

Logistic Regression

Koushik Khan

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Introduction to Logistic Regression A Beginner Friendly Tutorial

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The Binary Classification Problem

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- Imagine situations where we would like to know,
 - the *eligibility* of getting a bank loan given the value of credit score (x_{credit_score}) and monthly income (x_{income}) .
 - identifying a tumor as *benign* or *malignant* given its size (x_{tumor_size}) .
 - classifying an email as *promotional* given the no. of occurrences for some keywords like {'win', 'gift', 'discount'} $(x_{n \text{ win}}, x_{n \text{ qift}}, x_{n \text{ discount}})$.
 - finding a monetary transaction as *fraudulent* given the time of occurrence (x_{time_stamp}) and amount (x_{amount}) .
- These problems occur frequently in real life & can be dealt with machine learning.
- All such problems come under the umbrella of what is known as *Classification*.
- In each scenario, only one of the two possible outcomes can occur, hence these are specifically known as *Binary Classification* problems.



Understanding The Dataset

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- Any dataset containing *numerical* or *categorical* features can be used for classification.
- However, the target variable must be *categorical* in nature.
- Specifically, for binary classification, a target variable (Y) must take any one of the two distinct values like {'benign', 'malignant'}.
- To use with ML algorithms, values of the target variable are *encoded* into numeric representations e.g. {"benign": 0, "malignant": 1}, a.k.a 'class-0' & 'class-1' respectively.
- In literature, typically 'class-0' is tagged as *failure* and 'class-1' is tagged as *success*.



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MACHINE LEARNING FOR CLASSIFICATION



How Does A Machine Perform Classification?

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• During the inference, the goal is to have the ML model *predict the class* label for a given set of feature values.

- Specifically, a binary classification model estimates two probabilities \hat{p}_0 & \hat{p}_1 for 'class-0' and 'class-1' respectively where $\hat{p}_0 + \hat{p}_1 = 1$.
- The predicted label depends on $\max\{\hat{p}_0, \hat{p}_1\}$ i.e. it's the one which is most probable based on the given features.
- In logistic regression, \hat{p}_1 (i.e. success probability) is compared with a *predefined threshold* p^* to predict the class label like below:

predicted class = 1;
$$\hat{p}_1 \ge p^*$$

= 0; otherwise (1)

• \blacksquare To keep the notation simple and consistent, we will denote the *success* probability as p, and failure probability as (1-p) instead of p_0 and p_1 .



Why NOT Linear Regression?

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- Can't we really use linear regression to address classification? The answer is NO!
- Let's try to understand why:
 - \bullet To estimate p using linear regression, we would need:

$$\hat{p} = \hat{\alpha} + \hat{\beta} x_{tumor_size} \tag{2}$$

- Eqn. (2) doesn't seem to be feasible as the R.H.S, in principle, belongs to $\mathbb{R}(-\infty, +\infty)$ & the L.H.S belongs to (0,1).
- Can we convert $(\hat{\alpha} + \hat{\beta}x_{tumor_size})$ to something belonging to (0,1)? That may work as an estimate of a probability! The answer is YES!
- We need a converter (a function), say, g(.) that will connect $p \in (0,1)$ to $(\hat{\alpha} + \hat{\beta}x_{tumor_size}) \in \mathbb{R}$.
- Fortunately, such functions do exist and they are often referred to as *link* functions in this context.



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The Bernoulli Distribution

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- Following the definition, a link function connects the liner predictor like $(\alpha + \beta x_{tumor_size})$ to the expected value of the target variable.
- The *binary* target variable naturally suits the *Bernoulli Distribution* (Bernoulli's Trial) for explanation.

Bernoulli's Trial

A random experiment that results in one of the two possible outcomes, often called, a success and a failure, with a $constant\ probability$ of success, say, p.

Examples

- tossing a fair coin the coin shows either the 'HEAD' or the 'TAIL'
- performing COVID test result will be either '+ve' or '-ve'
- detecting a tumor as 'benign' or 'malignant' etc.



PMF And Expectation

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ullet Bernoulli's trial can be expressed mathematically with a random variable, say, Y as:

$$f(y|p) = P(Y = y) = p^{y} \times (1 - p)^{1 - y}; \ y \in \{0, 1\}$$
(3)

where eqn. (3) is known as Probability Mass Function (PMF).

• The PMF maps the values (y) taken by Y to probabilities e.g.

when
$$y = 1 \Rightarrow P(Y = 1) = p$$

when $y = 0 \Rightarrow P(Y = 0) = (1 - p)$

• The expected value of Y is calculated as:

$$E(Y) = 0 \times P(Y = 0) + 1 \times P(Y = 1)$$

$$= P(Y = 1)$$

$$= p$$

$$= Probability of 'success' or identifying a malignant case (4)$$



The Logit Link Function

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• The logit link function is defined as:

$$logit(E(Y)) = log\left(\frac{E(Y)}{1 - E(Y)}\right)$$
(5)

• And by the definition of a link function, it connects the linear predictor i.e. $(\alpha + \beta x_{tumor size})$ to E(Y) i.e. p as:

$$\log\left(\frac{p}{1-p}\right) = \alpha + \beta x_{tumor_size} \tag{6}$$

The eqn. (6) is formally called the *Logistic Regression* equation.

• For a linear regression, the link function is the *identity function* i.e. q(x) = x.

The Inverse of Logit - Sigmoid Function

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• In logistic regression, we try to model P(Y = 1) as:

$$\log\left(\frac{p}{1-p}\right) = \alpha + \beta x_{tumor_size}$$

where the L.H.S is also known as *log-odds*.

• Alternatively we can write:

$$p = \frac{e^{\alpha + \beta x_{tumor_size}}}{1 + e^{\alpha + \beta x_{tumor_size}}} \text{ (verify)}$$
 (7

• The R.H.S of eqn.(7) is referred to as *sigmoid* function, denoted by $\sigma(.)$ which is the *inverse* of the logit function.



Nature Of The Sigmoid

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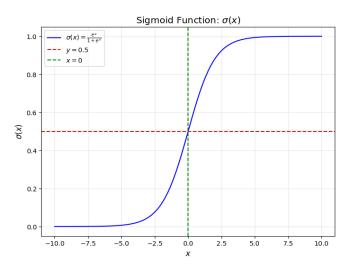


Figure: The sigmoid function

Math Behind The Sigmoid

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• The sigmoid function:

$$\sigma(z) = \frac{e^z}{1 + e^z} = \frac{1}{1 + e^{-z}}; \ z \in \mathbb{R}$$
 (8)

maps any $z \in \mathbb{R}$ to a number belonging to (0,1).

• It has a very nice & important property too:

$$\frac{d\sigma}{dz} = \sigma'(z) = \sigma(z) \times (1 - \sigma(z)) \text{ (verify)}$$
(9)

• \blacksquare It looks like an *elongated 'S'*, that is where it gets its name from.



Linear And Non-linear Classification

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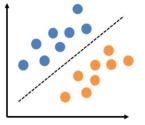
Parameter

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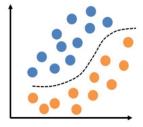


Figure: On the left, the classes are linearly separable as the boundary is a straight line, however they are not on the right



Logistic Regression Is A Linear Classifier

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- The logistic regression equation (6) is actually a *straight line*, like y = mx + c.
- Decall the prediction rule:

predicted class = 1;
$$\hat{p} \ge p^* \Rightarrow \hat{\alpha} + \hat{\beta} x_{tumor_size} \ge \log \left(\frac{p^*}{1 - p^*} \right)$$

= 0; otherwise

- A simple logistic regression (the one we discussed) predicts the class label by identifying the regions on either side of a *straight line* (or hyperplane in general), hence it's a *linear classifier*.
- Logistic regression works well for *linearly separable* classes.



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ESTIMATION OF PARAMETERS



The Likelihood Function

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- While defining PMF, it is assumed that the success probability (p) is known.
- However, in reality we $don't \ know$ the p conceptually, PMF is $not \ suitable$ for further use.
- The dataset involves N patients (say): $\{(x_{i,tumor_size}, y_i)\}_{i=1}^N$, $y_i \in \{0, 1\}$.
- Imagine the i^{th} patient, Y_i , has a probability p_i of developing a malignant case. Here $Y_i \sim Bernoulli(p_i)$.
- Interestingly, the expression for $P(Y_i = y_i)$ is same as the PMF, eqn. (3),

$$\ell_i(p_i|y_i) = P(Y_i = y_i) = p_i^{y_i} \times (1 - p_i)^{1 - y_i}; \ y_i \in \{0, 1\} \ \forall i = 1, 2, \dots, N$$
(10)

The eqn. (10) is known as the likelihood for Y_i taking a value y_i .

• In Likelihood we know the dataset (as we're observing it), but the p_i is unknown to us.



The Joint Likelihood Function

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- What's the likelihood for observing the entire dataset? Well, the joint likelihood gives that answer.
- It's computed as below:

$$L = P(Y_1 = y_1 \cap \ldots \cap Y_N = y_N) = \prod_{i=1}^N p_i^{y_i} \times (1 - p_i)^{1 - y_i}$$
 (11)

Eqn. (11) is called the *joint likelihood* (L).

• It is much easier to work with joint likelihood after a *log-transformation*, also called the *log-likelihood* (*LL*),

$$LL = \log(L) = \sum_{i=1}^{N} \left\{ y_i \log(p_i) + (1 - y_i) \log(1 - p_i) \right\}$$
 (12)

• \blacksquare Joint likelihood measures the probability of observing the underlying dataset i.e. having $\{Y_1 = y_1, \dots, Y_N = y_N\}$ for some unknown set of probabilities $\{p_1, \dots, p_N\}$.



Maximum Likelihood Estimation

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- Imagine that some process might have produced the observed dataset $\{(x_{i,tumor_size}, y_i)\}_{i=1}^{N}$
- We are NOT sure what values of p_i 's the process would have considered to produce the dataset.
- We can imagine *several potential candidates* for each p_i (say, all belong to \mathbb{P}) that might have been used to produce the dataset.
- In principle, the best candidate for each p_i would be the one that makes the joint likelihood (L) or log likelihood (LL) (both are functions of p_i 's) maximum.
- Mathematically, we would perform,

$$\underset{p_1, \dots, p_N \in \mathbb{P}}{\operatorname{arg max}} \sum_{i=1}^{N} \left\{ y_i \log(p_i) + (1 - y_i) \log(1 - p_i) \right\}$$
 (13)

to find $\hat{p}_1, \ldots, \hat{p}_N$.

Going Deeper

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• Finding $\hat{p}_1, \ldots, \hat{p}_N$ using eqn. (13) is basically an *optimization problem*.

• **9** Recall
$$p_i = \frac{e^{\alpha + \beta x_{i,tumor_size}}}{1 + e^{\alpha + \beta x_{i,tumor_size}}} = \sigma(\alpha + \beta x_{i,tumor_size})$$

• The eqn. (13) can be simplified as:

$$\underset{p_{1},\dots,p_{N}\in\mathbb{P}}{\operatorname{arg\,max}} \left[\sum_{i=1}^{N} \left\{ y_{i} \log \left(\sigma(\alpha + \beta x_{i,tumor_size}) \right) + \left(1 - y_{i} \right) \log \left(1 - \sigma(\alpha + \beta x_{i,tumor_size}) \right) \right\} \right]$$
(14)

- Here each p_i is a function of the parameters $\alpha \& \beta$ and the known data $x_{i,tumor_size}$.
- Finding $\hat{p}_1, \ldots, \hat{p}_N$ is *equivalent* to finding α and β with the help of eqn. (14).



Using Gradient Descent

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- Though this is a maximization problem, the ML community prefers minimizing the negative log-likelihood (NLL) using gradient descent.
- In practice, a scaled version of NLL $(\frac{1}{N}NLL)$ is used and that is known as Binary Cross Entropy (BCE) loss function.
- Now the problem boils down to:

$$\underset{p_1, \dots, p_N \in \mathbb{P}}{\operatorname{arg\,min}} \frac{1}{N} NLL = \underset{p_1, \dots, p_N \in \mathbb{P}}{\operatorname{arg\,min}} - \frac{1}{N} \sum_{i=1}^{N} \left\{ y_i \log(p_i) + (1 - y_i) \log(1 - p_i) \right\}$$
(15)

or equivalently,

$$\underset{\alpha,\beta}{\operatorname{arg\,min}} - \frac{1}{N} \sum_{i=1}^{N} \left\{ y_{i} \log \left(\sigma(\alpha + \beta x_{i,tumor_size}) \right) + (1 - y_{i}) \log (1 - \sigma(\alpha + \beta x_{i,tumor_size})) \right\}$$
(16)

Computing Derivatives

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- Gradient Descent computes *derivatives* of BCE w.r.t α and β .
- Here is how it works: Let us first consider $z_i = \alpha + \beta x_{i,tumor_size}$ we would like to compute:

$$\frac{\partial BCE}{\partial \alpha} = -\frac{1}{N} \frac{\partial}{\partial \alpha} \sum_{i=1}^{N} \left\{ y_i \log \left(\sigma(z_i) \right) + (1 - y_i) \log (1 - \sigma(z_i)) \right\}
= -\frac{1}{N} \sum_{i=1}^{N} \left\{ y_i \frac{\partial}{\partial \alpha} \log \left(\sigma(z_i) \right) + (1 - y_i) \frac{\partial}{\partial \alpha} \log \left(1 - \sigma(z_i) \right) \right\}$$
(17)

and similarly,

$$\frac{\partial BCE}{\partial \beta} = -\frac{1}{N} \sum_{i=1}^{N} \left\{ y_i \frac{\partial}{\partial \beta} \log \left(\sigma(z_i) \right) + (1 - y_i) \frac{\partial}{\partial \beta} \log \left(1 - \sigma(z_i) \right) \right\}$$
(18)



The Magic Of Using The Sigmoid

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- We will compute the derivatives one by one.
- Here is how we proceed for α :

 $\frac{\partial \log (\sigma(z_i))}{\partial \alpha} = \underbrace{\frac{\partial \log (\sigma(z_i))}{\partial \sigma(z_i)} \times \frac{\partial \sigma(z_i)}{\partial z_i} \times \frac{\partial z_i}{\partial \alpha}}_{\text{clain rule of differentiation}}$ $= \frac{1}{\sigma(z_i)} \times \sigma(z_i) \times (1 - \sigma(z_i)) \times 1$ $= (1 - \sigma(\alpha + \beta x_{i,tumor_size}))$ (19)

$$\frac{\partial \log (1 - \sigma(z_i))}{\partial \alpha} = \frac{\partial \log (1 - \sigma(z_i))}{\partial (1 - \sigma(z_i))} \times \frac{\partial (1 - \sigma(z_i))}{\partial \sigma(z_i)} \times \frac{\partial \sigma(z_i)}{\partial z_i} \times \frac{\partial z_i}{\partial \alpha}$$

$$= \frac{1}{(1 - \sigma(z_i))} \times (-1) \times \sigma(z_i) \times (1 - \sigma(z_i)) \times 1$$

$$= -\sigma(z_i) = -\sigma(\alpha + \beta x_{i,tumor_size})$$
(20)

Continuing...

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• And similarly for β :

chain rule of differentiation
$$\frac{\partial \log (\sigma(z_i))}{\partial \beta} = \underbrace{\frac{\partial \log (\sigma(z_i))}{\partial \sigma(z_i)} \times \frac{\partial \sigma(z_i)}{\partial z_i} \times \frac{\partial z_i}{\partial \beta}}_{\text{claim rule of differentiation}} (21)$$

$$= \frac{1}{\sigma(z_i)} \times \sigma(z_i) \times (1 - \sigma(z_i)) \times x_{i,tumor_size}$$

$$= (1 - \sigma(\alpha + \beta x_{i,tumor_size})) \times x_{i,tumor_size}$$

$$\frac{\partial \log (1 - \sigma(z_i))}{\partial \beta} = \frac{\partial \log (1 - \sigma(z_i))}{\partial (1 - \sigma(z_i))} \times \frac{\partial (1 - \sigma(z_i))}{\partial \sigma(z_i)} \times \frac{\partial \sigma(z_i)}{\partial z_i} \times \frac{\partial z_i}{\partial \beta}$$

$$= \frac{1}{(1 - \sigma(z_i))} \times (-1) \times \sigma(z_i) \times (1 - \sigma(z_i)) \times x_{i,tumor_size}$$

$$= -\sigma(\alpha + \beta x_{i,tumor_size}) \times x_{i,tumor_size}$$
(22)



The Iterative Rule Of Gradient Descent

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- We can now compute $\frac{\partial BCE}{\partial \alpha}$ and $\frac{\partial BCE}{\partial \beta}$.
- And finally make use of the Gradient Decent *update rule*:

estimates at
$$(t+1)^{th}$$
 step estimates at t^{th} step
$$\begin{pmatrix}
\hat{\alpha}^{(t+1)} \\
\hat{\beta}^{(t+1)}
\end{pmatrix} = \begin{pmatrix}
\hat{\alpha}^{(t)} \\
\hat{\beta}^{(t)}
\end{pmatrix} - \eta \cdot \begin{pmatrix}
\frac{\partial BCE}{\partial \alpha} \\
\hat{\alpha}^{(t)}
\end{pmatrix}$$

$$\frac{\partial BCE}{\partial \beta} \\
\hat{\beta}^{(t)}$$
(23)

Here η is the *learning rate*.



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EVALUATING MODEL PERFORMANCE

Confusion Matrix And Related Metrics

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Actual Values Positive (1) Negative (0) Positive (1) Predicted Values TP FP Negative (0) FN TN

Figure: Confusion Matrix

- All cases = TP + TN + FP + FN
- Correctly classified cases = TP + TN
- Misclassified cases = FP + FN
- Precision = $\frac{TP}{TP+FP}$
- Recall (Sensitivity) = $\frac{TP}{TP+FN}$
- Specificity = $\frac{TN}{TN+FP}$
- Accuracy = $\frac{TP+TN}{TP+TN+FP+FN}$
- F1-score = $\frac{2 \times Precision \times Recall}{Precision + Recall}$
- All the above metrics except the first one depends on the threshold p^* .



Understanding The Precision

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- Precision measures the *probability of predicting a true positive case* by a fitted model.
- According to the formula: $\frac{TP}{TP+FP}$, the *lesser* the number of false positive cases, the *higher* will be the precision.
- Precision is important where is making *false positive mistakes* is *risky*.
- In a email spam detection system, it's crucial that a *non-spam* email is *not* getting tagged as a spam email, otherwise an user may miss an important email. expecting a high precision
- In medical diagnosis (e.g. cancer detection) high precision gives confidence to the doctors to start treatment without further tests. expecting a high precision
- Precision is also important when the dataset is highly imbalanced (e.g. credit fraud detection, where getting a fraudulent transaction is rare). If the precision is low, even if the accuracy is very high, the model would probably raise many *false alarms*, which is misleading. *expecting a high precision*



Understanding The Recall

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- Recall (a.k.a **Sensitivity**) measures the probability of detecting a true positive case when it's actually positive.
- According to the formula: $\frac{TP}{TP+FN}$, the *lesser* the number of false negative cases, the *higher* will be the recall.
- Recall is important where making *false negative mistakes* is *risky*.
- In case of cancer detection diagnosis, it's very important that a cancer is getting detected in the body if it is actually there, otherwise it will be a life risk. expecting a high recall
- While using medical kit for detecting COVID, it's important that a person is *NOT tagged* as 'COVID -VE' when he is *actually* 'COVID +VE'. Having such cases will *infect* many other people. *expecting a high recall*
- Detecting as many threats as possible is important for an airport security system. Having higher values of recall will make sure almost all positive cases are identified. *expecting a high recall*



Understanding The F1 Score

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- The idea of using F1-score is to keep *a balance* between precision and recall.
- According to the formula: $\frac{2 \cdot Precision \cdot Recall}{Precision + Recall}$, F1-score calculates $\frac{harmonic}{mean}$ of precision and recall.
- Harmonic mean *penalizes* the extreme values of both precision and recall.
- For example, in case of credit fraud detection, actual fraudulent cases (positive) are very rare and this may make the model biased towards *legitimate cases* with a *very high accuracy*, however it *may not* make any sense.
- Here the model should actually:
 - have *high precision* i.e. lower chance of raising false alarm by identifying a legitimate case as fraudulent
 - have *high recall* i.e. lower chance of missing fraudulent transaction

The F1-score keeps a balance between these and gives a much more realistic evaluation of how well a model is performing in detecting the minority class (fraud).



The Receiver Operating Characteristic Curve (ROC)

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- It's a *graphical tool* depending on two metrics derived from the *confusion* matrix:
 - True Positive Rate = P(Predicted Positive | Actually Positive) = $\frac{TP}{TP+FN}$
 - False Positive Rate = P(Predicted Positive | Actually Negative) = $\frac{FP}{FP+TN}$
- By varying p^* within a permissible range a set of $\{(FPR_k, TPR_k)\}_{k=1}^K$ are obtained, and are plotted to form what is known as $ROC\ Curve$.
- The FPR varies along the X-axis and TPR varies along the Y-axis.
- Both TPR and FPR vary within [0,1] making the *total area* of the plotting canvas to be 1.
- The *diagonal line*, connecting the coordinates (0,0) and (1,1) indicates FPR = TPR, which is how a *random model* would behave.
- It's always good to have a model which produces *TPR values on the higher side* and *FPR values on the lower side*.



Area Under The [ROC] Curve (AUC)

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- It's one of the robust measures to compare different models or model configurations.
- The diagonal line (- -) divides
 the plotting canvas in to two halves
 having an area of 0.5 each. This
 line indicates a random classifier
 which is equally good and bad.
- Any model better than the random one will *cover an area* > 0.5.
- The *higher* the AUC the model achieves, the *better* its performance.

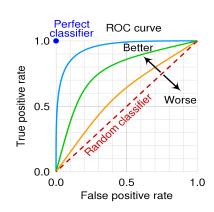


Figure: Comparing models with ROC curves and AUC values ¹

¹Image Source: Wikipedia



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- α : When x_{tumor_size} is zero, it's the value of the log-odds. It's often called the baseline log-odds.
- β : It's the change in log-odds for an unit change in $x_{tumor\ size}$.
- Interpretations are just like the *linear regression* as the alternate form of logistic regression (eqn. (6)) is exactly a linear regression w.r.t the log-odds.
- \square log-odds belongs to $(-\infty, +\infty)$ (verify)
- Even though logistic regression is used for *classification*, it actually *estimates a probability*, which is *continuous* within (0, 1) it's a *bridge* between *continuous modeling* and *discrete outcomes*, classification is just a *practical application* of this model.



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References

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- 🗗 Categorical Data Analysis by Alan Agresti 🗹
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Thanks

Thank You