How do drunk driving laws affect traffic deaths?

Introduction:

According to National Highway Traffic Safety Administration, almost 30 people in the United States die in drunk-driving crashes daily — that is one person every 50 minutes. These deaths have fallen by a third in the last three decades (from 1980s); however, drunk-driving crashes claim more than 10,000 lives per year. This project investigates the importance of beer tax and other state laws on motor vehicle fatality rate. Effects of economic and demographic factors are also considered in the regression models to understand the true effects.

Background and Literature Review:

Ruhm, C. (1996) explored the impact of beer taxes and a variety of alcohol-control policies on motor vehicle fatality rates. Special attention was paid to omitted variables biases resulting from failing to adequately control for grassroots efforts to reduce drunk driving, the enactment of other state laws which simultaneously operate to reduce highway fatalities, and the economic conditions existing at the time the legislation is passed. In the preferred models, most of the regulations had little or no impact on traffic mortality. By contrast, higher beer taxes were associated with reductions in crash deaths and the result was relatively robust across specifications.

Stock, J.H. and Watson, M.W. (2007) explains the importance of fixed effect models (both state and time effects) in determining the economic estimates and summarises Ruhm, C (1996) research.

However, demographic factors are not included in the econometric models proposed before. Consequently, such variables are included in the regression models.

Data:

The dataset contains vehicle fatality rate for 48 states (except Alaska and Hawaii) annually for 1982 through 1988. The fatality rates have highly significant state effects. Although the time effects are significant but since we have data for only seven years, the yearly fatality trends for various states are very slightly decreasing. New Mexico has historically high fatality rates with median 36.89 fatalities per 100,000 residents and on the other hand, Rhode Island has the least fatality rates with median 11.26 fatalities per 100,000 residents. The dataset also contains fatality parameters for three different age groups – 15-17 year old, 17-20 year old and 21-24 year old. Night-time fatality rates and single vehicle fatality rates are also included.

Various parameters can be divided into three different categories – state laws, economic conditions, and demographic factors.

State laws include Beer Tax, Minimum Legal Drinking Age, Mandatory Jail Sentence and Mandatory Community Service. Tax on a case of beer is highest for Georgia and lowest for Wyoming. Similar to fatality rate, state effects are very striking for beer tax as well. Surprisingly, the state beer taxes are decreasing for added year for many states. According to an article in New York Times, the Secretary of Transportation was required to withhold 5% of Federal highway construction funds from those

states that did not enact a minimum drinking age of 21 by October 1, 1986. The Secretary was required to withhold 10% of the funds for states that did not act by October 1, 1987. Therefore, all the states except South Dakota have set the minimum legal drinking age to 21 complying to federal recommendations by 1988. Regarding legal actions, only Colorado and Florida have just the mandatory community service. By 1988, only one third of US states have mandatory community service and/or mandatory jail sentence.

Three major underlying economic conditions are included in the regression models. These are Unemployment Rate, Per Capita Personal Income and Change in Gross State Product. According to Ruhm, C. (1996), ignoring economic parameters results in omitted variable biases. These variables vary across states but generally follow national trends annually. Unemployment rates decreased and per capita personal income increased in the given period. The variance of per capita income also increased and therefore, use of logarithm of the per capita income for the models is needed. Change in gross state product is stationary complying to AR (1) trend.

Various demographic factors are also included in the dataset such as Per Capita Pure Alcohol Consumption and Average Mile per Driver. There is a gradual decreasing trend in per capita pure alcohol consumption for added year for most of the states. Nevada and New Hampshire consume almost 250% more alcohol than the national average. People in Wyoming, on average, commute the most and New Yorkers commute the least on the other hand. In the year 1984, the average mile per driver recorded for Nebraska was abnormally extremely high. Other variables that come in this grouping are % of Southern Baptists, % of Mormons, % of Residents in Dry Counties and % of Young Drivers (aged 15-24) in the 48 states.

Empirical Approach:

The dataset resembles that of balanced panel with 48 entities (cross-sectional units) and 7 time periods. Most of the econometric estimates in this report comes from entities and time fixed effect model of the form:

$$Y_{it} = \beta_i X_{it} + \alpha_i + \lambda_t + u_{it}$$

where Y_{it} is the fatality rate for state i and year t. X_{it} are the state laws, prevailing economic conditions, and demographic factors for state i and year t. α_i are the individual effects and λ_t are the time effects. These economic estimates are compared with those of random effect model using Nerlove estimator (1971) where individual and time effects are not fixed, i.e., relaxing $(\alpha_i - \alpha_i hat) = 0$ and $(\lambda_i - \lambda_i hat) = 0$.

Empirical Results:

Dependent Variable: Vehicle Fatality Rate (per 100,000)

Effects of Beer Tax:

• Model-1: Pooled OLS, Independent Variable: Beer Tax

Model-2: Fixed Time Effect, Independent Variable: Beer Tax

• Model-3: Fixed Entity Effect, Independent Variable: Beer Tax

• Model-4: Fixed Time and Entity Effect, Independent Variable: Beer Tax

Increasing beer tax is expected to have negative effect on vehicle fatality rates. Ignoring state effects have led to positive relation between the independent and dependant variable as given by Model-1 and Model-2. Model-3 includes state effects and Model-4 includes both state and time effects. The coefficient for Beer Tax in both those models is negative and supporting our assumption. By conducting ANOVA to test the joint significance of time effects, the null hypothesis is not rejected with (Pr > F) = 0.0642 at 10% significance level. The models in the following sections include fixed time and entity effects. Including state effects mitigates the threat of omitted variable bias due to cultural attitudes towards drinking and driving. With time effects, potential factors like safety innovations in vehicles are mitigated which do not vary across states.

	Dependent variable:				
; ; ;	OLS		all panel linear		
: :	Model-1	Model-2	Model-3	Model-4	
beertax	3.646***	3.663***	-6.559**	-6.400*	
!	(0.529)	(1.190)	(2.888)	(3.501)	
‡					
Constant	18.533***				
‡	(0.471)				
‡					
Observations	336	336	336	336	
R2	0.093	0.095	0.041	0.036	
Adjusted R2	0.091	0.075	-0.120	-0.149	
Residual Std. Error					
F Statistic	34.394*** (df = 1; 334)	34.246*** (df = 1; 328)	12.190*** (df = 1; 287)) $10.513***$ (df = 1;	

Effects of State Laws:

- Model-5: Fixed Entity Effect, Independent Variable: Beer Tax and Minimum Legal Age Drinking
- Model-6: Fixed Time and Entity Effect, Independent Variable: Beer Tax and Minimum Legal Age Drinking
- Model-7: Fixed Time and Entity Effect, Independent Variable: State Laws with dummy Punishment Policies
- Model-8: Fixed Time and Entity Effect, Independent Variable: State Laws without dummy Punishment Policies

The minimum legal drinking age is precisely estimated to have a small effect on traffic fatalities in all the models given below. The joint hypothesis that the coefficients on the minimum legal drinking age variables are zero and cannot be rejected at 10% significance level: (Pr > F) = 0.202. Model-5 and Model-6 show that time effects have no significant effect on the fatality rates. The coefficient on the legal punishment variables are also estimated to be small and are not significantly different from zero at 10% significance level as given by both Model-7 and Model-8. Although states implementing mandatory jail sentence have desired effect on the fatality rates as compared to states implementing only community service, but both the variables are insignificant. Surprisingly all the models show that state laws have no significant impact on the fatality rates except for beer tax.

	Dependent variable:				
		mra			
	Model-5	Model-6	Model-7	Model-8	
beertax	-6.460**	-6.897*	-7.320**	-7.176**	
	(2.963)	(3.555)	(3.399)	(3.438)	
mldabin[19,20)	-0.796	-0.599	-0.642	-0.629	
m2dd211(23/23/	(0.622)	(0.649)	(0.668)	(0.663)	
mldabin[20,21)	-0.986	-0.752	-0.808	-0.796	
midabin(20,21)	(0.728)	(0.766)	(0.783)	(0.781)	
mldabin[21,22]	-0.209 (0.673)	0.155 (0.804)	0.129 (0.834)	0.119 (0.827)	
	(0.0.0)	(01001)	(0.001)	(01027)	
jaild1			-0.240		
			(0.181)		
comserd1			1.485		
			(1.388)		
jailcomserd				0.885	
J				(1.080)	
Observations	336	336	336	336	
R2	0.056	0.052	0.060	0.057	
	-0.114 4 189*** (df - 4: 284)	-0.142 3.806*** (df = 4; 278)	-0.141 2 927*** (df - 6: 276)		

Effects of State Laws, Existing Economic Conditions and Demographic Factors:

- Model-9: Fixed Time and Entity Effect, Independent Variable: State Laws with dummy Punishment Policies and Existing Economic Conditions
- Model-10: Fixed Time and Entity Effect, Independent Variable: State Laws without dummy Punishment Policies and Existing Economic Conditions
- Model-11: Fixed Time and Entity Effect, Independent Variable: State Laws, Existing Economic Conditions and Demographic Factors

The economic variables have considerable explanatory power for traffic fatalities. High unemployment rates are associated with fewer fatalities and similarly high values of per capital personal income are associated with higher fatality rates. According to the estimates given by all the models given below, good economic conditions are associated with higher fatality rates, perhaps due to increased traffic density when the unemployment rate is low or greater alcohol consumption when income is high. This inference was made by Stock, J.H. and Watson, M.W. (2007). The two economic variables are significant at 0.1% significance level. One interesting result from Model-9 and Model-10 is that the effect of beer tax has reduced and became insignificant at 10% significance level. A potential source of omitted variable bias is that the effect of beer tax could pick up the effect of a broader campaign to reduce drunk driving. Therefore, the economic estimates were overstated in the previous models. Ruhm, C. (1999) stated that the influence of groups such as Mothers Against Drunk Driving, which lobby for strict regulatory policies but also attempt to reduce drunk driving in other ways (e.g. designated driver programs), grew substantially during the period investigated.

Introducing demographic factors in the regression models has resulted in understating the estimates for beer tax and unemployment rate. It is imperative to include all the important factors and to remove irrelevant variables. To find the BLUE estimators for the model, we need to control for per capita pure alcohol for each state for given year. The variable is significant at 1% significance level and excluding it will lead to omitted variable bias. Number of people living in dry counties has also considerable

effect on the fatalities at 5% significance level. There is a positive effect which can be attributed to the fact that residents in dry counties might travel to places where alcohol is served/sold and consequentially contribute to higher fatality rates. Religious affiliations are insignificant and average mile per driver has no bearing on fatalities. The coefficient on % of young drivers is also estimated to be small and is not significantly different from zero at 10% significance level as given by Model-11.

‡		Dependent variable:	
‡ ‡		mrall	
‡	Model-9	Model-10	Model-11
beertax	-4.605	-4.569	-2.714
‡	(2.903)	(2.893)	(2.686)
‡ ‡ mldabin[19,20	0.446	-0.442	0.113
# mrd#bin(19,20	(0.584)	(0.583)	(0.629)
‡			
mldabin[20,21		0.099	0.459
‡ ‡	(0.667)	(0.665)	(0.729)
] -0.263	-0.265	0.263
‡	(0.694)	(0.689)	(0.759)
‡			
‡ jaildl	0.144		
‡ ‡	(0.163)		
≠ ‡ comserdl	0.340		
‡	(1.298)		
‡			
jailcomserd		0.401	0.365
‡ ‡		(0.998)	(0.973)
unrate	-0.639***	-0.641***	-0.441***
:	(0.128)	(0.129)	(0.126)
log(perinc)		18.409***	18.800***
: :	(6.268)	(6.253)	(5.423)
gspch	-0.618	-0.631	1.628
	(3.527)	(3.509)	(3.473)
spircons			7.909***
‡ ‡			(1.247)
vmiles			0.0001
			(0.0001)
‡			
sobapt			-0.455
‡ ‡			(0.614)
mormon			0.119
‡			(0.440)
.			
dry			0.197**
‡ ‡			(0.093)
yngdrv			-5.956
‡			(10.582)
‡			
	226	225	225
# Observations # R2	336 0.358	336 0.358	336 0.466
Adjusted R2		0.215	0.333
_	16.921*** (df = 9; 273)		

Final Models:

- Model-12: Fixed Time and Entity Effect, Independent Variable: State Laws and Important Variables
- Model-13: Fixed Time and Entity Effect, Independent Variable: Beer Tax and Important Variables
- Model-14: Fixed Entity Effect, Independent Variable: Beer Tax and Important Variables
- Model-15: Random Time and Entity Effect, Independent Variable: Beer Tax and Important Variables

The models include important economic and demographic factors such as unemployment rate, log of per capita personal income, per capita pure alcohol and % of residents in dry counties. State laws are not significant at 10% significance level as given by Model-12. By conducting ANOVA to test the joint significance of time effects on Model-13 and Model-14, the null hypothesis is rejected at 0.1% significance level. Hausman test on Model-13 and Model-15 rejects the null hypothesis of no endogeneity and concludes that fixed effect model should be chosen over random effect model (Chisq = 21.88; p-value = $5x10^{-4}$). Chow Test for joint significance on full specification model (Model-11) and chosen specification model (Model-13) has not rejected the null hypothesis with (Pr > F = 0.82) at 10% significance level. Therefore, the chosen model (Model-13) explains the model in a simper sense following Occam's Razor Principle.

##		egression Models of Traf		unk Driving		
##	Dependent variable:					
##			mrall			
##		Model-12	Model-13	Model-14	Model-15	
##	beertax	-3.629	-3.705*	-3.940*	-2.562	
##		(2.217)	(2.240)	(2.089)	(1.779)	
##						
	mldabin[19,20)					
##		(0.652)				
	mldabin[20,21)	0.578				
##	midabin[20,21)	(0.736)				
##		(
	mldabin[21,22]	0.304				
##		(0.760)				
##		0.405				
##	jailcomserd	0.406 (0.926)				
##		(0.920)				
	unrate	-0.443***	-0.445***	-0.166*	-0.478***	
##		(0.113)	(0.107)	(0.097)	(0.102)	
##						
	log(perinc)	19.240***	18.857***	20.534***	14.795***	
## ##		(5.061)	(4.864)	(4.726)	(4.917)	
	spircons	7.741***	7.702***	8.450***	6.818***	
##		(1.189)	(1.173)	(1.162)	(0.926)	
##						
	dry	0.187**	0.183**	0.243**	0.237***	
##		(0.086)	(0.084)	(0.109)	(0.071)	
##	Constant				-128.661***	
##	Constant				(47.249)	
##					(17.12.13)	
##						
##	Observations	336	336	336	336	
	R2	0.460	0.456	0.386	0.393	
	Adjusted R2 F Statistic	0.337	0.342 46.420*** (df = 5; 277)	0.274	0.384	
	F Statistic	(ul - 9; 2/3)	(ar - 5; 2//)	(a1 - 5; 20	3/ 213.022~^^	
	Note:			*p<0.1; **p<0	.05; ***p<0.01	
					•	

Conclusion:

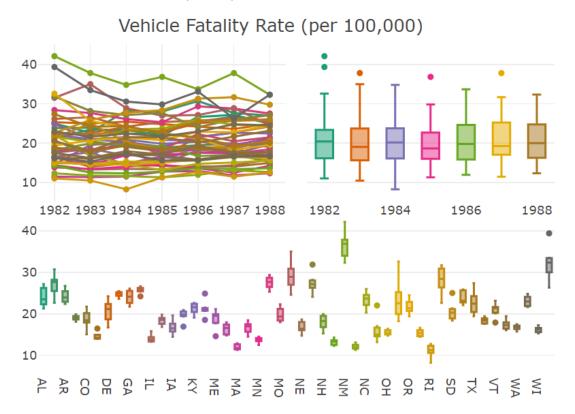
State effects play a major part in determining the fatality rates. The regression results show that neither stiff punishments nor increase in minimum legal drinking age have important effects on fatalities. In contrast, economic variables like increase in unemployment rates and decrease in per capita personal income have desired effect on fatalities. Controlling for net alcohol consumption and economic conditions have resulted in mitigating the omitted variable bias and reduced imprecision of the estimated beer tax. There is a significant impact of percentage of residents in dry counties on the fatalities. The estimates become insignificant while estimating night-time vehicle fatality rates except for unemployment rates and per capita alcohol consumption. Beer tax has increased impact on the underage fatality rates but is insignificant. Minimum legal drinking age variables have significant effect on underage fatalities, in addition to per capita personal income and per capita alcohol consumption. However, caution must be observed while drawing policy conclusions.

References:

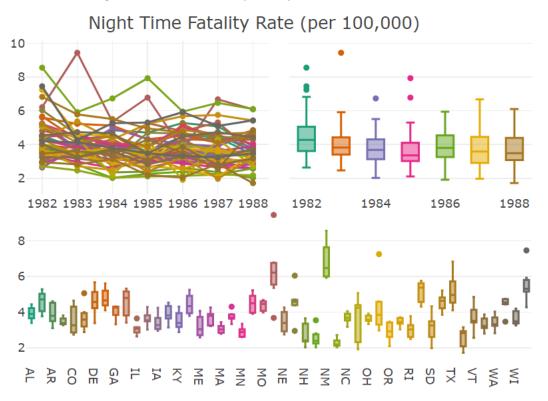
- 1. Drunk Driving. https://www.nhtsa.gov/risky-driving/drunk-driving
- 2. Ruhm, C. J. (1996). Alcohol Policies and Highway Vehicle Fatalities. Journal of Health Economics, 15, 435--454
- 3. Stock, J. H. and Watson, M. W. (2007). Introduction to Econometrics, 3rd ed. Boston: Addison Wesley
- 4. Reagan signs law linking federal aid to drinking age. https://www.nytimes.com/1984/07/18/us/reagan-signs-law-linking-federal-aid-to-drinking-age.html
- 5. Christoph Hanck, Martin Arnold, Alexander Gerber and Martin Schmelzer (2019). Introduction to Econometrics with R. https://www.econometrics-with-r.org/

Appendix:

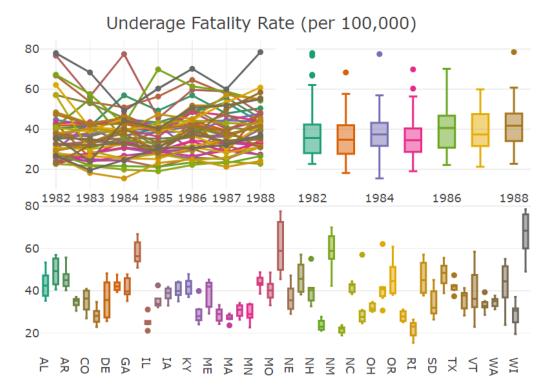
Dependent variable: Vehicle Fatality Rate (per 100,000)



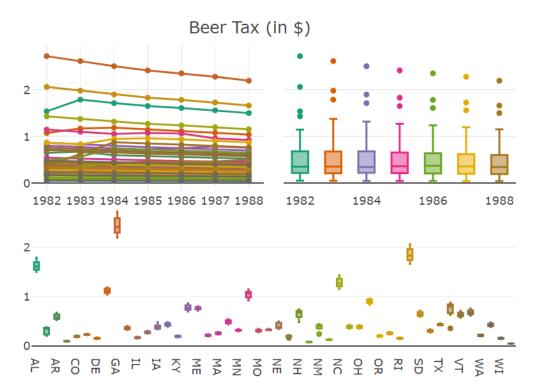
Dependent variable: Night-time Vehicle Fatality Rate (per 100,000)



Dependent variable: Underage Fatality Rate (per 100,000)



Independent variable: Beer Tax (in \$)



Independent variable: Minimum Legal Drinking Age



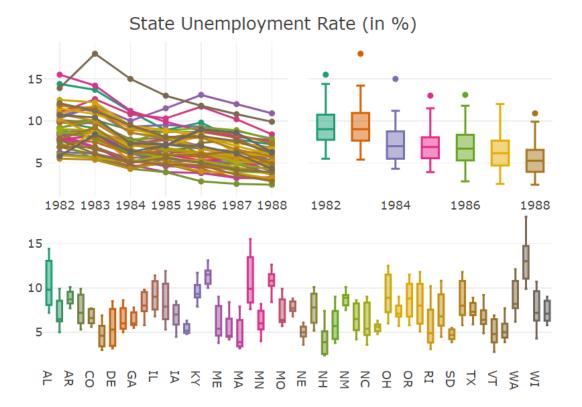
Independent variable: States with only Mandatory Community Service

Year	States Implementing Only Mandatory Community Service
1982	CO, FL
1983	CO, FL
1984	CO, FL
1985	CO, FL
1986	CO, FL
1987	CO, FL
1988	CO, FL

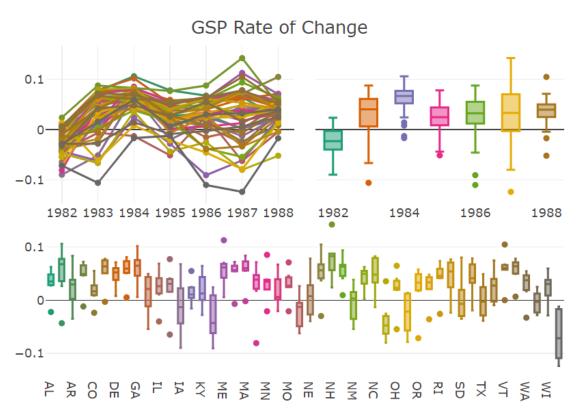
Independent variable: States with Mandatory Community Service or/and Mandatory Jail Sentence

Year	States Implementing either Community Service or Jail Sentence
1982	AZ, CO, FL, KS, LA, ME, MT, TN, WA, WV, WY
1983	AZ, CO, FL, KS, LA, ME, MT, NV, OH, SC, TN, UT, WA, WV, WY
1984	AZ, CO, FL, KS, LA, ME, MT, NV, OH, OR, SC, TN, UT, WA, WV, WY
1985	AZ, CO, CT, FL, KS, LA, ME, MT, NV, OH, OR, SC, TN, UT, WA, WV, WY
1986	AZ, CO, CT, FL, KS, LA, ME, MT, NV, OH, OR, SC, TN, UT, WA, WV, WY
1987	AZ, CO, CT, FL, KS, LA, ME, MT, NV, OR, SC, TN, UT, WA, WV, WY
1988	AZ, CO, CT, FL, KS, LA, ME, MT, NV, OR, SC, TN, UT, WA, WV, WY

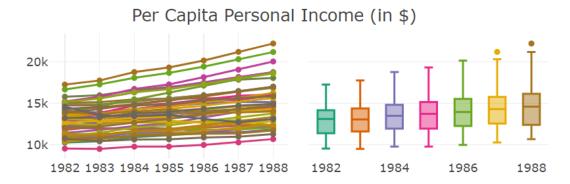
Independent variable: State Unemployment Rate (in %)

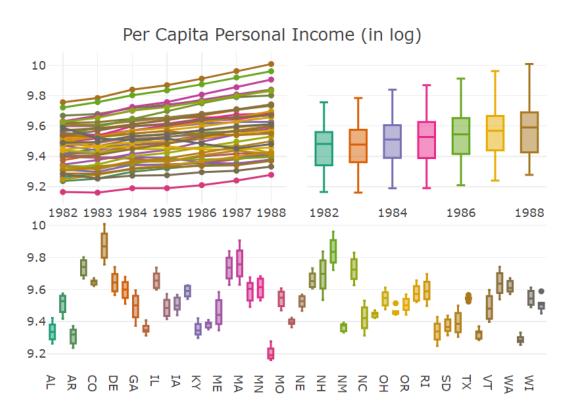


Independent variable: GSP Rate of Change

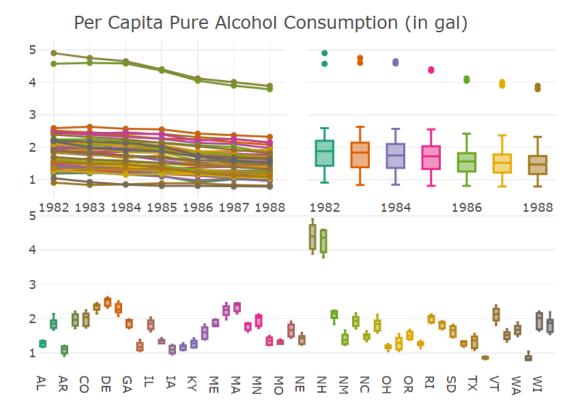


Independent variable: Per Capita Personal Income

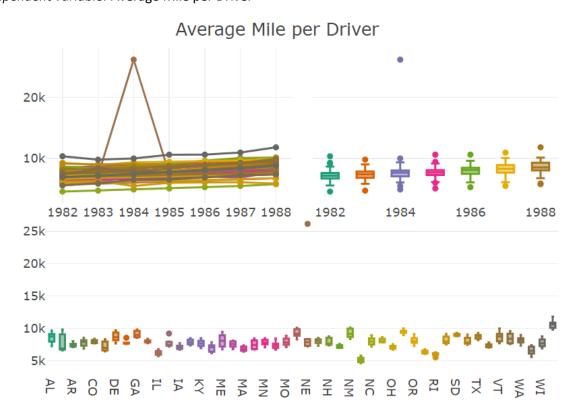




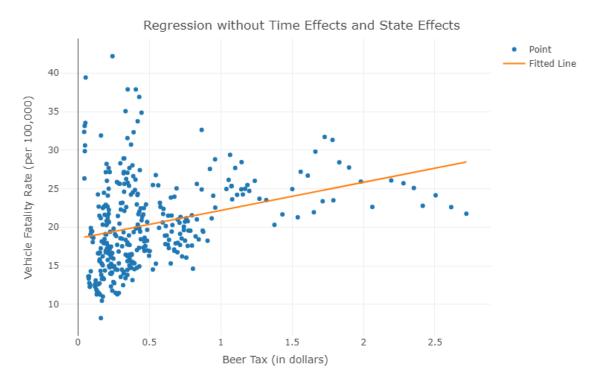
Independent variable: Per Capita Pure Alcohol Consumption (in Gallons)



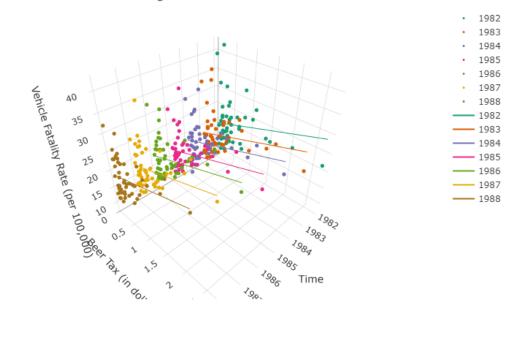
Independent variable: Average Mile per Driver



Regression: Dependent: Vehicle Fatality Rate (per 100,000); Independent: Beer Tax (\$)



Regression with Time Effects



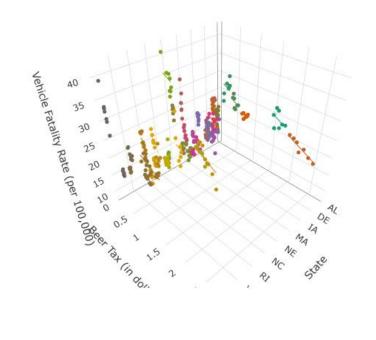
Regression with Individual Effects

CT DE

GA ID IL

IA KS KY LA

ME MA MI



Regression (Fixed Time and Entity Effects): Dependent: Night-Time Vehicle Fatality Rate (per 100,000)

	Dependent variable:					
‡ ‡	mralln					
‡	Model-16	Model-17	Model-18	Model-19		
beertax	-1.321	-1.322	-1.006	-1.105		
#	(0.805)	(0.829)	(0.814)	(0.898)		
‡				,		
mldabin[19,20))	-0.183	-0.185	-0.138		
ŧ		(0.150)	(0.186)	(0.181)		
:						
mldabin[20,21)		-0.073	0.019	0.077		
		(0.167)	(0.229)	(0.224)		
mldabin[21,22]	I	-0.080	-0.153	-0.078		
		(0.191)	(0.215)	(0.215)		
jailcomserd		-0.055	-0.125	-0.138		
jalicomserd		(0.274)	(0.271)	(0.273)		
		(0.2/3)	(0.2/2)	(0.270)		
unrate			-0.102**	-0.077*		
			(0.043)	(0.047)		
log(perinc)			1.728	2.251		
			(1.663)	(1.656)		
‡						
gspch			1.101	1.322		
			(2.396)	(2.438)		
•						
spircons				0.852**		
				(0.416)		
: 1				-0.00000		
vmiles				-0.00003 (0.00003)		
•				(0.00003)		
sobapt				0.055		
t sobapt				(0.223)		
				(0.220)		
mormon				-0.063		
				(0.150)		
				,		
dry				0.031		
				(0.045)		
yngdrv				-3.388		
:				(2.822)		
‡						
ŧ						
Observations	336	336	336	336		
R2	0.015	0.019	0.081	0.101		
Adjusted R2 F Statistic	-0.174	-0.187	-0.123	-0.124		
		OEG /JE - E. OEEL	3.029*** (df = 8; 274	1 2 14044 /JE - 14.		

Regression (Fixed Time and Entity Effects): Dependent: Underage Fatality Rate (per 100,000)

f ======= f	Dependent variable:					
*						
‡ ‡	Model-20	mr Model-21	Model-22	Model-23		
# # beertax	-16.054**	-12.372*	-6.556	-7.302		
ŧ	(6.738)	(6.954)	(7.446)	(7.056)		
# # mldabin[19,20)	10.415***	10.520***	11.852***	11.965***		
#	(1.977)	(1.927)	(1.950)	(1.883)		
#	0. 000+++	10.055444	10 000+++	10.055444		
# mldabin[20,21) #	9.792*** (2.286)	10.955*** (2.178)	12.000*** (2.272)	12.265*** (2.239)		
*	(2.200)	(2.170)	(2.272)	(2.235)		
	10.529***	9.808***	10.877***	11.117***		
ŧ	(2.424)	(2.235)	(2.257)	(2.176)		
f fisilcomeand	1.497	0.721	1.097			
# jailcomserd #	(2.460)	(2.337)	(2.298)			
 	(21100)	(21007)	(21230)			
# unrate		-1.009**	-0.499	-0.619*		
ŧ		(0.396)	(0.362)	(0.369)		
			00 5514	0.5 5484		
log(perinc) 		23.334 (16.514)	23.751* (14.372)	26.647* (15.014)		
,		(10.314)	(14.572)	(13.014)		
# gspch		8.591	15.919			
ŧ		(15.027)	(13.110)			
#						
# spircons			19.207***	18.671***		
f			(4.010)	(3.769)		
* # vmiles			0.0001			
ŧ			(0.0002)			
ŧ						
# sobapt			-0.664			
‡			(1.851)			
# # mormon			1.600*	1.384		
# MOIMON			(0.972)	(0.858)		
ŧ						
# dry			0.046			
!			(0.229)			
# #						
# Observations	336	336	336	336		
# R2	0.103	0.179	0.243	0.238		
# Adjusted R2	-0.084	-0.003	0.058	0.069		
		E 100444 (15 0 0E1)	C C40444 /46 10- 00/	9) 10.719*** (df = 8; 27		