

Priors

$$\mu \sim N(\mu, \nu)$$

$$1. \quad i=1, 2, \dots, k$$

$$\lambda \sim \text{Gamma}(a, b)$$

$$j=1, 2, \dots, n_i$$

$$\psi \sim \text{Gamma}(c, d)$$

$$\lambda = \frac{1}{6^2} \quad \psi = \frac{1}{\tau^2}$$

$$p(\mu | \lambda, \psi, \theta | \gamma) \propto \prod_{i=1}^k \prod_{j=1}^{n_i} p(y_{ij} | \theta_i, \lambda) \times \prod_{i=1}^k p(\theta_i | \mu, \lambda, \psi) \times p(\mu) p(\lambda) p(\psi)$$

$$\propto \prod_{i=1}^k \prod_{j=1}^{n_i} \lambda^{\frac{1}{2}} e^{-\frac{\lambda(y_{ij} - \theta_i)^2}{2}} \times \prod_{i=1}^k \lambda^{\frac{1}{2}} \psi^{\frac{1}{2}} e^{-\frac{\lambda \psi (\theta_i - \mu)^2}{2}} \times \frac{-\phi(\mu - \nu)^2}{e^{\frac{\phi(\mu - \nu)^2}{2}}} \times e^{\alpha_1 - \lambda b} \times \psi^{c-1} e^{-\psi d}$$

$$p(\mu | \cdot) \propto \prod_{i=1}^k \lambda^{\frac{1}{2}} \psi^{\frac{1}{2}} e^{-\frac{\lambda \psi (\theta_i - \mu)^2}{2}} \times \frac{-\phi(\mu - \nu)^2}{e^{\frac{\phi(\mu - \nu)^2}{2}}}$$

$$\propto e^{-\lambda \psi \sum_i (\theta_i - \mu)^2} \times \frac{-\phi(\mu - \nu)^2}{e^{\frac{\phi(\mu - \nu)^2}{2}}}$$

$$\propto N((K \lambda \psi + \phi)^{-1} (\lambda \psi \sum_i \theta_i + \phi \nu), (\lambda \psi \sum_i \theta_i + \phi \nu)^{-1}) \frac{\mu^2 (K \lambda \psi + \phi)}{-2 \mu (\lambda \psi \sum_i \theta_i + \phi \nu)}$$

$$p(\lambda|-\) \propto \prod_{i=1}^k \prod_{j=1}^{n_i} \frac{1}{\lambda^{\frac{n_i}{2}}} e^{-\frac{\lambda(y_{ij}-\theta_i)^2}{2}} \times \prod_{i=1}^k \frac{1}{\lambda^{\frac{1}{2}}} \psi^{\frac{1}{2}} e^{-\frac{\lambda \psi (\theta_i - \mu)^2}{2}} \times \lambda^{a-1} e^{-\lambda b}$$

$$= \lambda^{\sum n_i + k + a - 1} e^{-\lambda \left(\sum_j \frac{(y_{ij} - \theta_i)^2}{2} + \psi \sum_i \frac{(\theta_i - \mu)^2}{2} + b \right)}$$

$$\propto \text{Gamma} \left(\sum n_i + k + a, \sum_j \frac{(y_{ij} - \theta_i)^2}{2} + \psi \sum_i \frac{(\theta_i - \mu)^2}{2} + b \right)$$

$$p(\gamma|-) \propto \prod_{i=1}^k \frac{1}{\lambda^{\frac{1}{2}}} \psi^{\frac{1}{2}} e^{-\frac{\lambda \psi (\theta_i - \mu)^2}{2}} \psi^{c-1} e^{-\psi d}$$

$$\propto \psi^{\frac{k}{2} + c - 1} e^{-\psi \left(\lambda \sum_i \frac{(\theta_i - \mu)^2}{2} + d \right)}$$

$$\propto \text{Gamma} \left(\frac{k}{2} + c, \lambda \sum_i \frac{(\theta_i - \mu)^2}{2} + d \right)$$

$$p(\theta_m|-) \propto \prod_{i=1}^k \prod_{j=1}^{n_i} \frac{1}{\lambda^{\frac{n_i}{2}}} e^{-\frac{\lambda(y_{ij}-\theta_i)^2}{2}} \times \prod_{i=1}^k \frac{1}{\lambda^{\frac{1}{2}}} \psi^{\frac{1}{2}} e^{-\frac{\lambda \psi (\theta_i - \mu)^2}{2}}$$

$$\propto e^{-\lambda \sum_j \frac{(y_{mj} - \theta_m)^2}{2}} \times e^{-\lambda \frac{\psi (\theta_m - \mu)^2}{2}} - 2\theta_m \lambda \psi$$

$$\theta_m^2 (n_i \lambda + \lambda \psi) - 2\theta_m (\lambda \sum_j y_{mj} + \lambda \psi \mu)$$

$$\propto N((n_i \lambda + \lambda \psi)^{-1} (\lambda \sum_j y_{mj} + \lambda \psi \mu), (n_i \lambda + \lambda \psi)^{-1})$$

$$1. \quad i = 1, 2, \dots, K$$

$$j = 1, 2, \dots, n_i$$

$$p = 1, 2$$

$X_{i,n_i \times p}$ = design matrix of size
 $n_i \times p$

$$\tilde{Y}_i = \log Y_i$$

$$\beta_{i,p \times 1} \\ 1 \times p$$

$$\log Y_{ij} = \beta_{i0} + \beta_{i1} X_{i1} + \beta_{i2} X_{i2} + \beta_{i3} \log(X_{i1}) \times X_{i2}$$

$$\tilde{Y}_i = X_i \beta_i + \varepsilon_i$$

$$\theta \sim MVN(\mu_0, \Lambda_0^{-1})$$

$$\Sigma \sim InvWish(\eta_0, S_0)$$

$$\lambda \sim \text{Gamma}(a_0, b_0)$$

$$\beta_i \sim MVN(\theta, \Sigma)$$

$$\tilde{Y}_i | \beta_i, \lambda \sim MVN(X_i \beta_i, \lambda I_{n_i \times n_i})$$

$$p(\tilde{Y}_i, \beta_i, \lambda, \theta, \Sigma) \propto \prod_{i=1}^K [p(\tilde{Y}_i | \beta_i, \lambda) p(\beta_i | \theta, \Sigma)] \times p(\theta) \times p(\Sigma) \times p(\lambda)$$

$$\propto \left[\prod_{i=1}^K \frac{1}{2\pi} \frac{e^{-\frac{1}{2}(\tilde{Y}_i - X_i \beta_i)^T (\tilde{Y}_i - X_i \beta_i)}}{\lambda^{\frac{n_i}{2}}} \right] \times \frac{1}{2\pi} e^{-\frac{1}{2}(\beta_i - \theta)^T \Sigma^{-1} (\beta_i - \theta)}$$

$$\times e^{-\frac{1}{2}(\theta - \mu_0)^T \Lambda_0 (\theta - \mu_0)} \times \frac{1}{2\pi} e^{-\frac{1}{2}(\eta_0 + p + 1) \Sigma^{-1}} e^{-\frac{1}{2} \text{tr}(S_0 \Sigma^{-1})}$$

$$\times \lambda^{a_0 - 1} e^{-\lambda b_0}$$

$$p(\beta_i | \cdot) \propto p(\tilde{y}_i | \beta_i, \alpha) p(\beta_i | \theta, \Sigma)$$

$$\propto e^{-\frac{\sum_i (\tilde{y}_i - x_i^\top \beta_i)^2}{2} + \frac{(\beta_i - \theta)^\top \Sigma^{-1} (\beta_i - \theta)}{2}}$$

$$\beta_i^\top (\lambda x_i^\top x_i + \Sigma^{-1}) \beta_i - 2 \beta_i^\top (\lambda x_i^\top \tilde{y}_i + \Sigma^{-1} \theta)$$

$$\propto MVN((\lambda x_i^\top x_i + \Sigma^{-1})^{-1} (\lambda x_i^\top \tilde{y}_i + \Sigma^{-1} \theta), (\lambda x_i^\top x_i + \Sigma^{-1})^{-1})$$

$$p(\theta | \cdot) \propto \prod_{i=1}^k p(\beta_i | \theta, \Sigma) \times p(\theta)$$

$$\propto e^{-\frac{\sum_i (\beta_i - \theta)^\top \Sigma^{-1} (\beta_i - \theta)}{2} - \frac{(\theta - \mu_0)^\top \Lambda_0^{-1} (\theta - \mu_0)}{2}}$$

$$\theta^\top (K\Sigma^{-1} + \Lambda_0^{-1}) \theta - 2\theta^\top (\Sigma^{-1} \sum_i \beta_i + \Lambda_0^{-1} \mu_0)$$

$$\propto MVN((K\Sigma^{-1} + \Lambda_0^{-1})^{-1} (\sum_i \beta_i + \Lambda_0^{-1} \mu_0), (K\Sigma^{-1} + \Lambda_0^{-1})^{-1})$$

$$p(\Sigma | \cdot) \propto \prod_{i=1}^k p(\beta_i | \theta, \Sigma) \times p(\Sigma)$$

$$\propto |\Sigma|^{-\frac{k}{2}} e^{-\frac{\sum_i (\beta_i - \theta)^\top \Sigma^{-1} (\beta_i - \theta)}{2} - \frac{(\eta_0 + p + 1)}{2} - \frac{1}{2} \text{tr}(S_0 \Sigma^{-1})}$$

$$\propto |\Sigma|^{-\frac{[(k + \eta_0) + p + 1]}{2}} e^{-\frac{\sum_i (\beta_i - \theta)^\top (\beta_i - \theta)^\top \Sigma^{-1}}{2} - \frac{1}{2} \text{tr}(S_0 \Sigma^{-1})}$$

$$= |\Sigma|^{-\frac{[(k + \eta_0) + p + 1]}{2}} e^{-\frac{\sum_i (\beta_i - \theta)^\top (\beta_i - \theta)^\top \Sigma^{-1} + S_0}{2}}$$

$$\propto \text{InvWish}(k + \eta_0, \sum_i (\beta_i - \theta)^\top (\beta_i - \theta)^\top + S_0)$$

$$p(\lambda|I) \propto \prod_{i=1}^K p(\tilde{y}_i|\beta_i, \lambda) \times p(\lambda)$$

$$\propto \frac{\kappa}{\pi} \frac{1}{|\lambda|} \frac{e^{-\lambda(\tilde{y}_i - x_i \beta_i)^T (\tilde{y}_i - x_i \beta_i)}}{e^{a_0 - \lambda b_0}}$$

$$\propto \frac{\kappa}{\lambda} \frac{\sum_{i=1}^K e^{a_0 - \lambda(\tilde{y}_i - x_i \beta_i)^T (\tilde{y}_i - x_i \beta_i) + b_0}}{\lambda}$$

$$\propto \text{Gamma}\left(\frac{\sum_{i=1}^K e^{a_0 - \lambda(\tilde{y}_i - x_i \beta_i)^T (\tilde{y}_i - x_i \beta_i) + b_0}}{\lambda}, \frac{\kappa}{\lambda} \right)$$

Model

$$\theta \sim MVN(\mu_0, L_0)$$

$$\Sigma \sim InvWish(\eta_0, S_0)$$

$$\beta_i | \theta, \Sigma \sim MVN(\theta, \Sigma)$$

$$z_i | \beta_i \sim MVN(X_i \beta_i, I_{n_i})$$

$$Y_{ij} | z_{ij} \sim I(Y_{ij}=1)I(z_{ij} > 0) + I(Y_{ij}=0)I(z_{ij} \leq 0)$$

$$p(z, \beta, \theta, \Sigma | y) \propto p(y | z, \beta, \theta, \Sigma)$$

$$= \left[\prod_{i=1}^k \prod_{j=1}^{n_i} p(Y_{ij} | z_{ij}) \right] \times \prod_{i=1}^k p(z_i | \beta_i) \times \prod_{i=1}^k p(\beta_i | \theta, \Sigma) \times p(\theta) \times p(\Sigma)$$

Gibbs Sampling

$$p(z_{ij} | \cdot) \propto p(Y_{ij} | z_{ij}) p(z_{ij} | \beta_i)$$

$$\propto [I(Y_{ij}=1)I(z_{ij} > 0) + I(Y_{ij}=0)I(z_{ij} \leq 0)] N(z_{ij} | x_i^\top \beta_i, 1)$$

$$\propto \begin{cases} TruncN(x_{ij}^\top \beta_i, 1, 0, \infty) & \text{if } Y_{ij} = 1 \\ TruncN(x_{ij}^\top \beta_i, 1, -\infty, 0) & \text{if } Y_{ij} = 0 \end{cases}$$

$$\begin{aligned}
p(\beta_i | -) &\propto \prod_{i=1}^k p(z_i | \beta_i) \times \prod_{i=1}^k p(\beta_i | \theta, \Sigma) \\
&\propto e^{-\frac{(z_i - x_i^\top \beta_i)^T (z_i - x_i^\top \beta_i)}{2}} \times e^{-\frac{(\beta_i - \theta)^\top \Sigma^{-1} (\beta_i - \theta)}{2}} \\
&\propto \text{MVN}((x_i^\top x_i + \Sigma^{-1})^{-1} (x_i^\top z_i + \Sigma^{-1} \theta), (x_i^\top x_i + \Sigma^{-1})^{-1})
\end{aligned}$$

$$\begin{aligned}
p(\theta | -) &\propto \prod_{i=1}^k p(\beta_i | \theta, \Sigma) \times p(\theta) \\
&\propto \prod_{i=1}^k e^{-\frac{(\beta_i - \theta)^\top \Sigma^{-1} (\beta_i - \theta)}{2}} \times e^{-\frac{(\theta - \mu_0)^\top L_0^{-1} (\theta - \mu_0)}{2}}
\end{aligned}$$

$$\begin{aligned}
&\theta^\top (K\Sigma^{-1} + L_0^{-1}) \theta - 2\theta^\top (\sum_{i=1}^k \beta_i + L_0^{-1} \mu_0) \\
&\propto \text{MVN}((K\Sigma^{-1} + L_0^{-1})^{-1}, \sum_{i=1}^k \beta_i + L_0^{-1} \mu_0, (K\Sigma^{-1} + L_0^{-1})^{-1})
\end{aligned}$$

$$\begin{aligned}
p(\Sigma | -) &\propto \prod_{i=1}^k p(\beta_i | \theta, \Sigma) \times p(\Sigma) \\
&\propto \prod_{i=1}^k |\Sigma|^{-\frac{1}{2}} e^{-\frac{(\beta_i - \theta)^\top \Sigma^{-1} (\beta_i - \theta)}{2}} \times |\Sigma|^{\frac{-(n_0 + p + 1)}{2}} e^{-\frac{1}{2} \text{tr}(S_0 \Sigma^{-1})} \\
&\propto |\Sigma|^{-\frac{1}{2} \frac{(K + n_0 + p + 1)}{2}} e^{-\text{tr}((\sum_{i=1}^k (\beta_i - \theta) (\beta_i - \theta)^\top + S_0) \Sigma^{-1})} \\
&\propto \text{InvWish}(\ k + n_0, (\sum_{i=1}^k (\beta_i - \theta) (\beta_i - \theta)^\top + S_0))
\end{aligned}$$

$$\frac{\sum (x_i - \bar{x})^2}{(n-1)}$$

$$\frac{\sum x_i^2 - \bar{x}^2}{(n-1)}$$

$$\frac{\sum x_i^2}{n} - \bar{x}^2$$

$$\sum_{i=1}^n \left(\frac{\sum x_i^2}{n} - \bar{x}^2 \right)$$

$$X \rightarrow E(X) \rightarrow \bar{X}^2 \rightarrow E(\bar{X}^2)$$

$$\frac{\sum x^2}{n} \rightarrow E(X^2)$$

$$\lim_{n \rightarrow \infty} P(|Y_n - \theta| > \epsilon) = 0$$

$$\frac{\sum x_i^2}{n} - \bar{x}^2 \rightarrow E(X^2) - E(X)^2 = \sigma^2$$

$$Y_n = \max \{x_1, \dots, x_n\}$$

$$\frac{n}{n-1} \rightarrow 1$$

$$P(Y_n \leq y) = P(X_1 \leq y, X_2 \leq y, \dots, X_n \leq y)$$

$$= P(X_1 \leq y)^n$$

$$= F_{X_1}(y)^n$$

$$F_{X_1}(y) = \begin{cases} 0 & y < 0 \\ \frac{y}{\theta} & 0 < y \leq \theta \\ 1 & y > \theta \end{cases}$$

$$F_Y(y) = \begin{cases} 0 & y < 0 \\ \left(\frac{y}{\theta}\right)^n & 0 < y \leq \theta \\ 1 & y > \theta \end{cases}$$

$$\begin{aligned} P(\theta - \varepsilon \leq Y_n \leq \theta + \varepsilon) &= F_Y(\theta + \varepsilon) - F_Y(\theta - \varepsilon) \\ &= 1 - \left(\frac{\theta - \varepsilon}{\theta}\right)^n \quad \varepsilon < \theta \\ &\rightarrow 1 \end{aligned}$$

$$P(\theta - \varepsilon \leq Y_n \leq \theta + \varepsilon) \approx 1 \quad \varepsilon > \theta$$

$$P(|Y_n - \theta| > \varepsilon)$$

$$P(Y_n - \theta \geq \varepsilon) + P(Y_n - \theta \leq -\varepsilon)$$

$$1 - F_Y(\theta + \varepsilon) + P(Y_n \leq \theta - \varepsilon)$$

$$1 - F_Y(\theta + \varepsilon) + F_{Y_n}(\theta - \varepsilon)$$

$$1 - 1 + F_{Y_n}(\theta - \varepsilon)$$

$$f(x) = \begin{cases} e^{-(x-\theta)} & x > \theta \\ 0 & 0 \leq x \end{cases}$$

$$F(x) = \begin{cases} 0 & x \leq \theta \\ 1 - e^{-(x-\theta)} & \theta < x \end{cases}$$

$$\int_{\theta}^{\infty} e^{-(t-\theta)} dt$$

$$Y_1 = \min \{X_1, \dots, X_n\}$$

$$\int_{\theta}^{x-\theta} e^{-u} du \quad u = t-\theta \quad du = dt$$

$$P(Y_1 \leq y) = 1 - P(Y_1 > y)$$

$$= 1 - P(X_1 > y, \dots, X_n > y)$$

$$= 1 - \prod_{i=1}^n P(X_i > y)$$

$$= 1 - P(X_1 > y)^n$$

$$= 1 - [1 - F_X(y)]^n$$

$$(-e^{-u}) \Big|_0^{x-\theta}$$

$$-e^{-(x-\theta)} + 1$$

$$1 - e^{-n(y-\theta)}$$

$$P(Y_1 \leq y) = \begin{cases} 0 & y \leq \theta \\ 1 - e^{-n(y-\theta)} & y > \theta \end{cases}$$

$$F_Y(\theta + \varepsilon) - F_Y(\theta - \varepsilon)$$

$$1 - e^{-n\varepsilon} - 0$$

$$\left[\frac{s_1^2 + s_2^2}{n_1 + n_2} \right]$$

$$\left[\frac{n_2 s_1^2 + n_1 s_2^2}{n_2 b_1^2 + n_1 b_2^2} \right] \quad s_1^2 \xrightarrow{P} b_1^2 \\ s_2^2 \xrightarrow{P} b_2^2$$

$$F_{X_n}(x) = \begin{cases} 1 & x \geq \frac{1}{n} \\ 0 & x < \frac{1}{n} \end{cases}$$

$$P(|X_n - X| \geq \varepsilon)$$

$$P(|-2X| \geq \varepsilon)$$

$$P(|X| \geq \xi)$$

$$x_n - \varepsilon \leq x \leq x_n + \varepsilon \\ x_n \leq x + \varepsilon$$

$$x_n \leq x + \varepsilon$$

$$\begin{aligned} & \left[(1-p) + p e^+ \right]^n \\ & \left[1 - \frac{m}{n} + \frac{m e^+}{n} \right]^n \\ & \left[1 + \frac{n(e^+ - 1)}{n} \right]^n \end{aligned}$$

$$\sum_{k=0}^{\infty} \frac{\lambda^k e^{-\lambda}}{k!}$$

$$e^{-\lambda} \sum_{k=0}^{\infty} \frac{(e^+ \lambda)^k}{k!}$$

$$\begin{matrix} e^{-\lambda} & e^{+\lambda} \\ e^- & e^+ \\ \hline e & e \end{matrix}$$

$$n(e^+ - 1)$$

$$\frac{1}{\Gamma(p)} \left(\frac{1}{2} \right)^p \int_0^\infty e^{-x} x^{p-1} e^{-\frac{x}{2}} dx$$

$$\begin{aligned} & \frac{1}{\Gamma(p)} \frac{1}{2^p} \int_0^\infty x^{p-1} e^{-x} \left(\frac{1}{2} - \frac{1}{x} \right)^p dx \\ & \frac{1}{\Gamma(p)} \frac{\Gamma(p)}{\left(\frac{1}{2} - 1 \right)^p} \end{aligned}$$

$$\frac{1}{2} \rightarrow 0$$

$$\overbrace{(1-\frac{t}{\sqrt{n}}+1)^{\frac{p}{n}}}^{\int} > +$$

$$M_Y(t) = E(e^{tY}) = E(e^{t(\frac{Z-n}{\sqrt{n}})})$$

$$\approx E(e^{\frac{tZ}{\sqrt{n}}} e^{-\frac{tn}{\sqrt{n}}})$$

$$\approx e^{-\frac{tn}{\sqrt{n}}} E(e^{\frac{tZ}{\sqrt{n}}})$$

$$\approx e^{-\frac{tn}{\sqrt{n}}} \underbrace{(1-\sqrt{n}+1)}_{\sqrt{n}}^{-\frac{n}{2}}$$

$$\approx \left(e^{\frac{tn}{\sqrt{n}}} \underbrace{(1-\sqrt{n}+1)}_{\sqrt{n}} \right)^{-\frac{n}{2}}$$

$$\approx 1 \left(1 + \frac{\sqrt{n}t}{\sqrt{n}} + \frac{t^2}{n} \right) \underbrace{\left(1 - \frac{\sqrt{n}t}{\sqrt{n}} + 1 \right)}_{\sqrt{n}}^{-\frac{n}{2}}$$

$$\approx 1 - \cancel{\frac{\sqrt{n}t}{\sqrt{n}}} + \cancel{\frac{\sqrt{n}t}{\sqrt{n}}} \cdot -\frac{2t^2}{n} + \frac{t^2}{n} - \frac{\sqrt{n}t^3}{n^2}$$

$$\approx \left[1 - \frac{t^2}{n} + \Psi(n) \right]^{-\frac{n}{2}}$$

$$\lim_{n \rightarrow \infty} M_Y(t) \approx e^{\frac{t^2}{2}(1-\frac{1}{2})} = e^{\frac{t^2}{2}}$$

$$\overline{z_n}$$

$$\frac{\bar{z} - n}{\sqrt{n}}$$

$$\sqrt{n} \left(\frac{\bar{z} - n}{\sqrt{n}} \right)$$

$$\sqrt{n} \left(\frac{\bar{z}}{\sqrt{n}} - \frac{1}{\sqrt{n}} \right) \rightarrow N(0, 1)$$

$$g'(t) = \frac{1}{2(t^{\frac{1}{4}})}$$

$$\sqrt{n} \left(\sqrt{\frac{\bar{z}}{\sqrt{n}}} - \sqrt{\frac{1}{\sqrt{n}}} \right) \rightarrow N(0, i^{\frac{3}{2}}) \quad g'(i^{\frac{1}{2}}) = \frac{1}{2(i^{\frac{1}{4}})}$$

$$(g')^2 = i^{-\frac{3}{4}} = \frac{1}{i^{\frac{3}{4}}}$$

$$\sqrt{n} \left(\frac{\bar{z}}{\sqrt{n}} - \frac{1}{\sqrt{n}} \right) \rightarrow N(0, 1)$$

$$\sqrt{n} \left(\sqrt{\frac{\bar{z}}{\sqrt{n}}} - \sqrt{\frac{1}{\sqrt{n}}} \right) \rightarrow N(0, i^{\frac{3}{2}})$$

$$g(t) = \sqrt{t}$$

$$g'(t) = \frac{1}{2\sqrt{t}}$$

$$g'(i^{\frac{1}{2}}) = \frac{1}{2(i^{\frac{1}{4}})}$$

$$= \frac{1}{i^{\frac{3}{4}}} = i^{\frac{3}{4}}$$

$$[g'(i^{\frac{1}{2}})]^2 = i^{\frac{3}{2}}$$

$$n \left[\frac{Y_n}{n} + (1 - \frac{Y_n}{n}) \right]$$

$$\frac{Y_n - np}{\sqrt{np(1-p)}} \sim N(0, 1)$$

$$\frac{\sum_{i=1}^n I(Y_i \leq f(x_i))}{n}$$

WLLN

$$\rightarrow E(I(Y_i \leq f(x_i)))$$

$$= P(Y \leq f(X))$$

$$= \int_0^1 \int_0^{f(x)} dy dx$$

$$= \int_0^1 f(x) dx$$

$$E_{x,y} \{ g(x,y) \}$$

$$E_x E_{y|x} \{ g(x,y) \}$$

$$\int \int f_{x,y}(x,y) g(x,y) dx dy$$

$$\int f_x \int f_{y|x}(y|x) g(x,y) dy dx$$

$$E(Y)$$

$$E_y E(x|y)$$

$$\begin{aligned} & \frac{1}{n^2} \operatorname{Var}\left(\sum_{i=1}^n I(y_i < f(x_i))\right) \\ &= \frac{1}{n^2} \sum_{i=1}^n \operatorname{Var}(I(y_i < f(x_i))) \end{aligned}$$

$$\begin{aligned} & \operatorname{Var}(I(y < f(x))) \\ &= P(Y < f(x))(1 - P(Y < f(x))) \\ &= \left(\int_0^1 f(x) dx\right) \left(1 - \int_0^1 f(x) dx\right) \end{aligned}$$

$$\begin{aligned} & \frac{1}{n^2} \operatorname{Var}\left(\sum_{i=1}^n f(x_i)\right) \\ &= \int_0^1 f(x) dx - \left(\int_0^1 f(x) dx\right)^2 \end{aligned}$$

$$\operatorname{Var}(f(x_i)) = E(f(x_i))^2 - E(f(x_i))^2$$

$$= \int_0^1 f(x)^2 dx - \left(\int_0^1 f(x) dx\right)^2$$

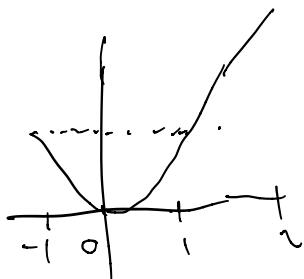
$$f(x)^2 < f(x)$$

$$E[Y] = E[E[Y|X]]$$

$$\int \int y f_{Y|X}(y|x) dy f_X(x) dx$$

$$\int_Y \int f_{Y|X}(y|x) dx dy$$

$$\int y f_Y(y) dy$$



$$f(x) = \begin{cases} \frac{1}{3} & -1 \leq x < 2 \\ 0 & \text{o.w.} \end{cases}$$

$$-1 < X < 0$$

$$0 < X < 1$$

$$1 < X < 2$$

$$u = x^2$$

$$x = \sqrt{u}$$

$$|x| = \sqrt{u}$$

$$x = -\sqrt{u}$$

$$\frac{dx}{du} = \frac{1}{2\sqrt{u}}$$

$$\frac{dx}{du} = -\frac{1}{2\sqrt{u}}$$

$$f_u(u) = \frac{1}{3} \frac{1}{\sqrt{u}} + \frac{1}{2\sqrt{u}} \frac{1}{3}$$

$$= \frac{1}{3\sqrt{u}} \quad 0 < u < 1$$

$$f_u(u) = \frac{1}{6\sqrt{u}} \quad 1 < u < 4$$

$$\int_0^1 \frac{1}{3\sqrt{u}} du + \int_1^4 \frac{1}{6\sqrt{u}} du$$

$$\left(\frac{2u^{\frac{1}{2}}}{3} \Big|_0^1 + \left(\frac{1}{3}u^{\frac{1}{2}} \right)_1^4 \right)$$

$$\frac{2}{3} + \left(\frac{2}{3} - \frac{1}{3} \right)$$

1

$$P(X=x) = \frac{1}{3} \quad x = -1, 1, 0$$

$$Y = X^2$$

$$Y = \begin{cases} 1 & \text{P.P. } \frac{1}{3} + \frac{1}{3} = \frac{2}{3} \\ 0 & \cdot \frac{1}{3} \end{cases}$$

$$\operatorname{Cov}(Y, X)$$

$$\operatorname{Cov}(X, X^2)$$

$$= E(X^3) - E(X)E(X^2)$$

$$= 0 - 0$$

$$-\frac{1}{3} + \frac{1}{3}$$

$$\operatorname{Cov}(X_1, X_2)$$

$$\operatorname{Cov}(Z, Z^2)$$

$$E(Z^3) - E(Z)E(Z^2)$$

$$P(Z \leq t, Z^2 \leq s)$$

$$P(Z$$

$$P(0 < X < 1, Y > 1) = 0$$

$$\neq P(0 < X < 1) P(Y > 1)$$

$$X \approx u^{\frac{1}{n}}$$

$$\left| \frac{\partial X}{\partial u} \right| = \left| \frac{1}{n} u^{\frac{1}{n}-1} \right|$$

$$\begin{aligned} f_u(u) &= f_X(x) \left| \frac{\partial X}{\partial u} \right| \\ &= \frac{1}{\alpha} x^{u^{\frac{n-1}{n}}} e^{-\frac{u}{\alpha}} \frac{1}{n} u^{\frac{1}{n}-1} \\ &\therefore \frac{1}{\alpha} e^{-\frac{u}{\alpha}} \end{aligned}$$

$$\frac{b}{a} x^{b-1} e^{-\frac{x}{a}}$$

$$\int \left(\frac{b}{a} \right) x^a \left(x^{b-1} \right) e^{-\frac{x}{a}} dx$$

$$\int_0^\infty (at)^b e^{-t} dt$$

$$\frac{x^b}{a} = t \quad \frac{dx}{dt} = \frac{b x^{b-1}}{a}$$

$$a^{\frac{k}{b}} \int_a^{\infty} t^{\left(\frac{k}{b}+1\right)-1} e^{-t} dt \quad x = (at)^{\frac{1}{b}} \quad dt = b x^{\frac{b-1}{b}} dx$$

$$a^{\frac{k}{b}} \Gamma\left(\frac{k}{b}+1\right)$$