Lecture 5: Determining Sample Size

Montgomery: Section 3.7 and 13.4

Choice of Sample Size: Fixed Effects

- Can determine the sample size for
  - Overall F test
  - Contrasts of interest
- For simplicity, typically assume n_i's constant, i.e.,
  $n_1 = n_2 = \cdots = n_a = n$
- Recall
  - Type I error rate: $\alpha =$ P(Reject $H_0 | H_0$)
  - Type II error rate: $\beta =$ P(Accept $H_0 | H_1$)
  - Power = P(Reject $H_0 | H_1) = 1 - \beta$
- Need to know
  - Test Statistics
  - Distr. of test statistics under $H_0 \Longrightarrow$ Reject Region (for given α)
  - Distr. of test statistics under $H_1 \Longrightarrow$ power = P(Reject Region $\mid H_1$)

Determining Power for F Test

- $\alpha =$ Pr($F_0 > F_{\alpha,a-1,N-a}|H_0$)
- $\beta =$ Pr($F_0 < F_{\alpha,a-1,N-a}|H_1$)
- Need to know distribution of F_0 when H_1 is true
  - Can show $F_0 = \mathrm{MS}_{\mathrm{Trt}}/\mathrm{MS}_{\mathrm{E}} \sim F_{a-1,N-a}(\delta)$
  - $\delta = n\sum\tau_i^2/\sigma^2$ (non-centrality parameter)
- Recall E(MS$_{\mathrm{Trt}}$)=$\sigma^2 + n\sum\tau_i^2/(a-1)$
  - $\delta = \{E(\mathrm{MS}_{\mathrm{Trt}}) - E(\mathrm{MS}_{\mathrm{E}})\} \times df_{\mathrm{Trt}}/E(\mathrm{MS}_{\mathrm{E}})$
- Need to specify $\{\tau_i\}$ (Note the zero-sum constraint: $\sum_{i=1}^{a}\tau_i = 0$)
- Power will vary for different choices

Power Calculation for F Test

- Given α, a, and n, can determine $F_{\alpha,a-1,N-a}$
- Given some value of δ, can use noncentral F to compute power
  - In SAS, use function PROBF
  - Power=1-PROBF($F_{\alpha,a-1,N-a}$,$a-1$,$N-a$,δ)
- Montgomery: OCC given in Chart V
  - Plots β vs Φ
  - $\Phi^2 = \delta/a = n\sum\tau_i^2/(a\sigma^2)$
  - Can use charts to determine power or sample size

Methods to Determine δ or Φ^2

1. Choose treatment means $(\mu + \tau_i)$
   - Solve for $\{\tau_i\}$ and compute Φ^2 or δ
   - Difficult to know what means to select
2. Take a mimimum difference approach
   - Suppose there exists a pair of (i, j) such that $|\tau_i - \tau_j| \geq D$
   - The minimum value: $\Phi^2 = nD^2/(2a\sigma^2)$ (e.g., $\{\tau_i\} = \{-D/2, 0, \ldots, 0, D/2\}$)
   - Power of test is at least $1 - \beta$
3. Specify a standard deviation increase in percentage (P)
   - Under H_1, variance of a randomly chosen y_i is $\sigma_y^2 = \sigma^2 + \sum \tau_i^2/a$
   - Randomly chosen τ_i has mean 0 and variance $\sum \tau_i^2/a$
   - $P = \left(\sqrt{\sigma^2 + \sum \tau_i^2/a}/\sigma - 1\right) \times 100$
   - $\delta = an\{(1 + .01P)^2 - 1\}$
   - $\Phi^2 = n\{(1 + .01P)^2 - 1\}$

Power Calculation for Specific Contrast

- Often with an experiment, a researcher is primarily interested in just a few comparisons or contrasts. In these cases, it can be preferable to determine sample size for these rather than the overall F test.
- This reduces problem back to the t test situation
- Need to determine
  - Difference of importance
  - Standard error of comparison
- May want/need to adjust for multiple comparisons
- Montgomery describes confidence interval approach
  - Consider pairwise difference in treatment means
  - Specify length of $(1-\alpha)\times 100\%$ confidence interval
  - Length/2 = $t_{\alpha/2,N-a}\sqrt{\frac{2\mathrm{MS_E}}{n}}$
  - Based on the choice of $\mathrm{MS_E}$, find n

Example 3.1 – Etch Rate (Page 75)

- Consider new experiment to investigate 5 RF power settings equally spaced between 180 and 200 W
- Wants to determine sample size to detect a mean difference of D=30 (Å/min) with 80% power
- Will use Example 3.1 estimates to determine new sample size

  $\hat{\sigma}^2 = 333.7$, $D = 30$, and $\alpha = .05$
- Using Table V : $\Phi^2 = 900 \times n/(2 \times 5 \times 333.7) \approx .27 \times n$

n	9	10	11
Φ	$\sqrt{2.43} \approx 1.56$	$\sqrt{2.70} \approx 1.64$	$\sqrt{3.0} \approx 1.72$
df$_{\mathrm{E}}$	40	45	50
β	26%	20%	15%
power	74%	80%	85%

Using SAS : $\delta = a\Phi^2$

```
data new; a=5; alpha=.05; d=30; var=333.7;
    do n=5 to 15;
        df = a*(n-1);
        nc = n*d*d/(2*var);
        fcut = finv(1-alpha,a-1,df);
        beta = probf(fcut,a-1,df,nc);
        power = 1-beta;   output;
    end;
proc print;
    var n df nc beta power; run;
```

```
Obs      n     df        nc         beta       power
  1      5     20      6.7426     0.57654     0.42346
  2      6     25      8.0911     0.47884     0.52116
  3      7     30      9.4396     0.39034     0.60966
  4      8     35     10.7881     0.31289     0.68711
  5      9     40     12.1366     0.24703     0.75297
  6     10     45     13.4852     0.19234     0.80766 ***n=10 needed
  7     11     50     14.8337     0.14788     0.85212
  8     12     55     16.1822     0.11239     0.88761
  9     13     60     17.5307     0.08451     0.91549
 10     14     65     18.8792     0.06292     0.93708
 11     15     70     20.2277     0.04641     0.95359
```

- Compare all pairs and detect any difference more than 30 (Å/min) with 80% power
- A pairwise comparison takes $t_0 = (\bar{Y}_{i.} - \bar{Y}_{k.})/\sqrt{2\text{MSE}/n}$
- Consider Tukey's adjustment: reject when $|t_0| > q_{.05}(5, df_{\rm E})/\sqrt{2}$
- $t_0 \overset{H_1}{\sim} t_{df_{\rm E}}(\delta)$ with $\delta = (\mu_i - \mu_k)/\sqrt{2\sigma^2/n}$,
- Use PROBMC in SAS to get the quantile for multiple comparisons

```
data new1; a=5; alpha=.05; var=333.7; d=30;
    do n=8 to 12;
        df = a*(n-1); nc = d/sqrt(var*2/n);
        /* crit = tinv(1-alpha/2,df); /* LSD approach*/
        crit = probmc("range",.,1-alpha,df,a)/sqrt(2); /*Tukey*/
        power=1-probt(crit,df,nc)+probt(-crit,df,nc);  output;
    end;
proc print;  var n df power; run;
```

```
Obs      n     df     power
  1      8     35    0.65814
  2      9     40    0.73085
  3     10     45    0.79139
  4     11     50    0.84057   *** 11 replicates needed here
  5     12     55    0.87971
```

- Interested in comparing each setting to 200 W, and detect a difference more than 30 (Å/min) with 80% power
- A pairwise comparison takes $t_0 = (\bar{Y}_{i.} - \bar{Y}_{c.})/\sqrt{2\text{MSE}/n}$
- Consider using Dunnett's adjustment: reject when $|t_0| > d_{.05}(4, df_{\rm E})$
- $t_0 \overset{H_1}{\sim} t_{df_{\rm E}}(\delta)$ with $\delta = (\mu_i - \mu_k)/\sqrt{2\sigma^2/n}$,

```
data new1; a=5; alpha=.05; var=333.7; d=30;
    do n=7 to 12;
        df = a*(n-1); nc = d/sqrt(var*2/n);
        crit = probmc("dunnett2",.,1-alpha,df,a-1); /*Two-Sided Dunnett*/
        power=1-probt(crit,df,nc)+probt(-crit,df,nc);  output;
    end;
proc print;
    var n df power; run;
```

```
Obs      n     df      power
  1      7     30    0.68794
  2      8     35    0.76201
  3      9     40    0.82136  ***Only 9 replicates needed here
  4     10     45    0.86780
  5     11     50    0.90341
  6     12     55    0.93024
```

Use of PROC POWER in SAS

- PROC POWER and PROC GLMPOWER both calculate $\sum_{i=1}^{a} \tau_i^2$ using pre-specified treatment means μ_i

```
proc power;
  onewayanova test=overall power=.80 npergroup=. stddev=18.27
       groupmeans = -15|0|0|0|15;
run; quit;
```

```
   Overall F Test for One-Way ANOVA

       Fixed Scenario Elements
Method                              Exact
Group Means                  -15 0 0 0 15
Standard Deviation                  18.27
Nominal Power                         0.8
Alpha                                0.05

   Computed N Per Group
Actual Power     N Per Group
       0.808              10
```

- Neither `PROC POWER` nor `PROC GLMPOWER` can easily do multiple comparison adjustment

```
proc power;
  onewayanova test=contrast power=.80 npergroup=. stddev=18.27
       groupmeans = -15|0|0|0|15
       contrast = (1 0 0 0 -1);
run; quit;
```

```
  Single DF Contrast in One-Way ANOVA

       Fixed Scenario Elements
Method                                Exact
Contrast Coefficients             1 0 0 0 -1
Group Means                     -15 0 0 0 15
Standard Deviation                    18.27
Nominal Power                           0.8
Number of Sides                           2
Null Contrast Value                       0
Alpha                                  0.05

Computed N Per Group
Actual Power      N Per Group
       0.844                7
```

Use of PROC GLMPOWER in SAS

```
data trtmeans; input trt resp @@; datalines;
1 -15 2 0 3 0 4 0 5 15
;

proc glmpower data=trtmeans;
  class trt;
  model resp = trt;
  contrast 'trt1 vs. trt5' trt 1 0 0 0 -1;
  power stddev=18.27 alpha=0.05 ntotal=. power=0.8;
run; quit;
```

```
                    Fixed Scenario Elements

              Dependent Variable          resp
              Alpha                       0.05
              Error Standard Deviation    18.27
              Nominal Power               0.8

                      Computed N Total

Index      Type          Source  Test DF  Error DF Actual Power N Total
    1    Effect             trt        4        45        0.808      50
    2  Contrast  trt1 vs. trt5        1        30        0.844      35
```

Choice of Sample Size: Random Effects

- Can use central F distribution
  - $(N-a)\text{MS}_{\text{E}}/\sigma^2 \sim \chi^2_{N-a}$
  - $(a-1)\text{MS}_{\text{Trt}}/(\sigma^2 + n\sigma^2_\tau) \sim \chi^2_{a-1}$
  - Thus $F_0/\lambda^2 \sim F_{a-1,N-a}$, where $\lambda^2 = \frac{E(\text{MS}_{\text{Trt}})}{E(\text{MS}_{\text{E}})} = 1 + n\sigma^2_\tau/\sigma^2$
  - Power: $P(F_0 > F_{\alpha,a-1,N-a}|\sigma^2_\tau > 0) = P(F > F_{\alpha,a-1,N-a}/\lambda^2)$
  - Can specify ratio of σ^2_τ/σ^2
  - Can specify percentage increase $P = (\sqrt{\sigma^2 + \sigma^2_\tau}/\sigma - 1) \times 100$
- OCC given in Chart VI
  - Plots β vs λ
- Use SAS function `PROBF`
  - power = 1-`PROBF`($F_{\alpha,a-1,N-a}/\lambda^2$,$a-1$,$N-a$)

Example: Batch Example on Slide 8 of Lecture 4

- Consider new experiment with 5 batches: a random effects problem
  - The variance estimate is $\sigma^2 = 1.8$
  - Desire to detect situation when $\sigma_\tau^2 \geq 3.6 = 2\sigma^2$
  - Set power at 80% and $\alpha = .05$
- Using Table VI : $\lambda = \sqrt{1+2n}$

n	3	4	5
λ	$\sqrt{7} \approx 2.65$	$\sqrt{9} = 3$	$\sqrt{11} \approx 3.32$
df_{E}	10	15	20
β	28%	18%	15%
power	72%	82%	85%

- Appears $n = 4$ gives appropriate power

Using SAS

```
data new; a = 5; alpha=.05; ratiovar=2.0;
    do n=2 to 10;
        df = a*(n-1);
        lambdasq = 1+ratiovar*n;
        fcut  = finv(1-alpha,a-1,df);
        beta=probf(fcut/lambdasq,a-1,df);
        power = 1-beta;   output;
    end;

proc print;
    var n beta power; run;
```

```
Obs      n       beta        power
 1       2    0.52933      0.47067
 2       3    0.26112      0.73888
 3       4    0.15292      0.84708   ** n=4 gives the power
 4       5    0.10027      0.89973
 5       6    0.07081      0.92919
 6       7    0.05267      0.94733
 7       8    0.04072      0.95928
 8       9    0.03242      0.96758
 9      10    0.02643      0.97357
```