CAP 5610: Machine Learning

Lecture 5:

Regression:

Logistic and Linear

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Bayes Classifier: A Generative model

- Model the posterior distribution P(Y|X)
 - Estimate class-conditional distribution P(X|Y) for each Y
 - Estimate prior distribution P(Y)
 - Predict the target value Y by Maximum A Posteriori (MAP)

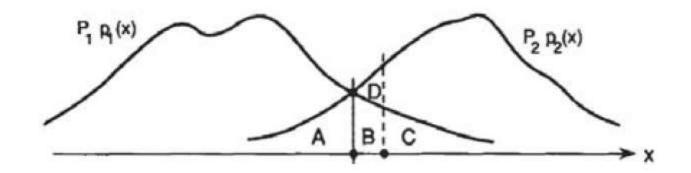
$$Y^* = \operatorname{argmax}_{Y} P(Y|X) = \operatorname{argmax}_{Y} P(X|Y) P(Y)$$

- It is a generative model, because
 - Given each class Y, we can draw (generate) a sample from P(Y|X)

Bayes Classifier is the optimal classifier

 Any classifier can never yield smaller prediction error than Bayes classifier

$$error_{true}(h_{Bayes})) \le error_{true}(h), \ \forall h(\mathbf{x})$$



Drawback of a generative model

- It is an indirect method
 - We need to estimate P(Y) and P(X|Y) from the training set at first
 - For a classification problem, we do not need any generative information contained in P(X|Y)
 - Unnecessary assumptions are usually needed to estimate P(X|Y)
 - Naive Bayes independence between features conditional on the label Y
- How about directly learn P(Y|X)?

Discriminative model

- Directly model the posterior distribution P(Y|X)
 - Assume some functional forms for P(Y|X), e.g., a linear model
 - Estimate the parameter of the proposed model directly from training set,
 - Option 1: Maximum likelihood estimation (MLE)
 - Option 2: Maximum A Posteriori (MAP)

Hints from generative model

- Suppose
 - X is a vector of continuous features: X=(X₁, ..., X_N)
 - Y is a binary random variable {0,1}
 - Naive Bayesian classifier: Xi's are independent given Y
 - P(Y) is a Bernoulli distribution
- Find the form for P(Y=1|X), by Bayes rule

$$P(Y = 1|X) = \frac{P(Y = 1)P(X|Y = 1)}{P(Y = 1)P(X|Y = 1) + P(Y = 0)P(X|Y = 0)}$$

Find P(Y=1|X)

$$\begin{split} P(Y=1|X) &= \frac{P(Y=1)P(X|Y=1)}{P(Y=1)P(X|Y=1) + P(Y=0)P(X|Y=0)} \quad \text{Bayes rule} \\ &= \frac{1}{1 + \frac{P(Y=0)P(X|Y=0)}{P(Y=1)P(X|Y=1)}} \\ &= \frac{1}{1 + \exp(\ln\frac{P(Y=0)P(X|Y=0)}{P(Y=1)P(X|Y=1)})} \quad \text{exp and ln will cancel out.} \\ &= \frac{1}{1 + \exp(\ln\frac{P(Y=0)}{P(Y=1)} + \sum_{i=1}^{N} \ln\frac{P(X_i|Y=0)}{P(X_i|Y=1)})} \quad \text{Assume } P(X_i|Y) \text{ are Gaussian with the same variance but different mean} \\ &= \frac{1}{1 + \exp(\ln\frac{P(Y=0)}{P(Y=1)} + \sum_{i=1}^{N} \ln\frac{N(X_i;\mu_{i0},\sigma_i^2)}{N(X_i;\mu_{i1},\sigma_i^2)})} = \frac{1}{1 + \exp(\ln\frac{P(Y=0)}{P(Y=1)} + \sum_{i=1}^{N} \left(\frac{\mu_{i0} - \mu_{i1}}{\sigma_i^2} X_i + \frac{\mu_{i1}^2 - \mu_{i0}^2}{2\sigma_i^2}\right))} \end{split}$$

Find P(Y|X=1)

Logistic function

$$P(Y = 1 \mid X) = \frac{1}{1 + \exp(\sum_{i=1}^{N} \frac{\mu_{i0} - \mu_{i1}}{\sigma_{i}^{2}} X_{i} + \underbrace{\frac{\mu_{i1}^{2} - \mu_{i0}^{2}}{2\sigma_{i}^{2}} + \ln \frac{P(Y = 0)}{P(Y = 1)}}_{\omega_{0}})}$$

$$= \frac{1}{1 + \exp(\sum_{i=1}^{N} \omega_{i} X_{i} + \omega_{0})}$$

Linear classifier

• When Y=0

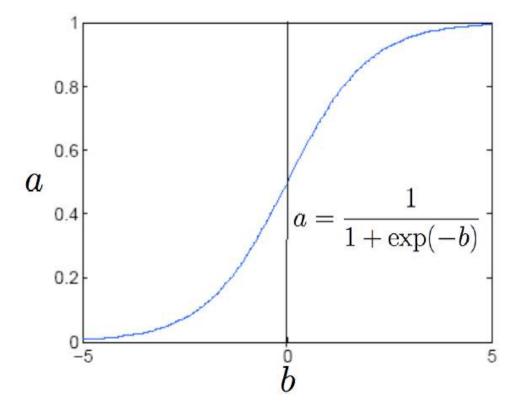
$$P(Y = 0 \mid X) = 1 - P(Y = 1 \mid X) = \frac{\exp(\sum_{i=1}^{N} \omega_i X_i + \omega_0)}{1 + \exp(\sum_{i=1}^{N} \omega_i X_i + \omega_0)}$$

Log likelihood ratio: linear in X and w

$$\log \frac{P(Y=1 \mid X)}{P(Y=0 \mid X)} = \log \frac{1}{\exp(\sum_{i=1}^{N} \omega_i X_i + \omega_0)} = -\sum_{i=1}^{N} \omega_i X_i + \omega_0$$

Logistic function

- Smooth in b: differentiable and continuously differentiable
- The logarithm of logistic function is concave.



Extend to multiple classes

- Y={0,1,...,C}, e.g., C=10 in MNIST dataset
- Assume $P(Y = c \mid X) \propto \exp(\sum_{i=1}^{N} w_{ci} X_i + w_{c0})$
- Normalizing P(Y|X) such that the summation over C classes is one.
 - Softmax: the maximum exponential function (above) for class c will have the largest posterior probability

$$P(Y = c \mid X) = \frac{\exp(\sum_{i=1}^{N} w_{ci} X_i + w_{c0})}{\sum_{c'=1}^{C} \exp(\sum_{i=1}^{N} w_{c'i} X_i + w_{c'0})}$$

• Binary logistic regression is a special case when C=2, and $w_{Ii}=0$

Learning logistic regression model

Given a set of training example {(X^(m),Y^(m)| m=1,..., M}, MLE corresponds to

$$\begin{split} L(\mathbf{w}) &= \sum_{m=1}^{M} \log P(\mathbf{Y}^{(m)} \mid \mathbf{X}^{(m)}, \mathbf{w}) \\ &= \sum_{m=1}^{M} \mathbf{Y}^{(m)} \log P(\mathbf{Y}^{(m)} = 1 \mid \mathbf{X}^{(m)}, \mathbf{w}) + (1 - \mathbf{Y}^{(m)}) \log P(\mathbf{Y}^{(m)} = 0 \mid \mathbf{X}^{(m)}, \mathbf{w}) \\ &= \sum_{m=1}^{M} \left\{ \mathbf{Y}^{(m)} \log \frac{P(\mathbf{Y}^{(m)} = 1 \mid \mathbf{X}^{(m)}, \mathbf{w})}{P(\mathbf{Y}^{(m)} = 0 \mid \mathbf{X}^{(m)}, \mathbf{w})} + \log P(\mathbf{Y}^{(m)} = 0 \mid \mathbf{X}^{(m)}, \mathbf{w}) \right\} \\ &= -\sum_{m=1}^{M} \mathbf{Y}^{(m)} \left(\sum_{i=1}^{N} w_{i} X_{i}^{(m)} + w_{0} \right) + \log \left\{ 1 + \exp(-\sum_{i=1}^{N} w_{i} X_{i}^{(m)} - w_{0}) \right\} \end{split}$$

LR Optimization problem

The optimal solution to w is

$$\mathbf{w}^* = \arg\max_{w} -\sum_{m=1}^{M} \mathbf{Y}^{(m)} \left(\sum_{i=1}^{N} w_i X_i^{(m)} + w_0 \right) + \log \left\{ 1 + \exp(-\sum_{i=1}^{N} w_i X_i^{(m)} - w_0) \right\}$$

20 15 0 2 1 1 0 2

- The objective function is concave.
 - Logistic function is concave in its logarithmic form.
 - It has global optimal point, avoiding local optimum.
 - Unfortunately, LR optimization problem does not have closed-form solution.

Gradient Ascent Method

- Gradient descent method is an iterative algorithm
 - hill climbing method to find the peak point of a "mountain"
 - At each point, compute its gradient

$$\nabla L = \left[\frac{\partial L}{\partial w_0}, \frac{\partial L}{\partial w_1}, ..., \frac{\partial L}{\partial w_N} \right]$$

- Gradient is a vector that points to the steepest direction climbing up the mountain.
- At each point, w is updated so it moves a size of step λ in the gradient direction

$$w \leftarrow w + \lambda \nabla L(w)$$

Gradient Ascent Method

• LR Log likelihood:

$$L(W) = -\sum_{m=1}^{M} \mathbf{Y}^{(m)} \left(\sum_{i=1}^{N} w_i X_i^{(m)} + w_0 \right) + \log \left\{ 1 + \exp(-\sum_{i=1}^{N} w_i X_i^{(m)} - w_0) \right\}$$

• Derivative to each weight w_i

$$\frac{\partial L}{w_0} = -\sum_{m=1}^{M} \mathbf{Y}^{(m)} - P(Y = 1 \mid X^{(m)})$$

$$\frac{\partial L}{w_i} = -\sum_{m=1}^{M} \mathbf{Y}^{(m)} X_i^{(m)} - X_i^{(m)} P(Y = 1 \mid X^{(m)})$$

Update rule

$$w_0 \leftarrow w_0 - \lambda (\sum_{m=1}^{M} \mathbf{Y}^{(m)} + P(Y = 1 \mid X^{(m)})) \qquad \qquad w_i \leftarrow w_i - \lambda (\sum_{m=1}^{M} \mathbf{Y}^{(m)} X_i^{(m)} + X_i^{(m)} P(Y = 1 \mid X^{(m)}))$$

Training algorithm

- Input: a set of training example {(X^(m),Y^(m)|m=1,..., M}
- Initialize $w=[w_0, w_1, ..., w_N]$
- Repeat

$$w_0 \leftarrow w_0 - \lambda (\sum_{m=1}^{M} \mathbf{Y}^{(m)} + P(Y = 1 \mid X^{(m)})) \qquad \qquad w_i \leftarrow w_i - \lambda (\sum_{m=1}^{M} \mathbf{Y}^{(m)} X_i^{(m)} + X_i^{(m)} P(Y = 1 \mid X^{(m)}))$$

- (optional) check if the loglikelihood increases after each update. If not, shrink the step size λ
- Until convergence

Training algorithm

- Input: a set of training example {(X^(m),Y^(m)|m=1,..., M}
- Initialize $w=[w_0, w_1, ..., w_N]$
- Repeat

$$w_0 \leftarrow w_0 - \lambda (\sum_{m=1}^{M} \mathbf{Y}^{(m)} + P(Y = 1 \mid X^{(m)})) \qquad \qquad w_i \leftarrow w_i - \lambda (\sum_{m=1}^{M} \mathbf{Y}^{(m)} X_i^{(m)} + X_i^{(m)} P(Y = 1 \mid X^{(m)}))$$

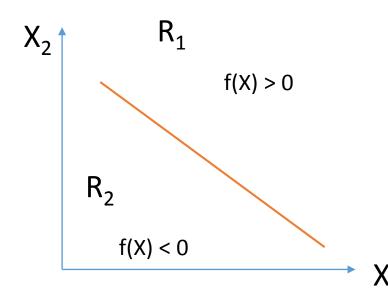
- (optional) check if the log likelihood increases after each update. If not, shrink the step size λ
- Until convergence

Test algorithm

Computing the log likelihood ratio

$$f(X) = \log \frac{P(Y = 1 \mid X)}{P(Y = 0 \mid X)} = \log \frac{1}{\exp(\sum_{i=1}^{N} \omega_i X_i + \omega_0)} = -\sum_{i=1}^{N} \omega_i X_i + \omega_0$$

- If >0, X is positive (Y=1)
- If <0, X is negative (Y=0)



Linear decision bounda

Stochastic Gradient Ascent Method

- Making the learning algorithm scalable to big data
- Typical gradient ascent method
 - Each update needs to go over all M examples to compute the derivatives

$$\frac{\partial L}{w_0} = -\sum_{m=1}^{M} \mathbf{Y}^{(m)} - P(Y = 1 \mid X^{(m)})$$

$$\frac{\partial L}{w_i} = -\sum_{m=1}^{M} \mathbf{Y}^{(m)} X_i^{(m)} - X_i^{(m)} P(Y = 1 \mid X^{(m)})$$

 It will save us a huge amount of computations if only a small portion of examples are used

Stochastic Gradient Ascent Method

ullet A random subset Ω of examples are drawn from training set, which are used to compute derivatives

$$\frac{\partial L}{w_0} = -\sum_{m=1}^{M} \mathbf{Y}^{(m)} - P(Y = 1 \mid X^{(m)})$$

$$\frac{\partial L}{w_0} = -\sum_{n=0}^{M} \mathbf{Y}^{(m)} - P(Y = 1 \mid X^{(m)})$$

$$\frac{\partial L}{w_0} = -\sum_{m=1}^{M} \mathbf{Y}^{(m)} - P(Y = 1 \mid X^{(m)})$$

$$\frac{\partial L}{w_i} = -\sum_{m=1}^{M} \mathbf{Y}^{(m)} X_i^{(m)} - X_i^{(m)} P(Y = 1 \mid X^{(m)})$$



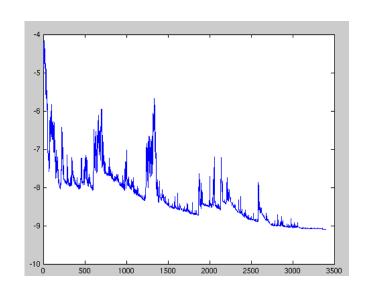
$$\frac{\partial L}{w_0} = -\sum_{m \in \Omega} Y^{(m)} - P(Y = 1 \mid X^{(m)}) \qquad \frac{\partial L}{w_i} = -\sum_{m \in \Omega} Y^{(m)} X_i^{(m)} - X_i^{(m)} P(Y = 1 \mid X^{(m)})$$

 Especially, only one example is drawn from training set to compute the derivative

Stochastic Gradient Ascent Method

- Pros: Each round of update only needs very small number of examples
 - Time complexity: More rapid update
 - Space complexity: Only upload the examples drawn in each round to the memory
- Cons: Cannot guarantee convergence

Practice: very efficient



Option 2: Maximum A Posteriori (MAP)

- Add a prior distribution on w
 - $P(w) = N(0, \sigma^2 I)$
- Posterior distribution on w can be given by

$$\log P(w | D) \propto \log P(w) + \sum_{m=1}^{M} \log P(Y^{(m)} | X^{(m)}, w)$$

$$= \log P(w) + L(w)$$

The optimal solution

$$w^* = \operatorname{argmax}_{w} L(w) + \log P(w) = \operatorname{argmax}_{w} L(w) - \frac{1}{2\sigma^2} ||w||_2^2$$

L₂ regularizer

• L₂ Regularized loglikelihood objective function

$$w^* = \operatorname{argmax}_{w} L(w) + \log P(w) = \operatorname{argmax}_{w} L(w) - \frac{1}{2\sigma^2} ||w||_2^2$$

- Role of a regularizer
 - Reduce the overfitting, by imposing a prior knowledge on w
 - keeping the weights closer to zero

Philosophy behind L₂ regularizer

- Why should we prefer a zero weight w in LR?
 - Maximum entropy principle: do NOT make a bias towards one of two outcomes unless you have enough information (least bias).
 - A coin flipping: how do you predict the outcome if you do not know any attributes about the coin?
- When w = 0, P(Y=1|X)=P(Y=0|X)=0.5 (equal probability)

$$P(Y = 1 | X) = \frac{1}{1 + \exp(\sum_{i=1}^{N} \omega_i X_i + \omega_0)}$$

Comparison between Naive Bayes and LR

Number of model parameters

- Naive Bayes: for each attribute, two parameters (mean and variance), for two classes; two class prior distribution – 4N+2
- LR: N weight coefficients
- More parameters usually mean more training examples to estimate them.

• Assumptions:

- Naive Bayes: conditional independence given the class label
- LR: no such independence assumption (why?)
- With weaker assumption, LR may outperform Naive Bayes if the assumption does not hold.

Summary: Logistic Regression

- Derive LR from Naive Bayes with Gaussian class-conditional distribution
- MLE: Estimate the coefficient weights in LR
 - Gradient ascent method
 - Stochastic gradient ascent method
- MAP: add a prior on w
 - A L₂ regularizer

Linear vs. Logistic Regression

- Logistic regression is used when the dependent variable is categorical to establish the probability of the event.
 - Logistic loss function causes errors to be penalized to an asymptotic constant.
- Linear regression is used when the dependent variable is continuous.
 - Solved by minimizing least squares error so large errors are penalized quadratically.
- Reading: PRML Chapter 3

What's a regression problem?

- Given an input vector X, predict the target value associated with X.
 - Input is size of a house, target variable is its price.
 - Predict the stock prices based their history performances.
 - Predict the person's age from the face image.



Regression

- Learn a predictor $f: X \mapsto Y$ where X is a feature vector and Y is the target variable
- Outline (compared with discrete case)
 - Design an optimal predictor (Bayes classifier)
 - Linear model based on Gaussian assumption
 - Estimate the parameters of linear model
 - MLE
 - MAP
 - Bias-variance decomposition

Optimal regression model

- Given a joint distribution over (X,Y), what's the optimal predictor?
 - Minimizing the expected least-square loss wrt P(X,Y)

$$E|Y - f(X)|^2 = \iint_{Y,X} |Y - f(X)|^2 P(X,Y) dX dY$$

• Compared with classification problem, where we minimize the expected prediction errors.

Optimal regression model

 Variational optimization method that minimizes wrt a function f(X) rather than a single variable (Appendix D in PRML)

$$E|Y - f(X)|^2 = \iint_{Y,X} |Y - f(X)|^2 P(X,Y) dX dY$$

Set the differential wrt f(X) to 0

$$\frac{\partial E|Y - f(X)|^2}{\partial f(X)} = 2\int (f(X) - Y) P(X, Y) dY = 0$$
$$f(X)P(X) - P(X)E(Y|X) = 0$$
$$f(X) = E(Y|X) = \int YP(Y|X) dY$$

Optimal regression model

Optimal regression model is the expectation of Y conditional on the input vector X

$$f(X) = E(Y|X) = \int YP(Y|X) \, dY$$

Compared with optimal classification model (Bayes Classifier)

$$Y = \operatorname{argmax}_{Y} P(Y|X)$$

Practical Issue

• We do not know the true posterior distribution P(Y|X), just as in the classification problem

 How about using multivariate Gaussian distribution to estimate the joint distribution P(X,Y)?

The optimal Gaussian regression model

Joint distribution

$$P(X,Y) = N(\boldsymbol{m}_X, m_Y; C)$$

- Conditional distribution of Y given X
- An important property: the conditional distribution P(Y|X) from a multivariate Gaussian distribution is still Gaussian.

$$P(Y|X) = N(m_Y + C_{YX}C_{XX}^{-1}(X - mX), C_{Y|X})$$

Optimal predictor

$$E(Y|X) = mY + C_{YX}C_{XX}^{-1}(X - mX)$$

- This is a linear predictor, although we do not know these parameters.
- That's why we often assume data are generated by Gaussian distribution, because the regression model is very simple.

Drawbacks

- It is an indirect method
 - We have to estimate the joint distribution before finding the optimal predictor (then computing expected Y conditional on X).
 - Too many parameters to estimate
 - For N attributes in X and target variable Y, there are (N+1)(N+2)/2 parameters in covariance matrix, along with N+1 parameters for the mean of X and Y.
 - We do not care about the covariance and mean of the attributes;
 - We only need to know the covariance **between** X and Y (N parameters) and the mean of Y (1 parameter): N+1 parameters totally that are of true interest to the regression problem.
 - A direct method, please!

Discriminative method

- Directly estimate P(Y|X) with a specified form, e.g., Gaussian again. $P(Y|X) = N(f(X), \sigma^2)$
- Assume f(X) is linear in X

$$f(X) = \sum_{i=1}^{N} W_i X_i + W_0 = W^T X$$

where we denote by X an extended vector with an additional entry $X=[1, X_1, X_2, ..., X_N]$, $W=[W_0, W_1, W_2, ..., W_N]$.

- It is discriminative since we only estimate the posterior distribution
 - No generating model will be estimated from which we can draw X given a Y.

Two Methods

- Option 1: MLE (Maximum Likelihood Estimation)
- Option 2: MAP (Maximum A Posteriori)

MLE

• Maximizing the log likelihood on a training set $\{(X^l,Y^l) \mid l=1,...,L\}$

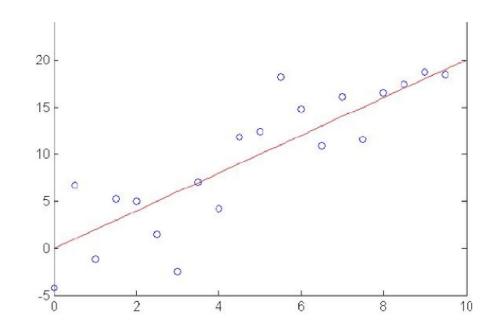
$$W^* = \operatorname{argmax}_W \sum_{l=1}^{L} \log P(Y^l | X^l; W)$$

Where
$$P(Y|X) = N(f(X), \sigma^2) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left\{-\frac{(Y-f(X))^2}{2\sigma^2}\right\}$$
.

MLE

The objective is equivalent to

$$W^* = \operatorname{argmax}_W \sum_{l=1}^{L} (Y - f(X))^2$$



where f(X) is a linear function, so it finds a line in hyperspace which best fits the training data, i.e., linear regression.

Solving linear regression problem

Take derivative of each square term to W, by a chain rule

$$\frac{\partial \sum_{l} (y - f(x; W))^{2}}{\partial w_{i}} = \sum_{l} 2(y - f(x; W)) \frac{\partial (y - f(x; W))}{\partial w_{i}}$$
$$= \sum_{l} -2(y - f(x; W)) \frac{\partial f(x; W)}{\partial w_{i}}$$

Solving linear regression problem

Set the derivative to zero,

$$\sum_{l=1}^{L} (Y^l - W^T X^l) X^l = 0$$

An alternative matrix representation

$$\mathbf{X}\mathbf{Y} - \mathbf{X}\mathbf{X}^{\mathrm{T}}\mathbf{W} = 0$$

Where X is data matrix, with each column being a training example, Y is the vector of target variables.

$$\mathbf{X} = [X^1, X^2, ..., XL], \mathbf{Y} = [Y^1, Y^2, ..., YL]^T$$

Pseudo Inverse

• Solving $\mathbf{X}\mathbf{Y} - \mathbf{X}\mathbf{X}^{\mathrm{T}}\mathbf{W} = 0$ $W = (\mathbf{X}\mathbf{X}^{T})^{-1}\mathbf{X}\mathbf{Y}$

Where $X^+ = (XX^T)^{-1}X$ is called pseudo inverse. When X is invertible, then it is the inverse of the matrix.

 Compared with logistic regression, we have closed-form solution to the linear regression problem.

(pseudo inverse can be calculated for a rectangular matrix)

MAP solution

• Add a prior distribution on W, e.g., Gaussian.

$$P(w) = N(0, \sigma^2 I)$$

Maximizing the posterior distribution on W over the training set D

$$egin{aligned} \log P(w|D) & \propto log P(w) + \sum_{m=1}^M \log(P^l|X^l,w) \ &= \log P(w) + L(w) \end{aligned}$$

The optimal solution

$$w^* = rg \max_w L(W) + \log P(W) = rg \max_w \sum_{l=1}^L (Y^l - W^T X^l)^2 + \gamma ||w||_2^2$$

MAP solution

Closed-form solution

$$W = (XX^T + \gamma I)^{-1}XY$$

- Improved performance
 - avoiding overfitting with prior information about W
- Improved numerical stability
 - avoiding ill-conditioned inverse matrix

Bias-Variance Decomposition

- Where does a prediction error come?
 - Suppose (X,Y) is drawn from a joint (unknown true) distribution
 - Perfect predictor is the conditional expectation E(Y|X)

$$E|Y - f(X)|^2 = \iint_{Y,X} |Y - f(X)|^2 P(X,Y) dX dY$$

$$= \iint_{Y,X} |Y - E(Y|X)|^2 P(X,Y) dX dY + \iint_{Y,X} |f(X) - E(Y|X)|^2 P(X) dX dY$$

- The first term is unavoidable, causing by the uncertainty of data.
 - Given a X, its true variable is uncertain.
 - E.g., for a house, its "true" price is bargainable

Bias-Variance Decomposition

• Suppose a training set D, and f(X;D) is learned from this set.

$$\iint_{Y,X} |f(X) - E(Y|X)|^{2} P(X) dX dY$$

$$= \int \{ E_{D}[f(X;D)] - E(Y|X) \}^{2} P(X) dX + \int E_{D}\{f(X;D) - E_{D}[f(X;D)] \}^{2} P(X) dX$$

where $E_D[f(X;D)]$ is an average predictor over training sets, e.g., cross-validation.

- The first term is called bias: the deviation of the average predictor from the optimal one.
- The second term is called variance: how stable the learned predictors are over different training sets?

What we learn from bias-variance decomposition?

- A good predictor shall have
 - Small bias: be as close to the optimal predictor E[Y|X] as possible
 - Small variance: Be stable to the change of training set.
 - No matter how the training examples are drawn, the learned predictors shall not change too much.
- Unfortunately the two goals are not consistent
 - Flexible model usually can yield smaller bias but larger variance
 - Smaller γ in MAP leads to flexible model
 - Flexible enough to capture the subtle changes in training set
 - Rigid model usually has larger bias but smaller variance
 - Larger γ in MAP leads to rigid model
 - Stronger prior causes the model less flexible

More examples: KNN

- KNN becomes more rigid when K increases
 - As K increases, the neighborhood of a test example becomes larger and the prediction result will become less sensitive to the change of few neighbors.
- When K is small, KNN can be better adapted to the training set
 - A small change in training set will affect the predictions made by KNN
 - In this case, KNN is more flexible
- Making trade-off between larger K (rigid) and smaller K (flexible).

Summary: Linear Regession

- Optimal predictor in expected least square error sense
- Gaussian assumption leads to optimal linear predictor
- Learning linear regression model
 - MLE
 - MAP
- Bias-variance Decomposition

End of Fundamentals

Now you know all the fundamentals for machine learning! Reading

- Chap 1-4 in Bishop (check the appendices for useful info)
- Chap 1-8 in Murphy
- Chap 1 in Marsland

Next lecture will start talking about specific classifiers: kernel methods (support vector machines) and neural networks