

BAYES-STEIN ESTIMATION FOR PORTFOLIO ANALYSIS

PHILIPPE JORION, JFQA, SEP 1986

Krishnan Raman

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- ▶ Individual averages that exceed grand mean shrink, those that are smaller inflate.

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► **Data, Distribution, Parameter**

$$X_i \sim N(\theta_i, 1), i = 1..p, p > 2, \theta = (\theta_1, \theta_2, \dots, \theta_p)'$$

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$$\text{MLE \& UMVUE: } \hat{\theta}^0(X) = (X_1, X_2, \dots, X_p)'$$

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If $R(\hat{\theta}^1, \theta) \leq R(\hat{\theta}^2, \theta) \forall \theta$, we say $\hat{\theta}^2$ is **inadmissible**

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► **Stein's Paradox**

$$R(\hat{\theta}^{JS}, \theta) \leq R(\hat{\theta}^0, \theta) \forall \theta \Rightarrow \hat{\theta}^0(X) \text{ is inadmissible!}$$

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- ▶ Roy's idea: Maximize $\frac{\mu_p - d}{\sigma_p^2}$, d = disaster return

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- ▶ Output of the Markowitz Problem: The x_j , optimal proportion of weights that sum to 1.

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- ▶ Step 2: Solve *Markowitz problem* using moments computed above.

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 Integrate out the unknown parameter θ
 Integrating out will explicitly takes into account uncertainty about θ
- ▶
$$\begin{aligned} \max E_{\theta}[E_{y|\theta}[U(z)|\theta]] \\ &= \int \int U(z)p(z|\theta)dz p(\theta|y, l_0)d\theta \\ &= \int U(z)[\int p(z|\theta) p(\theta|y, l_0)d\theta]dz \end{aligned}$$
- ▶ Predictive Density Function of $z = q'r = p(z|\theta) p(\theta|y, l_0)d\theta$
- ▶ Posterior Density of $\theta = p(\theta|y, l_0)$

ESTIMATION RISK

JORION: SECTION II

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- ▶ A reasonable minimum requirement for any estimator is admissibility
- ▶ *Central thesis of this paper is that the usual sample mean is not admissible for portfolio estimation*

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- ▶ In real-life (finance), standard deviations of stocks tend to not change as much (Merton)

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- ▶ What about Normality assumption ?
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STEIN ESTIMATION

JORION: SECTION III

- ▶ Let $\hat{\mu}$ = vector of N normal means, with covariance matrix assumed known.
- ▶ $y_t \sim N(\hat{\mu}, \Sigma)$, $t = 1..T$ are the stock returns.
- ▶ Stein's Paradox: If $N > 2$, MLE = vector of sample means = is inadmissible.
- ▶ Inadmissibility is wrt quadratic loss $L(\mu, \hat{\mu}(y)) = (\mu - \hat{\mu}(y))' \sigma^{-1} (\mu - \hat{\mu}(y))$
- ▶ If $N \leq 2$, MLE is admissible & the best estimator.
- ▶ Shrinkage Estimator: $\hat{\mu}_{JS}(y) = (1 - \hat{w})Y + \hat{w}Y_0$, Y_0 is any point.
- ▶ $\hat{w} = \min\left(1, \frac{(N-2)/T}{(Y - Y_0)' \Sigma^{-1} (Y - Y_0)}\right)$
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EMPIRICAL BAYES

JORION: SECTION IV

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 - Special Case $\lambda = 0$, corresponds to Diffuse Prior
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 - In this special case Bayes rule won't be admissible (Berger: Statistical Decision Theory)

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- ▶ Another interpretation of above prior: Means μ vary around the common grand mean η

EMPIRICAL BAYES

JORION: SECTION IV: DERIVATION

▶ $p(r, \mu, \eta | y, \Sigma, \lambda) \propto \exp((-1/2)G)$

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- ▶ After integration over η, μ , we find the predictive density is Normal with mean vector & covariance matrix
 Mean Vector $E(r) = (1 - w)Y + w1Y_0$
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 Covariance matrix $V = \Sigma(1 + \frac{1}{T + \lambda}) + \frac{\lambda}{T(T + 1 + \lambda)} \frac{11'}{1'\Sigma^{-1}1}$

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 Σ : variation of y_t around the mean μ
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EMPIRICAL BAYES

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KRISHNAN: MARKOWITZ FOR FAANG STOCKS

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- ▶ Unified View of Shrinkage: Stein Estimator if $\dim(\text{response}) > 2$
vs Ridge/Lasso/ElasticNet if $\dim(\text{predictors}) > p$ (model selection)

CONCLUSION

FUTURE DIRECTIONS

- ▶ We looked at 2 different estimators. There are many more...
- ▶ Certainty Equivalence(Markowitz): Y, S - use sample mean, sample cov
- ▶ Bayes Diffuse Prior(Brown): $\lambda = 0, Y, \hat{V} = \hat{\Sigma} \cdot (1 + 1/T)$
- ▶ Minimum Variance(Jobson): $w = 1, 1Y_0, \hat{V} = \Sigma + (1/T)11'(1/1'\Sigma^{-1}1)$
Based entirely on grand mean
Matrix added to Σ reflects uncertainty in this grand mean.
- ▶ Bayes-Stein Estimate(Jorion): compromise between grand mean & sample mean
- ▶ Unified View of Shrinkage: Stein Estimator if $\dim(\text{response}) > 2$
vs Ridge/Lasso/ElasticNet if $\dim(\text{predictors}) > p$ (model selection)
- ▶ Ledger (Peer Reviewed Blockchain Journal) & BTC Assets

