

Bayes Classifier with Multivariate Normal Distribution

Slides from Machine Learning by Ethem Alpaydin
Chapter 4 and 5

Expanded by some slides from Gutierrez-Osuna

Overview

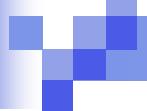
- We learned how to use the Bayesian approach for classification *if* we had the probability distribution of the underlying classes ($p(\mathbf{x}|C_i)$).
- **Naive Bayes** approached the problem by simplifying the task, with the assumption of:

$$P(\mathbf{x} | C_i) = \prod_{j=1..d} P(x_j | C_i)$$

- Now we are going to look at how to use the **multivariate Normal distribution** to model $P(\mathbf{x} | C_i)$

Overview

- Gaussian Bayes classifier ...
 - is a **Bayesian classifier** so it assigns the input \mathbf{x} to the class C_j for which $P(C_j|\mathbf{x})$ is largest;
 - assumes that \mathbf{x} is **multivariate normal**.
 - $\mathbf{x} \sim N(\mathbf{x} | \mu, \Sigma)$
 - I.e. $p(\mathbf{x} | C_j) = N(\mathbf{x} | \mu_i, \Sigma_i)$
 - It computes $p(\mathbf{x} | C_j)$ by estimating the parameters of the **multivariate Normal distribution** from the data.
 - For for 1D-Normal distribution:
 - Estimate μ, σ^2 using Maximum Likelihood Estimation
 - Use $p(\mathbf{x} | C_j) = N(\mathbf{x} | \mu_i, \Sigma_i)$ in Bayesian decision process.



Self study - [Skip](#)

REVIEW: EXPECTATION

- A discrete random variable X takes on values $x_1..x_m$ with probabilities $p(x_1)...p(x_m)$.

- Its expected value is defined as

$$E(X) = \sum_i x_i P(x_i)$$

- E.g. In a TV game, you will lose -1000 TL with 0.6 probability; and gain 10,000 TL with 0.4 probability. What is the expected reward?

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- E.g. In a TV game, you will lose -1000 TL with 0.6 probability; and gain 10,000 TL with 0.4 probability. What is the expected reward?

$$E(X) = \sum_i x_i P(x_i)$$

$$E[\text{reward}] = -1,000 \times 0.6 + 10,000 \times 0.4 = -600 + 1000 = 400 \text{ TL}$$

Expected Value and Variance of a Function

The **average value of a function $f(x)$ under a probability distribution $p(x)$** is called the **expectation** of $f(x)$.

- Weighted average, where weighting is done according to the probabilities of different values of x .

$$\mathbb{E}[f] = \sum_x p(x)f(x) \quad \mathbb{E}[f] = \int p(x)f(x) dx$$

The **variance of a function $f(x)$ provides a measure of how much $f(x)$ varies around its mean $\mathbb{E}[f(x)]$** .

$$\text{var}[f] = \mathbb{E} \left[(f(x) - \mathbb{E}[f(x)])^2 \right] = \mathbb{E}[f(x)^2] - \mathbb{E}[f(x)]^2$$

Expectations

$$\mathbb{E}[f] = \sum_x p(x)f(x) \quad \mathbb{E}[f] = \int p(x)f(x) dx$$

Approximate expectation:
(discrete and continuous)

$$\mathbb{E}[f] \simeq \frac{1}{N} \sum_{n=1}^N f(x_n)$$

In other words, if we have some samples x from a given distribution along with $f(x)$ values, we can calculate the approximate estimate by computing the average of $f(x)$ obtained for these samples obtained from this distribution.

Note that there will be more samples in the data set where $p(x)$ is larger and fewer samples where $p(x)$ is smaller; hence we do not need $p(x)$ anymore!

Now we are going to look at concepts: variance and co-variance , of one or more random variables, using the concept of expectation.

Variance and Covariance

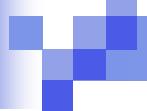
The **variance** of a random variable x provides a measure for how much x varies around its mean $E[x]$.

$$var[x] = \mathbb{E}[(x - \mathbb{E}[x])^2]$$

Co-variance of two random variables x and y measures the extent to which they vary together.

$$\text{cov}[x, y] = \mathbb{E}_{x,y} [\{x - \mathbb{E}[x]\} \{y - \mathbb{E}[y]\}]$$

Here the subscript x,y indicates that the expectation is taken over both x and y

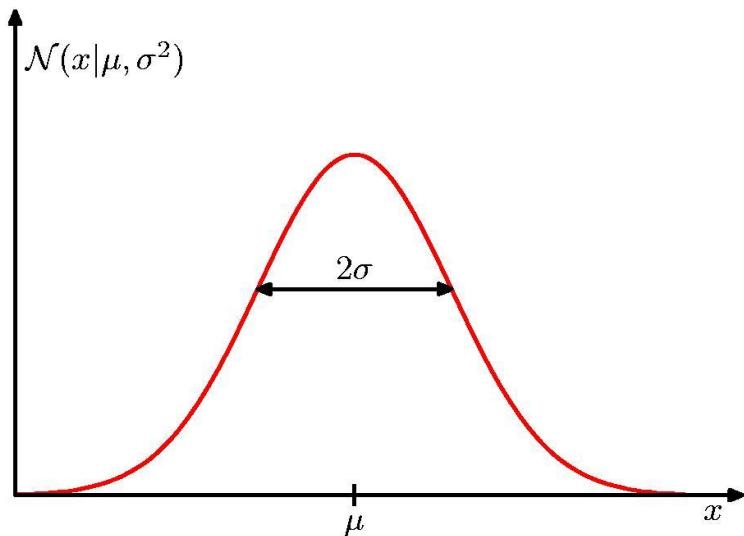


MULTIVARIATE NORMAL DISTRIBUTION

1D Normal (Gaussian) Distribution

1D Normal distribution is defined by its mean μ and variance σ^2 denoted as $N(x|\mu,\sigma^2)$ – it gives the probability density at any point x under Normal distribution with the given two parameters.

$$N(x|\mu,\sigma^2) = \frac{1}{(2\pi\sigma^2)^{1/2}} \exp\left\{-\frac{1}{2\sigma^2}(x-\mu)^2\right\}$$



$$N(x|\mu, \sigma^2) > 0$$

$$\int_{-\infty}^{\infty} N(x|\mu, \sigma^2) dx = 1$$

Expectations

Definition of expectation: $\mathbb{E}[f] = \int p(x)f(x) dx$

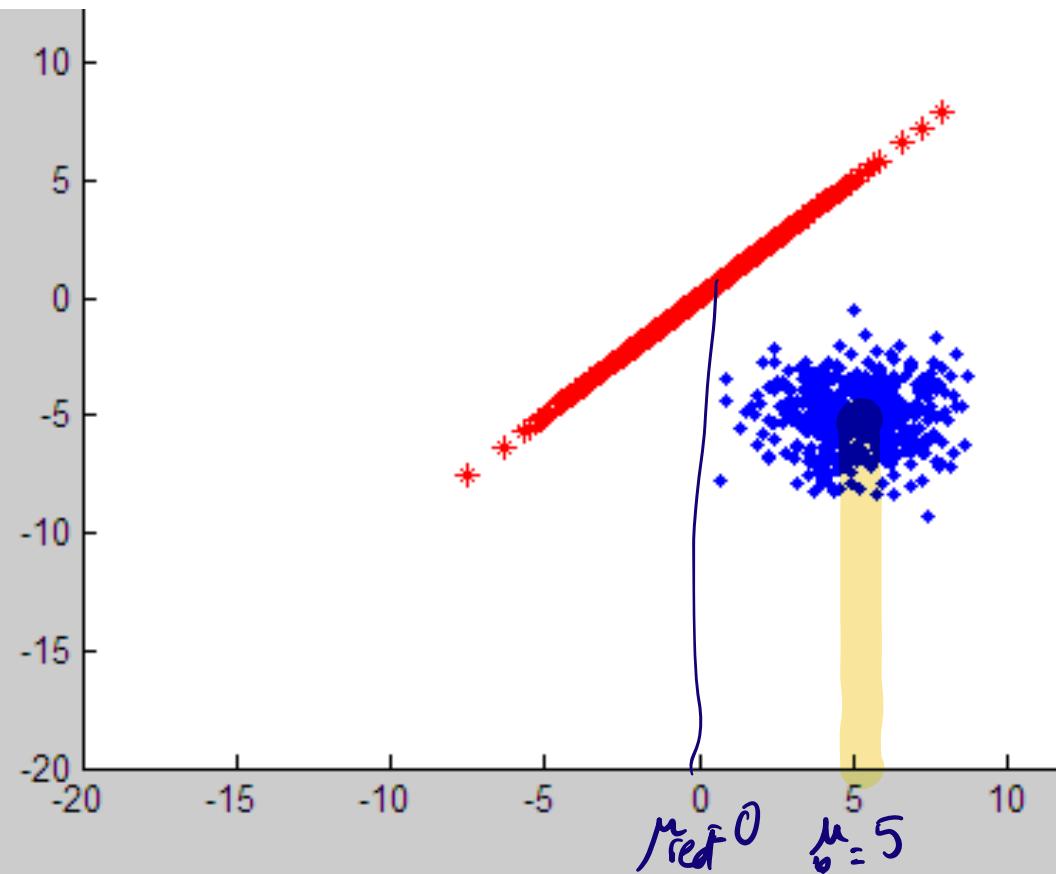
For normally distributed x :

$$\mathbb{E}[x] = \int_{-\infty}^{\infty} \mathcal{N}(x|\mu, \sigma^2) x dx = \mu$$

$$\mathbb{E}[x^2] = \int_{-\infty}^{\infty} \mathcal{N}(x|\mu, \sigma^2) x^2 dx = \mu^2 + \sigma^2$$

$$\text{var}[x] = \mathbb{E}[x^2] - \mathbb{E}[x]^2 = \sigma^2$$

- We will see how we can characterize $p(\mathbf{x})$, assuming \mathbf{x} is normally distributed.
 - For 1-dimension, we need the **mean** (μ) and **variance** (σ^2)
 - For d-dimensions, we need
 - the d-dimensional **mean** vector μ
 - $d \times d$ dimensional **covariance** matrix Σ



Multivariate Normal Distribution

- For a single variable, the normal density function is:

$$p(x) = \frac{1}{(2\pi\sigma^2)^{1/2}} \exp\left\{-\frac{(x - \mu)^2}{2\sigma^2}\right\}$$

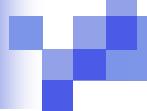
- For variables in higher dimensions, this generalizes to:

$$p(\mathbf{x}) = \frac{1}{(2\pi)^{d/2} |\Sigma|^{1/2}} \exp\left\{-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu})\right\}$$

where the mean $\boldsymbol{\mu}$ is now a d-dimensional vector,

$$\boldsymbol{\mu} = \mathcal{E}[\mathbf{x}]$$

$$\boldsymbol{\Sigma} = \mathcal{E}[(\mathbf{x} - \boldsymbol{\mu})(\mathbf{x} - \boldsymbol{\mu})^T].$$



UNDERSTANDING THE MULTIVARIATE NORMAL DISTRIBUTION

Multivariate Parameters: Mean, Covariance

$$\text{Mean : } E[\mathbf{x}] = \boldsymbol{\mu} = [\mu_1, \dots, \mu_d]^T$$

$$\Sigma \equiv \text{Cov}(\mathbf{x}) = \begin{bmatrix} \sigma_1^2 & \sigma_{12} & \cdots & \sigma_{1d} \\ \sigma_{21} & \sigma_2^2 & \cdots & \sigma_{2d} \\ \vdots & & & \\ \sigma_{d1} & \sigma_{d2} & \cdots & \sigma_d^2 \end{bmatrix}$$

$$\text{where } \sigma_{ij} \equiv \text{Cov}(x_i, x_j) \equiv E[(x_i - \mu_i)(x_j - \mu_j)]$$

and σ_i^2 is the variance of the i th feature.

Multivariate Parameters: Mean, Covariance

$$\text{Mean : } E[\mathbf{x}] = \boldsymbol{\mu} = [\mu_1, \dots, \mu_d]^T$$

$$\Sigma \equiv \text{Cov}(\mathbf{x}) = \begin{bmatrix} \sigma_1^2 & \sigma_{12} & \cdots & \sigma_{1d} \\ \sigma_{21} & \sigma_2^2 & \cdots & \sigma_{2d} \\ \vdots & & & \\ \sigma_{d1} & \sigma_{d2} & \cdots & \sigma_d^2 \end{bmatrix} = E[(\mathbf{x} - \boldsymbol{\mu})(\mathbf{x} - \boldsymbol{\mu})^T] = E[\mathbf{x}\mathbf{x}^T] - \boldsymbol{\mu}\boldsymbol{\mu}^T$$

In matrix notation:

$$\text{where } \sigma_{ij} \equiv \text{Cov}(x_i, x_j) \equiv E[(x_i - \mu_i)(x_j - \mu_j)]$$

$$\sigma_{ij} = \frac{1}{N} \sum_{n=1}^N (x_i^{(n)} - \mu_i)(x_j^{(n)} - \mu_j)$$

Matlab code

- close all;
- rand('twister', 1987) % seed

%Define the parameters of two 2d-Normal distribution

- **mu1 = [5 -5];**
- **mu2 = [0 0];**
- **sigma1 = [2 0; 0 2];**
- **sigma2 = [5 5; 5 5];**
- N=500; %Number of samples we want to generate from this distribution
- **samp1 = mvnrnd(mu1,sigma1, N);**
- **samp2 = mvnrnd(mu2, sigma2, N);**
-
- figure; clf;
- plot(samp1(:,1), samp1(:,2),'.', 'MarkerEdgeColor', 'b');
- hold on;
- plot(samp2(:,1), samp2(:,2),'*', 'MarkerEdgeColor', 'r');
- axis([-20 20 -20 20]); legend('d1', 'd2');

blue
 $\mu_1 = [5 \ -5];$
 $\mu_2 = [0 \ 0];$
red

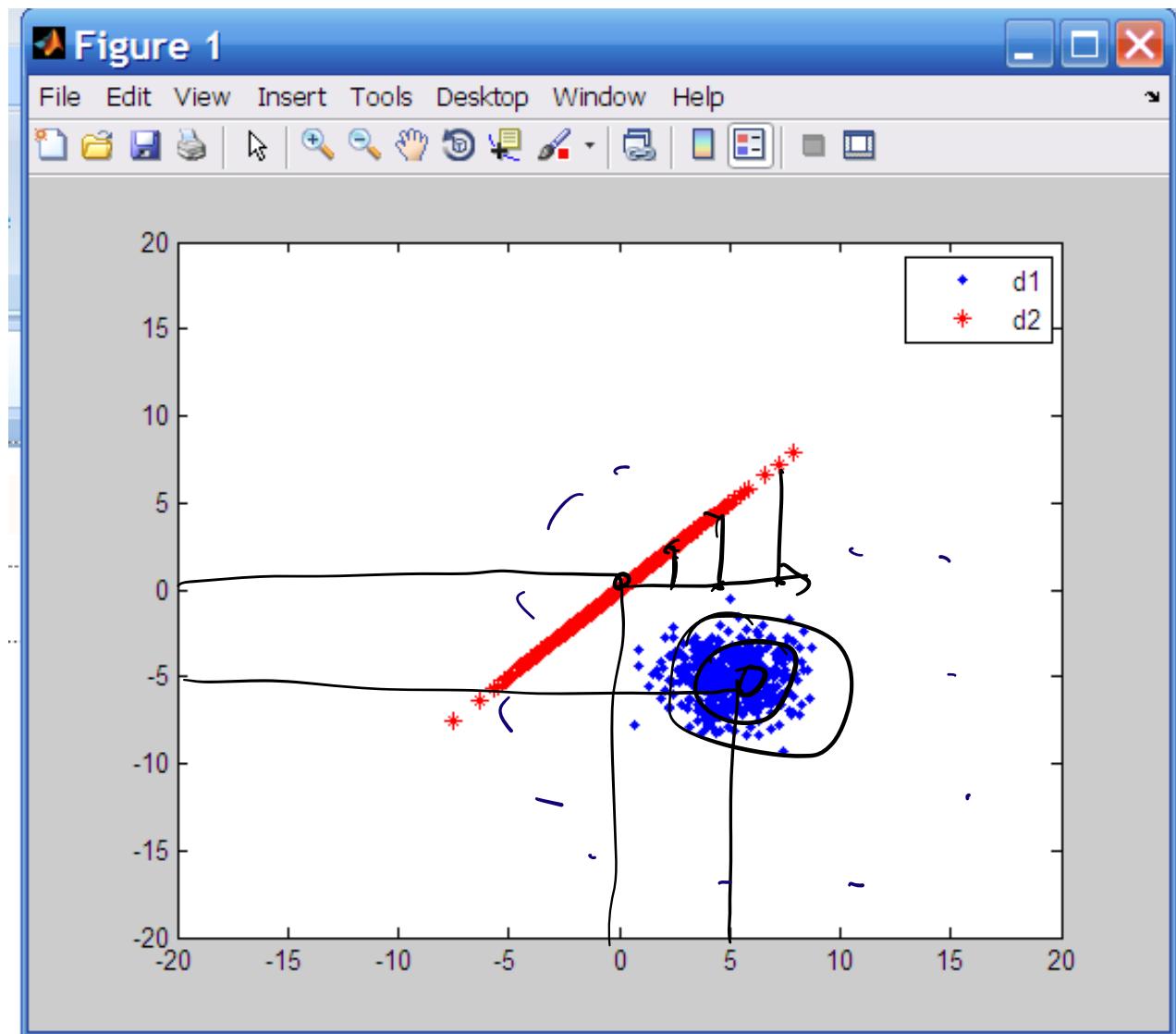
$\sigma_1 = [2 \ 0; 0 \ 2];$
 $\sigma_2 = [5 \ 5; 5 \ 5];$

$$\Sigma_{\text{blue}} = \begin{bmatrix} 2 & 0 \\ 0 & 2 \end{bmatrix}$$

$$\Sigma_{\text{red}} = \begin{bmatrix} 5 & 5 \\ 5 & 5 \end{bmatrix}$$

$$f_{12} = \frac{\sigma_{12}}{\sigma_1 \cdot \sigma_2}$$

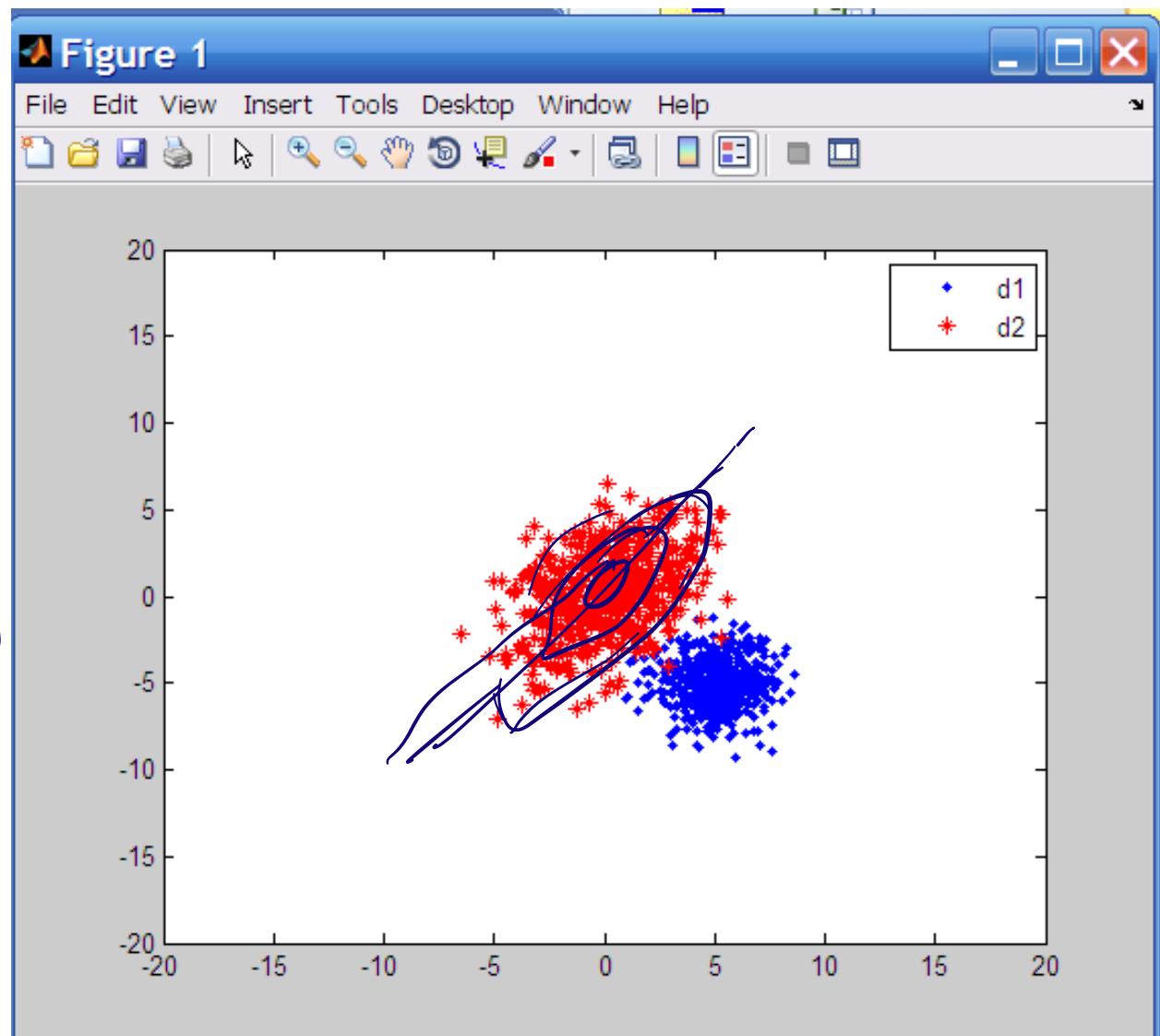
correlation



$\mu_1 = [5 \ -5];$
 $\mu_2 = [0 \ 0];$

$\sigma_1 = [2 \ 0; 0 \ 2];$
 $\sigma_2 = [5 \ 2; 2 \ 5];$

$$\Sigma_{\text{fed}} = \begin{bmatrix} 5 & 2 \\ 2 & 5 \end{bmatrix}$$

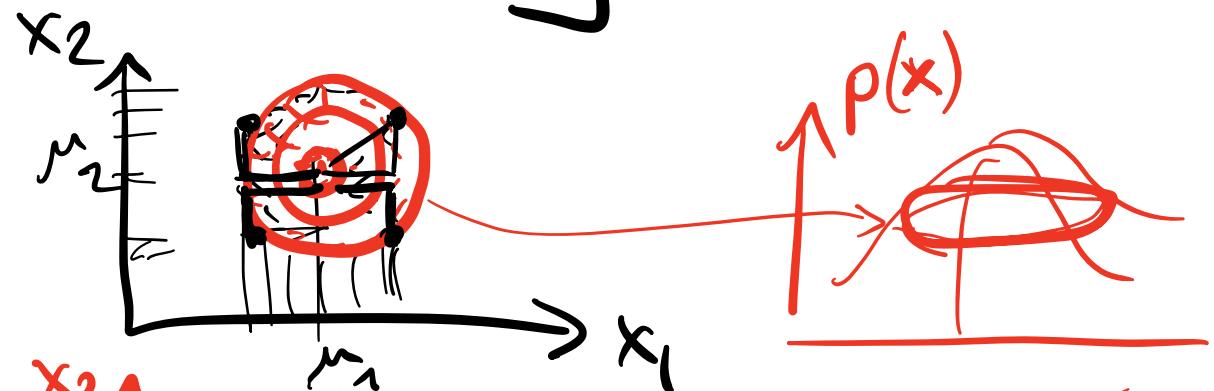


Matlab sample cont.

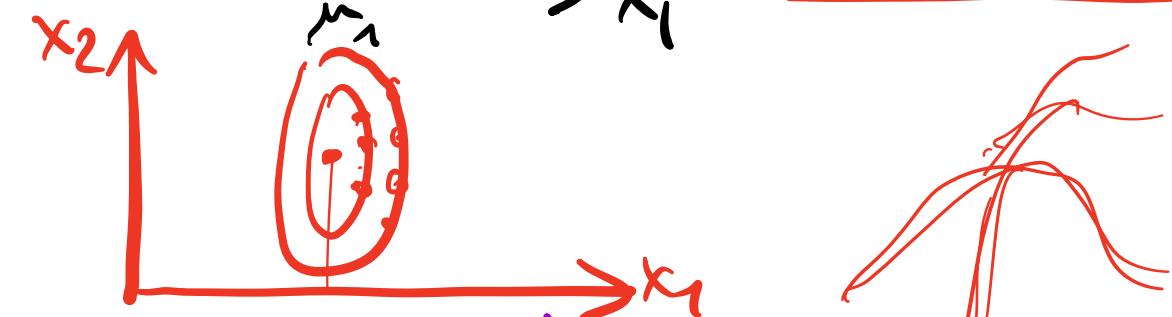
```
■ % Lets compute the mean and covariance as if we are given this data  
■ sampmu1 = sum(samp1)/N;  
■ sampmu2 = sum(samp2)/N;  
■ sampcov1 = zeros(2,2);  
■ sampcov2 = zeros(2,2);  
■ for i =1:N  
■     sampcov1 = sampcov1 + (samp1(i,:)-sampmu1)' * (samp1(i,:)-sampmu1);  
■     sampcov2 = sampcov2 + (samp2(i,:)-sampmu2)' * (samp2(i,:)-sampmu2);  
■ End  
■ sampcov1 = sampcov1 /N;  
■ sampcov2 = sampcov2 /N;  
■ %%%%%%  
■ % Lets compute the mean and covariance as if we are given this data USING MATRIX OPERATIONS  
■ % Notice that in samp1, samples are given in ROWS – but for this multiplication, columns * rows is req.  
■ sampcov1 = (samp1*samp1)/N - sampmu1*sampmu1;  
■ %Or simply  
■ mu=mean(samp1);  
■ cov=cov(samp1);
```

Some cases:

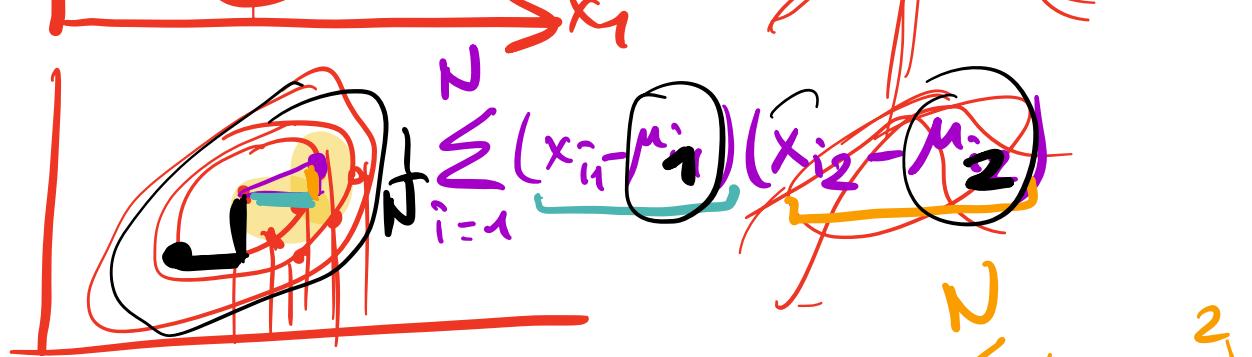
- A) $\Sigma = \begin{bmatrix} 4 & 0 \\ 0 & 4 \end{bmatrix}$



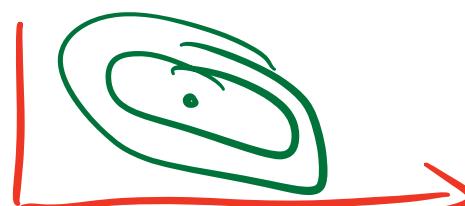
- B) $\Sigma = \begin{bmatrix} 4 & 0 \\ 0 & 9 \end{bmatrix}$



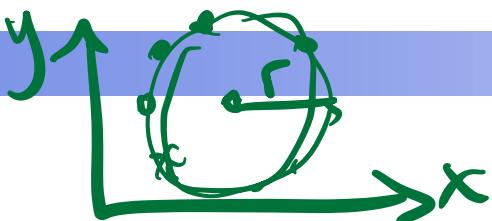
C) $\Sigma = \begin{bmatrix} 4 & 2 \\ 2 & 9 \end{bmatrix}$



D) $\Sigma = \begin{bmatrix} 4 & -2 \\ -2 & 9 \end{bmatrix}$



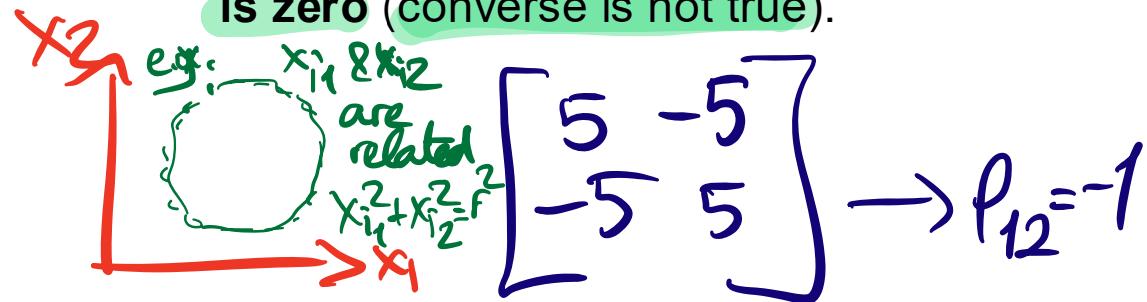
Variance of x_1 = $\frac{1}{N} \sum_{i=1}^N (x_{i1} - \mu_{11})^2$



$$x^2 + y^2 = r^2$$



- **Variance**: How much X varies around its expected value
 - **Covariance** is the measure the strength of the **linear relationship** between two random variables
 - covariance becomes more positive for each pair of values which differ from their mean in the same direction
 - covariance becomes more negative with each pair of values which differ from their mean in opposite directions.
 - if two variables are independent, then their covariance/correlation is zero (converse is not true).



$$\text{Covariance: } \sigma_{ij} \equiv \text{Cov}(x_i, x_j) \equiv E[(x_i - \mu_i)(x_j - \mu_j)]$$

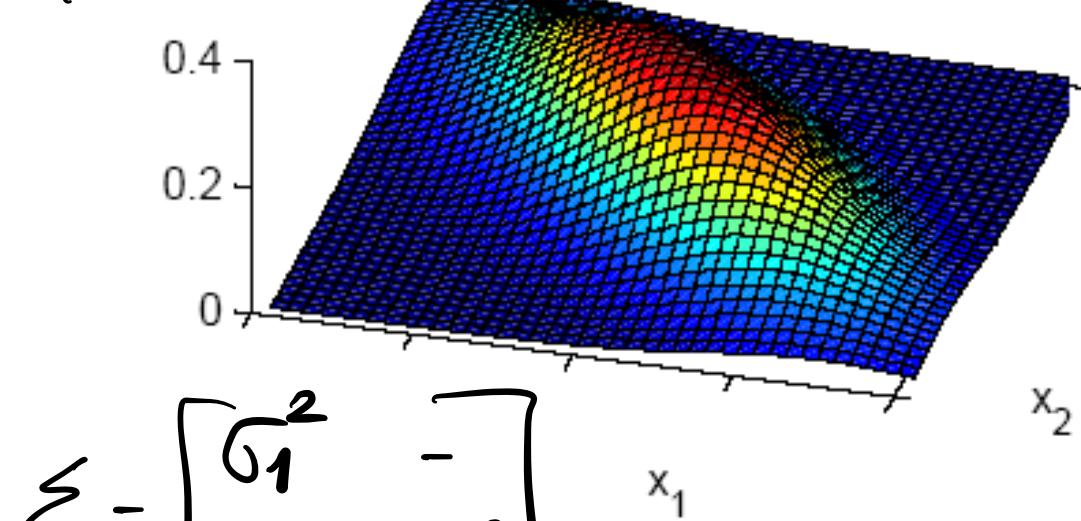
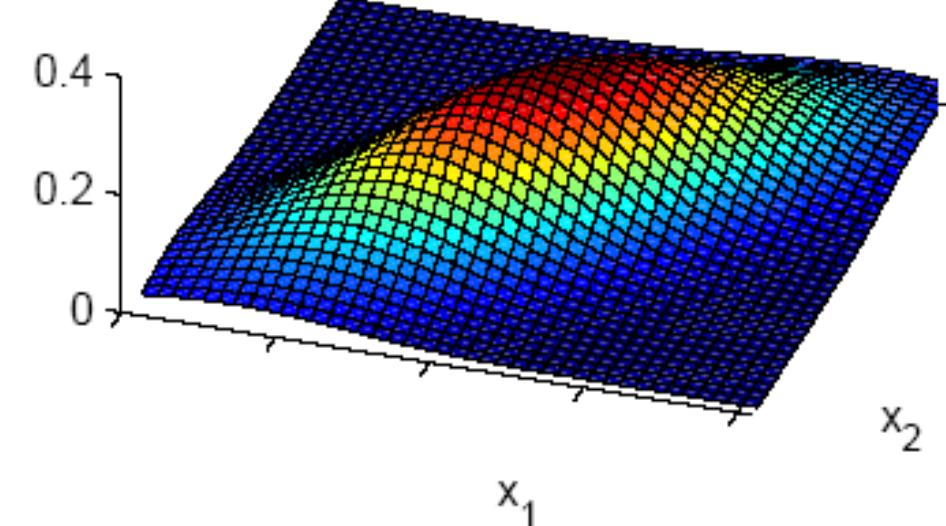
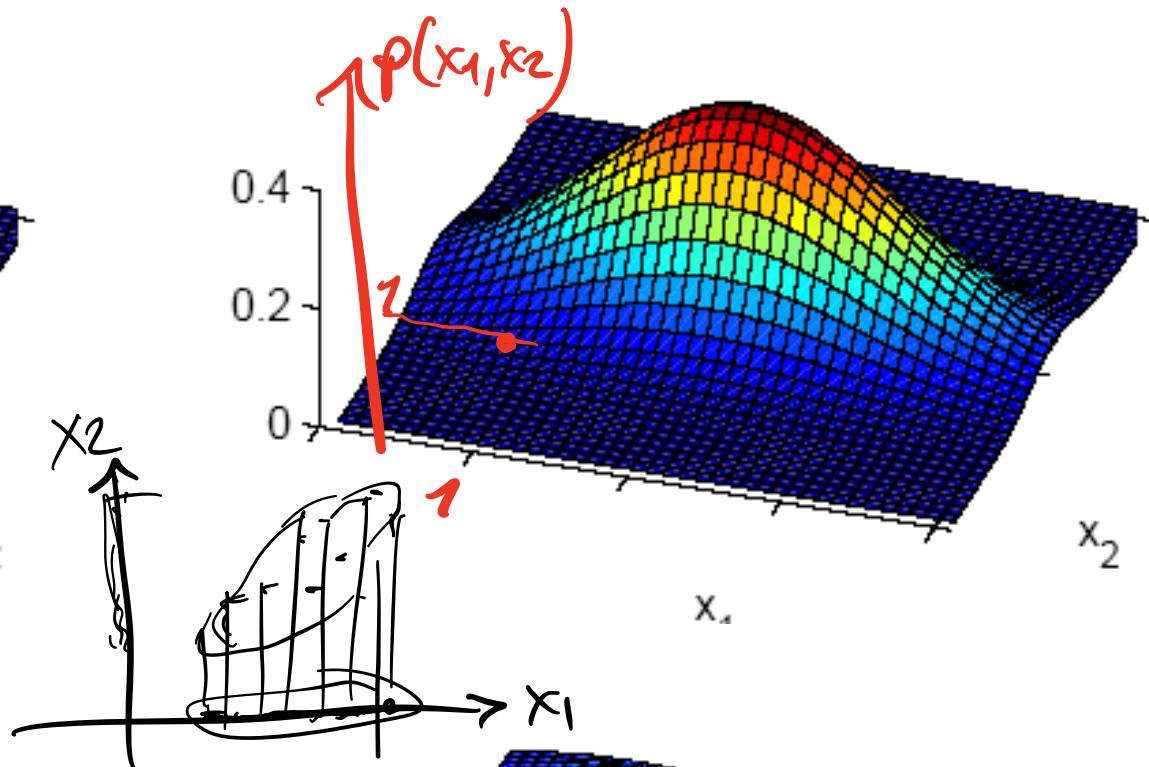
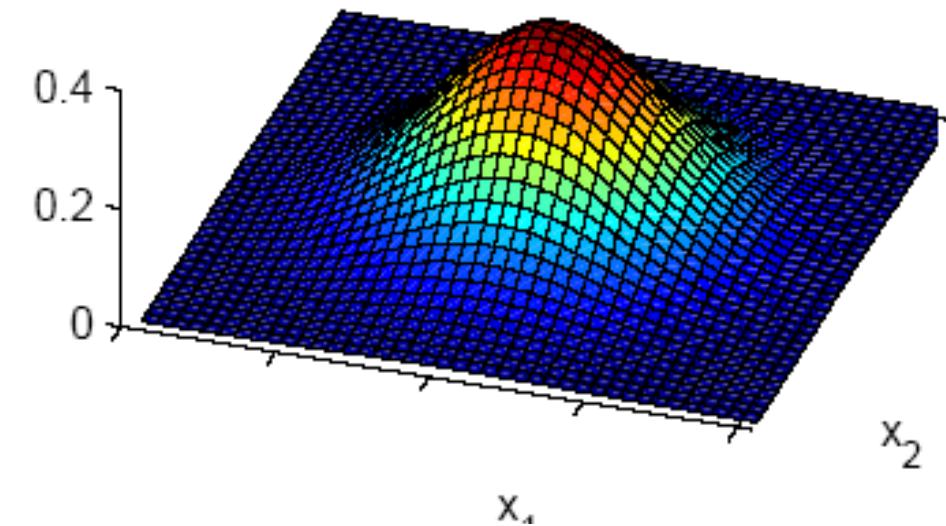
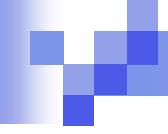
$$\text{Correlation : } \text{Corr}(x_i, x_j) \equiv \rho_{ij} = \frac{\sigma_{ij}}{\sigma_i \sigma_j}$$

- **Correlation** is a dimensionless measure of linear dependence.
 - range between -1 and $+1$

$$P_{ij} = \frac{G_{ij}}{G_i \cdot G_j} = \frac{5}{\sqrt{5} \cdot \sqrt{5}} = +1.$$

$$\sum_{bad} = \left[\begin{matrix} 5 & 5 \\ 5 & 5 \end{matrix} \right]$$

How to characterize differences between these distributions



$$\Sigma = \begin{bmatrix} \sigma_1^2 & - \\ - & \sigma_2^2 \end{bmatrix}$$



Shape and orientation of the hyper-ellipsoid centered at μ is defined by Σ

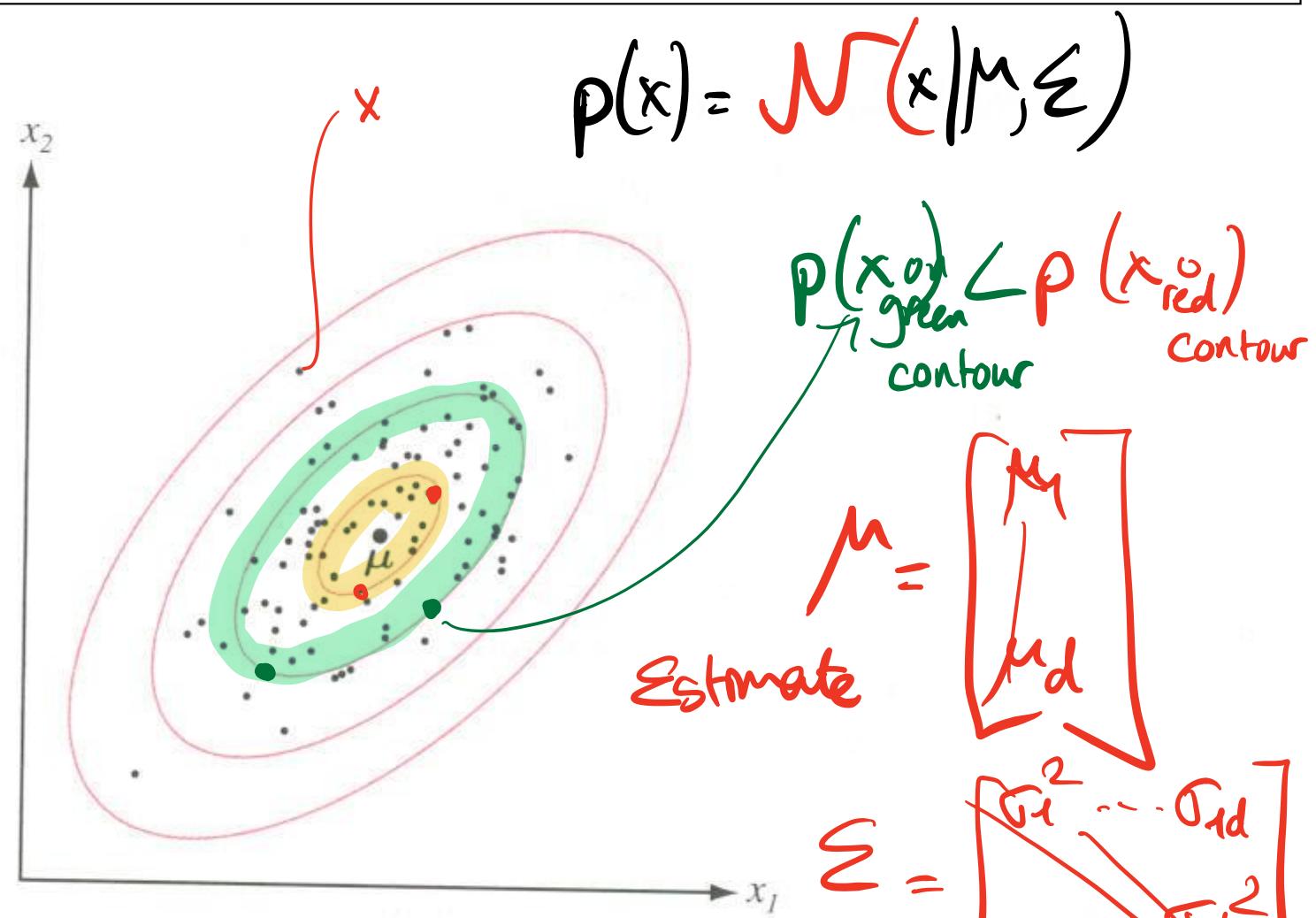
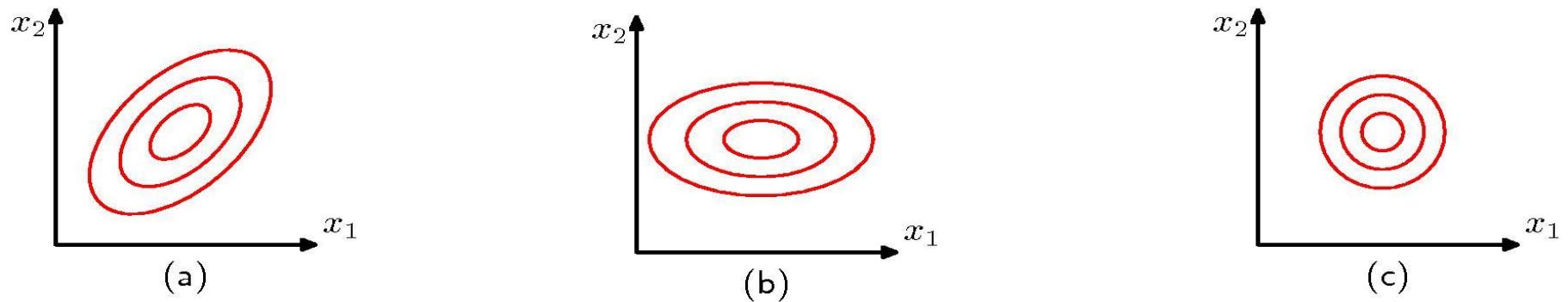
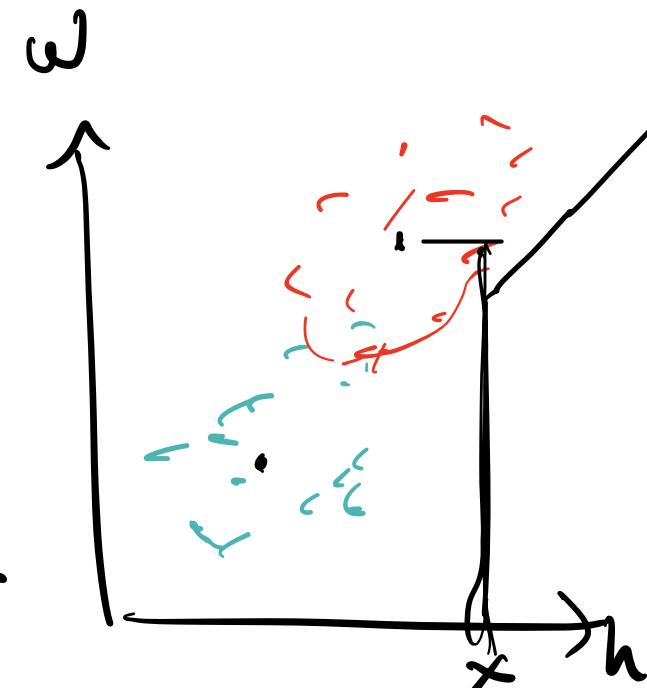
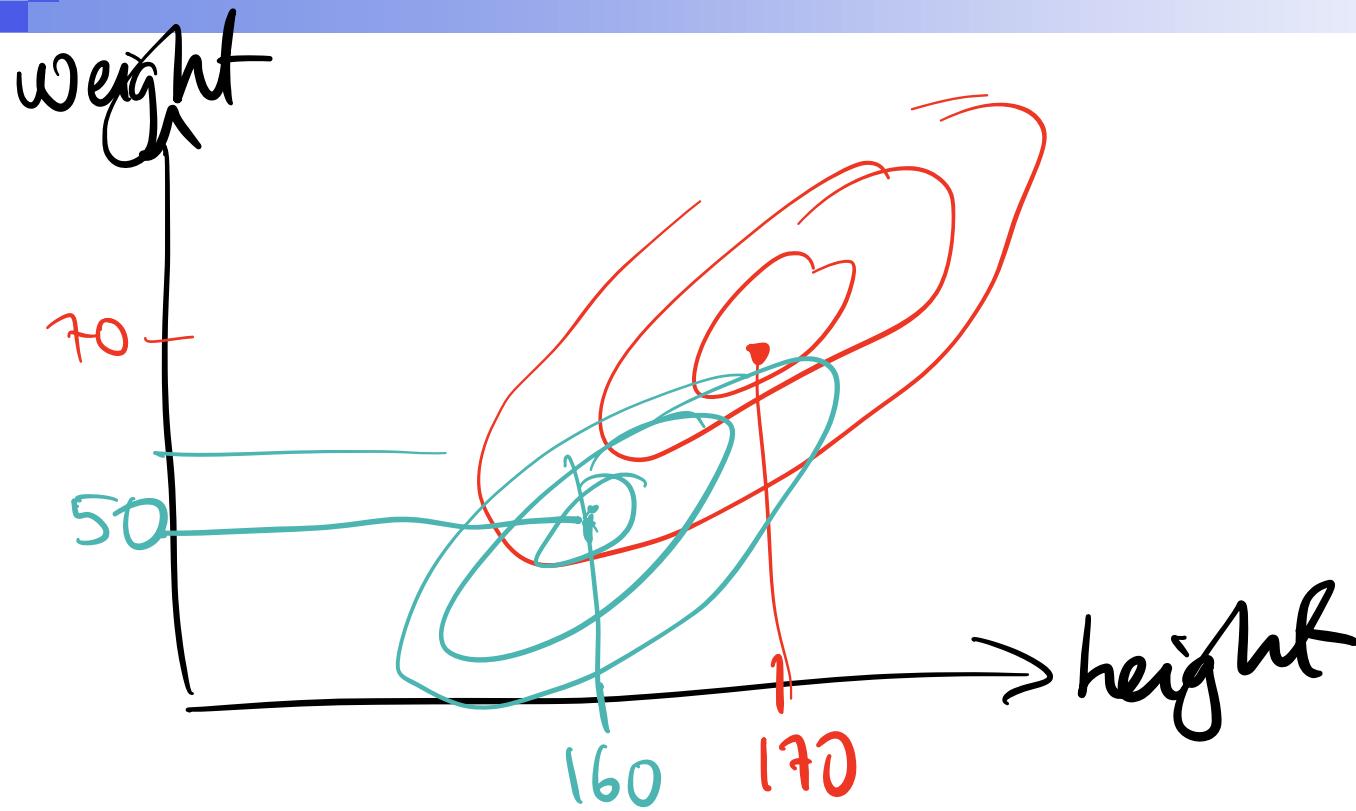


FIGURE 2.9. Samples drawn from a two-dimensional Gaussian lie in a cloud centered on the mean μ . The ellipses show lines of equal probability density of the Gaussian.



Contours of constant probability density for a 2D Gaussian distributions with

- general covariance matrix
- diagonal covariance matrix (covariance of x_1, x_2 is 0)
- Σ proportional to the identity matrix (covariances are 0, variances of each dimension are the same)



Mahalanobis Distance

estimate $\mu_{\text{male}}, \Sigma_{\text{male}}$
to estimate $p(x|\text{male})$

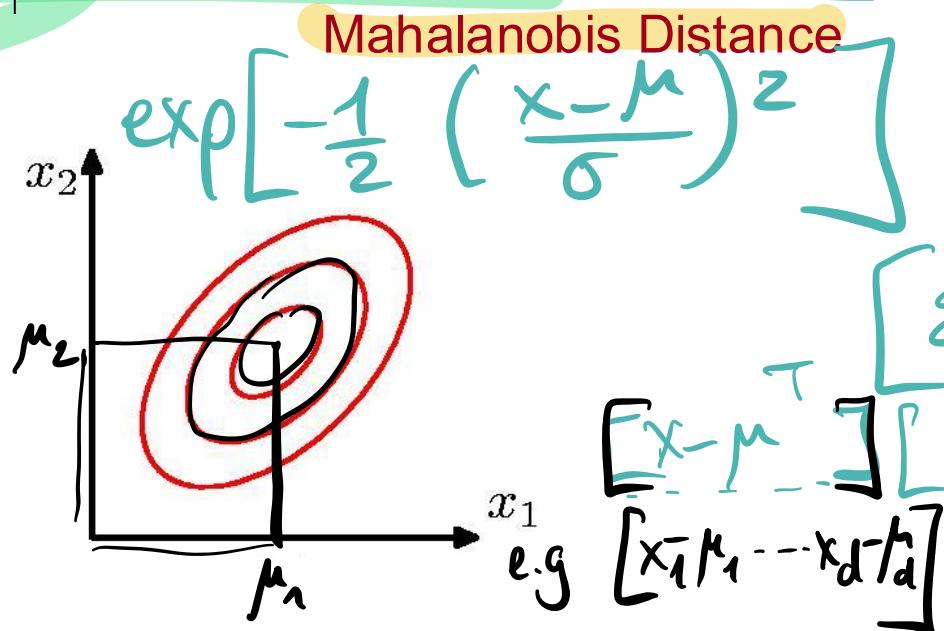
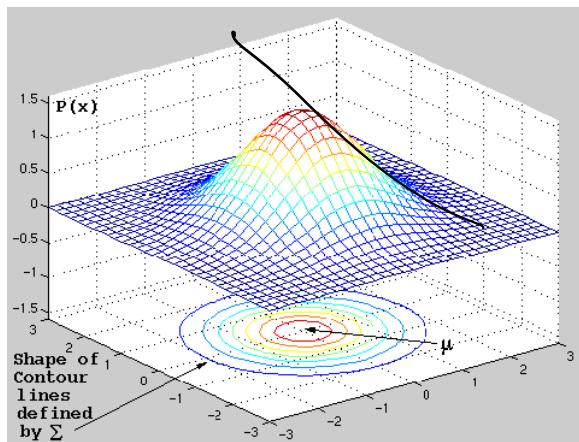
$$x = \begin{bmatrix} x_1 \\ x_2 \end{bmatrix} = \begin{bmatrix} 180 \\ 70 \end{bmatrix}$$

$$p(\text{male} | x) \propto p(x | \text{male})$$

$$p(\text{female} | x)$$

$$p(\mathbf{x}) = \frac{1}{(2\pi)^{d/2} |\Sigma|^{1/2}} \exp \left[-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^t \Sigma^{-1} (\mathbf{x} - \boldsymbol{\mu}) \right]$$

one sample



From the equation for the normal density, it is apparent that points which have the **same density** must have the same constant term

$$\boldsymbol{\mu} = \begin{bmatrix} \boldsymbol{\mu}_1 \\ \vdots \\ \boldsymbol{\mu}_d \end{bmatrix}$$

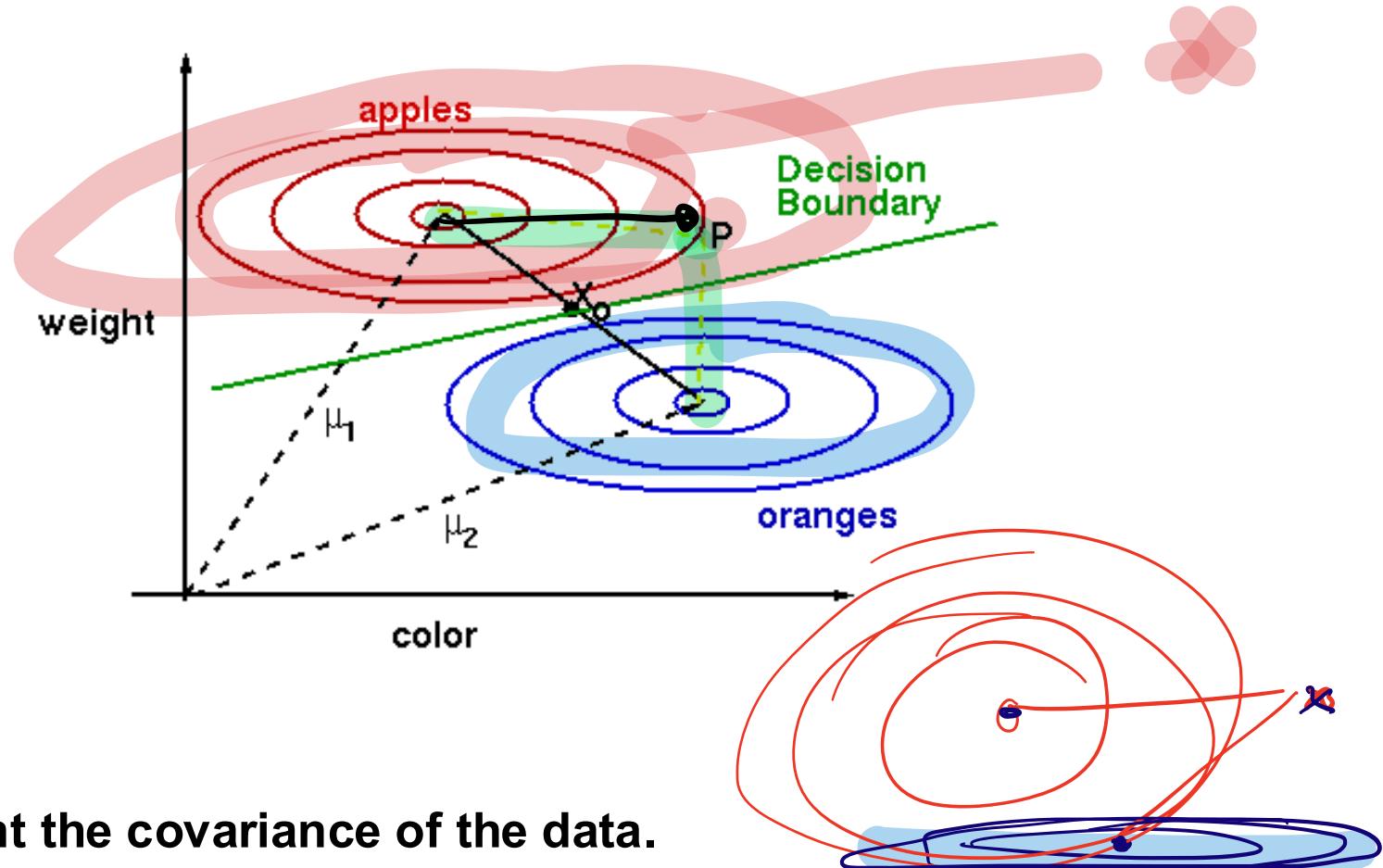
$$(\mathbf{x} - \boldsymbol{\mu})^t \Sigma^{-1} (\mathbf{x} - \boldsymbol{\mu})$$

e.g.

$$\mathbf{x} = \begin{bmatrix} x_1 = 180 \\ \vdots \\ x_d = 70 \end{bmatrix} \quad \mathbf{x} - \boldsymbol{\mu} = \begin{bmatrix} x_1 - \boldsymbol{\mu}_1 \\ \vdots \\ x_d - \boldsymbol{\mu}_d \end{bmatrix}$$

Mahalanobis distance measures the distance from \mathbf{X} to $\boldsymbol{\mu}$ in terms of Σ

Why Mahalanobis Distance



It takes into account the covariance of the data.

- Point P is at closer (Euclidean) to the mean for the orange class, but using the Mahalanobis distance, it is found to be closer to 'apple' class.

Points that are the same distance from μ

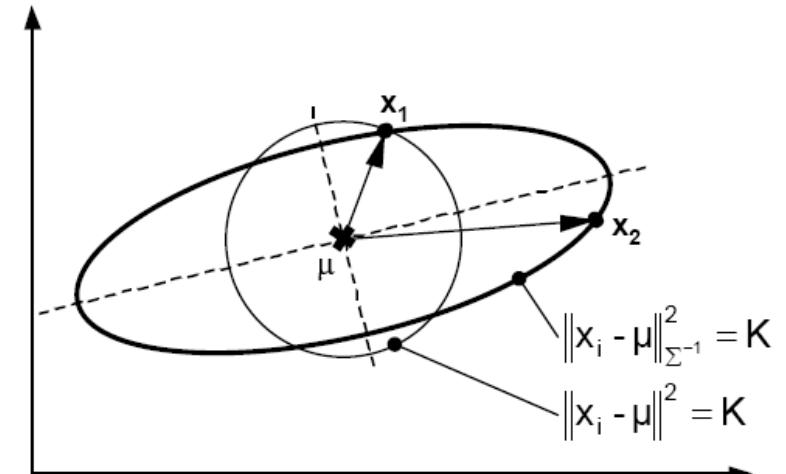
- The quadratic term is called the Mahalanobis distance, a very important distance in Statistical PR

Mahalanobis Distance

$$\|x - y\|_{\Sigma^{-1}}^2 = (x - y)^T \Sigma^{-1} (x - y)$$

- The Mahalanobis distance is a vector distance that uses a Σ^{-1} norm**

- Σ^{-1} can be thought of as a stretching factor on the space
- Note that for an identity covariance matrix ($\Sigma=I$), the Mahalanobis distance becomes the familiar Euclidean distance



- The ellipse consists of points that are equi-distant to the center w.r.t. Mahalanobis distance.
- The circle consists of points that are equi-distant to the center w.r.t. The Euclidian distance.

Overview

■ Gaussian Bayes classifier ...

- is a **Bayesian classifier** so it assigns the input \mathbf{x} to the class C_j for which $P(C_j|\mathbf{x})$ is largest;

- assumes that \mathbf{x} is **multivariate normal**.

$$p(\mathbf{x} | C_j) = N(\mathbf{x} | \mu_j, \Sigma_j)$$

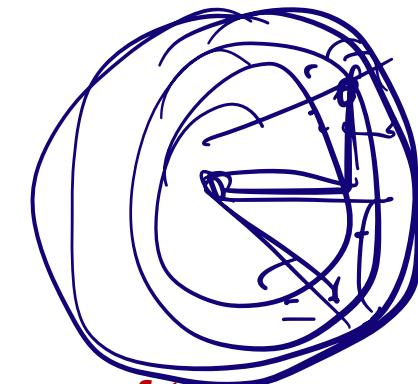
- It obtains $p(\mathbf{x} | C_j)$ by estimating the parameters of the multivariate Normal distribution from the data.

$$\mu_j \text{ and } \Sigma_j$$

- typically using the Maximum Likelihood Estimation

- Uses Bayes rule for decision

$$C_i^* = \underset{C_i}{\operatorname{argmax}} \ p(\mathbf{x}|C_i) \times P(C_i)$$



Example: Gaussian case, μ unknown

- Assume a dataset $X = \{x^{(1)}, x^{(2)}, \dots, x^{(N)}\}$ and a density of the form $p(x) = N(\mu, \sigma)$ where the standard deviation σ is known
- What is the Maximum Likelihood estimate of the mean?

$$\begin{aligned}\theta = \mu \Rightarrow \hat{\theta} &= \operatorname{argmax} \sum_{k=1}^N \log p(x^{(k)} | \theta) \\ &= \operatorname{argmax} \sum_{k=1}^N \log \left(\frac{1}{\sqrt{2\pi}\sigma} \exp \left(-\frac{1}{2\sigma^2} (x^{(k)} - \mu)^2 \right) \right) \\ &= \operatorname{argmax} \sum_{k=1}^N \left\{ \log \left(\frac{1}{\sqrt{2\pi}\sigma} \right) - \frac{1}{2\sigma^2} (x^{(k)} - \mu)^2 \right\}\end{aligned}$$

- The maxima (or minima) of a function are defined by the zeros of its derivative:

$$\frac{\partial \sum_{k=1}^N \log p(x^{(k)} | \theta)}{\partial \theta} = \frac{\partial}{\partial \mu} \sum_{k=1}^N \{\bullet\} = 0 \Rightarrow \boxed{\mu = \frac{1}{n} \sum_{k=1}^N x^{(k)}}$$

- So the ML estimate of the mean is the average value of the training data, a very intuitive result!

Maximum (Log) Likelihood for a 1D-Gaussian

$$\mathcal{N}(x|\mu, \sigma^2) = \frac{1}{(2\pi\sigma^2)^{1/2}} \exp\left\{-\frac{1}{2\sigma^2}(x - \mu)^2\right\}$$

We find that the estimates found by ML estimation, namely μ_{ML} and σ_{ML} are the **sample mean** and **sample variance**.

The actual μ and σ are called **population mean** and **variance**.

When you tell scikit to generate samples from a 1D normal distribution with μ and σ , you give population params; when you estimate them from those samples using the maximum likelihood method, you obtain the sample mean and var – which coincide with the ML estimates for the population params.

$$\mu_{\text{ML}} = \frac{1}{N} \sum_{n=1}^N x_n \quad \sigma_{\text{ML}}^2 = \frac{1}{N} \sum_{n=1}^N (x_n - \mu_{\text{ML}})^2$$

Maximum Likelihood Estimation for Multivariate Normal Distribution

Using the notation in Ethem book:

$$\hat{P}(C_i) = \frac{\sum_t r_i^t}{N}$$

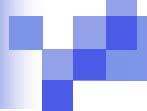
$$\mathbf{m}_i = \frac{\sum_t r_i^t \mathbf{x}^t}{\sum_t r_i^t}$$

$$\mathbf{S}_i = \frac{\sum_t r_i^t (\mathbf{x}^t - \mathbf{m}_i)(\mathbf{x}^t - \mathbf{m}_i)^T}{\sum_t r_i^t}$$

where r_i^t is 1 if the t^{th} sample belongs to class i

Using r_i^t is equivalent to saying:

Take into consideration the sample \mathbf{x}^t only if it belongs to class i



Summary

- Naïve Bayes Classifier with **Categorical Attributes**
 - Estimate $P(x_j | C_i)$ using frequencies and Laplace smoothing
- Naïve Bayes Classifier with a **Single Continuous Attribute x_1**
 - Estimate μ and σ for the single attribute
 - Use $p(x_1 | C_i) = N(x_1 | \mu, \sigma)$ in the Bayes formula
- Naïve Bayes Classifier with **Multiple Continuous Attributes**
 - Estimate μ_i (d-dimensional) and Σ_i (dxd dimensional) for all classes C_i
 - Use $p(\mathbf{x} | C_i) = N(\mathbf{x} | \mu_i, \Sigma_i)$
- Further simplifications tomorrow...