# Lecture 3: Deep generative models

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## Outline

- 1 Introduction
- 2 Variational autoencoder
- 3 Normalizing flow
- 4 Generative adversarial networks
- 5 Diffusion models

#### Introduction

- Suppose that  $\mathbf{X}_1, \ldots, \mathbf{X}_n \stackrel{\text{iid}}{\sim} P_0$ .
  - D-dimensional observations
- The goal is to estimate  $P_0$  or certain functionals of it, such as:
  - The density of  $P_0$
  - The expectation under  $P_0$
  - The support of  $P_0$  (e.g., a manifold)
- Several approaches are available, including:
  - Kernel density estimation
  - Wavelet-based methods
  - Nonparametric Bayesian methods

## Deep generative models

- A deep generative model refers to an indirect method for estimating  $P_0$  (or related quantities), which involves:
  - 1 enabling efficient sampling from an estimated distribution  $\hat{P}$  instead of directly estimating  $P_0$ ,
  - 2 parameterizing unknown functions using DNNs.
- Recently, deep generative models have achieved remarkable success in modeling high-dimensional data.

# Functions parameterized by DNNs

- In deep generative models, the distribution  $P_0$  is parameterized by a function.
- This function is, in turn, parameterized by a DNN.
- Common parameterization approaches include:
  - Generator-based methods
  - Score-based methods
  - Vector fields (defining probability flows)

## Generator-based methods

• One can model X as

$$\mathbf{X} = \mathbf{g}(\mathbf{Z}),$$

#### where

- **Z**: a d-dimensional latent vector with a known distribution  $P_Z$
- **g**: the generator, a map from  $\mathbb{R}^d$  to  $\mathbb{R}^D$
- Two typical regimes:
  - -d < D (low-dimensional latent space)
  - d = D (dimension-preserving mapping)

## Generator-based methods (cont.)

- Given a generator  $\hat{\mathbf{g}}$ , one can generate a sample by
  - 1 drawing **Z** from  $P_Z$ , and
  - 2 computing  $\hat{\mathbf{g}}(\mathbf{Z})$ .
- The distribution of  $\hat{\mathbf{g}}(\mathbf{Z})$  serves as an estimator of  $P_0$ .
- Popular approaches for estimating  $\hat{\mathbf{g}}$  include:
  - Variational autoencoders (VAEs)
  - Normalizing flows (NFs)
  - Generative adversarial networks (GANs)

## Score-based methods

- Suppose that  $P_0$  possesses a Lebesgue density  $p_0$ .
- The function  $\mathbf{s}_0(\mathbf{x}) = \nabla(\log p_0)(\mathbf{x})$  is called the score function of  $p_0$ .
- Score-based methods aim to estimate  $s_0(x)$ .
- Once an estimator  $\hat{\mathbf{s}}$  is obtained, samples can be generated, for example, via Langevin diffusion.

## Score-based methods (cont.)

• Langevin diffusion:

$$egin{aligned} \mathrm{d}\mathbf{X}_t &= rac{1}{2} 
abla (\log p_0)(\mathbf{X}_t) \, \mathrm{d}t + \mathrm{d}\mathbf{B}_t, \ \mathrm{d}\hat{\mathbf{X}}_t &= rac{1}{2} \hat{\mathbf{s}}(\hat{\mathbf{X}}_t) \, \mathrm{d}t + \mathrm{d}\mathbf{B}_t, \end{aligned}$$

where  $\mathbf{B}_t$  denotes a standard Brownian motion.

- The distribution  $\hat{P}$  can be defined as the stationary distribution (limit law) of  $\hat{\mathbf{X}}_t$ .
- One can discretize the estimated Langevin diffusion as

$$\mathbf{X}_i = \mathbf{X}_{i-1} + \frac{\epsilon}{2}\hat{\mathbf{s}}(\mathbf{X}_{i-1}) + \sqrt{\epsilon}\mathbf{Z}_i,$$

where  $\mathbf{Z}_i \sim \mathcal{N}(\mathbf{0}_D, \mathbb{I}_D)$ .

#### Remark

- Compared to score-based methods, generator-based methods are much easier for sample generation.
- On the other hand, learning the generator function is generally more challenging than estimating the score function.
- We first study the generator-based methods, with a particular focus on VAE and GAN.
- Then, we turn to score-based methods.

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#### **Notations**

- Let  $P_{\mathbf{g}}$  denote the law of  $\mathbf{g}(\mathbf{Z})$  (i.e., the pushforward of  $P_{\mathbf{Z}}$ ).
- Let  $P_{\mathbf{g},\sigma}$  be the law of  $\mathbf{g}(\mathbf{Z}) + \epsilon$  with  $\epsilon \sim \mathcal{N}(\mathbf{0}_D, \sigma^2 \mathbb{I}_D)$ , i.e.,

$$P_{\mathbf{g},\sigma} = P_{\mathbf{g}} * \mathcal{N}(\mathbf{0}_D, \sigma^2 \mathbb{I}_D).$$

• For  $\sigma > 0$ ,  $P_{\mathbf{g},\sigma}$  has the density

$$p_{\mathbf{g},\sigma}(\mathbf{x}) = \int \phi_{\sigma}(\mathbf{x} - \mathbf{z}) \, dP_{\mathbf{g}}(\mathbf{z})$$
$$= \int \phi_{\sigma}(\mathbf{x} - \mathbf{g}(\mathbf{z})) \, dP_{Z}(\mathbf{z}),$$

where  $\phi_{\sigma}$  is the density of  $\mathcal{N}(\mathbf{0}_D, \sigma^2 \mathbb{I}_D)$ .

## Variational autoencoders

ullet For a given class  ${\cal G}$  of DNN functions, consider the model class

$$\mathcal{P} = \Big\{ P_{\mathbf{g},\sigma} : \mathbf{g} \in \mathcal{G}, \ \sigma \in [\sigma_{\min}, \sigma_{\max}] \Big\},$$

which is a Gaussian mixture model where the mixing distribution is parametrized by g.

• One may estimate **g** via maximum likelihood:

$$(\hat{\mathbf{g}}, \hat{\sigma}) = \underset{(\mathbf{g}, \sigma): P_{\mathbf{g}, \sigma} \in \mathcal{P}}{\operatorname{argmax}} \sum_{i=1}^{n} \log p_{\mathbf{g}, \sigma}(\mathbf{X}_i).$$

- A variational autoencoder (VAE) is a specific algorithm for approximating the MLE.
  - It employs a variational inference approach.
- In VAE, the log-likelihood is replaced by the evidence lower bound (ELBO),
  - which is computationally more tractable.
- There exist several other computational methods for approximating the MLE,
  - e.g., expectation–maximization (EM) based methods.

Kingma, D. P. & Welling, M. "Auto-encoding variational Bayes". Proc. ICLR. 2014

- Suppose that **g** is parameterized by a DNN, and let  $\theta$  denote the network parameters (possibly including  $\sigma$ ).
- For simplicity, let  $p_{\theta}(\cdot) := p_{\mathbf{g},\sigma}(\cdot)$ .
- Let  $q_{\psi}(\mathbf{z} \mid \mathbf{x})$  be a conditional density parameterized by  $\psi$ , and define

$$\mathcal{L}(\theta, \psi; \mathbf{x}) := \mathbb{E}_{q_{\psi}(\cdot \mid \mathbf{x})} \left[ \log \left( \frac{p_{\theta}(\mathbf{x}, \mathbf{Z})}{q_{\psi}(\mathbf{Z} \mid \mathbf{x})} \right) \right],$$

where  $p_{\theta}(\cdot, \cdot)$  denotes the joint density of  $(\mathbf{X}, \mathbf{Z})$ .

•  $\mathcal{L}(\theta, \psi; \mathbf{x})$  is often referred to as the ELBO (evidence lower bound) because

$$\log p_{\theta}(\mathbf{x}) = \mathcal{L}(\theta, \psi; \mathbf{x}) + K(q_{\psi}(\cdot \mid \mathbf{x}), p_{\theta}(\cdot \mid \mathbf{x}))$$
  
 
$$\geq \mathcal{L}(\theta, \psi; \mathbf{x}),$$

where  $K(\cdot, \cdot)$  denotes the Kullback–Leibler divergence.

- If  $q_{\psi}(\cdot \mid \mathbf{x})$  is sufficiently close to  $p_{\theta}(\cdot \mid \mathbf{x})$ , then the ELBO closely approximates  $\log p_{\theta}(\mathbf{x})$ .
- If the class  $\{q_{\psi}(\cdot \mid \mathbf{x}) : \psi \in \Psi\}$  is rich enough to approximate  $p_{\theta}(\cdot \mid \mathbf{x})$  for all  $\theta \in \Theta$ , then maximizing the ELBO yields an approximate MLE.

An approximate MLE can be obtained by

$$(\hat{\theta}, \hat{\psi}) = \underset{(\theta, \psi) \in \Theta \times \Psi}{\operatorname{argmax}} \sum_{i=1}^{n} \mathcal{L}(\theta, \psi; \mathbf{X}_i).$$

• In VAE, the variational distribution  $q_{\psi}(\cdot \mid \mathbf{x})$  is typically modeled as a Gaussian:

$$q_{\psi}(\mathbf{z} \mid \mathbf{x}) = \phi_{\sigma_{\psi}(\mathbf{x})}(\mathbf{z} - \mu_{\psi}(\mathbf{x})),$$

where  $\phi_{\sigma}$  denotes the density of a Gaussian distribution with standard deviation  $\sigma$ .

- Both  $\mu_{\psi}$  and  $\sigma_{\psi}$  are implemented as DNNs in practice.

# Computation of ELBO

- Gradient-based optimization of the ELBO is not straightforward.
- The key step is computing the gradient  $\nabla_{(\theta,\psi)} \mathcal{L}(\theta,\psi;\mathbf{x})$ .
- Observe that

$$\mathcal{L}(\theta, \psi; \mathbf{x}) = \mathbb{E}_{q_{\psi}(\cdot \mid \mathbf{x})} \log p_{\theta}(\mathbf{x} \mid \cdot) - K(q_{\psi}(\cdot \mid \mathbf{x}), p_{Z}),$$

where  $p_Z$  is the prior density of **Z**.

• When  $p_Z$  is a Gaussian density, the gradient of the KL term with respect to  $\psi$  can be computed analytically.

# Computation of ELBO (cont.)

• Note that

$$\mathbb{E}_{q_{\psi}(\cdot \mid \mathbf{x})} \log p_{\theta}(\mathbf{x} \mid \cdot) = \mathbb{E}_{\mathbf{Y} \sim \mathcal{N}(\mathbf{0}, \mathbb{I})} \log p_{\theta}(\mathbf{x} \mid \mu_{\psi}(\mathbf{x}) + \sigma_{\psi}(\mathbf{x}) \mathbf{Y}),$$

where Y is a standard Gaussian random variable.

• This reparameterization allows us to approximate the gradient of the expectation via Monte Carlo.

#### Remark

- The variational posterior  $q_{\hat{\psi}}(\cdot \mid \mathbf{x})$  can be used for data compression.
- In practice, the design of the network architecture is as important as the optimization algorithm.
  - This aspect is not covered in this lecture.
- Once an estimator  $(\hat{\mathbf{g}}, \hat{\sigma})$  (parametrized by  $\theta$  in previous slides) is obtained, we define the implicit density and distribution estimators as

$$\hat{p} = p_{\hat{\mathbf{g}},\hat{\sigma}}, \quad \hat{P} = P_{\hat{\mathbf{g}}}.$$

 There is a minor notational inconsistency in this lecture regarding the use of uppercase vs lowercase letters.

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# Normalizing flows

• For diffeomorphic **g** with an invertible Jacobian,  $P_{\mathbf{g}}$  admits a density:

$$p_{\mathbf{g}}(\mathbf{x}) = p_{\mathbf{Z}}(\mathbf{g}^{-1}(\mathbf{x}))\det(\nabla \mathbf{g}^{-1}(\mathbf{x})).$$

• For a class G of such generators, the MLE is defined as

$$\hat{\mathbf{g}} = \underset{\mathbf{g} \in \mathcal{G}}{\operatorname{argmax}} \sum_{i=1}^{n} \log p_{\mathbf{g}}(\mathbf{X}_i).$$

• Normalizing flows (NF) refer to methods for modeling  $p_g$  in a computationally tractable way.

## Normalizing flows (cont.)

- For MLE to be computationally tractable, it is essential that  $det(\nabla \mathbf{g}^{-1}(\mathbf{x}))$  be easy to compute.
- This can be achieved by composing simple diffeomorphisms.
- If  $\mathbf{g} = \mathbf{g}_2 \circ \mathbf{g}_1$  with diffeomorphic  $\mathbf{g}_1$  and  $\mathbf{g}_2$ , then

$$\det(\nabla \mathbf{g}^{-1}(\mathbf{x})) = \det(\nabla \mathbf{g}_1^{-1}(\mathbf{g}_2^{-1}(\mathbf{x})))\det(\nabla \mathbf{g}_2^{-1}(\mathbf{x})).$$

# Normalizing flows (cont.)

- Commonly used simple flows include:
  - Radial flows
  - Coupling flows
  - Autoregressive flows
  - Residual flows
- A key advantage of (some) normalizing flows is that they allow explicit evaluation of the density  $q_{\mathbf{g}}(\mathbf{x})$ .

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## Generative adversarial networks

- Let  $\mathbf{Z}_1, \dots, \mathbf{Z}_n$  be an i.i.d. sample from  $P_Z$ .
- For a class  $\mathcal{F}$  of functions from  $\mathbb{R}^D$  to (0,1), generative adversarial networks (GANs) construct  $\hat{\mathbf{g}}$  by solving :

$$\underset{\mathbf{g} \in \mathcal{G}}{\operatorname{minimize}} \sup_{f \in \mathcal{F}} \frac{1}{n} \sum_{i=1}^{n} \left\{ \log f(\mathbf{X}_i) + \log \left( 1 - f(\mathbf{g}(\mathbf{Z}_i)) \right) \right\}.$$

- In the GAN framework, G and F are referred to as the generator and discriminator classes, respectively.
  - In practice, both are parametrized by DNNs.

Goodfellow, I., Pouget-Abadie, J., Mirza, M., Xu, B., Warde-Farley, D., Ozair, S., Courville, A. & Bengio, Y. "Generative adversarial nets". Proc. NIPS. 2014

## Generative adversarial networks (cont.)

• At the population level, GANs aim to solve

$$\begin{aligned} & \underset{\mathbf{g} \in \mathcal{G}}{\text{minimize}} \sup_{f \in \mathcal{F}} \mathbb{E} \left[ \log f(\mathbf{X}) + \log \left( 1 - f(\mathbf{g}(\mathbf{Z})) \right) \right] \\ \iff & \underset{\mathbf{g} \in \mathcal{G}}{\text{minimize}} \sup_{f \in \mathcal{F}} \left\{ \int \log f \, \mathrm{d}P_0 + \int \log (1 - f) \, \mathrm{d}P_\mathbf{g} \right\} \end{aligned}$$

 This objective is closely related to minimizing the Jensen–Shannon divergence:

$$\underset{\mathbf{g} \in \mathcal{G}}{\operatorname{minimize}} \operatorname{JS}(P_0, P_{\mathbf{g}}),$$

where

$$JS(P,Q) = \frac{1}{2} \left\{ K(P, (P+Q)/2) + K(Q, (P+Q)/2) \right\}.$$

#### Variations of GANs

 The adversarial training framework of GANs motivates several variations, which can be interpreted as solving the following problem at the population level:

$$\underset{\mathbf{g}\in\mathcal{G}}{\operatorname{minimize}}\,d(P_0,P_{\mathbf{g}}),$$

where d is a discrepancy measure of the form

$$d(P,Q) = \sup_{f \in \mathcal{F}} \left\{ \int f \, dP - \int h(f) \, dQ \right\},\,$$

with a fixed function  $h : \mathbb{R} \to \mathbb{R}$ .

## f-GAN

• For a convex function  $\phi$ , the f-divergence (also known as the Csiszár divergence or Ali–Silvey distance) is defined as

$$D_{\phi}(P,Q) = \int \phi\left(\frac{\mathrm{d}P}{\mathrm{d}Q}\right) \mathrm{d}Q.$$

- Examples:
  - Kullback–Leibler (KL) divergence (both directions)
  - Total variation distance
  - $\alpha$ -divergence (e.g., squared Hellinger, Pearson  $\chi^2$ )
  - Jensen-Shannon (JS) divergence
- At the population level, f-GAN aims to solve

$$\underset{\mathbf{g}\in\mathcal{G}}{\operatorname{minimize}}\,D_{\phi}(P_0,P_{\mathbf{g}}).$$

## f-GAN (cont.)

• It is well known that for any function class  $\mathcal{F}$ ,

$$D_{\phi}(P,Q) \ge \sup_{f \in \mathcal{F}} \left\{ \int f \, \mathrm{d}P - \int \phi^*(f) \, \mathrm{d}Q \right\},\,$$

where  $\phi^*$  is the convex conjugate of  $\phi$ , defined as

$$\phi^*(t) = \sup_{u \in \text{dom}(\phi)} \{ut - \phi(u)\}.$$

• Equality holds if and only if  $\partial \phi(p/q) \cap \mathcal{F} \neq \emptyset$ .

Nguyen, X., Wainwright, M. J. & Jordan, M. I. "Estimating divergence functionals and the likelihood ratio by convex risk minimization". *IEEE Trans. Inform. Theory.* 2010

## f-GAN (cont.)

• The variational representation of *f*-divergence motivates the *f*-GAN objective:

$$\underset{\mathbf{g} \in \mathcal{G}}{\operatorname{minimize}} \sup_{f \in \mathcal{F}} \frac{1}{n} \sum_{i=1}^{n} \left\{ f(\mathbf{X}_i) - \phi^*(f(\mathbf{g}(\mathbf{Z}_i))) \right\}.$$

• In practice, both  $\mathcal{G}$  and  $\mathcal{F}$  are parametrized by DNNs.

## **IPM GAN**

• For a class  $\mathcal{F}$  of real-valued functions on  $\mathbb{R}^D$ , the  $\mathcal{F}$ -IPM (Integral Probability Metric) is defined as

$$d_{\mathcal{F}}(P,Q) = \sup_{f \in \mathcal{F}} \left| \int f \, dP - \int f \, dQ \right|.$$

• IPM GAN aims to solve

$$\underset{\mathbf{g}\in\mathcal{G}}{\operatorname{minimize}}\,d_{\mathcal{F}}(P_0,P_{\mathbf{g}}).$$

• This leads to the following empirical objective:

$$\underset{\mathbf{g} \in \mathcal{G}}{\operatorname{minimize}} \sup_{f \in \mathcal{F}} \frac{1}{n} \sum_{i=1}^{n} \left\{ f(\mathbf{X}_i) - f(\mathbf{g}(\mathbf{Z}_i)) \right\}.$$

## IPM GAN (cont.)

- Example 1: Wasserstein GAN (WGAN)
  - $\mathcal{F}$  = {f : Lip(f) ≤ 1} (1-Lipschitz functions).
- Example 2: Maximum Mean Discrepancy (MMD) GAN
  - $-\mathcal{F}$  is the unit ball of a reproducing kernel Hilbert space (RKHS).
- In practice, both G and F are parametrized by DNNs.
   However, different choices of F result in different optimization algorithms.
- A key advantage: IPMs are well-defined even when probability densities do not exist.

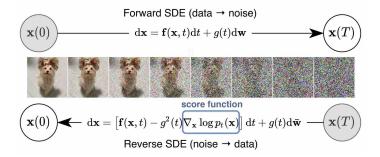
Arjovsky, M., Chintala, S. & Bottou, L. "Wasserstein generative adversarial networks". Proc. ICML. 2017

Dziugaite, G. K., Roy, D. M. & Ghahramani, Z. "Training generative neural networks via maximum mean discrepancy optimization". Proc. Conference on Uncertainty in Artificial Intelligence. 2015

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#### Introduction



Song, Y., Sohl-Dickstein, J., Kingma, D. P., Kumar, A., Ermon, S. & Poole, B. "Score-based generative modeling through stochastic differential equations". *Proc. ICLR*, 2021

# Score matching

- The estimation of the score function dates back to Hyvärinen (2005).
- For a function  $\mathbf{s}: \mathbb{R}^D \to \mathbb{R}^D$ , under mild regularity conditions,

$$\begin{split} \frac{1}{2}\mathbb{E}\Big[\|\boldsymbol{s}(\boldsymbol{X}) - \boldsymbol{s}_0(\boldsymbol{X})\|_2^2\Big] &= \mathbb{E}\left[\text{tr}\big(\nabla \boldsymbol{s}(\boldsymbol{X})\big) + \frac{1}{2}\|\boldsymbol{s}(\boldsymbol{X})\|_2^2\right] \\ &- \frac{1}{2}\mathbb{E}\|\boldsymbol{s}_0(\boldsymbol{X})\|_2^2. \end{split}$$

Hyvärinen, A. "Estimation of non-normalized statistical models by score matching". J. Mach. Learn. Res. 2005

# Score matching (cont.)

• For  $\mathbf{s}: \mathbb{R}^D \to \mathbb{R}^D$ , the score matching loss is defined by

$$\ell_{\mathbf{s}}(\mathbf{x}) = \operatorname{tr}(\nabla \mathbf{s}(\mathbf{x})) + \frac{1}{2} \|\mathbf{s}(\mathbf{x})\|_{2}^{2},$$

which leads to an M-estimator (or ERM):

$$\hat{\mathbf{s}} = \underset{\mathbf{s} \in \mathcal{S}}{\operatorname{argmin}} \int \ell_{\mathbf{s}}(\mathbf{x}) \, \mathbb{P}_{n}(d\mathbf{x})$$
$$= \underset{\mathbf{s} \in \mathcal{S}}{\operatorname{argmin}} \left[ \frac{1}{n} \sum_{i=1}^{n} \ell_{\mathbf{s}}(\mathbf{X}_{i}) \right],$$

where S denotes a class of DNNs.

## Score matching (cont.)

- The computation of  $tr(\nabla \mathbf{s}(\mathbf{x}))$  is not scalable for large D.
- Several scalable alternatives have been proposed:
  - Sliced score matching (Song et al., 2020)
  - Denoising score matching (Vincent, 2011)

Song, Y., Garg, S., Shi, J. & Ermon, S. "Sliced score matching: A scalable approach to density and score estimation". Uncertainty in Artificial Intelligence. 2020

#### Remarks

- In real-world applications, noise is often injected at multiple levels, and the score functions of the resulting noise-perturbed distributions are jointly modeled and estimated.
- A variety of approaches have been proposed, many of which can be understood within a unified framework presented by Song et al. (2021).

Song, Y., Sohl-Dickstein, J., Kingma, D. P., Kumar, A., Ermon, S. & Poole, B. "Score-based generative modeling through stochastic differential equations". *Proc. ICLR*. 2021

# Score-based methods through SDE

• Consider the Ornstein-Uhlenbeck (OU) process

$$\mathrm{d}\mathbf{X}_t = -\frac{1}{2}\mathbf{X}_t\mathrm{d}t + \mathrm{d}\mathbf{B}_t, \quad t \ge 0,$$

with initial distribution  $\mathbf{X}_0 \sim P_0$ .

• For the OU process, the transition kernels are explicitly given by

$$\mathbf{X}_{t+s} \mid \mathbf{X}_s = \mathbf{x} \sim \mathcal{N}\left(\mu_t \mathbf{x}, \sigma_t^2 \mathbb{I}_D\right),$$

where 
$$\mu_t = e^{-t/2}$$
 and  $\sigma_t^2 = 1 - \mu_t^2$ .

• Let  $p_t(\cdot \mid \cdot)$  be the transition kernel density of a diffusion process, and define

$$p_t(\mathbf{x}) = \int p_t(\mathbf{x} \mid \mathbf{X}_0 = \mathbf{x}_0) dP_0(\mathbf{x}_0)$$

as the marginal density of  $X_t$ .

- $p_t$  converges rapidly to the standard Gaussian density as  $t \to \infty$ .
- With slight abuse of notation, define the score function of  $p_t$  by

$$\mathbf{s}_0(t,\mathbf{x}) = \nabla(\log p_t)(\mathbf{x}).$$

- For a large enough T > 0, define the reverse process  $\mathbf{Y}_t := \mathbf{X}_{T-t}$ .
- Under mild regularity conditions, the reverse process satisfies the SDE (Anderson, 1982)

$$d\mathbf{Y}_t = \frac{1}{2}\mathbf{Y}_t dt + \mathbf{s}_0(T - t, \mathbf{Y}_t) dt + d\widetilde{\mathbf{B}}_t.$$

• Once an estimator  $\hat{\mathbf{s}}(\cdot,\cdot)$  for  $\mathbf{s}_0(\cdot,\cdot)$  is obtained, one can simulate the reverse process starting from a standard Gaussian sample.

- The function  $s_0(\cdot, \cdot)$  can be estimated via score matching.
- Alternatively, one may estimate the conditional expectation

$$\mathbf{m}_0(t,\mathbf{x}) = \mathbb{E}[\mathbf{X}_0 \mid \mathbf{X}_t = \mathbf{x}]$$

using (weighted) least squares regression.

 These two approaches are equivalent under reparametrization, since

$$\mathbf{s}_0(t,\mathbf{x}) = -\frac{\mathbf{x} - \mu_t \mathbf{m}_0(t,\mathbf{x})}{\sigma_t^2}.$$

• For any weight function  $\lambda(\cdot)$ ,

$$\mathbf{m}_0 = \underset{\mathbf{m}}{\operatorname{argmin}} \mathbb{E} \left[ \int_0^T \lambda(t) \|\mathbf{X}_0 - \mathbf{m}(t, \mathbf{X}_t)\|^2 dt \right].$$

- By modeling **m** using a DNN, an estimator  $\hat{\mathbf{m}}$  can be constructed via empirical risk minimization.
- Let  $\hat{\mathbf{s}}(\cdot,\cdot)$  be the corresponding estimator of the score function.

• Recall that the reverse process  $\mathbf{Y}_t = \mathbf{X}_{T-t}$  satisfies

$$d\mathbf{Y}_t = \frac{1}{2}\mathbf{Y}_t dt + \mathbf{s}_0(T - t, \mathbf{Y}_t) dt + d\widetilde{\mathbf{B}}_t, \quad \mathbf{Y}_0 \sim P_T.$$

• Define a process  $(\hat{\mathbf{Y}}_t)$  using the estimated score  $\hat{\mathbf{s}}$  as

$$d\hat{\mathbf{Y}}_t = \frac{1}{2}\hat{\mathbf{Y}}_t dt + \hat{\mathbf{s}}(T - t, \hat{\mathbf{Y}}_t) dt + d\widetilde{\mathbf{B}}_t, \quad \hat{\mathbf{Y}}_0 \sim \mathcal{N}(\mathbf{0}_D, \mathbb{I}_D).$$

- Define  $\hat{P}$  (or  $\hat{p}$ ) as the distribution (or density) of  $\hat{\mathbf{Y}}_T$ .
  - Algorithmic issues are not discussed in this talk.

#### Remarks

- Score-based generative models represent the current state of the art in generative modeling.
- Various approaches are available for estimating the score function.
- A key practical concern is the computational cost of sample generation.

# Thank you for attention!