## Topic 24 - Multifactor Studies

STAT 525 - Fall 2013

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### Outline

- 3-factor studies
  - Data
  - Model
  - Parameter Estimates
  - Inference
- Unequal sample size

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## Data for Three Factor ANOVA

- $\bullet$  Y is the response variable
- Factor A has levels i = 1, 2, ..., a
- Factor B has levels j = 1, 2, ..., b
- Factor C has levels k = 1, 2, ..., c
- $Y_{ijkl}$  is the  $l^{th}$  observation from cell (i, j, k)
- Now  $l = 1, 2, ..., n_{ijk}$

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## Example Page 1005

- Influence of several factors on exercise tolerance
- Considered age range: 25-35 years old
- Three factors were
  - Gender (a=2)
  - Percent body fat (b=2)
  - Smoking history (c=2)
- ullet Y is exercise tolerance (minutes until fatigue) when doing a bicycle test
- Set up as balanced design (n = 3)

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### General Plan

- Construct scatterplot / interaction plots
- Run full model
- Check assumptions
  - Residual plots
  - Histogram / QQplot
  - Ordered residuals plot
- Check significance of interaction

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```
data a1; infile 'u:\.www\datasets525\CH24TA04.txt';
   input extol gender fat smoke;
proc print;
data a1: set a1:
   if (gender eq 1)*(fat eq 1)*(smoke eq 1) then gfs='1_Mfs';
   if (gender eq 1)*(fat eq 2)*(smoke eq 1) then gfs='2_MFs';
   if (gender eq 1)*(fat eq 1)*(smoke eq 2) then gfs='3_MfS';
   if (gender eq 1)*(fat eq 2)*(smoke eq 2) then gfs='4_MFS';
   if (gender eq 2)*(fat eq 1)*(smoke eq 1) then gfs='5_Ffs';
   if (gender eq 2)*(fat eq 2)*(smoke eq 1) then gfs='6_FFs';
   if (gender eq 2)*(fat eq 1)*(smoke eq 2) then gfs='7_FfS';
   if (gender eq 2)*(fat eq 2)*(smoke eq 2) then gfs='8_FFS';
title1 'Plot of the data';
symbol1 v=circle i=none c=black;
proc gplot data=a1;
   plot extol*gfs/frame;
run;
```

**SAS** Commands

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## Output

```
        Obs
        extol
        gender
        fat
        smoke

        1
        24.1
        1
        1
        1

        2
        29.2
        1
        1
        1

        3
        24.6
        1
        1
        1

        4
        20.0
        2
        1
        1

        5
        21.9
        2
        1
        1

        6
        17.6
        2
        1
        1

        7
        14.6
        1
        2
        1

        8
        15.3
        1
        2
        1

        9
        12.3
        1
        2
        1

        10
        16.1
        2
        2
        1

        11
        9.3
        2
        2
        1

        12
        10.8
        2
        2
        1

        13
        17.6
        1
        1
        2

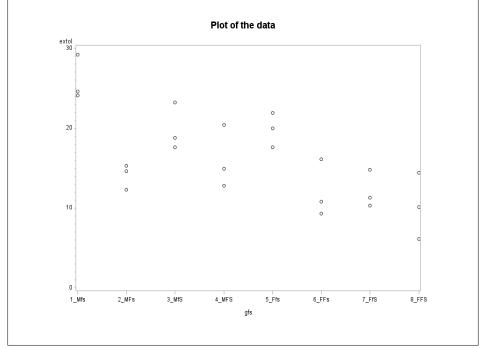
        14
        18.8
        1
        1
        2

        15
        23.2
        1
        1
        2

        16
        14.8
        2
        1
        2

        17
```

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## **SAS** Commands

```
proc sort data=a1; by gender fat smoke;

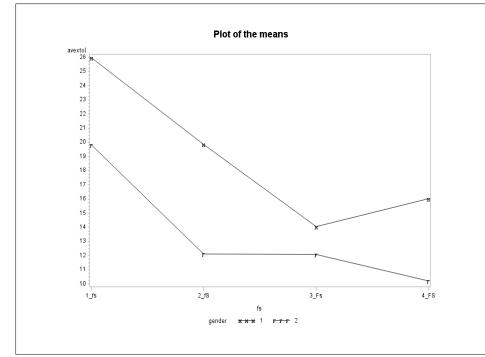
proc means data=a1;
  output out=a2 mean=avextol;
  by gender fat smoke;

data a2; set a2;
  if (fat eq 1)*(smoke eq 1) then fs='1_fs';
  if (fat eq 1)*(smoke eq 2) then fs='2_fS';
  if (fat eq 2)*(smoke eq 1) then fs='3_Fs';
  if (fat eq 2)*(smoke eq 2) then fs='4_FS';
run;

proc sort data=a2; by fs; title1 'Plot of the means';
symbol1 v='M' i=join c=black; symbol2 v='F' i=join c=black;
proc gplot data=a2;
  plot avextol*fs=gender/frame;
run;
```

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### The Cell Means Model

• Expressed numerically

$$Y_{ijkl} = \mu_{ijk} + \varepsilon_{ijkl}$$

where  $\mu_{ijk}$  is the theoretical mean or expected value of all observations in cell (i, j, k)

- The  $\varepsilon_{ijkl}$  are iid  $N(0, \sigma^2)$  which implies the  $Y_{ijkl}$  are independent  $N(\mu_{ijk}, \sigma^2)$
- Parameters

$$- \ \big\{ \mu_{ijk} \big\}, \ i = 1, 2, ..., a, \ j = 1, 2, ..., b, \ k = 1, 2, ..., c \\ - \ \sigma^2$$

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## **Estimates**

• Estimate  $\mu_{ijk}$  by the sample mean of the observations in cell (i, j, k)

$$\hat{\mu}_{ijk} = \overline{Y}_{ijk}.$$

• For each cell (i, j, k), also estimate of the variance

$$s_{ijk}^2 = \sum (Y_{ijkl} - \overline{Y}_{ijk.})^2 / (n_{ijk} - 1)$$

• These  $s_{ij}^2$  are pooled to estimate  $\sigma^2$ 

error

• If balanced

### Factor Effects Model

• Statistical model is

$$Y_{ijkl} = \mu + \alpha_i + \beta_j + \gamma_k + (\alpha\beta)_{ij} + (\alpha\gamma)_{ik} + (\beta\gamma)_{jk} + (\alpha\beta\gamma)_{ijk} + \varepsilon_{ijkl}$$

 $\mu$  - grand mean

 $\alpha_i, \beta_i, \gamma_k$  - main effects of A, B, and C

 $(\alpha\beta)_{ij}, (\alpha\gamma)_{ik}, (\beta\gamma)_{jk}$  are the two-factor (first-order) interactions

 $(\alpha\beta\gamma)_{ijk}$  is the three-factor (second-order) interaction

- Over-parameterized model.
- Extension of usual model constraints.

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### SAS Commands

```
proc glm data=a1;
   class gender fat smoke;
   model extol=gender fat smoke
     gender*fat gender*smoke fat*smoke
     gender*fat*smoke;
   means gender*fat*smoke;
run;
```

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## Output

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ANOVA Table

• Sources of variation are three main effects, three first-

order interactions, one second-order interaction, and

Sum of

Source Squares Mean Square F Value Pr > F Model 588.5829167 84.0832738 9.01 0.0002

Error 149.3666667 9.3354167 16

- SS add up to model SS

• Each effect tested over MSE

- Type I and Type III the same

Corrected Total 23 737.9495833

R-Square Coeff Var Root MSE extol Mean 0.797592 18.77833 3.055391 16.27083

Type I SS Mean Square F Value Source Pr > F176.5837500 gender 1 176.5837500 18.92 0.0005 242.5704167 242.5704167 25.98 0.0001 fat 70.3837500 7.54 0.0144 smoke 70.3837500 13.6504167 13.6504167 1.46 0.2441 gender\*fat 11.0704167 11.0704167 0.2923 gender\*smoke fat\*smoke 72.4537500 72.4537500 7.76 0.0132 1.8704167 0.20 0.6604 gender\*fat\*smoke 1.8704167

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data a1: set a1:

proc glm data=a1;
 class gender fs;

model extol=gender fs;

means gender fs/tukey;

run;

run:

## **Comments**

- First examine interactions
- Some options when one or more interactions significant
  - Interpret the plot of means
  - Run analyses for each level of one factor (slice)
  - Run as one-way with abc levels
  - Run as two-way with a and bc levels
  - Use contrasts
- If no interactions
  - Use contrasts
  - Multiple comparison procedure

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## Output

```
Sum of
                        Squares Mean Square F Value Pr > F
Source
Model
                 4 561.9916667 140.4979167
                                               15.17 <.0001
Error
                19 175.9579167
                                  9.2609430
Corrected Total 23 737.9495833
            Coeff Var
                           Root MSE
R-Square
                                      extol Mean
0.761558
             18.70328
                           3.043180
                                        16.27083
                      Type I SS Mean Square F Value Pr > F
Source
                 1 176.5837500 176.5837500
gender
                                               19.07 0.0003
fs
                 3 385.4079167 128.4693056
                                               13.87 <.0001
Source
                DF Type III SS Mean Square F Value Pr > F
gender
                 1 176.5837500 176.5837500
                                               19.07 0.0003
fs
                 3 385.4079167 128.4693056
                                               13.87 <.0001
```

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# Output

SAS Commands

if (fat eq 1)\*(smoke eq 1) then fs='1\_fs';

if (fat eq 1)\*(smoke eq 2) then fs='2\_fS';
if (fat eq 2)\*(smoke eq 1) then fs='3\_Fs';

if (fat eq 2)\*(smoke eq 2) then fs='4\_FS';

Tukey's Studentized Range (HSD) Test for extol Alpha 0.05 Error Degrees of Freedom 19 Error Mean Square 9.260943 Critical Value of Studentized Range 3.97655 Minimum Significant Difference 4.9404 Mean fs 22,900  $1_fs$ 16.000 2\_fS В В В 13.117 4\_FS 13.067 3\_Fs

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# Unequal Sample Size

- Similar approach as two-way ANOVA
- $\bullet\,$  Type I and Type III SS different
- $\bullet\,$  Type III more commonly used
- $\bullet\,$  ls means used for comparsions

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# **Background Reading**

- KNNL Chapter 24
- knnl1005.sas
- KNNL Section 25.1