2018 Intelligent Systems (知的システム構成論)

Subspace Identification and Spectral Learning of Dynamical Systems

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Outline

- What is subspace identification?
- Least square linear regression and singular value decomposition
- Deterministic and stochastic subspace identification
 - Extended state space representation
 - Canonical correlation analysis
- Spectral learning of dynamical systems

(Review) Learning LDS

Given :

- Observation sequence : $y_{1:T}$
- Input sequence : $oldsymbol{u}_{1:T}$

• Find:

- System matrices: A, B, C,
- Noise covariance matrices: Q, R
- Initial state distribution: $p(x_1) \sim N(m_0, V_0)$
- State sequence: $x_{1:T}$

(Review) EM Algorithm for LDS

Initialize model parameters

$$-\Theta^{(0)} = \left\{ \boldsymbol{A}^{(0)}, \boldsymbol{B}^{(0)}, \boldsymbol{C}^{(0)}, \boldsymbol{Q}^{(0)}, \boldsymbol{R}^{(0)}, \boldsymbol{m}_0^{(0)}, \boldsymbol{V}_0^{(0)} \right\}$$

- Repeat until convergence:
 - [E-step] Inference by Kalman (RTS) Smoothing
 - Filtering : Compute $p(\mathbf{x}_t|\mathbf{y}_{1:t},\mathbf{u}_{1:t},\Theta^{(t)}) = N(\mathbf{x}_t|\mathbf{m}_t,\mathbf{V}_t)$
 - Smoothing: Compute $p(x_t|y_{1:T}, u_{1:T}, \Theta^{(t)}) = N(x_t|\widehat{m}_t, \widehat{V}_t)$
 - Cross-covariance of x_t and x_{t+1} is also necessary
 - [M-step] Model update
 - $\Theta^{(t+1)} \leftarrow \underset{\Theta}{\operatorname{argmax}} E_{\boldsymbol{x}_{1:T}}[\log p(\boldsymbol{x}_{1:T}, \boldsymbol{y}_{1:T} | \boldsymbol{u}_{1:t}, \Theta)]$

Pros:

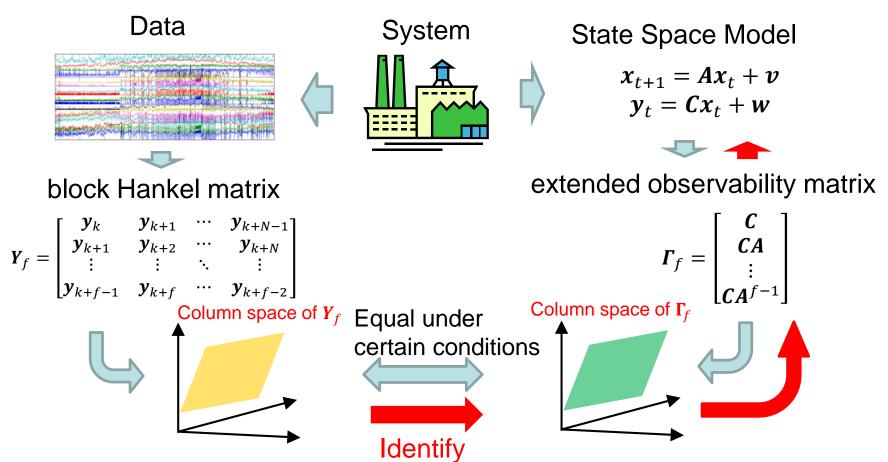
Exact (for LDS), highly applicable (to complicated models)

Cons:

4

Prone to local optimal many iterations (slow)

What is "subspace" identification method?



- Advantages:
 - Global optimum is obtained without iteration by linear algebraic operations

Least Squares Linear Regression, and Singular Value Decomposition

Linear Regression

Linear regression model (with vector output)

$$y = \Gamma x + v$$
 error (residual)

Output (target) Input (exploratory) variable (vectors) variable

Collected data

$$\begin{bmatrix} \mathbf{y}_1 & \mathbf{y}_2 & \cdots & \mathbf{y}_N \end{bmatrix} = \mathbf{\Gamma} \begin{bmatrix} \mathbf{x}_1 & \mathbf{x}_2 & \cdots & \mathbf{x}_N \end{bmatrix} + \mathbf{V}$$

$$\mathbf{Y} = \mathbf{\Gamma} \mathbf{X} + \mathbf{V}$$
Estimate

Least Squares Solution

Minimize sum of squares of residuals:

$$Loss = \|V\|_F^2 = \|Y - \Gamma X\|_F^2$$

$$= Tr[(Y - \Gamma X)^T (Y - \Gamma X)]$$

$$\frac{\partial}{\partial \Gamma} Loss = -2YX^T + 2\Gamma XX^T = \mathbf{0}$$
Gradient is zero at the minimum

$$\widehat{\Gamma} = YX^T (XX^T)^{-1}$$
Pseudo inverse of X

Least square solution

$$\widehat{Y} = \widehat{\Gamma}X = YX^T(XX^T)^{-1}X$$
 Prediction

$$V = Y - \widehat{Y} = Y \left(I - X^T (XX^T)^{-1} X \right)$$
 Residual

Geometric Interpretation of Least Squares

Prediction

$$\widehat{Y} = \widehat{\Gamma}X = YX^T(XX^T)^{-1}X$$

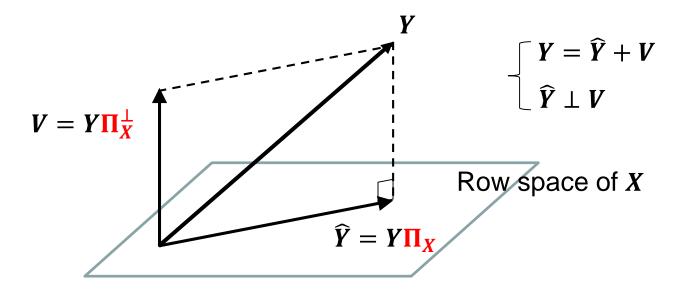
$$\equiv Y\Pi_X$$

Orthogonal projection of *Y* onto the space spanned by the rows of *X* (row space)

Redidual

$$V = Y - \widehat{Y} = Y \left(I - X^T (XX^T)^{-1} X \right)$$
$$= Y(I - \Pi_X) \equiv Y \Pi_X^{\perp}$$

Orthogonal projection of *Y* onto the complement of row space of *X*



When Two Sets of Explanatory Variables are given

Consider there are two sets of inputs X and U:

$$Y = \Gamma X + HU + V$$
Known (data) Residual

- Assume V is independent of both X and U
- Multiply Π_{U}^{\perp} from the right side

$$Y\Pi_{U}^{\perp} = \Gamma X\Pi_{U}^{\perp} + \underbrace{HU\Pi_{U}^{\perp}}_{0} + \underbrace{V\Pi_{U}^{\perp}}_{V}$$

$$\Rightarrow Y\Pi_{II}^{\perp} = \Gamma X\Pi_{II}^{\perp} + V$$

$$\Rightarrow \widehat{\Gamma} = Y \Pi_U^{\perp} X^T (X \Pi_U^{\perp} X^T)^{-1}$$

Remarks:

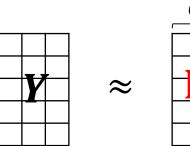
- *Ĥ* can also be similarly obtained
- Known as oblique 10 projection

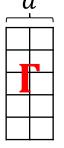
What if *X* is not given?

Consider both Γ and X are unknown

$$Y = \Gamma X + W$$

But assume the number of columns of Γ is known







This is the problem of low-rank approximation of a matrix



(Truncated) Singular Value Decomposition (SVD)

$$Y = U_d S_d V_d^T + W$$

where,
$$\boldsymbol{S}_d = \begin{bmatrix} \sigma_1 & 0 & 0 \\ 0 & \ddots & 0 \\ 0 & 0 & \sigma_d \end{bmatrix}$$

$$\int \widehat{\boldsymbol{\Gamma}} = \boldsymbol{U}_d \boldsymbol{S}_d^{1/2}$$

$$\widehat{\boldsymbol{X}} = \boldsymbol{S}_d^{1/2} \boldsymbol{V}_d^T$$

(Note that solution is not unique)

Innovation Form of State Space Model

Get back to the state space model

$$\begin{cases} x_{k+1} = Ax_k + Bu_k + v_k \\ y_k = Cx_k + w_k \end{cases}$$

Kalman filter representation

$$\widehat{\boldsymbol{x}}_{k+1} = A\widehat{\boldsymbol{x}}_k + B\boldsymbol{u}_k + K(\boldsymbol{y}_k - C\widehat{\boldsymbol{x}}_k)$$

• Define the *innovation* : $\boldsymbol{e}_k = \boldsymbol{y}_k - \boldsymbol{C}\widehat{\boldsymbol{x}}_k$

$$\begin{cases}
\widehat{x}_{k+1} = A\widehat{x}_k + Bu_k + Ke_k \\
y_k = C\widehat{x}_k + e_k
\end{cases}$$

Future Outputs

Innovation form

$$\widehat{x}_{k+1} = A\widehat{x}_k + Bu_k + Ke_k$$

$$y_k = C\widehat{x}_k + e_k$$
 1-step ahead
$$y_{k+1} = C\widehat{x}_{k+1} + e_{k+1} = CA\widehat{x}_k + CBu_k + CKe_k + e_{k+1}$$

$$\vdots$$

j-step
$$y_{k+j} = CA^j \widehat{x}_k + [CA^{j-1}B \quad \cdots \quad CB \quad 0] \begin{bmatrix} u_k \\ \vdots \\ u_{k+j} \end{bmatrix}$$
 ahead

$$+[CA^{j-1}K \quad \cdots \quad CK \quad I]\begin{bmatrix} e_k \\ \vdots \\ e_{k+j} \end{bmatrix}$$



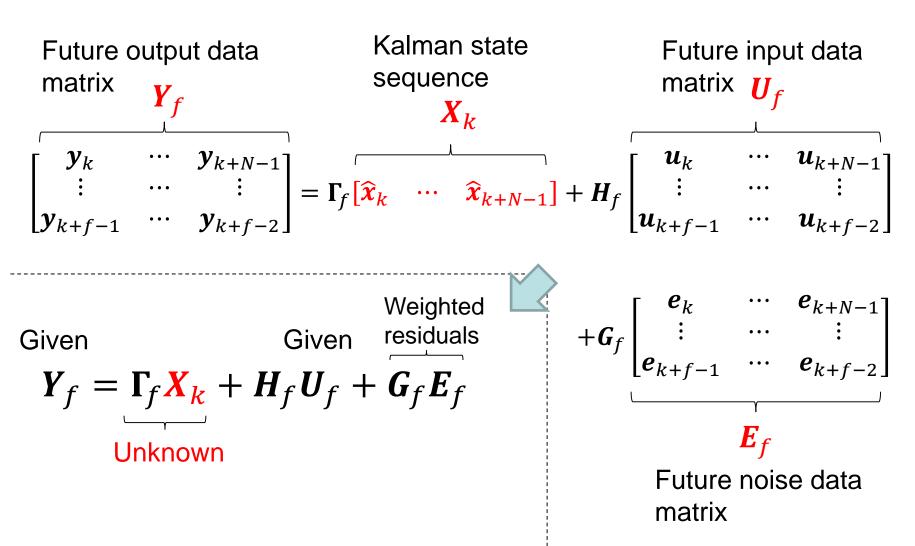
Extended observability matrix
$$\begin{matrix} \mathbf{y}_k \\ \mathbf{y}_{k+1} \\ \vdots \\ \mathbf{y}_{k+f-1} \end{matrix} = \begin{bmatrix} \mathbf{C} \\ \mathbf{C}\mathbf{A} \\ \vdots \\ \mathbf{C}\mathbf{A}^{f-1} \end{bmatrix} \widehat{\mathbf{x}}_k + \begin{bmatrix} \mathbf{0} & & & & & \\ \mathbf{C}\mathbf{B} & \mathbf{0} & & & \\ \vdots & \mathbf{C}\mathbf{B} & \ddots & & \\ \mathbf{C}\mathbf{A}^{f-2}\mathbf{B} & \cdots & \mathbf{C}\mathbf{B} & \mathbf{0} \end{bmatrix} \begin{bmatrix} \mathbf{u}_k \\ \mathbf{u}_{k+1} \\ \vdots \\ \mathbf{u}_{k+f-1} \end{bmatrix}$$

$$+\begin{bmatrix} I & & & & \\ CK & I & & \\ \vdots & CK & \ddots & \\ CA^{f-2}K & \cdots & CK & I \end{bmatrix} \begin{bmatrix} e_k \\ e_{k+1} \\ \vdots \\ e_{k+f-1} \end{bmatrix}$$

Relation between the current state (estimate) \hat{x}_k and future measurements $y_{k:k+f-1}$, given future inputs $u_{k:k+f-1}$

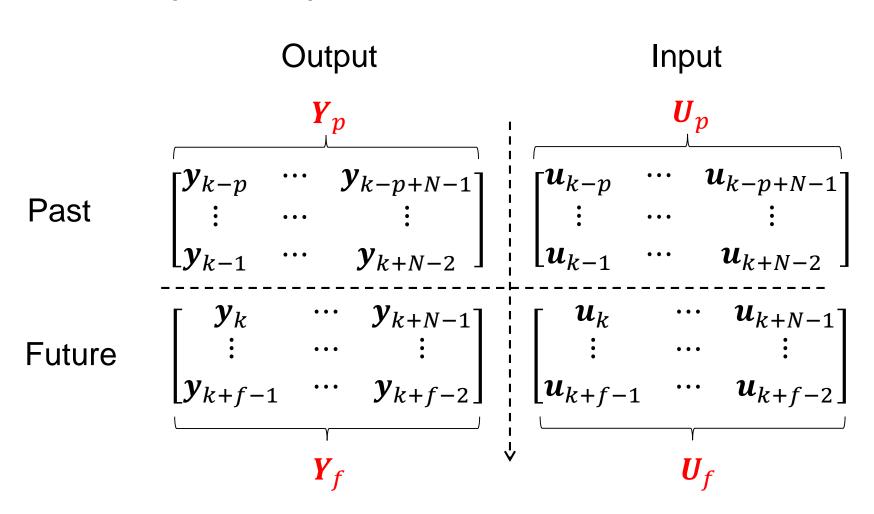
 H_f and G_f are Toeplitz matrices

stack horizontally



Past and Future Data Matrices

Past input / output data matrices are also defined



Kalman State Sequence

- Kalman state (sequence) X_k is unknown
- But, it is known it can be estimated from past inputs and outputs!

$$m{X}_k = m{L}_u m{U}_p + m{L}_y m{Y}_p = [m{L}_u \quad m{L}_y] egin{bmatrix} m{U}_p \ m{Y}_p \end{bmatrix} \equiv m{L}_z m{Z}_p \ \text{Past inputs} \ \text{and outputs} \end{cases}$$

$$Y_f = \Gamma_f X_k + H_f U_f + G_f E_f$$

$$Y_f = \Gamma_f L_z Z_p + H_f U_f + G_f E_f$$

Open Loop Assumption

- We assume the system is open loop
- Future innovation is uncorrelated to future inputs

$$E_f U_f^T = \mathbf{0} \Leftrightarrow E_f \Pi_{U_f}^{\perp} = E_f$$

 Future innovation is uncorrelated to past inputs and outputs

$$\boldsymbol{E}_f \boldsymbol{Z}_p^T = \boldsymbol{0}$$

Subspace Identification 1 (N4SID approach) [Qin 2004][Qin 2006]

• Data matrices Y_f , U_f and Z_p are given

$$Y_f = \Gamma_f L_z Z_p + H_f U_f + G_f E_f$$

• Multiply $\Pi_{U_f}^{\perp}$ from the right side $Y_f\Pi_{U_f}^{\perp} = \Gamma_f L_z Z_p \Pi_{U_f}^{\perp} + H_f U_f \Pi_{U_f}^{\perp} + G_f E_f \Pi_{U_f}^{\perp}$ $= \Gamma_f L_z Z_p \Pi_{U_f}^{\perp} + G_f E_f$

• Multiply \mathbf{Z}_{p}^{T} from the right side

$$egin{aligned} oldsymbol{Zero} oldsymbol{Zero} oldsymbol{Y}_f oldsymbol{\Pi}_{oldsymbol{U}_f}^oldsymbol{Z}_p^T = oldsymbol{\Gamma}_f oldsymbol{L}_z oldsymbol{Z}_p oldsymbol{\Pi}_{oldsymbol{U}_f}^oldsymbol{Z}_p^T + oldsymbol{G}_f oldsymbol{E}_f oldsymbol{Z}_p^T \ egin{aligned} oldsymbol{Known} & oldsymbol{\Gamma}_f oldsymbol{L}_z oldsymbol{Z}_p oldsymbol{\Pi}_{oldsymbol{U}_f}^oldsymbol{Z}_p^T \end{aligned}$$

N4SID Approach (Continued)

• Perform (truncated) SVD on $\mathbf{Y}_f \mathbf{\Pi}_{\mathbf{U}_f}^{\perp} \mathbf{Z}_p^{T}$

$$Y_f \Pi_{U_f}^{\perp} Z_p^T \approx U_d S_d V_d^T \left(\approx \Gamma_f L_z Z_p \Pi_{U_f}^{\perp} Z_p^T \right)$$

Determine the extended observability matrix as:

$$\hat{\mathbf{\Gamma}}_f \leftarrow \boldsymbol{U}_d \boldsymbol{S}_d^{-1/2} \\ - \text{Remember} \qquad \boldsymbol{\Gamma}_f = \begin{bmatrix} \boldsymbol{C} \\ \boldsymbol{C}\boldsymbol{A} \\ \vdots \\ \boldsymbol{C}\boldsymbol{A}^{f-1} \end{bmatrix}$$

System matrices A, C are obtained

Subspace Identification 2 : Regression Approach

• Multiply $\Pi_{U_f}^{\perp}$ to $Y_f = \Gamma_f L_z Z_p + H_f U_f + G_f E_f$ $Y_f \Pi_{U_f}^{\perp} = \Gamma_f L_z Z_p \Pi_{U_f}^{\perp} + G_f E_f \qquad \qquad \text{Same as the first approach}$ Known Unknown Known Residual Smaller is better

Target variables

Explanatory variables

Obtained by (multi-variate) linear regression:

$$\widehat{\Gamma_{f}L_{z}} = Y_{f}\Pi_{U_{f}}^{\perp} \left(\Pi_{U_{f}}^{\perp} Z_{p}^{T}\right) \left(Z_{p}\Pi_{U_{f}}^{\perp} \Pi_{U_{f}}^{\perp} Z_{p}^{T}\right)^{-1}$$

$$= Y_{f}\Pi_{U_{f}}^{\perp} Z_{p}^{T} \left(Z_{p}\Pi_{U_{f}}^{\perp} Z_{p}^{T}\right)^{-1} \approx U_{d}S_{d}V_{d}^{T} \quad (SVD)$$

$$\hat{\mathbf{\Gamma}}_f \leftarrow \boldsymbol{U}_d \boldsymbol{S}_d^{1/2}$$

Subspace Identification 3 : CCA Approach*

$$Y_f\Pi_{U_f}^\perp = \Gamma_f L_Z Z_p \Pi_{U_f}^\perp + G_f E_f$$
Known Unknown Known Residual
Variable set 1

Obtained by Canonical Correlation Analysis (CCA):

Define
$$W_r = \left(Y_f \Pi_{U_f}^{\perp} Y_f^{T}\right)^{-1/2}$$
 and $W_c = \left(Z_p \Pi_{U_f}^{\perp} Z_p^{T}\right)^{-1/2}$

Perform SVD on $W_r Y_f \Pi_{U_f}^{\perp} Z_p^T W_c \approx U_d S_d V_d^T$

Then,
$$\hat{\Gamma}_f \leftarrow W_r^{-1} U_d S_d^{1/2}$$

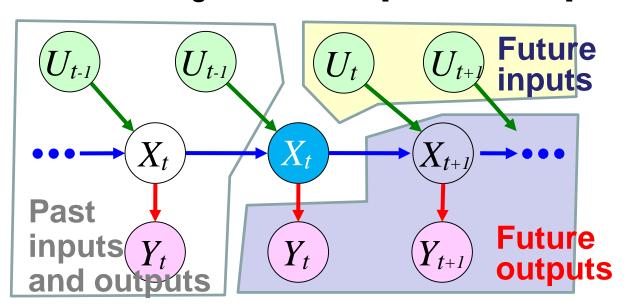
(*) Also known as canonical variate analysis (CVA)

CCA-based Subspace Identification (Continued)

Additionally, state vectors are obtained by

$$\widehat{\boldsymbol{X}}_k = \boldsymbol{S}_d^{1/2} \boldsymbol{V}_d^T \boldsymbol{W}_c \boldsymbol{Z}_p$$

- This result is very interesting, because the "state" is obtained by CCA between "past" and "future" data sets.
- This idea goes back to [Akaike 1975].



"The state vector carries information necessary to predict the future output based on the past."

[Katayama 2005]

Reference

- [Qin 2006] S. Joe Qin, "An overview of subspace identification", Computers & Chemical Engineering, Volume 30, Issues 10–12, 2006, Pages 1502-1513
 - Presentation file [Qin 2004] is also available (http://people.duke.edu/~hpgavin/SystemID/References/Qin-SubspaceID-2004.pdf)
- [Overschee 1996] Peter VAN OVERSCHEE & Bart DE MOOR, "SUBSPACE IDENTIFICATION FOR LINEAR SYSTEMS - Theory - Implementation - Applications", Springer, 1996
- [Katayama 2005] Toru Katayama, "Subspace Methods for System Identification", Springer, 2005
- [Akaike 1975] H.Akaike, "Markovian Representation of Stochastic Processes by Canonical Variables", SIAM J. CONTROL, Vol. 13, No. 1, pp.162-173, January 1975

Appendix: Canonical Correlation Analysis (CCA)

Canonical Correlation Analysis (1)

There are two random variables x and y

$$\mathbf{x} \in \mathbb{R}^m$$
 (m-dim.) $\mathbf{y} \in \mathbb{R}^n$ (n-dim.)

For simplicity, both x and y are "centered", i.e., their means are zero vectors.

$$E[\mathbf{x}] = \mathbf{0}_m \qquad E[\mathbf{y}] = \mathbf{0}_n$$

Define covariance matrices of / between x and y

$$\operatorname{var}(\mathbf{x}) = \operatorname{E}[\mathbf{x}\mathbf{x}^T] = \mathbf{\Sigma}_{xx} \quad \operatorname{var}(\mathbf{y}) = \operatorname{E}[\mathbf{y}\mathbf{y}^T] = \mathbf{\Sigma}_{yy}$$

$$\operatorname{cov}(\mathbf{x}, \mathbf{y}) = \operatorname{E}[\mathbf{x}\mathbf{y}^T] = \mathbf{\Sigma}_{xy}$$

Canonical Correlation Analysis (2)

Consider to construct synthetic variable u and v by linear combinations of x and y, respectively.

$$u = \mathbf{a}^T \mathbf{x} = a_1 x_1 + a_2 x_2 + \dots + a_m x_m$$

 $v = \mathbf{b}^T \mathbf{y} = b_1 y_1 + b_2 y_2 + \dots + b_n y_n$

Problem: Find a and b so that the correlation between u and v is maximized

$$\rho = \mathbf{cor}(u, v) = \frac{\mathbf{cov}(u, v)}{\sqrt{\mathbf{var}(u)}\sqrt{\mathbf{var}(v)}}$$
$$= \frac{\mathbf{a}^{T} \mathbf{\Sigma}_{xy} \mathbf{b}}{\sqrt{\mathbf{a}^{T} \mathbf{\Sigma}_{xx} \mathbf{a}} \sqrt{\mathbf{b}^{T} \mathbf{\Sigma}_{yy} \mathbf{b}}}$$

Canonical Correlation Analysis (3)

Impose constraints on a and b, so that the variances of u and v become 1

$$\operatorname{var}(u) = \operatorname{var}(\mathbf{a}^T \mathbf{x}) = \mathbf{a}^T \mathbf{\Sigma}_{xx} \mathbf{a} = 1$$

$$\operatorname{var}(v) = \operatorname{var}(\mathbf{b}^T \mathbf{y}) = \mathbf{b}^T \mathbf{\Sigma}_{yy} \mathbf{b} = 1$$

The goal of CCA is to solve:

$$(\boldsymbol{a}_1, \boldsymbol{b}_1) = \underset{\boldsymbol{a}^T \boldsymbol{\Sigma}_{xx} \boldsymbol{a} = 1, \boldsymbol{b}^T \boldsymbol{\Sigma}_{yy} \boldsymbol{b} = 1}{\operatorname{argmax}} \boldsymbol{a}^T \boldsymbol{\Sigma}_{xy} \boldsymbol{b}$$

It can be solved by Lagrange multiplier, but there is a more elegant method.

Canonical Correlation Analysis (4)

Let square root matrices of Σ_{xx} and Σ_{vv} be $\Sigma_{xx}^{1/2}$ and $\Sigma_{vv}^{1/2}$, respectively. Then, define c and d as,

$$c = {oldsymbol \Sigma_{\chi\chi}}^{1/2} {oldsymbol a}$$
 , and ${oldsymbol d} = {oldsymbol \Sigma_{\chi\chi}}^{1/2} {oldsymbol b}$

The problem turns to be

$$(\boldsymbol{c}_1, \boldsymbol{d}_1) = \underset{\|\boldsymbol{c}\|=1}{\operatorname{argmax}} \boldsymbol{c}^T \left(\boldsymbol{\Sigma}_{xx}^{-T/2} \boldsymbol{\Sigma}_{xy} \boldsymbol{\Sigma}_{xx}^{-1/2} \right) \boldsymbol{d}$$

This problem can be solved by SVD!

$$\boldsymbol{\Sigma}_{xx}^{-T/2} \boldsymbol{\Sigma}_{xy} \boldsymbol{\Sigma}_{xx}^{-1/2} = \boldsymbol{U} \boldsymbol{S} \boldsymbol{V}^T \quad \Rightarrow \quad \boldsymbol{c}_1 = \boldsymbol{u}_1, \boldsymbol{d}_1 = \boldsymbol{v}_1$$

$$\Rightarrow$$
 $a_1 = \Sigma_{\chi\chi}^{1/2} u_1$, and $b_1 = \Sigma_{\chi\chi}^{1/2} v_1$