

Multilevel Models with Varying Intercepts and Slopes

Nambari Short Course

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Overview

- Continue with radon data
 - ▶ Elaborate model to allow varying slopes
 - ▶ Interpretation of variance components
 - ▶ Inclusion and interpretation of individual- and group-level predictors
- Different representations of the multilevel model

Varying slopes in a multilevel model

Radon data

First recall the multilevel model with varying intercepts

Level 1: Distribution of household radon level within county j , regressing on where measurement was taken (X_{ij})

$$Y_{ij} \sim \mathcal{N}(\alpha_j + X_{ij}\beta, \sigma^2)$$

Level 2: Distribution of county-specific mean radon levels

$$\alpha_j \sim \mathcal{N}(\mu, \tau^2)$$

In this model, the slope on X_{ij} is considered constant across counties

Varying slopes in a multilevel model

Radon data

It is possible to allow both the intercept and slope to vary across counties

Level 1: Distribution of household radon level within county j , regressing on where measurement was taken (X_{ij})

$$Y_{ij} \sim \mathcal{N}(\alpha_j + X_{ij}\beta_j, \sigma^2)$$

Level 2: Distribution of county-specific mean radon levels

$$\begin{pmatrix} \alpha_j \\ \beta_j \end{pmatrix} \sim \mathcal{N} \left[\begin{pmatrix} \mu_\alpha \\ \mu_\beta \end{pmatrix}, \begin{pmatrix} \tau_\alpha^2 & \rho\tau_\alpha\tau_\beta \\ \rho\tau_\alpha\tau_\beta & \tau_\beta^2 \end{pmatrix} \right]$$

In this model, the slope on X_{ij} varies across counties

Varying slopes in a multilevel model

Radon data

A closer look at level 2:

Level 2: Distribution of county-specific mean radon levels

$$\begin{pmatrix} \alpha_j \\ \beta_j \end{pmatrix} \sim \mathcal{N} \left[\begin{pmatrix} \mu_\alpha \\ \mu_\beta \end{pmatrix}, \begin{pmatrix} \tau_\alpha^2 & \rho\tau_\alpha\tau_\beta \\ \rho\tau_\alpha\tau_\beta & \tau_\beta^2 \end{pmatrix} \right]$$

In this set up,

τ_α^2 = variance of intercepts

τ_β^2 = variance of slopes

ρ = $\text{corr}(\alpha_j, \beta_j)$

Fitting the model using lmer

```
> ## Varying intercept & slopes w/ no group level predictors  
> M3 <- lmer (y ~ x + (1 + x | county))
```

```
> display (M3)  
lmer(formula = y ~ x + (1 + x | county))  
      coef.est coef.se  
(Intercept)  1.46     0.05  
x            -0.68     0.09
```

Error terms:

Groups	Name	Std.Dev.	Corr
county	(Intercept)	0.35	
	x	0.34	-0.34
Residual		0.75	

```
---  
number of obs: 919, groups: county, 85  
AIC = 2180.3, DIC = 2153.9  
deviance = 2161.1
```

Fitting the model using lmer

```
> coef (M3)
$county
      (Intercept)           x
1      1.1445320 -0.5406088
2      0.9333792 -0.7708179
3      1.4716908 -0.6688827
4      1.5354379 -0.7525587
5      1.4270363 -0.6206673
6      1.4826565 -0.6877101
7      1.8193559 -0.4747484
[...]
```

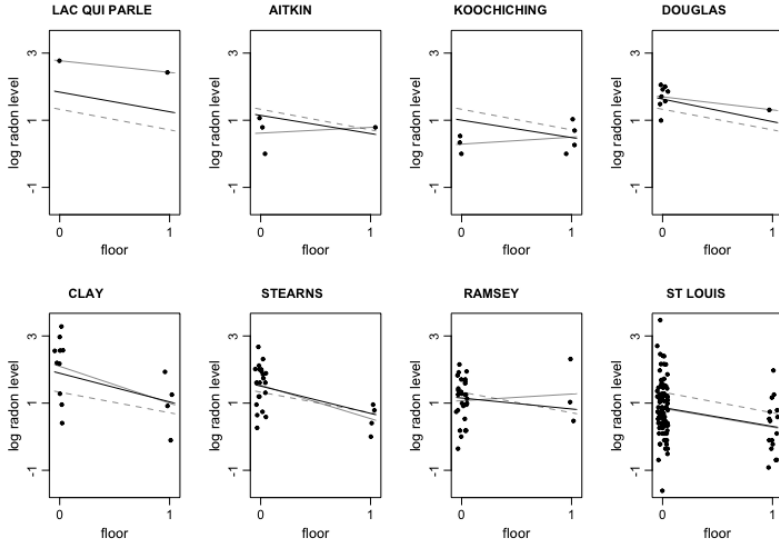
81	1.8441153	-0.5360210
82	1.6002912	-0.7268221
83	1.6942411	-1.1511024
84	1.5991199	-0.7327255
85	1.3787929	-0.6531768

Fitting the model using lmer

```
> fixef (M3)
(Intercept)                x
  1.4627703   -0.6810982

> ranef (M3)
$county
  (Intercept)                x
1  -0.318238351   0.140489364
2  -0.529391140  -0.089719737
3   0.008920474   0.012215455
4   0.072667555  -0.071460543
5  -0.035734045   0.060430874
6   0.019886162  -0.006611892
7   0.356585564   0.206349823
[...]
```

81	0.381344973	0.145077244
82	0.137520843	-0.045723902
83	0.231470744	-0.470004198
84	0.136349565	-0.051627299
85	-0.083977464	0.027921421



Including group-level predictor

Recall that U_j is uranium level in county j . Including group-level predictors entails expanding the level 2 model as follows:

$$\begin{pmatrix} \alpha_j \\ \beta_j \end{pmatrix} \sim \mathcal{N} \left[\begin{pmatrix} \gamma_0^\alpha + \gamma_1^\alpha U_j \\ \gamma_0^\beta + \gamma_1^\beta U_j \end{pmatrix}, \begin{pmatrix} \tau_\alpha^2 & \rho \tau_\alpha \tau_\beta \\ \rho \tau_\alpha \tau_\beta & \tau_\beta^2 \end{pmatrix} \right]$$

γ_1^α = uranium effect on intercept

γ_1^β = uranium effect on slope

Fitting the model in lmer

```
> ## Including group level predictors
> M4 <- lmer (y ~ x + u.full + x:u.full + (1 + x | county))

> display (M4)
lmer(formula = y ~ x + u.full + x:u.full + (1 + x | county))
      coef.est coef.se
(Intercept)  1.47    0.04
x            -0.67    0.08
u.full        0.81    0.09
x:u.full     -0.42    0.23
```

Error terms:

Groups	Name	Std.Dev.	Corr
county	(Intercept)	0.12	
	x	0.31	0.41
Residual		0.75	

```
---
number of obs: 919, groups: county, 85
AIC = 2142.6, DIC = 2101.9
deviance = 2114.2
```

Fitting the model in lmer

```
> coef (M4)
$county
      (Intercept)           x      u.full    x:u.full
1      1.458668 -0.6468782  0.8081345 -0.4195149
2      1.495642 -0.8889935  0.8081345 -0.4195149
3      1.476934 -0.6465966  0.8081345 -0.4195149
4      1.527552 -0.5989640  0.8081345 -0.4195149
5      1.479225 -0.6232842  0.8081345 -0.4195149
6      1.448104 -0.6916993  0.8081345 -0.4195149
7      1.581426 -0.3212303  0.8081345 -0.4195149
8      1.504498 -0.5967115  0.8081345 -0.4195149
```

Fitting the model in lmer

```
> fixef (M4)
(Intercept)          x          u.full      x:u.full
  1.4685942  -0.6709437   0.8081345  -0.4195149
```

```
> ranef (M4)
$county
  (Intercept)          x
1  -0.009926311  0.024065524
2   0.027047629 -0.218049771
3   0.008339473  0.024347096
4   0.058957543  0.071979706
5   0.010630259  0.047659481
6  -0.020490584 -0.020755637
7   0.112831703  0.349713432
8   0.035903697  0.074232254
```

