Inference in Multi-parameter Models, Conditional Posterior, Local Conjugacy

CS698X: Topics in Probabilistic Modeling and Inference
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Plan

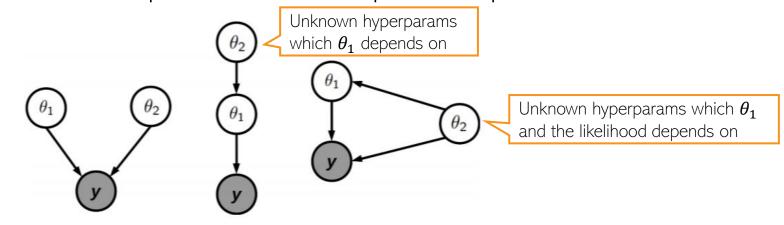
- Foray into models with several parameters
- Goal: Infer the posterior over all of them (not posterior for some, MLE-II for others)
- Idea of conditional/local posteriors in such problems
- Local conjugacy (which helps in computing conditional posteriors)
- Gibbs sampling (an algorithm that infer the joint posterior via conditional posteriors)
- An example: Bayesian matrix factorization (model with many parameters)
- Conditional/local posterior, local conjugacy, etc are important ideas (will appear in many inference algorithms that we will see later)

Moving Beyond Simple Models..

- So far, our models usually had a "main" parameter and maybe a few hyperparams, e.g.,
 - For a Gaussian, infer the mean assuming variance known (or vice-versa)
 - Bayesian linear regression with weight vector \boldsymbol{w} and noise/prior precision hyperparams β , λ
 - GP regression with one function to be learned, and fixed hyperparams
- Easy posterior inference if the likelihood and prior are conjugate to each other
- Hyperparams usually estimated via MLE-II (full posterior much harder)
- For non-conjugate models or models with many parameters, need posterior approx
 - Can use Laplace approx but it has limitations (unimodality posterior, model should be differentiable)
- We will now look at more general inference schemes for "difficult" cases
 - Difficult = Models with many params/hyperparams, non-conjugacy, non-differ., etc
 - Will be intractable in general. We will study approx. inference methods to handle such cases

Multiparameter Models

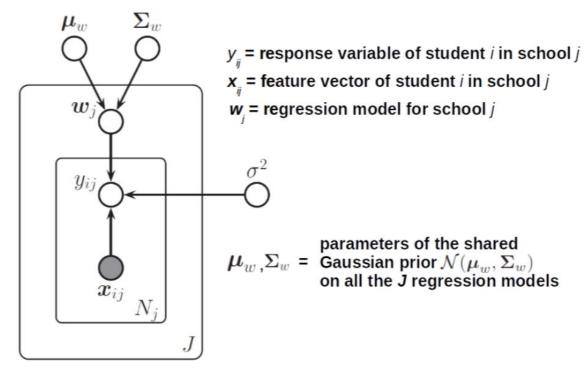
- lacktriangle Multiparameter models consist of two or more unknowns, say $heta_1$ and $heta_2$
- \blacksquare Given some data y, some examples for the simple two parameter case



- lacktriangle Assume the likelihood model to be of the form $p(m{y}|m{ heta}_1,m{ heta}_2)$
- Assume a joint prior distribution $p(\theta_1, \theta_2)$ This prior may still be conditioned on some fixed hyperparams
- The joint posterior $p(\theta_1, \theta_2 | \mathbf{y}) \propto p(\theta_1, \theta_2) p(\mathbf{y} | \theta_1, \theta_2)$
 - Easy the joint prior is conjugate to the likelihood (e.g., NIW prior for Gaussian likelihood)
 - Otherwise needs more work, e.g., MLE-II, MCMC, VB, etc. (already saw MLE-II)

Multiparameter Models: Some Examples

- Multiparameter models arise in many situations, e.g.,
 - Probabilistic models with unknown hyperparams (e.g., Bayesian linear regression)
 - Joint analysis of data from multiple (and possibly related) groups: Hierarchical models



.. and in fact, pretty much in any non-toy example of probabilistic model



Another Example: Matrix Factorization/Completion®

- Given: Data $\mathbf{R} = \{r_{ij}\}$ of "interactions" (e.g., ratings)
 - Here i = 1, 2, ..., N denotes users, j = 1, 2, ..., M denotes items

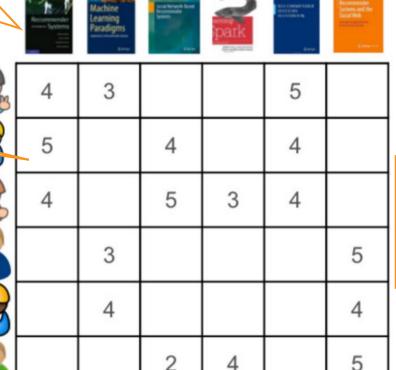
Assume each item j=1,2,...,M modeled by an unknown parameter vector $\boldsymbol{v}_j \in \mathbb{R}^K$

Only a small fraction of user-item ratings are observed (say user-item index set Ω)

Need to predict the rest using these

Many methods exist

Assume each user i=1,2,...,N modeled by an unknown parameter vector $\boldsymbol{u}_i \in \mathbb{R}^K$ for some small K



Similar to assuming a low-rank structure on the ratings matrix

Can assume zero mean Gaussian noise $\mathcal{N}(0, \beta^{-1})$ for real-valued ratings

$$r_{ij} = oldsymbol{u}_i^{\mathsf{T}} oldsymbol{v}_j + oldsymbol{\epsilon}_{ij}$$
 Likelihood

$$p(r_{ij}|\boldsymbol{u}_i,\boldsymbol{v}_j) = \mathcal{N}(r_{ij}|\boldsymbol{u}_i^{\mathsf{T}}\boldsymbol{v}_j,\beta^{-1})$$

Individual priors for each user and item

Individual priors for
$$p(u_i) = \mathcal{N}(u_i|\mathbf{0},\lambda_u^{-1})$$

$$p(\mathbf{v}_j) = \mathcal{N}(\mathbf{v}_j | \mathbf{0}, \lambda_v^{-1})$$

$$p(\{{m u}_i\}_{i=1}^N, \{{m v}_j\}_{j=1}^M | {m R})$$

Posterior

For simplicity, hyperparams β , λ_u , λ_v are fixed or unknow but we may only want to do point estimation for them

Bayesian Matrix Factorization (BMF): The Posterior

Our target posterior distribution for this model is

$$p(\mathbf{U}, \mathbf{V}|\mathbf{R}) = \frac{p(\mathbf{R}|\mathbf{U}, \mathbf{V})p(\mathbf{U}, \mathbf{V})}{\int \int p(\mathbf{R}|\mathbf{U}, \mathbf{V})p(\mathbf{U}, \mathbf{V})d\mathbf{U}d\mathbf{V}}$$

$$= \frac{\prod_{(i,j)\in\Omega} p(r_{ij}|\mathbf{u}_i, \mathbf{v}_j)\prod_{i} p(\mathbf{u}_i) \prod_{j} p(\mathbf{v}_j)}{\int \dots \int \prod_{(i,j)\in\Omega} p(r_{ij}|\mathbf{u}_i, \mathbf{v}_j) \prod_{i} p(\mathbf{u}_i) \prod_{j} p(\mathbf{v}_j) d\mathbf{u}_1 \dots d\mathbf{u}_N d\mathbf{v}_1 \dots d\mathbf{v}_M}$$

$$= \frac{\prod_{(i,j)\in\Omega} p(r_{ij}|\mathbf{u}_i, \mathbf{v}_j) \prod_{i} p(\mathbf{u}_i) \prod_{j} p(\mathbf{v}_j) d\mathbf{u}_1 \dots d\mathbf{u}_N d\mathbf{v}_1 \dots d\mathbf{v}_M}{\prod_{i} p(\mathbf{u}_i) \prod_{j} p(\mathbf{v}_j) d\mathbf{u}_1 \dots d\mathbf{v}_M}$$

- Posterior still tractable since integrals here are intractable
- Need to approx. the posterior. One way is via conditional posteriors (CP), e.g.,

$$p(u_i|R,V,U_{-i})$$
 All of V except u_i $p(v_j|R,U,V_{-j})$

- CP of each unknown is conditioned on fixed value of all other unknowns
 - The different CPs can be computed in an alternating fashion (like ALT-OPT/EM)
- Note: CP individually won't give us joint posterior. Need to combine them

E.g., using MCMC, variational inference, EM, etc

Conditional Posterior and Local Conjugacy

- Conditional Posteriors are easy to compute for model that are locally conjugacy
 - Note: Some researchers also call each CP as Complete Conditional or Local Posterior
- lacktriangle Consider a model with data $m{X}$ and $m{K}$ unknown params/h.p. $\Theta = (\theta_1, \theta_2, ..., \theta_K)$
- Suppose the joint posterior $p(\Theta|X) = \frac{p(\Theta)p(X|\Theta)}{p(X)}$ is intractable (like BMF)
- Local Conjugacy: If we can compute each CP tractably

 Θ_{-k} is assumed known while computing this CP

Possible if the likelihood $p(X|\theta_k, \Theta_{-k})$ and prior $p(\theta_k)$ are conjugate to each other

This is the notion of local conjugacy as opposed to full/joint conjugacy

$$p(\theta_k|\mathbf{X},\Theta_{-k}) = \frac{p(\mathbf{X}|\theta_k,\Theta_{-k})p(\theta_k)}{\int p(\mathbf{X}|\theta_k,\Theta_{-k})p(\theta_k)d\theta_k} \propto p(\mathbf{X}|\theta_k,\Theta_{-k})p(\theta_k)$$

- Important: In the above context, when considering the likelihood $p(X|\theta_k,\Theta_{-k})$
 - lacktriangledown X actually refers to only that part of data X that depends on $heta_k$

In the likelihood model

lacksquare Θ_{-k} refers to only those unknowns that "interact" with θ_k in generating that part of data

Approximating Joint Posterior via CPs

■ With the conditional posterior based approximation, the target joint posterior

$$p(\Theta|\mathbf{X}) = \frac{p(\mathbf{X}|\Theta)p(\Theta)}{p(\mathbf{X})}$$

... is represented by several conditional posteriors $p(\theta_k, | X, \Theta_{-k})$, k = 1, 2, ..., K

- Each CP is a distribution over one unknown $heta_k$, given all other unknowns
- Need a way to "combine" these CPs to get the overall posterior
 - MCMC (e.g., Gibbs sampling): Based on generating samples from the CPs
 - Variational Inference (VI): Based on cyclic estimation of each CP
 - Note: Expectation Maximization also computes CP of latent variables in its E step
- More on this when we discuss MCMC, VI, EM, etc

Will revisit Gibbs sampling again when discussing MCMC algos

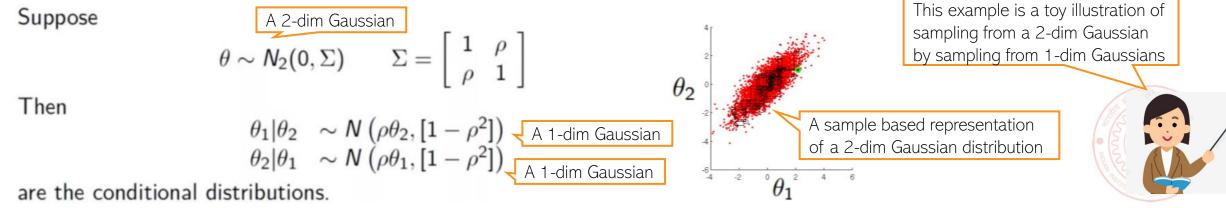
But let's look at Gibbs sampling (an MCMC algo) right away as it is fairly simple.

Gibbs Sampling (Geman and Geman, 1982)

- A general algo to generate samples from multivar. distr. one component at a time
 - Not limited to sampling from intractable posteriors only
 - Sometimes can be used even if we can draw from the multivar distr. directly

Note: If posterior, it will be conditioned on other stuff too (e.g., data, other param, etc)

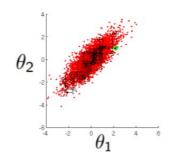
- Assume we want to sample from a joint distribution $p(\theta_1, \theta_2, ..., \theta_K)$
- It generates one component θ_k at a time using its conditional $p(\theta_k|\Theta_{-k})$
- Each conditional is assumed to be available in closed form. An example below:



Gibbs Sampling (Geman and Geman, 1982)

- Can be used to get a sampling-based approx. of a multiparam. posterior
- Gibbs sampler iteratively draws random samples from CPs
- When run long enough, the sampler produces samples from the joint posterior
- For the simple two-param case $\theta = (\theta_1, \theta_2)$, the Gibb sampler looks like this
 - Initialize $\theta_2^{(0)}$
 - For s = 1, 2, ..., S
 - Draw a random sample for θ_1 as $\theta_1^{(s)} \sim p(\theta_1|X,\theta_2^{(s-1)})$ This CP uses the most recent
 - Draw a random sample for θ_2 as $\theta_2^{(s)} \sim p(\theta_2|X, \theta_1^{(s)})$ value of θ_1

This CP uses the most recent value of θ_2 This CP uses the most recent point θ_2



- These S random samples $\left(\theta_1^{(s)}, \theta_1^{(s)}\right)_{s=1}^{S}$ represent joint posterior $p(\theta_1, \theta_2|X)$
- This is just a high-level idea. More on this when we discuss MCMC

Back to Bayesian Matrix Factorization



Bayesian Matrix Factorization: The CPs

- BMF with Gaussian likelihood and Gaussian prior on each user/item params is not fully conjugate but has local conjugacy
- lacktriangle To see this, note that the conditional posterior (CP) for user parameter $oldsymbol{u}_i$

$$p(u_i|\mathsf{R},\mathsf{V},\mathsf{U}_{-i}) \propto \prod_{j:(i,j)\in\Omega} p(r_{ij}|u_i,v_j)p(u_i)$$
 Only depends on the ratings of user i . Also doesn't depend on U_{-i} This is due to the structure of the problem and the u_i 's being independent a priori $=\prod_{j:(i,j)\in\Omega} \mathcal{N}(r_{ij}|u_i^\top v_j,\beta^{-1})\mathcal{N}(u_i|0,\lambda_u^{-1}\mathsf{I}_K)$

- lacktriangle The above is just like Bayesian linear regression, given $m{R}$ and fixed $m{V}$
 - lacktriangle Weight vector is $oldsymbol{u}_i$, training data is $\{(oldsymbol{v}_j, r_{ij})\}_{j:(i,j)\in\Omega}$, given
 - Also have local conjugacy since likelihood and prior are conjugate (assuming hyperparams fixed)
- lacktriangle Likewise, the CP for the item parameter $oldsymbol{v}_{i}$ can be computed as

$$p(\mathbf{v}_j|\mathbf{R},\mathbf{U}) \propto \prod_{i:(i,j)\in\Omega} \mathcal{N}(r_{ij}|\mathbf{u}_i^{\top}\mathbf{v}_j,\beta^{-1}) \mathcal{N}(\mathbf{v}_j|\mathbf{0},\lambda_v^{-1}\mathbf{I}_K)$$

Another Bayesian linear regression problem with weight vector $oldsymbol{v_j}$

Bayesian Matrix Factorization: The CPs

■ The CPs will have forms similar to solution of Bayesian linear regression

$$\rho(\mathbf{u}_{i}|\mathsf{R},\mathsf{V}) = \mathcal{N}(\mathbf{u}_{i}|\boldsymbol{\mu}_{u_{i}},\boldsymbol{\Sigma}_{u_{i}}) \qquad \qquad \rho(\mathbf{v}_{j}|\mathsf{R},\mathsf{U}) = \mathcal{N}(\mathbf{v}_{j}|\boldsymbol{\mu}_{v_{j}},\boldsymbol{\Sigma}_{v_{j}}) \\
\boldsymbol{\Sigma}_{u_{i}} = (\lambda_{u}\mathsf{I} + \beta \sum_{j:(i,j) \in \Omega} \mathbf{v}_{j}\mathbf{v}_{j}^{\top})^{-1} \qquad \qquad \boldsymbol{\Sigma}_{v_{j}} = (\lambda_{v}\mathsf{I} + \beta \sum_{i:(i,j) \in \Omega} \mathbf{u}_{i}\mathbf{u}_{i}^{\top})^{-1} \\
\boldsymbol{\mu}_{u_{i}} = \boldsymbol{\Sigma}_{u_{i}}(\beta \sum_{j:(i,j) \in \Omega} r_{ij}\mathbf{v}_{j}) \qquad \qquad \boldsymbol{\mu}_{v_{j}} = \boldsymbol{\Sigma}_{v_{j}}(\beta \sum_{i:(i,j) \in \Omega} r_{ij}\mathbf{u}_{i})$$

- These CPs can be updated in an alternating fashion until convergence
 - Many ways. One popular way is to use Gibbs sampling
 - Note: Hyperparameters can also be inferred by computing their CPs¹
- Can extend Gaussian BMF easily to other exp. family distr. while maintaining local conj.
 - Example: Poisson likelihood and gamma priors on user/item parameters²

¹ "Bayesian Probabilistic Matrix Factorization using Markov Chain Monte Carlo" by Salakhutdinov and Mnih (2008)

²"Scalable recommendation with Poisson factorization" by Gopalan et al (2013)

BMF: Making Predictions

PPD for each missing entry of the matrix (assuming hyperparams known)

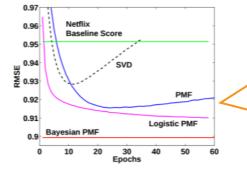
$$p(r_{ij}|\mathbf{R}) = \int \int p(r_{ij}|\boldsymbol{u}_i,\boldsymbol{v}_j)p(\boldsymbol{u}_i,\boldsymbol{v}_j|\mathbf{R})d\boldsymbol{u}_id\boldsymbol{v}_j$$

- In general, this is intractable and needs approximation
 - If using Gibbs sampling, we can use S samples $(u_i^{(s)}, v_i^{(s)})_{s=1}^S$ to compute mean of r_{ij}
 - For the Gaussian likelihood case, the mean can be computed as

Can also compute the variance in the predicted
$$r_{ij}$$
 using these S samples (think how)

$$\mathbb{E}[r_{ij}] pprox \frac{1}{S} \sum_{s=1}^{S} \boldsymbol{u}_{i}^{(s)^{\top}} \boldsymbol{v}_{j}^{(s)}$$
 (Monte-Carlo averaging)

Comparison of Bayesian MF with others (from Salakhutdinov and Mnih (2008))



All other baselines are optimization based or point estimation based probabilistic models (PMF = probabilistic matrix factorization with point estimation)



Summary

- Local conjugacy is helpful even for complex prob. models with many params
 - CPs will have a closed form
 - Easy to implement Gibbs sampling can be used to be (approx.) posterior
 - Many other approx. inference algos (like variational inference) benefit from local conjugacy
- Helps to choose likelihood and priors on each param as exp. family distr.
 - So if we can't get a globally conjugate model, we can still get a model with local conjugacy
- Even if we can't have local conjugacy, the notion of CPs is applicable
 - We can break the inference problem into estimating CPs (exactly if we have local conjugacy, or approximately if we don't have local conjugacy)
 - Almost all approx. algorithms work by estimating CPs exactly or approximately



Coming Up

- Latent Variable Models
- Expectation Maximization
- Variational Inference

