STA302A2

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SETUP

```
setwd("~/Google Drive/Uni/Winter 2018/STA302/A2")
dat <- read.table("Census.txt", sep = "" ,header=T)</pre>
LIFE <- dat$LIFE
MALE <- dat$MALE
BIRTH <- dat$BIRTH
DIVO <- dat*DIVO
BEDS <- dat$BEDS
EDUC <- dat$EDUC
INCO <- dat$INCO</pre>
y <- LIFE
x = as.matrix(dat)
x \leftarrow x[1:51, 2:7]
x \leftarrow cbind(1, x)
mode(x) = 'numeric'
n <- 51
p <-6
X<-x
Xt \leftarrow t(X)
XtX <- Xt %*% X
XtXinv <-solve(XtX)</pre>
Xty <- Xt %*% y
bhat <- XtXinv %*% Xty
1a)
y_hat <- X %*% bhat</pre>
y_hat
##
              [,1]
## [1,] 71.32712
## [2,] 69.44152
## [3,] 70.61014
## [4,] 70.29776
## [5,] 71.65898
## [6,] 71.97295
## [7,] 72.28677
## [8,] 66.87019
## [9,] 70.41924
## [10,] 71.38047
```

```
## [11,] 69.42376
## [12,] 71.53120
## [13,] 71.42614
## [14,] 71.04053
## [15,] 70.16589
## [16,] 70.29139
## [17,] 71.64995
## [18,] 70.78642
## [19,] 70.31884
## [20,] 70.92242
## [21,] 72.03916
## [22,] 70.25335
## [23,] 70.24288
## [24,] 71.16816
## [25,] 70.77073
## [26,] 68.72947
## [27,] 72.04422
## [28,] 70.81770
## [29,] 71.67050
## [30,] 71.22929
## [31,] 71.04496
## [32,] 72.13043
## [33,] 70.80619
## [34,] 68.76111
## [35,] 70.46957
## [36,] 70.31486
## [37,] 71.13894
## [38,] 72.51628
## [39,] 70.91515
## [40,] 71.43818
## [41,] 70.51633
## [42,] 71.11649
## [43,] 70.03452
## [44,] 70.18014
## [45,] 69.57848
## [46,] 71.56217
## [47,] 70.51127
## [48,] 72.50224
## [49,] 71.48537
## [50,] 69.92352
## [51,] 70.45665
e_hat = y - y_hat
e_hat
##
                [,1]
   [1,] -2.01712273
##
## [2,] -0.39152270
## [3,] 0.04985782
## [4,] 0.25224335
```

```
[5,] 0.05102320
##
   [6,] 0.08704553
##
   [7,] 0.19323411
##
   [8,] -1.16019333
## [9,] -0.35923989
## [10,] -0.72046876
## [11,] -0.88376476
## [12,] 2.06880340
## [13,] 1.13386172
## [14,] 0.82946681
## [15,] -0.02589499
## [16,] 0.58861254
## [17,] 0.93004669
## [18,] -0.68642174
## [19,] -1.55884433
## [20,] 0.90758262
## [21,] -1.81915643
## [22,] 0.67665009
## [23,] 0.38711962
## [24,] 1.79184021
## [25,] -0.08073436
## [26,] -0.63947073
## [27,] -1.48421996
## [28,] -1.60770155
## [29,] 1.11949999
## [30,] 1.37071164
## [31,] 0.18503566
## [32,] -1.20043418
## [33,] -0.48618775
## [34,] 0.26888640
## [35,] 0.08043332
## [36,] 0.50514224
## [37,] 0.28105954
## [38,] -0.38627710
## [39,] -0.48514830
## [40,] 0.46181996
## [41,] -2.55632758
## [42,] 0.96350641
## [43,] 0.07548104
## [44,] 0.71986346
## [45,] 3.32151698
## [46,] -1.48217033
## [47,] 1.12872551
## [48,] -0.78223737
## [49,] 0.99463448
## [50,] -0.44351769
## [51,] -0.16664779
```

```
## [,1]
## 70.5577812705
## MALE 0.1261018758
## BIRTH -0.5160557876
## DIVO -0.1965375074
## BEDS -0.0033392036
## EDUC 0.2368222541
## INCO -0.0003612011
```

Equation: LIFE = 70.5577812705 + 0.1261018758 (MALE) - 0.5160557876 (BIRTH) - 0.1965375074 (DIVO) - 0.0033392036 (BEDS) + 0.2368222541 (EDUC) - 0.0003612011 (INCO)

1b) - MALE

```
bhat[2] #b_1
## [1] 0.1261019
```

b_1 (MALE) corresponds to the Expected change (+ 0.1261019) in the Averafe Lifespan with 1 unit increase in the proportion of Males to Female.

BIRTH

```
bhat[3] #b_2
## [1] -0.5160558
```

b_2 (BIRTH) corresponds to the Expected change (-0.5160558) in the Averafe Lifespan with a unit increase in the birth rate per 1,000 people.

1c)

```
RSS <- (t(e_hat) %*% e_hat)
sigma_sq_hat = RSS / (n)
sigma_sq_hat
##     [,1]
## [1,] 1.192215

s_sq = RSS / (n-p-1)
s_sq
##     [,1]
## [1,] 1.381885</pre>
```

```
1d)
```

```
s < -c(s sq^{(1/2)})
se_bs = diag(s * (XtXinv)^(1/2))
se bs
##
                         MALE
                                      BIRTH
                                                     DIVO
                                                                   BEDS
## 4.2897471299 0.0472317551 0.1172774621 0.0739532971 0.0009795303
##
           EDUC
                         INCO
## 0.1110224835 0.0004597943
#Corresponding to B_0, B_1 ... B_6
1e)
y_bar = c(sum(y_hat)/n)
SST <- sum((LIFE - y_bar)^(2))
R_sq \leftarrow 1 - (RSS/SST)
R_sq
##
              [,1]
## [1,] 0.4684927
```

Explains the proportion of variation in the Average lifespan explained by the regression, which is 46.849%.

```
2) 2a)
```

```
MLR <- lm(formula = LIFE~ MALE + BIRTH + DIVO + BEDS + EDUC + INCO, data=dat)
summary(MLR)
##
## Call:
## lm(formula = LIFE ~ MALE + BIRTH + DIVO + BEDS + EDUC + INCO,
##
     data = dat)
##
## Residuals:
     Min
             10 Median
##
                          3Q
                               Max
## -2.5563 -0.6629 0.0755 0.6983 3.3215
##
## Coefficients:
##
             Estimate Std. Error t value Pr(>|t|)
## (Intercept) 70.5577813 4.2897471 16.448 < 2e-16 ***
## MALE
            0.1261019 0.0472318 2.670 0.01059 *
## BIRTH
            ## DIVO
## BEDS
            ## EDUC
            0.2368223 0.1110225
                               2.133 0.03853 *
## INCO
            -0.0003612 0.0004598 -0.786 0.43633
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

```
## Residual standard error: 1.176 on 44 degrees of freedom ## Multiple R-squared: 0.4685, Adjusted R-squared: 0.396 ## F-statistic: 6.464 on 6 and 44 DF, p-value: 6.112e-05  

2b) H_0: B_0 = B_1 = B_2 = B_3 = B_4 = B_5 = B_6 = 0 H_A: B_i \neq 0 [any i in range 0 to 6]  

SSReg = SST-RSS  
F_val = (SSReg/(p)/(RSS/(n-p-1)))  
F_val  

## [1,] 6.463905  

F_crt = qf(.95, df1=p, df2=n-p-1)  
F_crt  
## [1] 2.313264
```

 $F_{val} > F_{crit}$, hence we fail to accept H_0 . Then atleast one of the B_i 's [i = 1, ... 6] is not Zero. Our Model is significant.

2c)

MALE

```
H_0: B_1 = 0 \ H_A: B_1 \neq 0

t_{\text{value1}} = 2.670

t_{\text{crit}} = qt(c(.005, .995), df=n-p-1)

\#t_{\text{crit}} = [-2.692278, 2.692278]
```

Since $|t_crit| > |t_value1|$, we accept H_0 . Then $B_1 \neq 0$ and we can remove the predictor (MALE) from the model.

BIRTH

```
H_0: B_2 = 0 \ H_A: B_2! = 0

t_{\text{value2}} = 4.400

t_{\text{crit}} = qt(c(.005, .995), df=n-p-1)

\#t_{\text{crit}} = [-2.692278, 2.692278]
```

Since $|t_value 2| > |t_crit|$, we fail to accept H_0 . Then $B_2 \neq 0$ then we can't remove the predictor (BIRTH) variable from the model.

DIVO

```
H_0: B_3 = 0 \ H_A: B_3 \neq 0

t_value3 = 2.658

t_crit = qt(c(.005, .995), df=n-p-1)

\#t_crit = [-2.692278, 2.692278]
```

Since $|t_crit| > |t_value3|$, we accept H_0 . Then $B_3 \neq 0$ and we can remove the predictor (DIVO) from the model.

BEDS

```
H_0: B4 = 0 \ H_A: B4 \neq 0

t_value4 = 3.409

t_crit = qt(c(.005, .995), df=n-p-1)

\#t_crit = [-2.692278, 2.692278]
```

Since $|t_value 4| > |t_crit|$, we fail to accept H_0 . Then $B_4 \neq 0$ then we can't remove the predictor (BEDS) variable from the model.

EDUC

```
H_0: B5 = 0 \ H_A: B5 \neq 0

t_{\text{value5}} = 2.133

t_{\text{crit}} = qt(c(.005, .995), df=n-p-1)

\#t_{\text{crit}} = [-2.692278, 2.692278]
```

Since $|t_crit| > |t_value5|$, we accept H_0 . Then $B_3 \neq 0$ and we can remove the predictor (EDUC) from the model.

INCO

```
H_0: B6 = 0 \ H_A: B6 \neq 0

t_value6 = 0.786

t_crit = qt(c(.005, .995), df=n-p-1)

\#t_crit = [-2.692278, 2.692278]
```

Since $|t_{crit}| > |t_{value6}|$, we accept H_0 . Then $B_6 \neq 0$ and we can remove the predictor (INCO) from the model.

Yes, the results indicate that variables MALE (x1), DIVO (x3), EDUC (x5) and INCO x(6) should be removed.

2d)

```
MLR reduced <- lm(formula = LIFE~ BIRTH + BEDS, data=dat)
summary(MLR_reduced)
##
## Call:
## lm(formula = LIFE ~ BIRTH + BEDS, data = dat)
## Residuals:
##
      Min
               1Q Median
                               3Q
                                     Max
## -2.5627 -0.8180 -0.0819 0.9261 3.6202
## Coefficients:
                Estimate Std. Error t value Pr(>|t|)
##
## (Intercept) 79.1473186 2.2717401 34.840 < 2e-16 ***
            -0.3281679 0.1026214 -3.198 0.00245 **
## BIRTH
## BEDS
              ## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 1.352 on 48 degrees of freedom
## Multiple R-squared: 0.2329, Adjusted R-squared: 0.2009
## F-statistic: 7.286 on 2 and 48 DF, p-value: 0.001725
LIFE = 79.1473186 - 0.3281679 (BIRTH) - 0.0027415 (BEDS)
2e) H_0: B_1 = B_6 = 0 H_A: B_1, B_6 both not zero.
MLR_red_MALE_INCO <- lm(formula = LIFE~ BIRTH + BEDS + DIVO + EDUC, data=dat)
anova(MLR, MLR_red_MALE_INCO)
## Analysis of Variance Table
## Model 1: LIFE ~ MALE + BIRTH + DIVO + BEDS + EDUC + INCO
## Model 2: LIFE ~ BIRTH + BEDS + DIVO + EDUC
    Res.Df RSS Df Sum of Sq
                                   F Pr(>F)
## 1
        44 60.803
## 2
        46 70.654 -2 -9.8507 3.5642 0.03676 *
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
F val = 3.5642
F Crit = qf(.99, df1=2, df2=n-p-1)
\#F\_Crit = 5.122628
```

 $|F_val| > |F_crit|$ we fail to reject H_0 , hence the predictors MALE and INCO can be removed from the model.

```
MLR_MALE <- lm(LIFE ~ MALE, data=dat)</pre>
anova(MLR, MLR_MALE)
## Analysis of Variance Table
## Model 1: LIFE ~ MALE + BIRTH + DIVO + BEDS + EDUC + INCO
## Model 2: LIFE ~ MALE
    Res.Df
               RSS Df Sum of Sq
                                      F
                                          Pr(>F)
        44 60.803
## 1
        49 109.834 -5 -49.031 7.0963 6.099e-05 ***
## 2
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
F_{val} = 7.0963
```

We fail to accept H_0 , hence we infer that predictors BIRTH (x2), DIVO (x3), BEDS (x4), EDUC (x5) and INCO (x6) are significant after taking the predictor MALE (x1) into consideration.

```
2g) H_0: B_i = 0 [i = 1, 2] H_A: B_1, B_2, both not zero.
```

```
MLR B0 = lm(LIFE \sim 1, data=dat)
MLR_MALE_BIRTH <- lm(LIFE ~ MALE + BIRTH, data=dat)</pre>
anova(MLR_B0, MLR_MALE_BIRTH)
## Analysis of Variance Table
## Model 1: LIFE ~ 1
## Model 2: LIFE ~ MALE + BIRTH
     Res.Df
                RSS Df Sum of Sq
                                          Pr(>F)
## 1
         50 114.397
         48 85.424 2 28.973 8.14 0.0009036 ***
## 2
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
F \text{ val} = 8.14
```

We fail to accept the H_0 , we infer that (MALE) x1 and(BIRTH) x2 are significant/useful in predicting the response.

```
2h)
```

```
MLR B0 = lm(LIFE \sim 1, data=dat)
MLR B0B3 = lm(LIFE \sim DIVO)
MLR B0B2B3 = 1m(LIFE \sim BIRTH + DIVO)
MLR MALE BIRTH DIVO <- lm(LIFE~ MALE + BIRTH + DIVO, data=dat)
MLR_BIRTH_DIVO <- lm(LIFE ~ BIRTH + DIVO, data=dat)</pre>
SSRegB3B0 <- anova(MLR_B0)["Residuals", "Sum Sq"] - anova(MLR_B0B3)["Residual
s", "Sum Sq"]
SSRegB3B0
## [1] 3.307311
SSRegB2_B0B3 <- anova(MLR_B0B3)["Residuals", "Sum Sq"] - anova(MLR_B0B2B3)["R
esiduals", "Sum Sq"]
SSRegB2 B0B3
## [1] 8.914548
RSSB0B1B2B3 <- anova(MLR MALE BIRTH DIVO)["Residuals", "Sum Sq"]
RSS_B0B2B3 <- anova(MLR_BIRTH_DIVO)["Residuals", "Sum Sq"]</pre>
SSRegB1_B0B2B3 <- RSS_B0B2B3 - RSSB0B1B2B3
SSRegB1 B0B2B3
## [1] 21.42422
SSRegB1B2B3 B0 = SSRegB3B0 + SSRegB2 B0B3 + SSRegB1 B0B2B3
SSRegB1B2B3_B0
## [1] 33.64607
2i)
- 1
#MLR B0
#MLR MALE
MLR BIRTH <- 1m(LIFE ~ BIRTH, data=dat)
MLR_DIVO <- lm(LIFE ~ DIVO, data=dat)</pre>
MLR_INCO <- lm(LIFE ~ INCO, data=dat)</pre>
#MLR MALE BIRTH
MLR MALE DIVO <- 1m(LIFE~ MALE + DIVO, data=dat)
MLR_MALE_INCO <- lm(LIFE~ MALE + INCO, data=dat)</pre>
#MLR BIRTH DIVO
MLR_BIRTH_INCO <- lm(LIFE ~ BIRTH + INCO, data=dat)</pre>
MLR_DIVO_INCO <- lm(LIFE ~ DIVO + INCO, data=dat)</pre>
MLR MALE BIRTH DIVO
```

```
##
## Call:
## lm(formula = LIFE ~ MALE + BIRTH + DIVO, data = dat)
## Coefficients:
                                                      DIVO
## (Intercept)
                         MALE
                                      BIRTH
       62.3656
                       0.1689
                                    -0.3912
                                                   -0.1272
MLR_MALE_BIRTH_INCO <- lm(LIFE~ MALE + BIRTH + INCO, data=dat)</pre>
MLR BIRTH DIVO INCO <- lm(LIFE ~ BIRTH + DIVO + INCO, data=dat)
MLR MALE DIVO INCO <- lm(LIFE ~ MALE + DIVO + INCO, data=dat)
MLR_mbdi <- lm(LIFE ~ MALE + BIRTH + DIVO + INCO, data=dat)</pre>
#AIC
aic_b0 <- AIC(MLR_B0)</pre>
aic_b0
## [1] 189.9321
aic_m <- AIC(MLR_MALE)</pre>
aic_m
## [1] 189.8561
aic_d <- AIC(MLR_DIVO)</pre>
aic_d
## [1] 190.4359
aic_i <- AIC(MLR_INCO)</pre>
aic_i
## [1] 191.2477
aic_b <- AIC(MLR_BIRTH)</pre>
aic_b
## [1] 186.7525
aic_mb <- AIC(MLR_MALE_BIRTH)</pre>
aic_mb
## [1] 179.0377
aic_md <- AIC(MLR_MALE_DIVO)</pre>
aic md
## [1] 188.5535
aic_mi <- AIC(MLR_MALE_INCO)</pre>
aic_mi
## [1] 191.4429
```

```
aic_bd <- AIC(MLR_BIRTH_DIVO)</pre>
aic_bd
## [1] 188.1698
aic_bi <- AIC(MLR_BIRTH_INCO)</pre>
aic_bi
## [1] 188.5226
aic_di <- AIC(MLR_DIVO_INCO)</pre>
aic_di
## [1] 191.4755
aic_mbd <- AIC(MLR_MALE_BIRTH_DIVO)</pre>
aic_mbd
## [1] 178.1686
aic_mbi <- AIC(MLR_MALE_BIRTH_INCO)</pre>
aic_mbi
## [1] 180.932
aic_bdi <- AIC(MLR_BIRTH_DIVO_INCO)</pre>
aic_bdi
## [1] 189.8012
aic_mdi <- AIC(MLR_MALE_DIVO_INCO)</pre>
aic_mdi
## [1] 189.9273
aic_mbdi <- AIC(MLR_mbdi)</pre>
aic_mbdi
## [1] 180.1406
```

Lowest AIC is with model LIFE ~ BIRTH + MALE + DIVO

```
- 2
null<- MLR_B0
full<- MLR_mbdi
forwdAIC=step(null, scope=list(lower=null, upper=full), direction="forward")
## Start: AIC=43.2
## LIFE ~ 1
##
          Df Sum of Sq
##
                         RSS
                                AIC
## + BIRTH 1 11.0479 103.35 40.021
## + MALE 1 4.5632 109.83 43.124
## <none>
                       114.40 43.200
## + DIVO 1 3.3073 111.09 43.704
## + INCO 1 1.5249 112.87 44.516
##
## Step: AIC=40.02
## LIFE ~ BIRTH
##
         Df Sum of Sq
##
                          RSS
                                AIC
              17.9252 85.424 32.306
## + MALE 1
## <none>
                      103.349 40.021
## + DIVO 1
              1.1739 102.175 41.438
## + INCO 1
              0.4647 102.885 41.791
##
## Step: AIC=32.31
## LIFE ~ BIRTH + MALE
##
                         RSS
         Df Sum of Sq
##
                               AIC
## + DIVO 1 4.6730 80.751 31.437
## <none>
                      85.424 32.306
## + INCO 1 0.1768 85.247 34.200
##
## Step: AIC=31.44
## LIFE ~ BIRTH + MALE + DIVO
##
##
         Df Sum of Sq
                         RSS
                               AIC
## <none>
                      80.751 31.437
## + INCO 1 0.044334 80.707 33.409
```

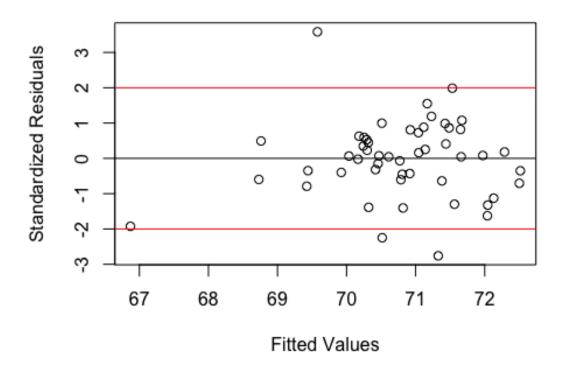
Best Model according to forwardAIC is LIFE ~ BIRTH + MALE + DIVO

```
- 3
backAIC=step(full, direction="backward", data=dat)
## Start: AIC=33.41
## LIFE ~ MALE + BIRTH + DIVO + INCO
##
##
          Df Sum of Sq
                           RSS
                                   AIC
## - INCO
                0.0443 80.751 31.437
                         80.707 33.409
## <none>
## - DIVO
          1
                4.5405 85.247 34.200
               20.7328 101.440 43.070
## - MALE 1
## - BIRTH 1
               20.9838 101.691 43.196
##
## Step: AIC=31.44
## LIFE ~ MALE + BIRTH + DIVO
##
          Df Sum of Sq
##
                           RSS
                                  AIC
## <none>
                         80.751 31.437
## - DIVO
           1
                 4.673 85.424 32.306
## - MALE
                21.424 102.175 41.438
           1
## - BIRTH 1
                22.196 102.947 41.822
```

Best Model according to backwardAIC is LIFE ~ BIRTH + MALE + DIVO

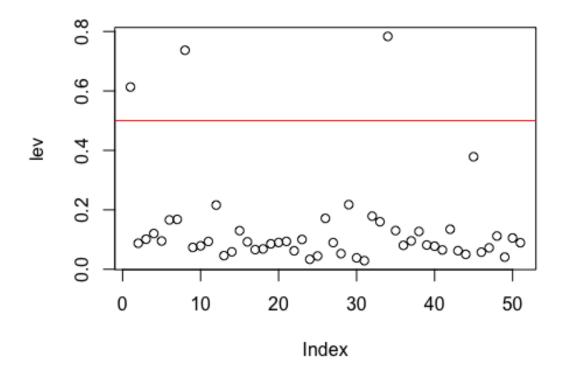
3a)

```
std_res <- rstandard(MLR)
plot(MLR$fitted.values, std_res, ylab="Standardized Residuals", xlab="Fitted
Values", main="Standardized Residuals")
abline(0, 0)
abline(2, 0, col=c("red"))
abline(-2, 0, col=c("red"))</pre>
```



Yes, there are 3 points with residuals greater than |2| and an outlier that may be influential with the lowest fitted value (\sim 67).

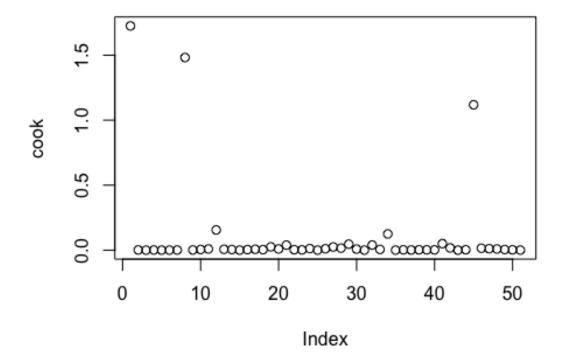
```
lev = hat(model.matrix(MLR))
plot(lev)
abline(0.5, 0, col=c("red"))
```



```
dat[lev >0.5,]
## STATE MALE BIRTH DIVO BEDS EDUC INCO LIFE
## 1    AK 119.1    24.8    5.6    603.3    14.1    4638    69.31
## 8    DC    86.8    20.1    3.0    1859.4    17.8    4644    65.71
## 34    NV 102.8    19.6    18.7    560.7    10.8    4583    69.03
```

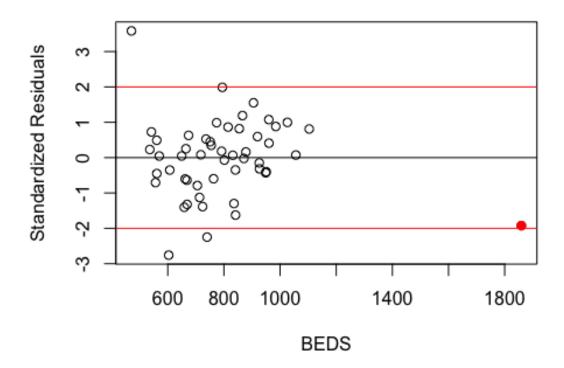
3 points have lev > 0.5. Indices 1, 8 and 34. Corresponding to States: AK, DC and NV.

```
cook = cooks.distance(MLR)
plot(cook)
```



```
dat[cook > 1,]
## STATE MALE BIRTH DIVO BEDS EDUC INCO LIFE
## 1     AK 119.1 24.8 5.6 603.3 14.1 4638 69.31
## 8     DC 86.8 20.1 3.0 1859.4 17.8 4644 65.71
## 45     UT 97.6 25.5 3.7 470.5 14.0 3169 72.90
```

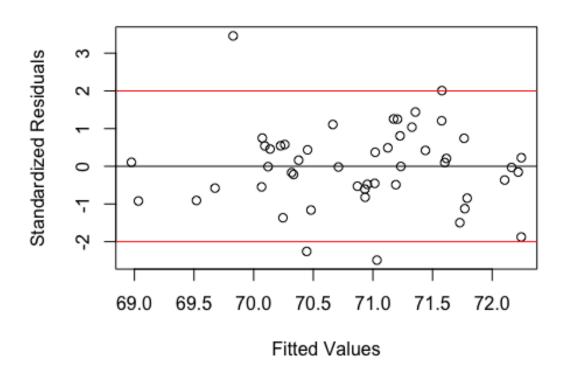
No, not all Observations are the same. We now see data corresponding to UT instead of NV.



There are 3 points that are outliers based on std_res (>2) and the outlier red point corresponding to DC (RED)

```
dat2 <- subset(dat, STATE != 'DC')</pre>
LIFE2 <- dat2$LIFE
MALE2 <- dat2$MALE
BIRTH2 <- dat2$BIRTH
DIVO2 <- dat2$DIVO
BEDS2 <- dat2$BEDS
EDUC2 <- dat2$EDUC
INCO2 <- dat2$INCO
MLR NoDC <- lm(formula = LIFE2 ~ MALE2 + BIRTH2 + DIVO2 + BEDS2 + EDUC2 + INC
02, data=dat2)
summary(MLR_NoDC)
##
## Call:
## lm(formula = LIFE2 ~ MALE2 + BIRTH2 + DIVO2 + BEDS2 + EDUC2 +
##
      INCO2, data = dat2)
##
## Residuals:
                 1Q
##
       Min
                     Median
                                  3Q
                                         Max
## -2.48448 -0.61603 -0.00768 0.58701 3.07199
## Coefficients:
                Estimate Std. Error t value Pr(>|t|)
## (Intercept) 69.8207566 4.1692564 16.747 < 2e-16 ***
## MALE2
               0.0922054 0.0487996 1.889 0.06558 .
## BIRTH2
              ## DIVO2
              -0.1377833 0.0774530 -1.779 0.08232 .
## BEDS2
              -0.0011637 0.0014481 -0.804 0.42604
## EDUC2
               0.3157699 0.1145801 2.756 0.00855 **
## INCO2
              -0.0004698 0.0004485 -1.048 0.30064
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 1.138 on 43 degrees of freedom
## Multiple R-squared: 0.3679, Adjusted R-squared: 0.2797
## F-statistic: 4.171 on 6 and 43 DF, p-value: 0.002171
```

```
std_res_NoDC <- rstandard(MLR_NoDC)
plot(MLR_NoDC$fitted.values, std_res_NoDC, ylab="Standardized Residuals", xla
b="Fitted Values", main="Standardized Residuals")
abline(0, 0)
abline(2, 0, col=c("red"))
abline(-2, 0, col=c("red"))</pre>
```



Yes there is a noticeable change in the Standardised residuals. The fitted line has changed. NEW: -

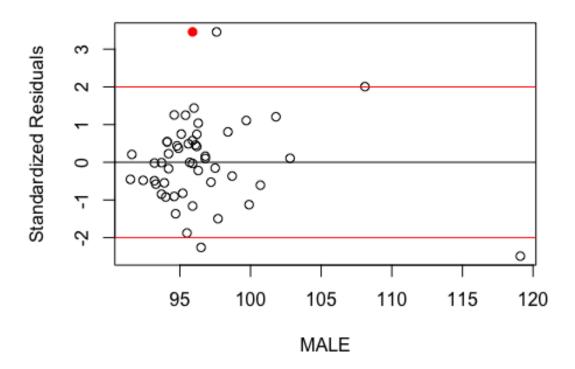
LIFE = 69.8207566 + 0.0922054 (MALE2) - 0.4261398 (BIRTH2) - 0.1377833 (DIVO2) - 0.0011637 (BEDS2) + 0.3157699 (EDUC2) - 0.0004698 (INCO2)

OLD: -

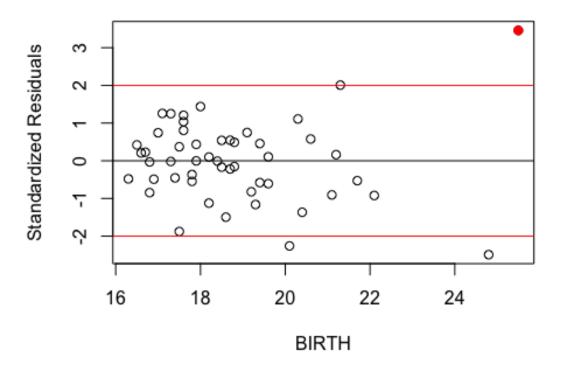
LIFE = 70.5577813 + 0.1261019 (MALE2) - 0.5160558 (BIRTH2) - 0.1965375 (DIVO2) - 0.0033392 (BEDS2) + 0.2368223 (EDUC2) - 0.0003612 (INCO2)

Also, R²s changed: New: Multiple R-squared: 0.3679 Old: Multiple R-squared: 0.4685

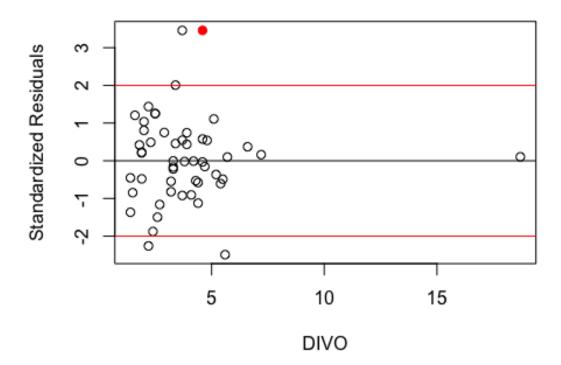
```
#MALE
plot(MALE2, std_res_NoDC, ylab="Standardized Residuals", xlab="MALE", main="S
tandardized Residuals")
abline(0, 0)
abline(2, 0, col=c("red"))
abline(-2, 0, col=c("red"))
points(MALE[44], std_res_NoDC[44], col='red', pch = 19)
```



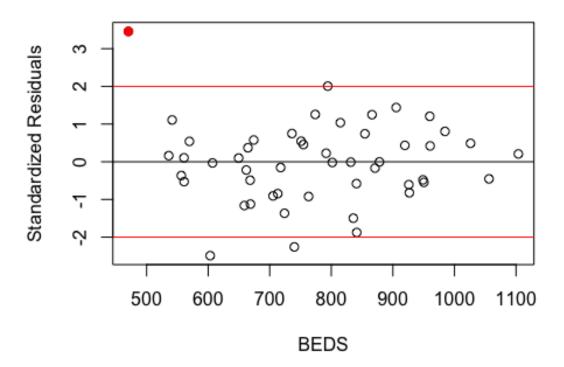
```
#BEDS
plot(BIRTH2, std_res_NoDC, ylab="Standardized Residuals", xlab="BIRTH", main=
"Standardized Residuals")
abline(0, 0)
abline(2, 0, col=c("red"))
abline(-2, 0, col=c("red"))
points(BIRTH2[44], std_res_NoDC[44], col='red', pch = 19)
```



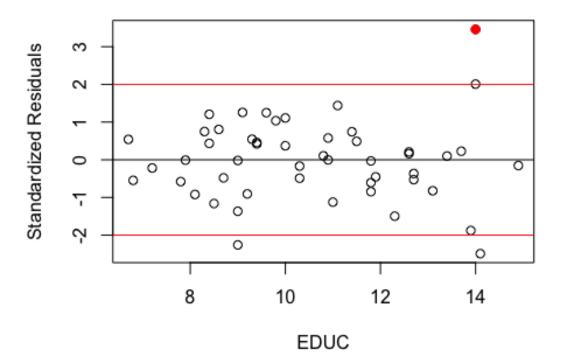
```
#DIVO
plot(DIVO2, std_res_NoDC, ylab="Standardized Residuals", xlab="DIVO", main="S
tandardized Residuals")
abline(0, 0)
abline(2, 0, col=c("red"))
abline(-2, 0, col=c("red"))
points(DIVO[44], std_res_NoDC[44], col='red', pch = 19)
```



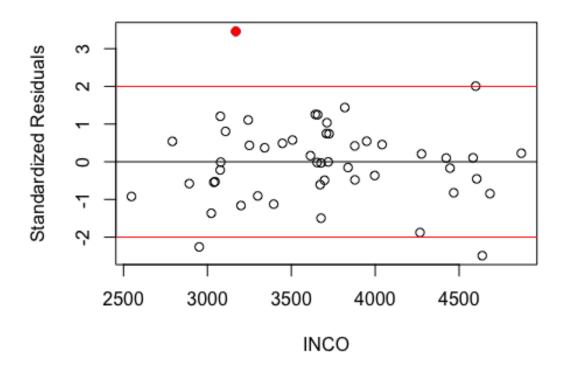
```
#BEDS
plot(BEDS2, std_res_NoDC, ylab="Standardized Residuals", xlab="BEDS", main="S
tandardized Residuals")
abline(0, 0)
abline(2, 0, col=c("red"))
abline(-2, 0, col=c("red"))
points(BEDS2[44], std_res_NoDC[44] , col='red', pch = 19)
```



```
#EDUC
plot(EDUC2, std_res_NoDC, ylab="Standardized Residuals", xlab="EDUC", main="S
tandardized Residuals")
abline(0, 0)
abline(2, 0, col=c("red"))
abline(-2, 0, col=c("red"))
points(EDUC2[44], std_res_NoDC[44], col='red', pch = 19)
```



```
#INCO
plot(INCO2, std_res_NoDC, ylab="Standardized Residuals", xlab="INCO", main="S
tandardized Residuals")
abline(0, 0)
abline(2, 0, col=c("red"))
abline(-2, 0, col=c("red"))
points(INCO2[44], std_res_NoDC[44], col='red', pch = 19)
```



ALL

Data corresponding to this STATE has very standardized residual values. Difference between observations and expected values is large.

```
dat3 <- subset(dat2, STATE != 'UT')
LIFE3 <- dat3$LIFE
MALE3 <- dat3$MALE
BIRTH3 <- dat3$BIRTH
DIVO3 <- dat3$DIVO
BEDS3 <- dat3$BEDS
EDUC3 <- dat3$EDUC
INCO3 <- dat3$INCO
MLR_NODC_NOUT <- lm(formula = LIFE3 ~ MALE3+ BIRTH3 + DIVO3 + BEDS3 + EDUC3 +
INCO3, data=dat3)</pre>
```

 R^2 is higher without UT data. Therefore this model explains the proportion of variability explained by the regression, better.

Predictor have coefficients changed:

```
MLR_NoDC_NoUT$coefficients

## (Intercept) MALE3 BIRTH3 DIV03 BEDS3

## 68.2344178920 0.1494306784 -0.6334939159 -0.1116666651 -0.0007407427

## EDUC3 INC03

## 0.2163469795 -0.0003390279
```

Vs

```
MLR$coefficients

## (Intercept) MALE BIRTH DIVO BEDS

## 70.5577812704 0.1261018758 -0.5160557876 -0.1965375074 -0.0033392036

## EDUC INCO

## 0.2368222541 -0.0003612011
```