

1 Signal Model

The signal model is as follows:

$$x[t] = A \cos(\omega_0(t - \nu) + \phi) + \epsilon[t] \quad (1)$$

for $t = 0, \dots, T - 1$ where A is the signal amplitude, $\omega_0 \in (0, \pi)$ is the true signal frequency, $\nu = \frac{T-1}{2}$, ϕ is the initial signal phase and is uniformly distributed $[0, 2\pi)$, and ϵ is an independent, identically normally distributed, zero mean process with variance σ_ϵ^2 .

The signal-to-noise ratio is given by

$$\text{SNR} = \frac{A^2}{2\sigma_\epsilon^2} \quad (2)$$

2 Maximum Likelihood Estimation

Much of this follows Kay's **Example 7.16** in [1].

Assuming that the noise ϵ is normally distributed, then the probability density function of

$$\mathbf{x} = [x[0], x[1], \dots, x[T-1]]^T \quad (3)$$

is

$$p(\mathbf{x}; \theta) = \frac{1}{(2\pi\sigma_\epsilon^2)^{T/2}} \exp\left(-\frac{1}{2\sigma_\epsilon^2} \sum_{t=0}^{T-1} (x[t] - \tilde{A} \cos(\tilde{\omega}_0(t - \nu) + \tilde{\phi}))^2\right) \quad (4)$$

where our parameters of interest are $\theta = [\tilde{A}, \tilde{\omega}_0, \tilde{\phi}]^T$.

Then we can take the negative logarithm to get

$$L(\theta) = \frac{T}{2} \log_e(2\pi\sigma_\epsilon^2) + \frac{1}{2\sigma_\epsilon^2} \sum_{t=0}^{T-1} (x[t] - \tilde{A} \cos(\tilde{\omega}_0(t - \nu) + \tilde{\phi}))^2 \quad (5)$$

and the maximum likelihood problem because one of *minimizing* L :

$$\hat{\theta} = [\hat{A}, \hat{\omega}_0, \hat{\phi}]^T = \arg \min_{\theta} L(\theta) \quad (6)$$

2.1 Linearizing

First, Kay [1] uses the trigonometric identity

$$\cos(A + B) = \cos A \cos B - \sin A \sin B \quad (7)$$

to rewrite

$$L(\theta) = \frac{T}{2} \log_e(2\pi\sigma_\epsilon^2) + \frac{1}{2\sigma_\epsilon^2} \sum_{t=0}^{T-1} (x[t] - \tilde{A} \cos(\tilde{\omega}_0(t - \nu)) \cos(\phi) + \tilde{A} \sin(\tilde{\omega}_0(t - \nu)) \sin(\phi))^2 \quad (8)$$

$$L(\theta') = \frac{T}{2} \log_e(2\pi\sigma_\epsilon^2) + \frac{1}{2\sigma_\epsilon^2} \sum_{t=0}^{T-1} (x[t] - \tilde{\alpha}_c \cos(\tilde{\omega}_0(t - \nu)) - \tilde{\alpha}_s \sin(\tilde{\omega}_0(t - \nu)))^2 \quad (9)$$

¹Do we want to add σ_ϵ^2 ?

where

$$\tilde{\alpha}_c = \tilde{A} \cos \phi \quad (10)$$

$$\tilde{\alpha}_s = -\tilde{A} \sin \phi \quad (11)$$

So now we've reparamaterized the problem to use

$$\theta' = [\tilde{\alpha}_c, \tilde{\alpha}_s, \tilde{\omega}_0]^T \quad (12)$$

2.2 Vectorizing

Kay [1] takes (9) and forms the vector version:

$$L(\theta') = \frac{T}{2} \log_e(2\pi\sigma_\epsilon^2) + (\mathbf{x} - \tilde{\alpha}_c \mathbf{c} - \tilde{\alpha}_s \mathbf{s})^T (\mathbf{x} - \tilde{\alpha}_c \mathbf{c} - \tilde{\alpha}_s \mathbf{s}) \quad (13)$$

$$= \frac{T}{2} \log_e(2\pi\sigma_\epsilon^2) + (\mathbf{x} - \mathbf{H}\underline{\alpha})^T (\mathbf{x} - \mathbf{H}\underline{\alpha}) \quad (14)$$

where

$$\mathbf{x} = [x[0], x[1], \dots, x[T-1]]^T \quad (15)$$

$$\mathbf{c} = \left[\cos\left(-\tilde{\omega}_0 \frac{T-1}{2}\right), \dots, \cos\left(-\frac{\tilde{\omega}_0}{2}\right), \cos\left(\frac{\tilde{\omega}_0}{2}\right), \dots, \cos\left(\tilde{\omega}_0 \frac{T-1}{2}\right) \right]^T \quad (16)$$

$$\mathbf{s} = \left[\sin\left(-\tilde{\omega}_0 \frac{T-1}{2}\right), \dots, \sin\left(-\frac{\tilde{\omega}_0}{2}\right), \sin\left(\frac{\tilde{\omega}_0}{2}\right), \dots, \sin\left(\tilde{\omega}_0 \frac{T-1}{2}\right) \right]^T \quad (17)$$

$$\underline{\alpha} = [\tilde{\alpha}_c \ \tilde{\alpha}_s]^T \quad (18)$$

$$\mathbf{H} = [\mathbf{c} \ \mathbf{s}] \quad (19)$$

assuming T is even.

Kay [1] then notes that this is a least squares problem which, for a given \mathbf{H} has a solution for $\hat{\underline{\alpha}}$ of

$$\hat{\underline{\alpha}} = (\mathbf{H}^T \mathbf{H})^{-1} \mathbf{H}^T \mathbf{x} \quad (20)$$

This means

$$L(\theta'|\hat{\underline{\alpha}}) = \frac{T}{2} \log_e(2\pi\sigma_\epsilon^2) + (\mathbf{x} - \mathbf{H}\hat{\underline{\alpha}})^T (\mathbf{x} - \mathbf{H}\hat{\underline{\alpha}}) \quad (21)$$

$$= \frac{T}{2} \log_e(2\pi\sigma_\epsilon^2) + (\mathbf{x} - \mathbf{H}(\mathbf{H}^T \mathbf{H})^{-1} \mathbf{H}^T \mathbf{x})^T (\mathbf{x} - \mathbf{H}(\mathbf{H}^T \mathbf{H})^{-1} \mathbf{H}^T \mathbf{x}) \quad (22)$$

$$= \mathbf{x}^T \left(\mathbf{I} - \mathbf{H}(\mathbf{H}^T \mathbf{H})^{-1} \mathbf{H}^T \right) \mathbf{x} \quad (23)$$

Which means that the maximum likelihood estimator of frequency can be found by maximizing:

$$\mathbf{x}^T \mathbf{H}(\mathbf{H}^T \mathbf{H})^{-1} \mathbf{H}^T \mathbf{x} = \begin{bmatrix} \mathbf{c}^T \mathbf{x} \\ \mathbf{s}^T \mathbf{x} \end{bmatrix}^T \begin{bmatrix} \mathbf{c}^T \mathbf{c} & \mathbf{c}^T \mathbf{s} \\ \mathbf{s}^T \mathbf{c} & \mathbf{s}^T \mathbf{s} \end{bmatrix} \begin{bmatrix} \mathbf{c}^T \mathbf{x} \\ \mathbf{s}^T \mathbf{x} \end{bmatrix} \quad (24)$$

3 What does this look like?

Kay [1] states near equation (7.66) that when ω_0 is not close to 0 or π and when $\mathbf{c}^T \mathbf{c}/N \approx \frac{1}{2}$ and $\mathbf{s}^T \mathbf{s}/N \approx \frac{1}{2}$ then the MLE of frequency is approximately the same as the periodogram maximizer:

$$\hat{\omega}_0 = \arg \max_{\tilde{\omega}_0} \left| \sum_{t=0}^{T-1} x[t] \exp(-j\tilde{\omega}_0 t) \right|^2 \quad (25)$$

So it might be interesting to plot (24) and (25) on the same scale. Figure 1 shows this for the signal

$$x[t] = \cos(2\pi 0.0123456789t + \pi/4) + \cos(2\pi(0.0123456789 + 1/200)t + \pi/4) + \epsilon[t]; \quad (26)$$

where $T = 32$, and $\epsilon[t]$ is independent and identically normally distributed with zero mean and standard deviation 0.5.

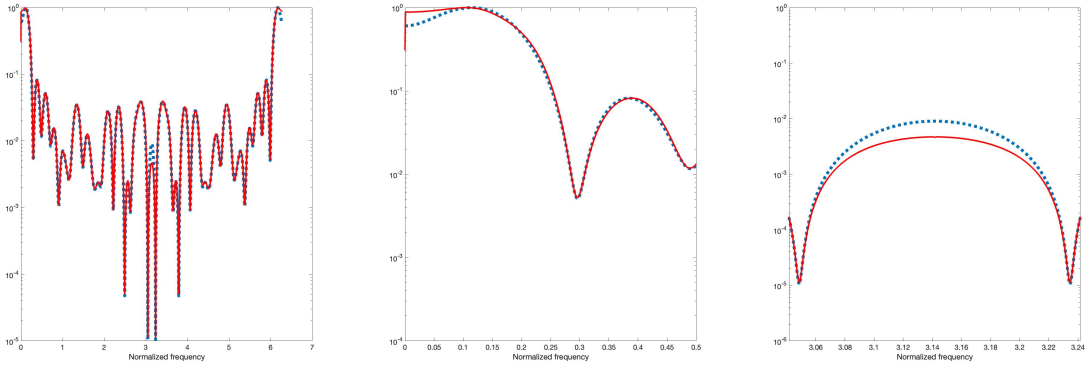


Figure 1: Full frequency range comparison, zoom in around 0, zoom in around π .

References

- [1] S. M. Kay, *Fundamentals of Statistical Signal Processing: Estimation Theory*. Prentice Hall, 1997.