

$Y = (y_1, \dots, y_n)$  and  $M = (m_1, \dots, m_n)$ ;  $M$  indicates missingness in  $Y$

$E(Y) = \mu$  population mean

$E(Y|M=0) = \bar{Y}$  sample mean;  $E[\bar{Y}] = E\left[\frac{1}{n} \sum_{i=1}^n Y_i\right] = \frac{1}{n} \sum E[Y_i] = E(Y) = \mu$

$M_i \in \{1, \dots, n\} \sim \text{Bern}(\exp(\beta_0 + \beta_1 y_i))$ ;  $P(M_i=1) = \exp(\beta_0 + \beta_1 y_i)$ ; note that  $M$  depends on  $Y$

Suppose we have  $k$  missing values where  $0 \leq k \leq n$ .  $Y^*$  indicates  $Y$  after imputed

$y_i^* = \begin{cases} y_i & (m_i=0) \text{ total } n-k \text{ entries} \\ \bar{Y} & (m_i=1) \text{ total } k \text{ entries; impute } \bar{Y} \text{ (sample mean) when } M=1 \end{cases}$

$$Y^* = \frac{1}{n} \sum_{i=1}^n y_i^* = \frac{1}{n} \left\{ \sum_{i=k+1}^n y_i + k \cdot \bar{Y} \right\} = \frac{1}{n} \{ (n-k) \cdot \bar{Y} + k \cdot \bar{Y} \} = \bar{Y}$$

Thus,  $E[\bar{Y}] = E[Y^*]$

$$\text{Bias}(Y^*) = E(Y^*) - Y$$

$E(Y^*) = E(\bar{Y}) = \mu$   $\therefore$  Thus, Bias in imputed for mean imputation is  $\text{Bias}(y_i^*) = \mu - y_i$

OR

$$\begin{aligned} E(\bar{Y}) &= E[Y|M=0] \text{ by definition} \\ &= \frac{E[Y; M=0]}{P(M=0)} \\ &= \frac{E[Y; M=0]}{P(M=0)} = \frac{E[Y \cdot I(M=0)]}{P(M=0)} \end{aligned}$$

By definition,  $E[X|A] = \frac{E[X \cdot I(A)]}{P(A)}$

Let  $R = (1-M)$  where  $R$  = non-missing indicator, then

$$\begin{aligned} E(\bar{Y}) &= \frac{E[Y \cdot I(R=1)]}{P(R=1)} = \frac{E[Y \cdot R]}{E(R)} = \frac{E_Y[E_{R|Y}[Y \cdot R_i | Y_i]]}{E_Y[E_{R|Y}[R_i | Y_i]]} = \frac{E_Y[Y_i \cdot \underbrace{E_{R|Y}[R_i | Y_i]}_{\text{note } E_{R|Y}[R_i | Y_i] = (1 - \exp(-\cdot))}]}{E_Y[\underbrace{E_{R|Y}[R_i | Y_i]}_{\text{note } E_{R|Y}[R_i | Y_i] = (1 - \exp(-\cdot))}]} \\ &= \frac{E_Y[Y_i \cdot \frac{1}{1 + \exp(\beta_0 + \beta_1 y_i)}]}{E_Y[\frac{1}{1 + \exp(\beta_0 + \beta_1 y_i)}]} \end{aligned}$$

$$Y_i^* = Y_i [M_i = 0] + \overset{\text{constant}}{c} [M_i = 1]$$

$$P(M_i = 1) = \text{expit}(\beta_0 + \beta_1 Y_i)$$

$$E(Y_i^*) = E[Y_i [M_i = 0] + c [M_i = 1]]$$

$$= E_Y [E_{M|Y} \{Y_i [M_i = 0] + c [M_i = 1]\}]$$

$$= E_Y [Y_i \cdot \underbrace{(1 - \text{expit}(\beta_0 + \beta_1 Y_i))}_{P(M=0)} + c \cdot \underbrace{\text{expit}(\beta_0 + \beta_1 Y_i)}_{P(M=1)}]$$

$$= \int y \cdot \{1 - \text{expit}(\beta_0 + \beta_1 y)\} \underbrace{f_Y(y)}_{f_Y(y)} dy + c \cdot \int \text{expit}(\beta_0 + \beta_1 y) f_Y(y) dy$$

$$\text{Assume } Y \sim \text{normal}(\mu, \sigma^2); f_Y(y) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2}(\frac{y-\mu}{\sigma})^2}$$

$$= \int_{-\infty}^{\infty} y \cdot \{1 - \text{expit}(\beta_0 + \beta_1 y)\} \cdot \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2}(\frac{y-\mu}{\sigma})^2} dy + c \cdot \int_{-\infty}^{\infty} \text{expit}(\beta_0 + \beta_1 y) \cdot \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2}(\frac{y-\mu}{\sigma})^2} dy$$

$$\text{Let } y = \mu + \sigma\sqrt{2}t, \quad dy = \sigma\sqrt{2} dt \quad \Rightarrow \quad \frac{y-\mu}{\sigma\sqrt{2}} = t, \quad \frac{1}{2}\left(\frac{y-\mu}{\sigma}\right)^2 = t^2$$

Also  $dy = \sigma\sqrt{2}dt$  so,  $\sigma\sqrt{2\pi}$  in the denominator would now be  $\sqrt{\pi}$

$$= \int_{-\infty}^{\infty} \underbrace{(\mu + \sigma\sqrt{2}t) \{1 - \text{expit}(\beta_0 + \beta_1(\mu + \sigma\sqrt{2}t))\}}_{f_1(t)} \cdot \frac{1}{\sqrt{\pi}} e^{-t^2} dt + c \int_{-\infty}^{\infty} \underbrace{\text{expit}(\beta_0 + \beta_1(\mu + \sigma\sqrt{2}t))}_{f_2(t)} \cdot \frac{1}{\sqrt{\pi}} e^{-t^2} dt$$

$\approx \sum_{i=1}^n w_i \cdot f(t_i)$  by Gauss-Hermite quadrature

$$\text{If we let } k = \beta_0 + \beta_1(\mu + \sigma\sqrt{2}t), \quad f_1(t_i) = (\mu + \sigma\sqrt{2}t_i) \{1 - \text{expit}(k)\} \frac{1}{\sqrt{\pi}}, \quad f_2(t_i) = c \cdot \text{expit}(k) \cdot \frac{1}{\sqrt{\pi}}$$

We can calculate numerical bias if we plug in parameter values  $\beta_0, \beta_1, \mu, \sigma, c$

Also,  $t_i$  and  $w_i$  can be computed with R

Note \*\*\*

If we let  $P[M_i = 1] = \pi_i$

$$\text{Then, } E(Y_i^*) = E_Y [Y_i \cdot (1 - \pi_i) + c \cdot \pi_i] = (1 - \pi_i) E_Y(Y_i) + \pi_i \cdot c \quad \text{if MCAR } (Y_i \perp \pi_i) \quad \begin{matrix} \pi_i \\ \text{Missingness not related to} \\ \text{any other values} \\ \text{(including } Y_i) \end{matrix}$$

$$= \mu - \mu \pi_i + c \pi_i = \mu \cdot (1 - \pi_i) + c \cdot \pi_i = \mu + (c - \mu) \cdot \pi_i$$

bias term added (under MCAR)

$\Rightarrow$  bias increases with the increase in absence rate  $\pi_i$

$\Rightarrow$  but, our study assumes MAR (Missingness depends on observed)