE
2	TRUE	TRUE	FALSE	TRUE	FALSE
3	TRUE	TRUE	FALSE	TRUE	TRUE
4	TRUE	TRUE	TRUE	TRUE	TRUE