

SUBTASK - 3

Given: For Gaussian discriminant analysis, joint probability distribution is given by:

$$p(y) = \begin{cases} \phi & \text{if } y=1 \\ 1-\phi & \text{if } y=0 \end{cases}$$

$$p(x|y=0) = \frac{1}{(2\pi)^{n/2} |\Sigma|^{1/2}} \exp\left(-\frac{1}{2}(x-\mu_0)^T \Sigma^{-1}(x-\mu_0)\right)$$

$$p(x|y=1) = \frac{1}{(2\pi)^{n/2} |\Sigma|^{1/2}} \exp\left(-\frac{1}{2}(x-\mu_1)^T \Sigma^{-1}(x-\mu_1)\right)$$

where ϕ , μ_0 , μ_1 and Σ are parameters of the model.

T.P.T: $p(y=1|x; \phi, \mu_0, \mu_1, \Sigma) = \frac{1}{1 + \exp(-(\theta^T x + \theta_0))}$

where $\theta \in \mathbb{R}^n$ and $\theta_0 \in \mathbb{R}$.

Proof:

By Bayes' Rule, which states that,

if B_1, B_2, \dots, B_n partition the sample space S and if A is an event with $P(A) > 0$, then for $j = 1, 2, \dots, n$, we have,

$$P(B_j|A) = \frac{P(A|B_j) P(B_j)}{P(A)}$$

Applying Bayes' Rule, we get that,

$$p(y=1|x; \phi, \mu_0, \mu_1, \Sigma) = \frac{p(x|y=1) p(y=1)}{p(x)}$$

Note: Omitting formal notation for now, eg $p(x|y=1; \phi, \mu_0, \mu_1, \Sigma)$

From law of total probability,

(using the same notation), $P(A) = P(A|B_1) P(B_1) + \dots + P(A|B_n) P(B_n)$

$$\therefore p(x) = p(x|y=1) \cdot p(y=1) + p(x|y=0) p(y=0)$$

Substituting in earlier eqⁿ, we get,

$$P(y=1/x; \phi, \mu_0, \mu_1, \Sigma) = \frac{P(x|y=1) \cdot P(y=1)}{P(x|y=1) P(y=1) + P(x|y=0) P(y=0)}$$

$$= \frac{1}{1 + \frac{P(x|y=0) P(y=0)}{P(x|y=1) P(y=1)}}$$

$$\frac{P(x|y=0) P(y=0)}{P(x|y=1) P(y=1)} = \text{substituting formulae given} = \left(\frac{1-\phi}{\phi} \right) \frac{\exp\left(-\frac{1}{2}(x-\mu_0)^T \Sigma^{-1}(x-\mu_0)\right)}{\exp\left(-\frac{1}{2}(x-\mu_1)^T \Sigma^{-1}(x-\mu_1)\right)}$$

$$= \left(\frac{1-\phi}{\phi} \right) \exp\left(-\frac{1}{2}(x-\mu_0)^T \Sigma^{-1}(x-\mu_0) + \frac{1}{2}(x-\mu_1)^T \Sigma^{-1}(x-\mu_1)\right)$$

Note: $(x-\mu_0)^T \Sigma^{-1}(x-\mu_0) = \begin{matrix} (1 \times n) & (n \times n) & (n \times 1) \end{matrix}$ \rightarrow let this be (1×1) dimensional = scalar

As it is a scalar, it can also be expressed as $\frac{(x-\mu_0)^2}{\Sigma}$.

$$\frac{P(x|y=0) P(y=0)}{P(x|y=1) P(y=1)} = \left(\frac{1-\phi}{\phi} \right) \cdot \exp\left(-\frac{1}{2} \frac{(x-\mu_0)^2}{\Sigma} + \frac{1}{2} \frac{(x-\mu_1)^2}{\Sigma}\right)$$

$$= \left(\frac{1-\phi}{\phi} \right) \exp\left(-\frac{1}{2\Sigma} (x^2 - 2\mu_0 x + \mu_0^2 - x^2 + 2\mu_1 x - \mu_1^2)\right)$$

$$= \left(\frac{1-\phi}{\phi} \right) \exp\left(-\frac{1}{2\Sigma} ((2\mu_1 - \mu_0 \times 2)x + \mu_0^2 - \mu_1^2)\right)$$

$$= \exp\left(-\left(\frac{\mu_1 - \mu_0}{\Sigma}\right)x - \left(\frac{\mu_0^2 - \mu_1^2}{2\Sigma} - \log_e\left(\frac{1-\phi}{\phi}\right)\right)\right)$$

\therefore We observe that, $P(y=1/x; \phi, \mu_0, \mu_1, \Sigma) = \frac{1}{1 + \exp(-(\theta^T x + \theta_0))}$

where $\theta_0 = \frac{\mu_0^2 - \mu_1^2}{2\Sigma} - \log_e\left(\frac{1-\phi}{\phi}\right)$

$$\theta^T = \frac{\mu_1 - \mu_0}{\Sigma}$$

$$\frac{(1-p) \sum_{i=1}^n (1-p(x)) \phi(x) + (1-p) \sum_{i=1}^n (1-p(x)) \phi(x)}{(1-p) \sum_{i=1}^n (1-p(x)) \phi(x) + (1-p) \sum_{i=1}^n (1-p(x)) \phi(x)}$$

Hence proved.

$$\frac{(1-p) \sum_{i=1}^n (1-p(x)) \phi(x) + 1}{(1-p) \sum_{i=1}^n (1-p(x)) \phi(x)}$$

$$\frac{((1-p)^T \sum_{i=1}^n (1-p(x)) \phi(x) + 1)}{((1-p)^T \sum_{i=1}^n (1-p(x)) \phi(x) + 1)} = \text{probability} = \frac{(1-p) \sum_{i=1}^n (1-p(x)) \phi(x)}{(1-p) \sum_{i=1}^n (1-p(x)) \phi(x)}$$