

# Particle Filtering: Solutions

ELG 5218 – Uncertainty Evaluation in Engineering Measurements and Machine Learning

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## 1 Resampling Techniques

### Problem 1: Multinomial Resampling

Given  $M$  particles  $\{\mathbf{z}^{(i)}, \tilde{w}^{(i)}\}_{i=1}^M$  where  $\sum_i \tilde{w}^{(i)} = 1$ :

1. Explain how multinomial resampling works.
2. If a particle has normalized weight  $\tilde{w}^{(i)} = 0.2$  and we draw  $M = 100$  resampled particles, what is the expected number of copies?
3. Show that the resampled distribution  $\bar{p}(\mathbf{z}) = \sum_{i=1}^M \frac{N^{(i)}}{M} \delta(\mathbf{z} - \mathbf{z}^{(i)})$  is an unbiased approximation.

#### Solution:

##### 1. How Multinomial Resampling Works:

Each **particle**  $\mathbf{z}^{(i)}$  is treated as a **category** in a multinomial distribution with **parameter**  $\tilde{w}^{(i)}$ . We draw  **$M$  samples (with replacement)** from this categorical distribution:

- Each draw selects particle  $i$  with probability  $\tilde{w}^{(i)}$
- Let  $N^{(i)} =$  number of times particle  $i$  is selected
- $(N^{(1)}, \dots, N^{(M)}) \sim \text{Multinomial}(M, \tilde{w}^{(1)}, \dots, \tilde{w}^{(M)})$
- The new population consists of  $N^{(i)}$  copies of particle  $\mathbf{z}^{(i)}$

Result: **High-weight** particles are **replicated multiple times**; **low-weight** particles are likely discarded.

##### 2. Expected Number of Copies: 照乘 : Weight X M copies = Mean

$$\mathbb{E}[N^{(i)}] = M \cdot \tilde{w}^{(i)} = 100 \times 0.2 = 20 \text{ expected copies}$$

The actual realization is random:  $N^{(i)} \sim \text{Binomial}(M, \tilde{w}^{(i)})$  marginally.

##### 3. Unbiasedness: The resampled distribution is:

$$\bar{p}(\mathbf{z}) = \sum_{i=1}^M \frac{N^{(i)}}{M} \delta(\mathbf{z} - \mathbf{z}^{(i)})$$

Taking expectation:

$$\begin{aligned}
\mathbb{E}[\bar{p}(\mathbf{z})] &= \sum_{i=1}^M \mathbb{E}\left[\frac{N^{(i)}}{M}\right] \delta(\mathbf{z} - \mathbf{z}^{(i)}) \\
&= \sum_{i=1}^M \frac{\mathbb{E}[N^{(i)}]}{M} \delta(\mathbf{z} - \mathbf{z}^{(i)}) \\
&= \sum_{i=1}^M \frac{M\tilde{w}^{(i)}}{M} \delta(\mathbf{z} - \mathbf{z}^{(i)}) \\
&= \sum_{i=1}^M \tilde{w}^{(i)} \delta(\mathbf{z} - \mathbf{z}^{(i)}) = \hat{p}(\mathbf{z})
\end{aligned}$$

Thus  $\bar{p}$  is an unbiased approximation of the original discrete distribution.

## Problem 2: State-Space Models and Optimal Proposal

For nonlinear state-space models:

**Latent State Equation:**  $\mathbf{z}_k = f_z(\mathbf{z}_{k-1}, \mathbf{u}_k)$

**Observation Equation:**  $\mathbf{y}_k = f_y(\mathbf{z}_k, \mathbf{v}_k)$

with densities:

- $p(\mathbf{z}_k | \mathbf{z}_{k-1})$ : state transition
  - $p(\mathbf{y}_k | \mathbf{z}_k)$ : observation likelihood
1. What is the optimal proposal distribution  $q^{\text{opt}}(\mathbf{z}_k | \mathbf{z}_{k-1}, \mathbf{y}_k)$ ?
  2. What are the resulting weights?
  3. When is this optimal proposal tractable?
  4. Suggest practical alternatives when intractable.

**Solution:**

### 1. Optimal Proposal:

From importance sampling theory (Problem 1), the optimal proposal minimizes weight variance. For sequential problems, this is:

$$q^{\text{opt}}(\mathbf{z}_k | \mathbf{z}_{1:k-1}, \mathbf{y}_{1:k}) = p(\mathbf{z}_k | \mathbf{z}_{1:k-1}, \mathbf{y}_{1:k})$$

Using the Markov structure of the state-space model (latent state depends only on previous state, observations depend only on current state):

$$q^{\text{opt}}(\mathbf{z}_k | \mathbf{z}_{k-1}, \mathbf{y}_k) = p(\mathbf{z}_k | \mathbf{z}_{k-1}, \mathbf{y}_k)$$

This incorporates both the state dynamics and the current observation.

## 2. Optimal Weights:

Using  $\phi_k(\mathbf{z}_{1:k}) = \prod_{\ell=1}^k p(\mathbf{y}_\ell | \mathbf{z}_\ell) p(\mathbf{z}_\ell | \mathbf{z}_{\ell-1})$  with the optimal proposal:

$$\begin{aligned} w_k &\propto \frac{p(\mathbf{y}_k | \mathbf{z}_k) p(\mathbf{z}_k | \mathbf{z}_{k-1})}{p(\mathbf{z}_k | \mathbf{z}_{k-1}, \mathbf{y}_k)} \\ &= \frac{p(\mathbf{y}_k | \mathbf{z}_k) p(\mathbf{z}_k | \mathbf{z}_{k-1})}{p(\mathbf{y}_k | \mathbf{z}_k) p(\mathbf{z}_k | \mathbf{z}_{k-1}) / p(\mathbf{y}_k | \mathbf{z}_{k-1})} \\ &= p(\mathbf{y}_k | \mathbf{z}_{k-1}) \end{aligned}$$

So:  $w_k(\mathbf{z}_{1:k}) = w_{k-1}(\mathbf{z}_{1:k-1}) p(\mathbf{y}_k | \mathbf{z}_{k-1})$ .

Key insight: With the optimal proposal, weights depend only on the \*\*previous\*\* state, not the newly sampled state.

## Problem 3: Particle Filter Algorithm - Prior Proposal

For state-space model with prior proposal  $q(\mathbf{z}_k | \mathbf{z}_{k-1}, \mathbf{y}_k) = p(\mathbf{z}_k | \mathbf{z}_{k-1})$ :

### Particle Filter Algorithm (Prior Proposal):

For each particle  $i = 1, \dots, M$ :

#### Step 1: Particle Generation

$$\mathbf{z}_k^{(i)} \sim p(\mathbf{z}_k | \mathbf{z}_{k-1}^{(i)})$$

#### Step 2a: Weight Computation

$$w_k^{*(i)} = \tilde{w}_{k-1}^{(i)} \cdot p(\mathbf{y}_k | \mathbf{z}_k^{(i)})$$

#### Step 2b: Weight Normalization

$$\tilde{w}_k^{(i)} = \frac{w_k^{*(i)}}{\sum_{j=1}^M w_k^{*(j)}}$$

#### Step 3: Estimate

$$\hat{\mathbb{E}}[g(\mathbf{z}_k) | \mathbf{y}_{1:k}] \approx \sum_{i=1}^M \tilde{w}_k^{(i)} g(\mathbf{z}_k^{(i)})$$

### Solution:

#### For a 1D random walk example:

#### Model:

$$\begin{aligned} \mathbf{z}_k &= \mathbf{z}_{k-1} + \mathbf{u}_k, \quad \mathbf{u}_k \sim \mathcal{N}(0, 1) \\ \mathbf{y}_k &= \mathbf{z}_k + \mathbf{v}_k, \quad \mathbf{v}_k \sim \mathcal{N}(0, 1) \end{aligned}$$

Densities:

$$\begin{aligned} p(\mathbf{z}_k | \mathbf{z}_{k-1}) &= \mathcal{N}(\mathbf{z}_k | \mathbf{z}_{k-1}, 1) \\ p(\mathbf{y}_k | \mathbf{z}_k) &= \mathcal{N}(\mathbf{y}_k | \mathbf{z}_k, 1) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{(\mathbf{y}_k - \mathbf{z}_k)^2}{2}\right) \end{aligned}$$

### Problem 4: Effective Sample Size

The effective number of samples is:

$$N_{\text{eff}} = \frac{1}{\sum_{i=1}^M (\tilde{w}^{(i)})^2}$$

1. When is  $N_{\text{eff}} = M$  (maximum)?
2. When is  $N_{\text{eff}} = 1$  (worst)?
3. Why trigger resampling at  $N_{\text{eff}} < M/2$ ?

**Solution:**

1. **Maximum  $N_{\text{eff}} = M$ :**

Occurs when all weights are equal:  $\tilde{w}^{(i)} = 1/M$  for all  $i$ .

$$N_{\text{eff}} = \frac{1}{\sum_{i=1}^M (1/M)^2} = \frac{1}{M \cdot 1/M^2} = M$$

All particles contribute equally to the posterior estimate.

2. **Minimum  $N_{\text{eff}} = 1$ :**

Occurs when one particle has weight 1 and all others have weight 0:

$$N_{\text{eff}} = \frac{1}{1^2 + 0 + \dots + 0} = 1$$

Only one particle carries information; all others are useless.

3. **Resampling Threshold  $N_{\text{eff}} < M/2$ :**

This threshold balances two competing objectives:

- **Prevent Degeneracy:** If  $N_{\text{eff}} < M/2$ , then weight concentration is severe (too much variance). Resampling discards particles in low-probability regions.
- **Preserve Diversity:** If resampling too frequently (higher threshold), we lose particle diversity unnecessarily. The threshold  $M/2$  is empirically found to be near-optimal.
- **Adaptive:** Don't resample every step (wastes diversity). Only resample when needed.

**Why  $M/2$ :** Empirical experience shows this threshold provides best balance between tracking accuracy and computational efficiency. Too early resampling increases sample impoverishment; too late resampling allows weight degeneracy to worsen.

## 2 Practical Applications

### Problem 5: Bearings-Only Tracking

**Problem:** A passive observer (e.g., sonar) measures only bearing angles to a moving target, not range.  
**Goal:** Estimate 2D target position and velocity.

**Latent State:**  $\mathbf{z}_k = [s_{x,k}, V_{x,k}, s_{y,k}, V_{y,k}]^\top$  (position and velocity)

**Latent State Equation:**

$$\mathbf{z}_{k+1} = F\mathbf{z}_k + \mathbf{u}_k$$

where:

$$F = \begin{bmatrix} 1 & \Delta t & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & \Delta t \\ 0 & 0 & 0 & 1 \end{bmatrix}$$

and  $\mathbf{u}_k \sim \mathcal{N}(0, Q)$  (process noise accounting for unmodeled acceleration, wind, etc.)

#### Observation Equation:

$$\mathbf{y}_k = f_y(\mathbf{z}_k, \mathbf{v}_k) = \arctan\left(\frac{s_{y,k}}{s_{x,k}}\right) + \mathbf{v}_k$$

where  $\mathbf{v}_k \sim \mathcal{N}(0, R)$  is measurement noise (bearing uncertainty).

Observation likelihood:

$$p(\mathbf{y}_k | \mathbf{z}_k) = \mathcal{N}\left(\mathbf{y}_k \mid \arctan(s_{y,k}/s_{x,k}), R\right)$$

#### Why Kalman Filter Fails:

- **Nonlinear Observation:** The function  $f_y(\mathbf{z}_k) = \arctan(s_{y,k}/s_{x,k})$  is highly nonlinear (not a linear function of  $\mathbf{z}_k$ )
- **Gaussian Assumption Violated:** If we pass a Gaussian distribution on position through arctan, the resulting distribution on bearing is NOT Gaussian
- **Standard Kalman Limitation:** Kalman filter assumes:
  - Linear dynamics:  $\mathbf{z}_{k+1} = F\mathbf{z}_k + \mathbf{u}_k$  (satisfied)
  - Linear observations:  $\mathbf{y}_k = H\mathbf{z}_k + \mathbf{v}_k$  (violated due to arctan)
  - Gaussian noise: Both  $\mathbf{u}_k, \mathbf{v}_k$  Gaussian (satisfied)
- **Posterior Non-Gaussian:** Even starting with Gaussian prior on position, the posterior given bearing measurements is non-Gaussian. Kalman's Gaussian updates are suboptimal.

#### Particle Filter Solution:

Apply the algorithm from Problem 8:

1. **Initialization:** Sample  $M = 100$  particles:  $\mathbf{z}_0^{(i)} \sim \mathcal{N}(\mathbf{z}_0^{\text{prior}}, P_0)$
2. **For each time step  $k$ :**
  - (a) **Prediction:**  $\mathbf{z}_k^{(i)} = F\mathbf{z}_{k-1}^{(i)} + \mathbf{u}_k^{(i)}$  where  $\mathbf{u}_k^{(i)} \sim \mathcal{N}(0, Q)$
  - (b) **Weight:**  $w_k^{*(i)} = \tilde{w}_{k-1}^{(i)} \cdot \mathcal{N}(\mathbf{y}_k | \arctan(s_{y,k}^{(i)}/s_{x,k}^{(i)}), R)$
  - (c) **Normalize:**  $\tilde{w}_k^{(i)} = w_k^{*(i)} / \sum_j w_k^{*(j)}$
  - (d) **Estimate:**  $\hat{\mathbf{z}}_k = \sum_i \tilde{w}_k^{(i)} \mathbf{z}_k^{(i)}$
  - (e) **Resample:** If  $N_{\text{eff}} < 50$ , use Algorithm 4 (systematic resampling)

#### Advantages:

- Handles nonlinearity exactly (no linearization approximation)
- Posterior naturally non-Gaussian (represents multimodal uncertainty)
- No divergence issues (unlike EKF)
- Can estimate full posterior, not just mean and covariance

## Problem 6: Visual Object Tracking with Non-Gaussian Noise

**Latent State Equation:**

$$\mathbf{z}_k = \mathbf{z}_{k-1} + \mathbf{u}_k, \quad \mathbf{u}_k \sim \mathcal{U}([-10, 10]^2)$$

where  $\mathbf{z}_k = [x_k, y_k]^\top$  is 2D object position.

**Observation Equation:**

$$\mathbf{y}_k = \mathbf{z}_k + \mathbf{v}_k, \quad \mathbf{v}_k \sim \mathcal{N}(0, \Sigma_v), \quad \Sigma_v = \begin{bmatrix} 2 & 0.5 \\ 0.5 & 2 \end{bmatrix}$$

**Why Kalman Filter Suboptimal:**

Kalman filter is optimal only for LINEAR dynamics with GAUSSIAN noise on both state and observation equations:

- **Linear dynamics:**  $\mathbf{z}_{k+1} = F\mathbf{z}_k + \mathbf{u}_k$  (satisfied)
- **Linear observations:**  $\mathbf{y}_k = H\mathbf{z}_k + \mathbf{v}_k$  (satisfied with  $H = I$ )
- **Gaussian process noise:**  $\mathbf{u}_k \sim \mathcal{N}(0, Q)$  (VIOLATED:  $\mathbf{u}_k$  is uniform!)
- **Gaussian measurement noise:**  $\mathbf{v}_k \sim \mathcal{N}(0, R)$  (satisfied)

Under uniform process noise, the posterior  $p(\mathbf{z}_k | \mathbf{y}_{1:k})$  is NOT Gaussian. Kalman's assumption of Gaussian distributions breaks down, making it suboptimal.

**Particle Filter Design:**

Use prior proposal from Problem 8:

1. **Particle Generation:** Sample from uniform distribution:

$$\mathbf{u}_k^{(i)} \sim \mathcal{U}([-10, 10]^2), \quad \mathbf{z}_k^{(i)} = \mathbf{z}_{k-1}^{(i)} + \mathbf{u}_k^{(i)}$$

2. **Weight Update:**

$$w_k^{*(i)} = \tilde{w}_{k-1}^{(i)} \cdot \mathcal{N}(\mathbf{y}_k | \mathbf{z}_k^{(i)}, \Sigma_v)$$

where the Gaussian likelihood is:

$$\mathcal{N}(\mathbf{y}_k | \mathbf{z}_k^{(i)}, \Sigma_v) = \frac{1}{2\pi\sqrt{\det(\Sigma_v)}} \exp\left(-\frac{1}{2}(\mathbf{y}_k - \mathbf{z}_k^{(i)})^\top \Sigma_v^{-1} (\mathbf{y}_k - \mathbf{z}_k^{(i)})\right)$$

3. **Normalization and Resampling:** Standard steps from Problem 8

**Comparison: With vs Without Resampling**

- **With resampling:**

- Fewer particles have significant weight (concentrated on likely region)
- Computational efficiency: fewer useless particles
- Better tracking quality over long horizons
- Trade-off: some loss of diversity

- **Without resampling:**

- More particles retain nonzero weight initially
- Wasted computation on unlikely particles
- Tracking degrades as time increases (weight degeneracy)
- Advantage: preserves full diversity of prior
- Not practical for long sequences

**Recommendation:** Always use resampling with threshold  $N_{\text{eff}} < M/2$ .

## Problem 7: Degeneracy vs Sample Impoverishment

### Definitions:

#### 1. Weight Degeneracy:

As time progresses, the weight distribution becomes increasingly skewed. After many steps:

- Few particles (e.g., 1-5 out of  $M = 100$ ) accumulate nearly all weight
- Most particles have negligible weight ( $\tilde{w}^{(i)} < 0.01$ )
- Computation is wasted on irrelevant particles
- Posterior approximation quality suffers (small effective sample size)

**Cause:** Weights are products of likelihoods; those are typically  $< 1$ , so  $w_k \approx L^k$  decays exponentially.

#### 2. Sample Impoverishment:

After resampling, the particle population loses diversity:

- High-weight particles are replicated many times
- Posterior approximation becomes coarse (many identical particles)
- Effective sample size drops due to redundancy
- Affects future predictions: correlated particles lead to poor uncertainty estimates

**Cause:** Resampling concentrates particles on high-probability regions; new diversity only emerges through noise in state evolution.

### Mechanism of Impoverishment:

Suppose at time  $k - 1$  one particle has weight 0.9 and we resample with  $M = 100$ :

- This particle is drawn  $\approx 90$  times
- Others divided among remaining 10 draws
- At time  $k$ : population has  $\sim 90$  copies of one particle, plus scattered copies of others
- If process noise is small, all 90 copies evolve nearly identically
- Posterior  $p(\mathbf{z}_k | \mathbf{y}_{1:k})$  becomes concentrated at discrete points (not smooth)
- For estimation, we effectively have  $\ll 100$  independent samples

### When Severe:

Sample impoverishment is worst when \*\*process noise is small\*\* (or near-zero):

- Small  $Q$  means state evolution is nearly deterministic
- Resampled particles don't diverge during prediction
- Diversity doesn't recover
- Over multiple steps, population collapses to few distinct values

**Example:**  $\mathbf{z}_k = \mathbf{z}_{k-1} + \mathbf{u}_k$  with  $\mathbf{u}_k \sim \mathcal{N}(0, 0.001)$  (tiny noise)

After resampling once, 90 copies of a single particle stay nearly identical through many steps.

### Solutions:

- **Regularized Particle Filters:** Instead of resampling to discrete points, approximate posterior with smooth kernels:

$$p(\mathbf{z}|\mathbf{y}) \approx \sum_{i=1}^M \tilde{w}^{(i)} K(\mathbf{z} - \mathbf{z}^{(i)})$$

where  $K$  is a smooth kernel (e.g., Gaussian). Maintains continuous support.

- **Better Proposal:** Using observation-dependent proposal (Problem 7) places particles in likely regions without pure resampling-induced collapse.
- **Jittering:** Add small random noise to resampled particles to encourage diversity.
- **Increase Process Noise Estimate:** If you have control over  $Q$ , larger  $Q$  helps particles diverge after resampling.