

Financial Anomalies

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Preface

The article is desgied to study financial anomalies

1 Introduction

START

Extreme Gradient Boosting (XGBoost) is a highly scalable, powerful and leading decision tree algorithm proposed by Chen and Guestrin (2016). Designed to handle the large datasets, XGBoost implements distributed gradient tree boosting learning strategy to train the model. That is, it sequentially learns from multiple weak classifiers (base learners or decision trees) and iteratively updates their weights (Friedman et al. (2000), Littlestone and Warmuth (1989)). Each combination of the update of the “weak learners” creates the powerful ensemble of “stronger learners”. The sequential ordering means tree generated in the later steps will consider the prediction results of the preceding steps. In other words, the strategy is to gradually instruct weak learners to correct the fault from the preceding step. It also accounts for the deviations when comparing between the former and later decision trees of each iteration. In summary, XGBoost trains model in stages as follows:

1. Initialize model with a single decision tree.
2. Calculate the residuals of the model for each iteration. Residuals also known as errors are the differences between the actual and predicted values.
3. Construct and fit a new decision tree (weak learners) to predict the residuals.
4. Add the decision tree to the ensemble model in a way that minimizes the loss function.
5. Update the predictions, then added to preceding predictions to effectively updating model’s approximation.

6. Calculate the new residuals based on the updated predictions.
7. Build a new tree and update the residuals repeated for predefined number of times.
8. Combine all predictions to give final predictions of the ensemble

Formally, in a given training dataset $\mathbb{D} = (X_i, y_i)$, where \mathbb{D} is organized in n rows and m columns, X_i denotes vector of covariates and y_i represents labels (if lawful = 1, else 0), and $|D| = n, x_i \in \mathbb{R}^m$ and $y_i \in \mathbb{R}$. The predicted value of \hat{y} for the i^{th} sample can be represented by a model comprising k decision trees, expressed as in Equation 1.

$$\hat{y}_i = \sum_{k=1}^K f_k(x_i), f_k \in \mathcal{F}, \quad (1.1)$$

where f_k represents the k^{th} decision tree, K is the number of trees and \mathcal{F} is the functional space denoting the collection of all possible decision trees.

In order to predict the tree in subsequent step, the algorithm computes the loss function given by 1.

$$L = \sum_{i=1}^n \ell(y_i, \hat{y}_i), \quad (1.2)$$

where L is the loss function, and respectively y_i and \hat{y}_i are the real and predicted value of sample i .

In addition to the loss function, the salient feature of the XGBoost is also to estimate the regularization term. Together loss function and regularization term results into the objective function (see Equation 1). The model estimates and find the best fitting parameters (θ) of the objective function during training.

$$Obj(\theta) = \sum_{i=1}^n \ell(y_i, \hat{y}_i) + \sum_{k=1}^K \Omega(f_k), \quad (1.3)$$

where θ are the parameters to be estimated, ℓ is the loss function, respectively y_i and \hat{y}_i are the real and predicted value of sample i , and Ω is the regularization term.

The loss function ℓ can be different type, for example, mean squared error, log loss, etc. The regularization $\Omega(f_k)$ component represents the penalty term for k^{th} tree estimated according to Equation 1

$$\Omega(f_k) = \gamma T_k + \frac{1}{2} \lambda \sum_{j=1}^{T_k} w_{k,j}^2, \quad (1.4)$$

where $w_{k,j}^2$ is the weight of the j^{th} leaf in the k^{th} tree, T is the number of leaf nodes, γ and λ are respectively minimum loss requires shrinkage (learning) parameters.

END

Fama and MacBeth (1973) show two-parameter regression model estimates average risk-return relationships based on efficient market portfolio (m), that is, the market prices fully reflect the available information. The asset are constructed based on Equation 1 for an asset (i) proposed by @Black (1972).

$$x_{im} \equiv \frac{\text{total market value of all units of assets } i}{\text{total market value of all assets}} \quad (1.5)$$

where asset(i) in the portfolio(m)

Expected return of a security (i) is $E(\tilde{R}_0)$, the expected return on a security that is riskless in the portfolio m , plus a risk premium that is β_i times the difference between expected return of the portfolio ($E(\tilde{R}_m)$) and riskless portfolio ($E(\tilde{R}_0)$). is calculated by Equation 1, β_i is the risk of the asset i of the portfolio m , measured relative to $\sigma^2(\tilde{R}_m)$

$$E(\tilde{R}_i) = [E(\tilde{R}_m) - S_m \sum \tilde{R}_m] + S_m \sigma(\tilde{R}_m) \beta_i,$$

where,

$$\beta_i \equiv \frac{\text{cov}(\tilde{R}_i, \tilde{R}_m)}{\sigma^2(\tilde{R}_m)} = \frac{\sigma_{j=1}^N x_{jm} \sigma_{ij}}{\sigma^2(\tilde{R}_m)} = \frac{\text{cov}(\tilde{R}_i, \tilde{R}_m) / \sigma(\tilde{R}_m)}{\sigma(\tilde{R}_m)}$$

$$S_m = \frac{E(\tilde{R}_m) - E(\tilde{R}_0)}{\sigma(\tilde{R}_m)}$$

hence

$$E(\tilde{R}_i) = E(\tilde{R}_0) + [E(\tilde{R}_m) - E(\tilde{R}_0)] \beta_i \quad (1.6)$$

For each period of t , the cross sectional regression is given by

$$R_{pt} = \tilde{\gamma}_{0t} + \tilde{\gamma}_{1t} \tilde{\beta}_{p,t-1} + \tilde{\gamma}_{2t} \tilde{\beta}_{p,t-1}^2 + \tilde{\gamma}_{3t} \bar{s}_{p,t-1} \tilde{\epsilon}_i + \tilde{\eta}_{pt}, \quad (1.7)$$

$$p = 1, 2, \dots, t$$

Equation 1 the independent variable $\tilde{\beta}_{p,t-1}$ is the average of the $\tilde{\beta}_i$ for securities in portfolio p , $\tilde{\beta}_{p,t-1}^2$ is the average of the squared values of these $\tilde{\beta}_i$, $\bar{s}_{p,t-1} \tilde{\epsilon}_i$ is the average of $s \tilde{\epsilon}_i$ for portfolio p_i

Gupta and Ofer (1975) examines investors growth expectations reflected in the stock prices. A change in the expectation is reflected in the price movement. The study defines the earnings

price ratio is a function of risk characteristics of the security and the expected growth in the earnings in Equation 1. The risk component are measured by: - the beta coefficient - the firm asset size (natural logarithm of total asset) - dividend payout ratio - leverage ratio of liabilities and preferred stocks to the common stock outstanding - earnings variability (standard deviation of earnings to price ratio calculated over period of seven years)

$$EP = f(RS, EG) \quad (1.8)$$

where:

EP = earnings price ratio- ,

RS = risk characteristics of the security

EG = the expected growth rate in the earnings

(1.9)

$$\Delta P_i^t = \frac{P_{it} - P_{it-1}}{P_{it-1}} \times 100 \quad (1.10)$$

where:

ΔP_i^t is the percent change of the security i during the period $t - 1$ to t

The average yearly percentage of the prices change for portfolio j (ΔP_j) is given by

$$\Delta P_j = \frac{\sum_{t=1}^{14} \frac{\sum_{r_t=10j-9}^{10j} \Delta P_{rt}^t}{14}}{14}, r_t = 1, \dots, 190$$

where:

r_t = the relative ranking of a security at time t according to its prediction error at that year. (1.11)

ΔP_t^{rt} = percentage price change during the period $t - 1$ to t for a security that has the rank of r_t at time t (1.12)

Basu (1977): to determine empirically whether the investment performance of the common stocks is related to the P/E ratios. P/E is the ratio of market value of the common stock (market price times the number of shares outstanding) as of December 31 to reporting annual earnings (before extraordinary items) available for common stockholders. According to the paper, due the exaggerated investor expectations, P/E ratio may be the indicator of future investment performance. The paper shows that P/E is not fully reflected in security prices in as rapid a manner as postulated by the semi-strong form of the efficient market hypothesis. Instead, it seems that disequilibrium persisted in the capital market during the period studied. The results suggests that market efficiency does not exist due to lags and frictions in the price adjustment process.

Litzenberger and Ramaswamy (1979):

$$\tilde{R} - r_{ft} = \gamma_0 + \gamma_1\beta_{it} + \gamma_2(d_{it} - r_{ft}) + \tilde{\epsilon}_{it}, i = 1, 2, \dots, N, t = 1, 2, \dots, T \quad (1.13)$$

where :

\tilde{R}_{it} Return of the security i in period t , β_{it} and d_{it} are the systemic risk and the dividend yield of security i in period t , $\tilde{\epsilon}_{it}$, a disturbance term is $\tilde{R} - E(\tilde{R}_{it})$, the deviation of the realized return from the expected value.
coefficients $\gamma_0, \gamma_1, \gamma_2$ (1.14)

Banz (1981): studies “size effect” by using market indices - Market Index, CRSP value weighted, CRSP equal weighted indexes, US Bond Index, Corporate Bond Index to estimate the relationship between the return and market value based on Equation 1. The study suggests the CAPM is misspecified.

$$R_{it} = \gamma_{0t} + \gamma_{1t}\beta_{1t} + \gamma_{2t}[(\phi_{1t} - \phi_{mt})] + \epsilon_{it}, i = 1, \dots, N \quad (1.15)$$

where:

R_i = return on security i

γ_0 = return on a zero-beta portfolio

γ_1 = return on market risk premium

ϕ_i = market value security i

ϕ_m = average market value

γ_2 constant measuring the contribution of ϕ_i to the expected return of a security

(1.16)

Figlewski (1981) estimates “The test methodology we will use is a familiar one. We construct

ten port? folios to which stocks are assigned according to their average short interest during the previous six months, and calculate the portfolio a's or excess returns."

Basu (1983): The empirical findings reported in this paper indicate that, at least during the 1963-80 time period, the returns on the common stock of NYSE firms appear to have been related to earnings' yield and firm size.

De Bondt and Thaler (1985): "This study of market efficiency investigates whether such behavior affects stock prices. The empirical evidence, based on CRSP monthly return data, is consistent with the overreaction hypothesis. Substantial weak form market inefficiencies are discovered. The results also shed new light on the January returns earned by prior "winners" and "losers." Portfolios of losers experience exceptionally large January returns as late as five years after portfolio formation. To repeat, our goal is to test whether the overreaction hypothesis is predictive. Specifically, two hypotheses are suggested: (1) Extreme movements in stock prices will be followed by subsequent price movements in the opposite direction. (2) The more extreme the initial price movement, the greater will be the subsequent adjustment. Both hypotheses imply a violation of weak-form market efficiency."

Bhandari (1988): "The expected returns on common stocks are positively related to the debt/equity ratio (DER), controlling for the beta and firm size, both including and excluding January. This relationship is not sensitive to variations in the market proxy, estimation technique, etc. The evidence suggests that the "premium" associated with the debt/equity ratio is not likely to be just some kind of "risk premium"

$$E(\tilde{r}_i) = E(\tilde{\gamma}_0) + E(\tilde{\gamma}_1)LTEQ_i + E(\tilde{\gamma}_2)BETA_i + E(\tilde{\gamma}_3) + DER_i, i = 1, \dots, N \quad (1.17)$$

where:

Natural Logarithm of Total Common Equity (LETQ) where total common equity

is the number of shares outstanding times (month-end) price per share

$$\text{Debt to Equity Ratio (DER)} = \frac{\text{book value of total assets} - \text{book value of common equity}}{\text{market value of common equity}}$$

BETA, the risk measure

Amihud and Mendelson (1986): “Illiquidity can be measured by the cost of immediate execution. An investor willing to transact faces a tradeoff: He may either wait to transact at a favorable price or insist on immediate execution at the current bid or ask price. The quoted ask (offer) price includes a premium for immediate buying, and the bid price similarly reflects a concession required for immediate sale. Thus, a natural measure of illiquidity is the spread between the bid and ask prices, which is the sum of the buying premium and the selling concession.’ Indeed, the relative spread on stocks has been found to be negatively correlated with liquidity characteristics such as the trading volume, the number of shareholders, the number of market makers trading the stock and the stock price continuity.”

Amihud and Mendelson (1989):

The percentage bid-ask spread (= dollar spread divided by the stock price)

$$R_{pm} = \gamma_0 + \gamma_1\beta_{pn} + \gamma_2\sigma_{pn} + \gamma_3SZ_{pn} + \gamma_4S_{pn} + \sum_{n=1}^{19} d_n DY_n + \epsilon_{pn} \quad \text{where:}$$

p is the portfolio

β_{pn} portfolio beta

σ_{pn} residual standard deviation

S_{pn} the portfolio average spread

SZ_{pn} average market value of the portfolio

(1.18)

Ou and Penman (1989): “Implements multivariate LOGIT analysis in financial statements. Fundamental analysis extracts value measures from financial statements and compares them to prices to identify mispriced stocks. The evidence here suggests that financial statements capture fundamentals that are not reflected in prices.”

Jegadeesh (1990): “This paper documents strong evidence of predictable behavior of security returns. The results here show that the monthly returns on individual stocks exhibit. significantly negative first-order serial correlation and significantly positive higher-order serial correlation. The pattern of serial correlation exhibits seasonality, with the pattern in January significantly different from that in the other months.”

Jegadeesh and Titman (1993):

Loughran and Ritter (1995) : “Investing in firms issuing stock is hazardous to your wealth. Firms issuing stock during 1970 to 1990, whether an IPO or an SEO, have been poor long-run investments for investors. s. The average annual return during the five years after issuing is only 5 percent for firms conducting IPOs, and only 7 percent for firms conducting SEOs. Investing an equal amount at the same time in a nonissuing firm with approximately the same

market capitalization, and holding it for an identical period, would have produced an average compound return of 12 percent per year for IPOs and 15 percent for SEOs. The magnitude of the underperformance is large: it implies that 44 percent more money would need to be invested in the issuers than in the nonissuers to be left with the same wealth five years later.”

Michaely, Thaler, and Womack (1995): “Market reaction to the dividend initiations.”

La Porta (1996): ” Contrarian strategies that use analysts’ expectations to form portfolios yield high returns. Specifically, when stocks are sorted by the expected growth rate in earnings, low $E\{g\}$ stocks beat high $E\{g\}$ stocks by twenty percentage points.. Finally, event study evidence suggests that the market was overly pessimistic about the earnings of the low $E\{g\}$ portfolio and excessively optimistic about the earnings of the high $E\{g\}$ portfolio”

Lev and Sougiannis (1996): “To address these concerns, we estimate the R&D capital of a large sample of public companies and find these estimates to be statistically reliable and economically meaningful. We then adjust the reported earnings and book values of sample firms for the R&D capitalization and find that such adjustments are value-relevant to investors.”

Sloan (1996): “This paper investigates whether stock prices reflect information about future earnings contained in the accrual and cash flow components of current earnings. The persistence of earnings performance is shown to depend on the relative magnitudes of the cash and accrual components of earnings. However, stock prices act as if investors fail to identify correctly the different properties of these two components of earnings”

Womack (1996): ” An analysis of new buy and sell recommendations of stocks by security analysts at major U.S. brokerage firms shows significant, systematic discrepancies between prerecommendation prices and eventual values. The initial return at the time of the recommendations is large, even though few recommendations coincide with new public news or provide previously unavailable facts. However, these initial price reactions are incomplete. For buy recommendations, the mean postevent drift is modest (+2.4%) and short-lived, but for sell recommendations, the drift is larger (-9.1%) and extends for six months. Analysts appear to have market timing and stock picking abilities”

Brennan, Chordia, and Subrahmanyam (1998): “We examine the relation between stock returns, measures of risk, and several non-risk security characteristics, including the book-to-market ratio, firm size, the stock price, the dividend yield, and lagged returns. Our primary objective is to determine whether non-risk characteristics have marginal explanatory power relative to the arbitrage pricing theory benchmark, with factors determined using, in turn, the Connor and Korajczyk (CK; 1988) and the Fama and French (FF; 1993b) approaches. Fama—MacBeth-type regressions using risk adjusted returns provide evidence of return momentum, size, and book-to-market effects, together with a significant and negative relation between returns and trading volume, even after accounting for the CK factors. When the analysis is repeated using the FF factors, we find that the size and book-to-market effects are attenuated, while the momentum and trading volume effects persist. In addition, Nasdaq stocks show significant underperformance after adjusting for risk using either method”

Brennan, Chordia, and Subrahmanyam (1998): Regardless of the method used to risk-adjust returns, we find a strong negative relation between average returns and trading volume, which is consistent with a liquidity premium in asset prices. In addition, the size and book-to-market ratio effects are strong in the CK method of risk-adjustment, while the FF factors attenuate both the magnitude and significance of these effects.

Dichev (1998): ” Several studies suggest that a firm distress risk factor could be behind the size and the book-to-market effects. A natural proxy for firm distress is bankruptcy risk. If bankruptcy risk is systematic, one would expect a positive association between bankruptcy risk and subsequent realized returns. However, results demonstrate that bankruptcy risk is not rewarded by higher returns. Thus, a distress factor is unlikely to account for the size and book-to-market effects. Surprisingly, firms with high bankruptcy risk earn lower than average returns since 1980. A risk-based explanation cannot fully explain the anomalous evidence.”

- MV is log of fiscal-year-end price times number of shares outstanding
- BIM is common equity divided by fiscal-year-end price times number of shares outstanding
- Z risk (bankruptcy risk, Altman (1968)) comprised of

- Working capital/Total assets
- Retained Earnings/Total assets
- Earnings before interest and taxes/Total assets
- Market value equity/Book value of total debt
- Sales/Total assets
- O risk
 - $SIZE = \log(\text{total assets}/\text{GNP price-level index})$.
 - * TLTA = Total liabilities divided by total assets.
 - WCTA Working capital divided by total assets.
 - CLCA Current liabilities divided by current assets.
 - OENEG = One if total liabilities exceeds total assets, zero otherwise.
 - NITA Net income divided by total assets.
 - FUTL = Funds provided by operations divided by total liabilities.
 - INTWO = One if net income was negative for the last two years, zero otherwise.
 - $CHIN = (NI_t - NI_{t-1})/(|NI_t| + |NI_{t-1}|)$ where NI_t is net income for the most recent period.

Datar, Naik, and Radcliffe (1998): “In this paper, we provide an alternative test of A&M’s model using the turnover rate as a proxy for liquidity and found strong support for A&M’s model. In particular, we find that the stock returns are strongly negatively related to their turnover rates confirming the notion that illiquid stocks provide higher average returns.” - Monthly Returns - turnover rate of every stock = monthly trading volume (the average number of shares traded during the previous three months, i.e., during months t- 3, t-2 and t-1) and divide it by the number of shares outstanding of that firm - turnover rate of every stock - book-to-market ratio, - firm size - firm beta

Moskowitz and Grinblatt (1999): ” This paper documents a strong and prevalent momentum effect in ind ponents of stock returns which accounts for much of the individual sto tum

anomaly. Specifically, momentum investment strategies, which buy past winning stocks and sell past losing stocks, are significantly less profitable once for industry momentum. By contrast, industry momentum investment which buy stocks from past winning industries and sell stocks from past losing industries, appear highly profitable, even after controlling for size, book equity, individual stock momentum, the cross-sectional dispersion in returns and potential microstructure

Lee and Swaminathan (2000): “This study shows that past trading volume provides an important link between “momentum” and “value” strategies. Specifically, we find that firms with high (low) past turnover ratios exhibit many glamour (value) characteristics, earn lower (higher) future returns, and have consistently more negative (positive) earnings surprises over the next eight quarters. Past trading volume also predicts both the magnitude and persistence of price momentum. Specifically, price momentum effects reverse over the next five years, and high (low) volume winners (losers) experience faster reversals. Collectively, our findings show that past volume helps to reconcile intermediate-horizon “underreaction” and long-horizon “overreaction” effects”

Asness, Porter, and Stevens (2000): “Better proxies for the information about future returns contained in firm characteristics such as size, book-to-market equity, cash flow-to-price, percent change in employees, and various past return measures are obtained by breaking these explanatory variables into two industry-related components.”

Piotroski (2000):

Chordia, Subrahmanyam, and Anshuman (2001): “A body of literature starting with Amihud and Mendelson (1986) has found that investors demand a premium for less liquid stocks, so that expected returns should be negatively related to the level of liquidity. In this paper, we document negative and significant cross-sectional relationship between average stock returns and the level as well as the second moment of measures of trading activity such as dollar volume and share turnover. Given the evidence that the level of liquidity affects asset returns, a reasonable hypothesis is that the second moment of liquidity should be positively related to asset returns, provided agents care about the risk associated with fluctuations in liquidity.

Motivated by this observation, we analyze the relation between expected equity returns and the level as well as the volatility of trading activity, a proxy for liquidity. We document a result contrary to our initial hypothesis, namely, a negative and surprisingly strong cross-sectional relationship between stock returns and the variability of dollar trading volume and share turnover, after controlling for size, book-to-market ratio, momentum, and the level of dollar volume or share turnover. This effect survives a number of robustness checks, and is statistically and economically significant. Our analysis demonstrates the importance of trading activity-related variables in the crosssection of expected stock returns.”

Lamont, Polk, and Saaá-Requejo (2001): “We test whether the impact of financial constraints on firm value is observable in stock returns. We form portfolios of firms based on observable characteristics related to financial constraints and test for common variation in stock returns. Financially constrained firms’ stock returns move together over time, suggesting that constrained firms are subject to common shocks. Constrained firms have low average stock returns in our 1968–1997 sample of growing manufacturing firms. We find no evidence that the relative performance of constrained firms reflects monetary policy, credit conditions, or business cycles. We construct various zero-cost portfolios that are long financially constrained firms and short less constrained firms and find three results. First, these portfolios capture common variation in stock returns not captured by other sources of return comovements. Thus we conclude that there is a financial constraints factor, an identifiable independent common source of economic shocks to firm value. The evidence suggests that financial constraints do affect firm value and that the severity of constraints varies over time. Second, our investigation of the role of financial constraints in asset pricing reveals the surprising result that constrained firms earn lower returns than unconstrained firms, a result not explainable using existing asset-pricing models. Third, financially constrained firms do not have returns that are significantly more cyclical than average. Thus, the source of the common economic shocks to financially constrained firms remains an open question. The proxies are constructed based on Kaplan and Zingales (1995)”

Elgers, Lo, and Pfeiffer Jr (2001): “The paper documents that the weighting of analysts

annual earnings forecasts implicit in security prices is lower than the historical relation between the financial analysts forecasts and realized earnings. Our evidence that analysts the beginning of the year annual earnings forecasts are associated with abnormal security returns subsequently accumulated over the earnings year is consistent with the delayed price reaction to the value-relevant information in the positions in the securities in the bottom (top) deciles of the cross-sectional distribution of the analysts earnings forecasts early in the earnings year, generates statistically significant trading profits in the year after portfolio formation for firms with relatively low analysts coverage.”

P. A. Gompers and Metrick (2001): “We analyze the investors preferences for the stock and the implications that these preferences have for stock-market prices and returns. We find that large institutional investors- a category including all managers with greater than \$100 million under discretionary control –have nearly doubled their share of the common stock from 1980 to 1996, with most of this increase driven by the growth in the holdings of the largest one hundred institutions.”

Griffin and Lemmon (2002): “This paper examines the relationship between book-to-market equity, distress risk, and stock returns. Among firms with the highest distress risk as proxied by Ohlson (1980) 0-score, the difference in returns between high and low book-to-market securities is more than twice as large as that in other firms. This large return differential cannot be explained by the three-factor model or by differences in economic fundamentals. Consistent with mispricing arguments, firms with high distress risk exhibit the largest return reversals around earnings announcements, and the book-to-market effect is largest in small firms with low analyst coverage.”

Diether, Malloy, and Scherbina (2002): “We provide evidence that stocks with higher dispersion in analysts’ earnings forecasts earn lower future returns than otherwise similar stocks. This effect is most pronounced in small stocks and stocks that have performed poorly over the past year. Interpreting dispersion in analysts’ forecasts as a proxy for differences in opinion about a stock, we show that this evidence is consistent with the hypothesis that prices will reflect the optimistic view whenever investors with the lowest valuations do not trade. By

contrast, our evidence is inconsistent with a view that dispersion in analysts' forecasts proxies for risk."

J. Chen, Hong, and Stein (2002): " In this paper, we bring new evidence to bear on an asset-pricing hypothesis which has been around for a long while, but which has thus far not recieved much empirical support. The idea, which dates back to Miller, has to do with the combined effects of short-sales constraints and differences of opinion on stock prices. We develop a model of stock prices in which there are both differences of opinion among investors as well as short-sales constraints. The key insights that emerge in that breadth of ownerwhip is a valuation indicator. When the breadth is low- when investors have long positions in the stock- this signals that hte short-sales constrarint in binding tightly, implying that prices are high relative to fundamentsls and that expected reutrns are therefore low."

P. Gompers, Ishii, and Metrick (2003): "Shareholder rights vary across firms. Using the incidence of 24 governance rules, we construct a "Governance Index" to proxy for the level of shareholder rights at about 1500 large firms during the 1990s. An investment strategy that bought firms in the lowest decile of the index (strongest rights) and sold firms in the highest decile of the index (weakest rights) would have earned abnormal returns of 8.5 percent per year during the sample period. We find that firms with stronger shareholder rights had higher firm value, higher profits, higher sales growth, lower capital expenditures, and made fewer corporate acquisitions."

Doyle, Lundholm, and Soliman (2003): "We investigate the informational properties of pro forma earnings. This increasingly popular measure of earnings excludes certain expenses that the company deems non-recurring, non-cash, or otherwise unimportant for understanding the future value of the firm. We find, however, that these expenses are far from unimportant. Higher levels of exclusions lead to predictably lower future cash flows. We also find that investors do not fully appreciate the lower cash flow implications at the time of the earnings announcement. A trading strategy based on the excluded expenses yields a large positive abnormal return in the years following the announcement, and persists after controlling for various risk factors and other anomalies."

Doyle, Lundholm, and Soliman (2003):

Watkins (2003): “I analyze the degree to which return consistency in the past predicts future returns. It is discovered here that consistency is a strong predictive measure for future stock returns. In a portfolio context, positively consistent stocks exhibit positive future risk-adjusted returns, and negatively consistent stocks exhibit negative future risk-adjusted returns. The results are economically and statistically significant over multiple sub-periods. Also, odd return behavior persists for nearly two years after portfolio formation. Stocks that have been consistently positive (negative) for longer time horizons have higher(lower) risk-adjusted returns during the following month than those which have been consistent for shorter time periods. Finally, it is determined that high consistency enhances momentum when the two factors are allowed to interact. Thus, there appears to be strong path dependence in the momentum effect, and consistency in stock returns appears to be an important component of return predictability.”

Eberhart, Maxwell, and Siddique (2004): “We examine a sample of 8,313 cases, between 1951 and 2001, where firms unexpectedly increase their research and development (R&D) expenditures by a significant amount. We find consistent evidence of a misreaction, as manifested in the significantly positive abnormal stock returns that our sample firms’ shareholders experience following these increases. We also find consistent evidence that our sample firm experience significantly positive long-term abnormal operating performance follow their R&D increases. Our findings suggest that R&D increases are beneficial in themselves, and that the market is slow to recognize the extent of this benefit (consistent with investor underreaction).” George and Hwang (2004): “When coupled with a stock’s current price, a readily available piece of information 52-week high price-explains a large portion of the profits from momentum investment. Nearness to the 52-week high dominates and improves upon the forecasting power past returns (both individual and industry returns) for future returns. Future return forecast using the 52-week high do not reverse in the long run. These results indicate that short-term momentum and long-term reversals are largely separate phenomena which presents a challenge to current theory that models these aspects of security returns as integrated components of the market’s response”

Jegadeesh et al. (2004)

Titman, Wei, and Xie (2004)

K. M. Cremers and Nair (2005)

Acharya and Pedersen (2005):

Measure of Illiquidity (based on Amihud, 2002):

$$ILLIQ_t^i = \frac{1}{Days_i} \sum_{d=1}^{Days_t^i} \frac{|R_{td}^i|}{V_{td}^i} \quad (1.19)$$

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2 Summary

In summary, this book has no content whatsoever.

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