

Multilevel Logistic Regression for Binary Data – Part 2

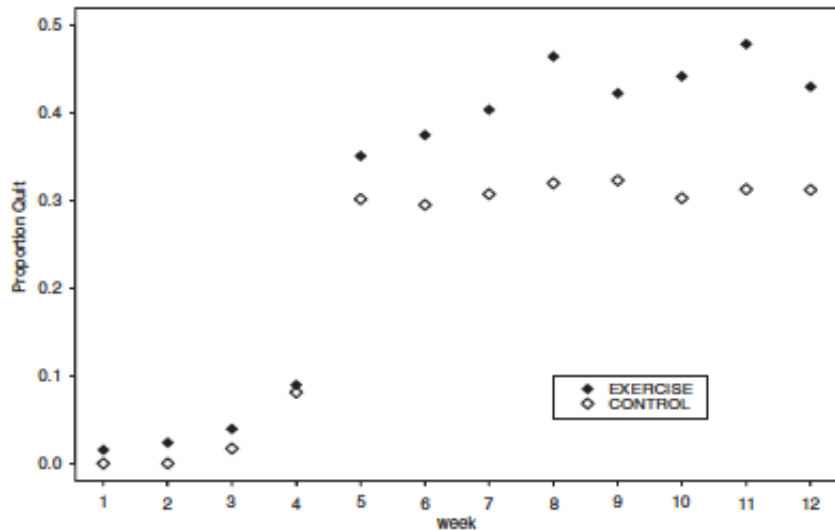
Nambari Short Course

16 July 2019

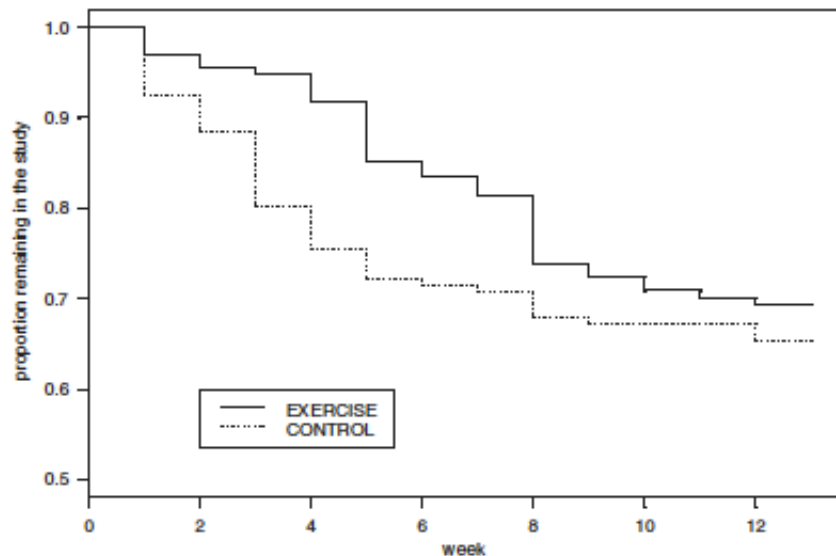
Overview

- Revisit data structure of CTQ data
- Model treatment effect using multilevel logit model
- Compare conditional (subject-specific) versus marginal (population-averaged) treatment effect

Smoking Cessation Study: Summaries



Smoking Cessation Study: Summaries



Data Analyses

- ➊ Multilevel logistic regression with intercept only
- ➋ Multilevel logistic regression on single covariate
- ➌ MLR of time trend and treatment effect

Variables used in this analysis

Y_{ij} = quit status for person i at time t_j
= 1 if quit, 0 if not

t_j = measurement time in weeks

Z_i = treatment group (1 = exercise, 0 = control)

X_i = baseline level of nicotine dependence (0 to 10)

Data excerpt

	ID	week	wk>4	Z	X	Y
[1,]	305	4	0	0	8	0
[2,]	305	5	1	0	8	1
[3,]	305	6	1	0	8	1
[4,]	305	7	1	0	8	1
[5,]	305	8	1	0	8	1
[6,]	305	9	1	0	8	1
[7,]	305	10	1	0	8	1
[8,]	305	11	1	0	8	1
[9,]	305	12	1	0	8	1
[10,]	309	4	0	1	6	0
[11,]	309	5	1	1	6	0
[12,]	309	6	1	1	6	0
[13,]	309	7	1	1	6	0
[14,]	309	8	1	1	6	0
[15,]	309	9	1	1	6	0
[16,]	309	10	1	1	6	0
[17,]	309	11	1	1	6	0
[18,]	309	12	1	1	6	0

Model 2: Include individual-level covariate

Covariate here is nicotine dependence score (0 to 10)

Level 1

$$\begin{aligned} Y_{ij} &\sim \text{Ber}(\pi_{ij}) \\ \text{logit}(\pi_{ij}) &= \alpha_i + \beta X_i \end{aligned}$$

Level 2

$$\alpha_i \sim N(0, \tau^2)$$

Interpretation:

- Coefficient β is the *subject-specific* effect of X ; i.e., the effect of X within an individual.
- Another interpretation is that it is the *conditional* effect of X (conditioning on α_i , the individual-level propensity to quit smoking)

Fitting the model in R

```
> M1 = glmer( Y ~ totfager + (1 | id), family=binomial(link=logit), data=ctq)
```

```
> display(M1)
```

```
glmer(formula = Y ~ totfager + (1 | id), data = ctq, family = binomial(link = logit))
```

	coef.est	coef.se
(Intercept)	-0.23	0.94
totfager	-0.44	0.15

Error terms:

Groups	Name	Std.Dev.
id	(Intercept)	3.68
Residual		1.00

number of obs: 1688, groups: id, 266

AIC = 1336.6, DIC = 257.7

deviance = 794.2

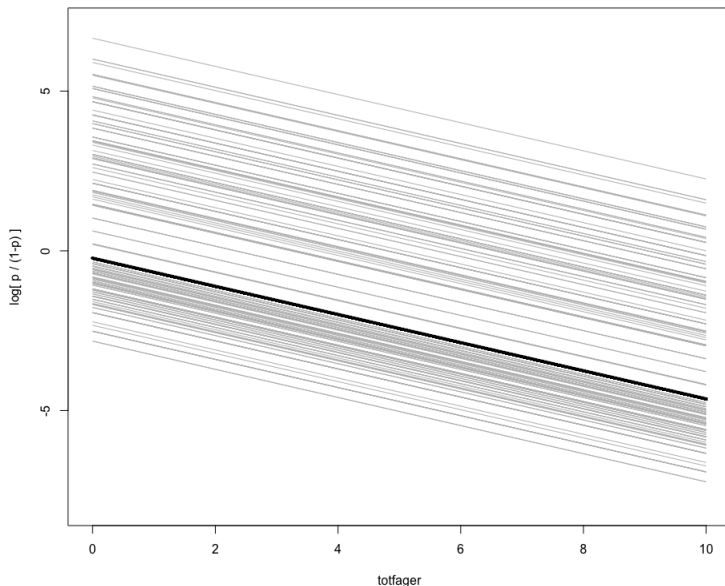
Fitting the model in R

```
> coef(M0)
```

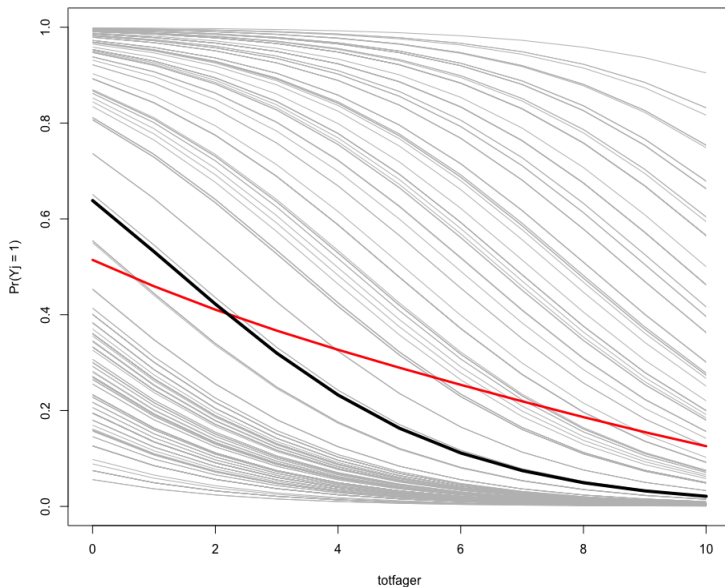
```
$id
```

	(Intercept)	totfager
305	5.0808334	-0.4404643
309	-1.6677924	-0.4404643
311	-0.9908914	-0.4404643
313	-1.4195797	-0.4404643
314	1.8312968	-0.4404643
317	-0.8460499	-0.4404643
321	-1.9349872	-0.4404643

Plot of individual-level effect of X on logit scale



Probability scale



Population-averaged versus subject-specific effect of X

Subject-specific effect

Can write model like this

$$P(Y_{ij} = 1 \mid \alpha_i, X_i) = \frac{\exp(\alpha_i + \beta X_i)}{1 + \exp(\alpha_i + \beta X_i)}$$

For each individual i , the effect of X is β

Population-averaged versus subject-specific effect of X

Population-averaged effect

- To calculate PA effect, need to integrate (average) over distribution of α_i
- Recall that this distribution is $\mathcal{N}(\mu, \tau^2)$.
- Denote by $f(\alpha | \mu, \tau)$

The population-averaged effect is

$$P(Y_{ij} = 1 | X_i) = \int \frac{\exp(\alpha + \beta X_i)}{1 + \exp(\alpha + \beta X_i)} f(\alpha | \mu, \tau) d\alpha$$

Population-averaged versus subject-specific effect of X

When α has a normal distribution, can show that the integral is approximately

$$P(Y_{ij} = 1 | X_i) \approx \frac{\exp(\mu^* + \beta^* X_i)}{1 + \exp(\mu^* + \beta^* X_i)}$$

where

$$\beta^* = \frac{\beta}{(1 + .346\tau^2)^{1/2}}$$

In our example,

$$\begin{aligned}\hat{\beta} &= -0.44 \\ \hat{\beta}^* &= \frac{-0.44}{(1 + .346(3.68^2))^{1/2}} \approx -0.18\end{aligned}$$

- At the individual level, nicotine dependence has a strong negative effect on probability of smoking cessation in a given week, associated with decreased log odds of quitting of -0.44
- Averaged over the population, the effect is attenuated toward zero. The population-averaged effect represents a between-group comparison of those who differ by one unit in dependence score.

Model of treatment effect

- The structure of the model is motivated by the longitudinal data patterns
- Three parameters:
 - ▶ Overall intercept (quit rate prior to week 5)
 - ▶ Quit rate for control from weeks 5-12 (constant)
 - ▶ Difference in quit rate for treatment from weeks 5-12 (constant)

Model of treatment effect

Level 1

$$\begin{aligned} Y_{ij} &\sim \text{Ber}(\pi_{ij}) \\ \text{logit}(\pi_{ij}) &= \alpha_i + \beta T_j + \theta T_j Z_i \end{aligned}$$

Level 2

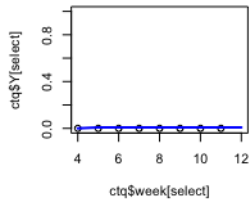
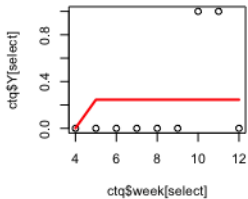
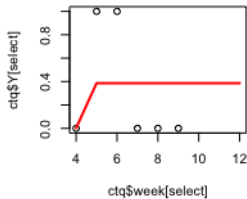
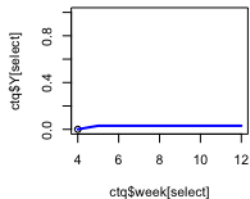
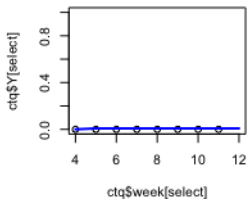
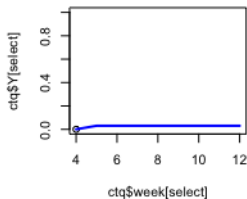
$$\alpha_i \sim \mathcal{N}(\mu, \tau^2)$$

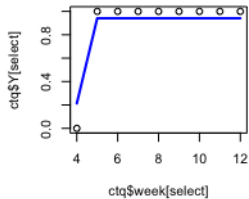
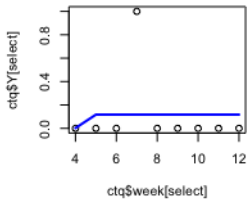
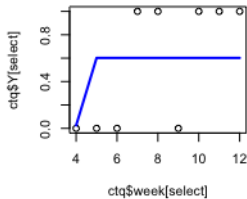
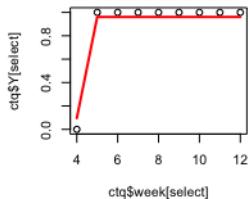
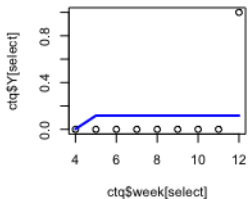
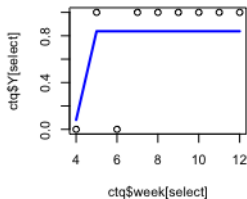
In this notation,

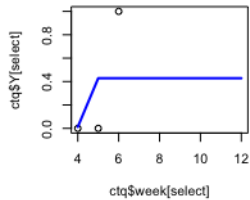
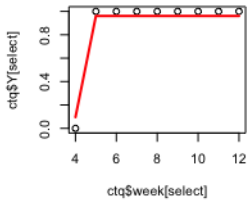
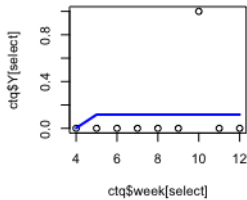
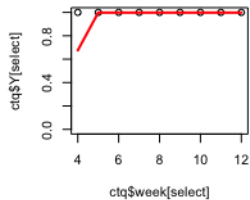
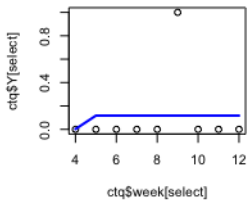
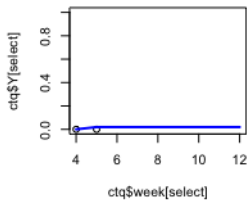
$$\begin{aligned} T_i &= 1 \text{ if week} \geq 5 \\ &= 0 \text{ if not} \end{aligned}$$

Fitting the model in R

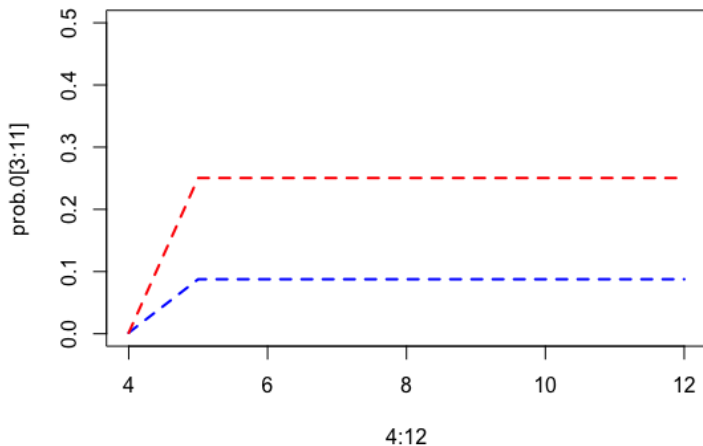
```
> M2 = glmer( Y ~ target_wk + target_wk:Z + (1 | id ), family=binomial(link=logit),  
> display(M2)  
glmer(formula = Y ~ target_wk + target_wk:Z + (1 | id), data = ctq,  
      family = binomial(link = logit))  
              coef.est coef.se  
(Intercept) -7.52      0.80  
target_wk      4.08      0.53  
target_wk:Z    1.36      0.66  
  
Error terms:  
  Groups      Name          Std.Dev.  
  id          (Intercept)  4.92  
  Residual                        1.00  
---  
number of obs: 1887, groups: id, 266  
AIC = 1220.6, DIC = 84.9  
deviance = 648.7
```







Subject-specific effect



Population-averaged effect

