

7/3/13

4. Correlation

In a bivariate data if the change in one variable create, any change in the other variable, the two variables are said to be correlated. It is of 2 types.

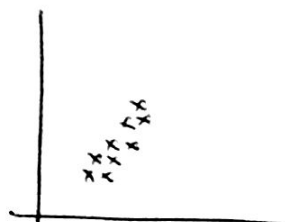
(i) Positive correlation: If the change in two variables is in the same direction, it is called as 've correlation.

Ex: $x \uparrow, y \uparrow$ & $x \downarrow, y \downarrow$. (Income & Expenditure)

Negative correlation: If the change is in opposite directions, then it is -ve correlation.

Ex: $x \uparrow, y \downarrow$; $x \downarrow, y \uparrow$ (Pressure & Volume)

Scattered diagrams: Used to get an idea about correlation.



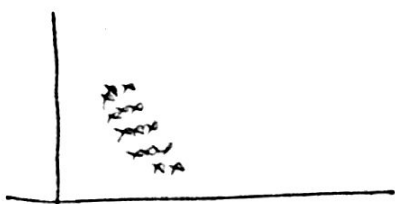
high +ve correlation.



positive.



perfect positive.



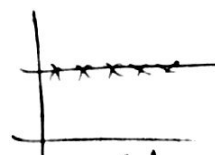
high -ve correlation



negative



perfect -ve.



null correlation

Karl-Pearson's correlation coefficient:

This is a numerical measurement for the linear relationship of x and y . It is denoted by r and is given as

$$r = \frac{\text{cov}(x, y)}{\sigma_x \sigma_y}$$

$$\frac{E[(X - E(X))(Y - E(Y))]}{\sqrt{E(X - E(X))^2} \sqrt{E(Y - E(Y))^2}}$$

$$= \frac{\frac{1}{n} \sum xy - \bar{x}\bar{y}}{\sqrt{\frac{1}{n} \sum x^2 - \bar{x}^2} \sqrt{\frac{1}{n} \sum y^2 - \bar{y}^2}}$$

Note:

$$-1 \leq r \leq +1$$

Find the correlation coefficient.

X	Y	XY	X ²	Y ²
80	44	3520	6400	1936
90	32	2880	8100	1024
65	75	4875	4225	5625

To find the correlation coefficient, we calculate the foll. table.

$$\sum x^2 = 18725$$

$$\sum xy = 11275$$

$$\sum y^2 = 8585$$

$$\bar{x} = 78.33$$

$$\text{Cov}(x, y) = \frac{1}{n} \sum xy - \bar{x}\bar{y}$$

$$\bar{y} = 50.33$$

$$= \frac{1}{3} (11275) - (78.33)(50.33) = -184.02$$

$$\sigma_x = \sqrt{\frac{1}{n} \sum x^2 - \bar{x}^2} = 10.30$$

$$\sigma_y = \sqrt{\frac{1}{n} \sum y^2 - \bar{y}^2} = 18.13$$

$$\therefore \text{correlation coeff} = \underline{\underline{-0.99}}$$

The two variables are negatively correlated and correlation is high.

correlation is high if (>0.5).

Properties of ρ
 * Correlation coeff. is independent of change of origin and scale.

Let x, y are two variables for which the correlation coefficient is $\rho_{x,y}$.

Consider $U = \frac{x-a}{h}$, $V = \frac{y-b}{k}$ where a, b, h, k are constants.
 we can write $x = a + Uh \Rightarrow E(x) = a + hE(U)$

$$\Rightarrow x - E(x) = a + Uh - a - hE(U) = h[U - E(U)].$$

Similarly $y - E(y) = k[V - E(V)]$.

We know that

$$\begin{aligned} \rho_{x,y} &= \frac{E[(x - E(x))(y - E(y))]}{\sqrt{E(x - E(x))^2} \sqrt{E(y - E(y))^2}} \\ &= \frac{E[h(U - E(U))k(V - E(V))]}{\sqrt{E(h(U - E(U)))^2} \sqrt{E(k(V - E(V)))^2}} \\ &= \frac{kk E[(U - E(U))(V - E(V))]}{k \sqrt{E(U - E(U))^2} k \sqrt{E(V - E(V))^2}} \\ &= \frac{\text{cov}(U, V)}{\sigma_U \sigma_V} = \rho_{UV}. \end{aligned}$$

* correlation coefficient lies b/w -1 and $+1$

$$\boxed{-1 \leq \rho \leq 1}$$

$$\rho_{x,y} = \frac{E[(x - E(x))(y - E(y))]}{\sqrt{E(x - E(x))^2} \sqrt{E(y - E(y))^2}}$$

$$\text{Let } (X - E(X)) = U, \quad Y - E(Y) = V$$

$$\rho_{XY} = \frac{E(UV)}{\sqrt{E(U)^2} \sqrt{E(V)^2}}$$

$$\rho_{XY}^2 = \frac{(E(UV))^2}{E(U)^2 E(V)^2}$$

from Cauchy-Schwarz inequality, we have

$$(E(XY))^2 \leq E(X^2) E(Y^2)$$

$$\rho_{XY}^2 \leq 1$$

$$\Rightarrow \boxed{-1 \leq \rho \leq 1}$$

* For independent random variable, correlation is 0 but the converse is not true.

$$\rho_{XY} = \frac{E[(X - E(X))(Y - E(Y))]}{\sqrt{E(X - E(X))^2} \sqrt{E(Y - E(Y))^2}}$$

$$E[(X - E(X))(Y - E(Y))]$$

$$= E[XY - XE(Y) - E(X)Y + E(X)E(Y)]$$

$$= E(XY) - E(X)E(Y) - E(X)E(Y) + E(X)E(Y)$$

$$= E(XY) - E(X)E(Y)$$

$$= 0 \quad \left[\text{if } X \text{ \& } Y \text{ are independent} \right]$$

$$\rho = 0$$

Though $\rho = 0$, here $Y = X^2$ i.e., there may exist other than linear relationship b/w X and Y .

Spearman's Rank Correlation coefficient:

It is useful to measure the correlation b/w 2 qualitative variables. It is denoted by ρ given as

$$\rho = 1 - \frac{6 \sum d_i^2}{n(n^2 - 1)}$$

$$\boxed{-1 \leq \rho \leq 1}$$

$$\rho = \frac{1 - 6 \left(\sum d_i^2 + \frac{m(m^2-1)}{12} \right)}{n(n^2-1)}$$

Regression: It is useful to estimate the value of one variable for a given value of other variable.

Regression of Y on X: It is useful to estimate of Y, for a given value of X, and is given as

$$(Y - \bar{Y}) = b_{YX} (X - \bar{X})$$

b_{YX} is regression coefficient, $b_{YX} = \frac{\text{cov}(X, Y)}{V(X)}$

$$b_{YX} = \frac{\frac{1}{n} \sum xy - \bar{x}\bar{y}}{\frac{1}{n} \sum x^2 - \bar{x}^2} = \frac{r \sigma_Y}{\sigma_X}$$

Regression line of X on Y is useful to estimate the value of X for a given value of Y and is given as

$$(X - \bar{X}) = b_{XY} (Y - \bar{Y})$$

$$b_{XY} = \frac{\text{cov}(X, Y)}{V(Y)} = \frac{\frac{1}{n} \sum xy - \bar{x}\bar{y}}{\frac{1}{n} \sum y^2 - \bar{y}^2} = \frac{r \sigma_X}{\sigma_Y}$$

$$= \frac{\text{cov}(X, Y)}{\sigma_Y^2}$$

14/8/18 Properties of Regression coefficients:

1. Correlation coefficient is the geometric mean of regression coefficients.

$$b_{xy} = \frac{\text{cov}(X, Y)}{V(Y)} \quad b_{yx} = \frac{\text{cov}(X, Y)}{V(X)}$$

$$\begin{aligned} \text{GM} &= \sqrt{b_{xy} b_{yx}} = \sqrt{\frac{\text{cov}(X, Y)}{V(Y)} \cdot \frac{\text{cov}(X, Y)}{V(X)}} \\ &= \frac{\text{cov}(X, Y)}{\sigma_Y \sigma_X} = r_{xy} \end{aligned}$$

2. If one of the regression coeff. is greater than unity, the other is less than unity.

$$\text{let } b_{xy} > 1$$

$$\begin{aligned} \text{wkt } r &= \sqrt{b_{xy} b_{yx}} \Rightarrow r^2 = b_{xy} b_{yx} \leq 1 \\ &\Rightarrow b_{xy} \leq \frac{1}{b_{yx}} < 1 \end{aligned}$$

3. The arithmetic mean of regression coefficient is greater than that of correlation coefficient

$$\begin{aligned} \text{Assume that } \frac{b_{xy} + b_{yx}}{2} &> r \\ b_{xy} + b_{yx} &> 2r \end{aligned}$$

$$\frac{\cancel{\sigma_X}}{\sigma_Y} + \frac{\cancel{\sigma_Y}}{\sigma_X} > 2r \Rightarrow \frac{\sigma_X^2 + \sigma_Y^2}{\sigma_X \sigma_Y} > 2$$

$$\sigma_X^2 + \sigma_Y^2 - 2\sigma_X \sigma_Y > 0$$

$$\Rightarrow (\sigma_X - \sigma_Y)^2 > 0$$

A squared quantity is always $> 0 \Rightarrow$ Assumption is true.

4. Regression coefficients are independent of change of origin but not change of scale.

$$U = \frac{X-a}{h}, \quad V = \frac{Y-b}{k}$$

$$\text{wkt } b_{xy} = \frac{\text{COV}(X,Y)}{V(Y)} = \frac{E((X-E(X))(Y-E(Y)))}{E(Y-E(Y))^2}$$

$$X = a + Uh \quad Y = kV + b$$

$$E(X) = a + hE(U)$$

$$\Rightarrow X - E(X) = h(U - E(U))$$

$$\text{III } Y - E(Y) = k(V - E(V))$$

$$\begin{aligned} b_{xy} &= \frac{E[h(U - E(U))k(V - E(V))]}{E(k(V - E(V)))^2} \\ &= \frac{h \cancel{k} E((U - E(U))(V - E(V)))}{k \cancel{k} E(V - E(V))^2} \\ &= \frac{h}{k} b_{uv} \end{aligned}$$

5. Angle b/w 2 regression lines.

$$\tan \theta = \frac{m_1 - m_2}{1 + m_1 m_2}$$

$$(y - \bar{y}) = b_{yx}^{m_1} (x - \bar{x})$$

$$(x - \bar{x}) = b_{xy} (y - \bar{y}) \Rightarrow (y - \bar{y}) = \left(\frac{1}{b_{xy}} \right) (x - \bar{x})$$

$$\tan \theta = \frac{b_{yx} - \frac{1}{b_{xy}}}{1 + b_{yx} \cdot \frac{1}{b_{xy}}}$$

$$\frac{\frac{\lambda \sigma_y}{\sigma_x} - \frac{1}{\lambda \frac{\sigma_x}{\sigma_y}}}{1 + \frac{\lambda \frac{\sigma_y}{\sigma_x} + \frac{\lambda \frac{\sigma_x}{\sigma_y}}{\lambda \frac{\sigma_x}{\sigma_y}}} = \frac{\lambda^2 - 1}{\frac{\sigma_x}{\sigma_y} \left[1 + \frac{\sigma_y^2}{\sigma_x^2} \right]} \left(\frac{1}{\lambda} \right)$$

$$\tan \theta = \left(\frac{\lambda^2 - 1}{\lambda} \right) \frac{\sigma_x \sigma_y}{\sigma_x^2 + \sigma_y^2}$$

$$\theta = \tan^{-1} \left[\left(\frac{\lambda^2 - 1}{\lambda} \right) \left(\frac{\sigma_x \sigma_y}{\sigma_x^2 + \sigma_y^2} \right) \right]$$

Case 1: If $\lambda = \pm 1, \Rightarrow \theta = \tan^{-1}(0) = 0$

\therefore The two regression lines completely coincide with each other.

Case 2: If $\lambda = 0 \Rightarrow \theta = \pi/2$ then the two regression lines are \perp to each other.

Small Samples.

(i) t -test for testing the significance of single mean.

(ii) Null hypothesis: The population mean is equal to given value.

$$H_0: \mu = \mu_0.$$

(iii) Alternative hypothesis: The population mean is not equal to given value.

(iv) Fix the level of significance, α .

(v) Under H_0 , compute the test statistic.

$$t = \frac{\bar{x} - \mu_0}{s/\sqrt{n}}$$

where $\bar{x} = \frac{1}{n} (\sum x_i) = \text{sample mean}$.

$s^2 = \text{sample mean sum of squares}$.

$$s^2 = \frac{1}{(n-1)} \sum (x_i - \bar{x})^2 = \frac{n s^2}{n-1}$$

(vi) $v = \text{degrees of freedom} = n-1$

If t_{cal} is in acceptance region, accept H_0 .

Or else reject H_0 .

(i) t test for 2 popⁿ means differences.

(i) Null hypothesis: two population means are equal
(No significant difference)

$$H_0: \mu_1 = \mu_2$$

(ii) Alternative hypothesis: The two population means are not equal.

$$H_1: \mu_1 \neq \mu_2$$

(iii) Fix the level of significance, α .

(iv) Under H_0 , compute t ,

$$t = \frac{(\bar{x}_1 - \bar{x}_2) - (\mu_1 - \mu_2)}{\sqrt{s^2 \left(\frac{1}{n_1} + \frac{1}{n_2} \right)}} \sim t_{n_1 + n_2 - 2}$$

$$\text{where } s^2 = \frac{n_1 s_1^2 + n_2 s_2^2}{n_1 + n_2 - 2}$$

$$= \frac{\sum (x_i - \bar{x})^2 + \sum (y_i - \bar{y})^2}{n_1 + n_2 - 2}$$

$\bar{x}_1 \rightarrow$ first sample mean.

$n_1 \rightarrow$ first sample size.

$\bar{x}_2 \rightarrow$ second sample mean.

$n_2 \rightarrow$ second sample size.

$v =$ degrees of freedom $= n_1 + n_2 - 2$.

If t_{cal} is in accepted region, accept H_0 .
otherwise reject H_0 .

(iii) Paired t-test or t-test for dependent samples.
The test statistic is $t = \bar{d} / (s/\sqrt{n}) \sim t_{n-1}$ where

$$d_i = x_i - y_i \Rightarrow \bar{d} = \frac{1}{n} \sum d_i$$

$$s^2 = \frac{1}{n-1} \sum (d_i - \bar{d})^2$$

iii t test for correlation coefficient

Null hypothesis: variables are null correlated.

$$H_0: \rho = 0$$

Alternative hypothesis: Variables are correlated

$$H_1: \rho \neq 0$$

Under H_0 , test statistic is

$$t = \frac{r - \rho}{\sqrt{\frac{1-r^2}{n-2}}} \sim t_{n-2}$$

If t_{cal} is in acceptance region, accept H_0 .
otherwise reject H_0 .

F-test for two population variances.

- (i) Null hypothesis: The two population variances are equal. $H_0: \sigma_1^2 = \sigma_2^2$
- (ii) Alternative hypothesis: The two population variances are not equal.
- (iii) Fix the level of significance (α).
- (iv) Compute the test statistic,

$$F = \frac{S_1^2}{S_2^2} \sim F_{(n_1-1, n_2-1)} \quad \text{if } S_1^2 > S_2^2$$

$(v_1) \quad (v_2)$

$$= \frac{S_2^2}{S_1^2} \sim F_{(n_2-1, n_1-1)} \quad \text{if } S_2^2 > S_1^2$$

$(v_1) \quad (v_2)$

$$\text{where } S_1^2 = \frac{1}{n_1-1} \sum (x_i - \bar{x})^2$$

$$S_2^2 = \frac{1}{n_2-1} \sum (y_i - \bar{y})^2$$

- (v) If F_{cal} is in acceptance region, accept H_0 .
otherwise reject H_0 .

Note: i) F must be always greater than unity.

ii) For a 2 tail test, let $\alpha =$ either 10% or 20%
for left tail critical value, $F_{\alpha}(v_1, v_2) = \frac{1}{F_{1-\alpha}(v_2, v_1)}$

Anova - I way.

The observations of the data are different with respect to one factor i.e. either row or column.

$E_x:$						R_i
	X	40	50	80	60	230
	Y	45	50	80		175
	Z	45	55	85		185
						<hr/>
						590 \rightarrow G.T (Grand Total)

$$\text{Correction factor} = \frac{G^2}{\text{no. obs.}} = \frac{(590)^2}{10} \\ = 34810$$

Sum of squares according to rows

$$(SSR) = \sum \frac{R_i^2}{n_i} - \text{c.f.}$$

$$= \frac{(230)^2}{4} + \frac{(175)^2}{3} + \frac{(185)^2}{3} - 34810 \\ = 31.667$$

Total sum of squares (TSS)

$$= \sum \sum y_{ij}^2 - \text{c.f.} \\ = 2490$$

Anova 1-way Table

Source of Variation	Sum of Squares (S.S)	degrees of freedom (d.f)	Mean sum of squares (S.S/d.f)	F_{cal}	F_{table}
1. Rows	31.66	$3 - 1 = 2$	15.83	22.182	99.9
2. Error	2458.34	$9 - 2 = 7$	351.143		
Total	2490	$10 - 1 = 9$	276.68		

Rows - Total = error

degrees of freedom = d.f = no. of rows - 1.

$$F_{cal} = \frac{351.14}{15.83} \sim F_{(7,2)} = 22.182$$

Ans

Anova 2 way table

	S	R	W	R_i
A	8	4	3	15
B	5	6	7	18
C	9	2	4	15
D	5	8	6	19
C_j	27	20	20	67 $\rightarrow G.T$

For testing the quality of salesman.
A, B, C, D.

H_{01} : The sales of all salesman are equal.

H_{11} : They are not equal.

H_{02} : For the quality of seasons
(Summer, Rainy, Winter).

H_{12} : All the these seasons are not equal.

Same process of Anova I way table is continued where as a new row of 'column' is added to the table.