

# Predictability and Reduced Order Modeling in Stochastic Reaction Networks

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## Outline

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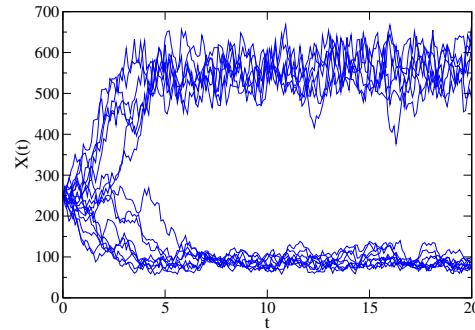
- Motivation: Stochastic Reaction Networks
- Predictability: Parametric Uncertainty Propagation using Polynomial Chaos Expansion
  - Polynomial Chaos
  - Bayesian Inference
  - Markov Chain Monte Carlo
  - Adaptive Domain Decomposition
- Dynamical Analysis: Reduced Order Modeling via Karhunen-Loève Decomposition
  - Karhunen-Loève Decomposition
  - Rosenblatt Transformation
  - Quadrature Integration
  - Adaptive Data Clustering
- Application: Schlögl Model (a benchmark bistable process)



## Motivation: Stochastic Reaction Networks (SRNs)

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- Reaction networks involving small number of molecules necessitate the use of *stochastic* modeling instead of the *deterministic* one. E.g.
  - Immune system signaling reactions
  - Microbial reactions
  - Surface catalytic reactions
- SRNs are modeled as Jump Markov Processes
  - Governed by Chemical Master Equation
$$\dot{P}(X(t) = n) = \sum_m A_{nm} P(X(t) = n)$$
  - Reduces to deterministic Rate Equations in the large volume limit
  - Numerically, Gillespie's Stochastic Simulation Algorithm (SSA, Gillespie, 1977)



## Problem Definition and Methods

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- Develop tools for *predictability*( $\lambda$ ) and *dynamical analysis*( $t$ ) of SRNs accounting for

- Inherent stochasticity ( $\theta$ )
  - Model/parameter variability ( $\lambda$ )
  - Limited data

$$\mathcal{D} = \{X_i\}_{i=1}^N$$

- Techniques employed:
  - Polynomial chaos expansion; Bayesian inference; Domain decomposition
  - Karhunen-Loève decomposition; Rosenblatt transformation; Data clustering

$$X(t, \theta, \lambda)$$

## Polynomial Chaos Expansion (PCE) - Intro

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- A second order random variable  $X(\boldsymbol{\theta})$  can be described by a PCE in terms of standard orthogonal polynomials  $\Psi_k$ , of associated standard random variables  $\{\eta_i\}_{i=1}^{\infty}$ , and spectral mode strengths  $c_k$ .

(Wiener, 1938)(Cameron & Martin, 1947)(Ghanem & Spanos, 1991)

- Truncated PCE: finite dimension  $n$  and order  $p$

$$X(\boldsymbol{\theta}) \simeq \sum_{k=0}^P c_k \Psi_k(\eta_1, \dots, \eta_n)$$

with the number of terms  $P + 1 = \frac{(n+p)!}{n!p!}$ .

- Most common standard Polynomial-Variable pairs:  
(continuous) Gauss-Hermite, Legendre-Uniform,  
(discrete) Poisson-Charlier.  
(Askey Scheme: Xiu & Karniadakis, 2002)



## Galerkin Projection is challenged by the intrinsic noise

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$$X(\boldsymbol{\theta}) \simeq \sum_{k=0}^P c_k \Psi_k(\boldsymbol{\eta}) = g_{\mathcal{D}}(\boldsymbol{\eta})$$
$$c_k = \frac{\langle X(\boldsymbol{\theta}) \Psi_k(\boldsymbol{\eta}) \rangle}{\langle \Psi_k^2(\boldsymbol{\eta}) \rangle}$$

- Intrusive Spectral Projection (ISP)
  - ★ Direct projection of governing equations
  - ★ Leads to deterministic equations for PC coefficients
  - \* No explicit governing equation for SRNs
- Non-intrusive Spectral Projection (NISP)
  - ★ Sampling based
  - ★ No explicit evolution equation for  $X$  needed
  - \* Galerkin projection not well-defined for SRNs



## Bayesian inference handles the intrinsic stochasticity well

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$$\overbrace{P(\mathbf{c}|\mathcal{D})}^{\text{Posterior}} \propto \overbrace{P(\mathcal{D}|\mathbf{c})}^{\text{Likelihood}} \overbrace{P(\mathbf{c})}^{\text{Prior}}$$
$$L(\mathbf{c}) = P(\mathcal{D}|\mathbf{c}) = \prod_{i=1}^N \text{pdf}_g(X_i)$$

$$X \simeq \sum_{k=0}^P c_k \Psi_k(\eta) = g_{\mathcal{D}}(\eta)$$

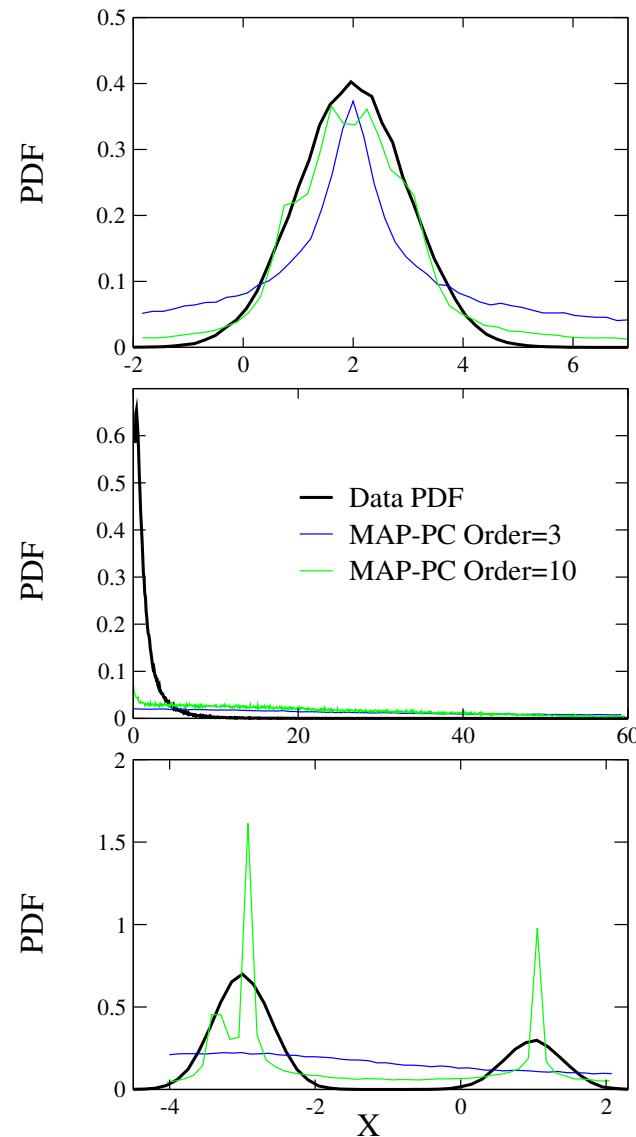
- Noise model is inherent in SSA data  $\mathcal{D} = \{X_i\}_{i=1}^N$
- Uniformly distributed priors
- Posterior exploration using Markov Chain Monte Carlo (MCMC)
- The whole posterior distribution is accessible
- Maximum a posteriori (MAP) estimate:  $\mathbf{c}^{MAP} = \operatorname{argmax}_{\mathbf{c}} P(\mathbf{c}|\mathcal{D})$



## Global PCE can fail for strongly non-linear or bimodal variables

### Legendre-Uniform PC

- Normal
- Lognormal
- Bimodal

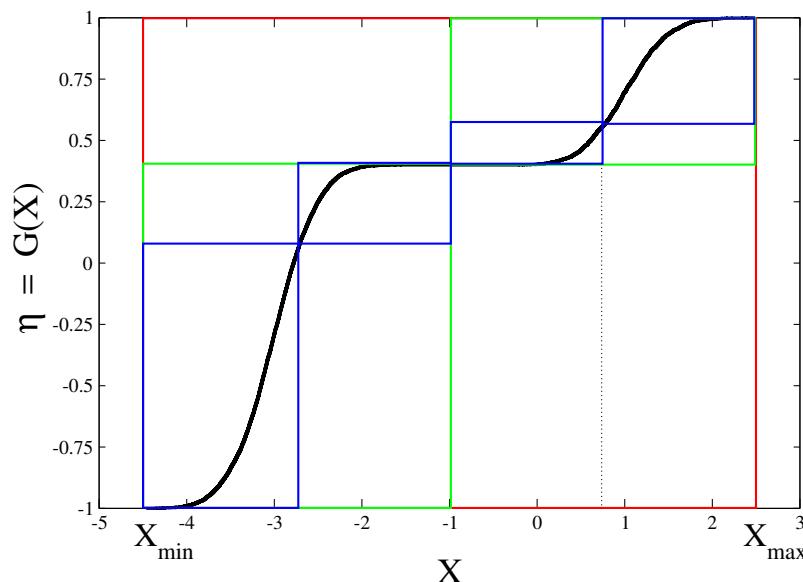


## Two domain decomposition regimes available

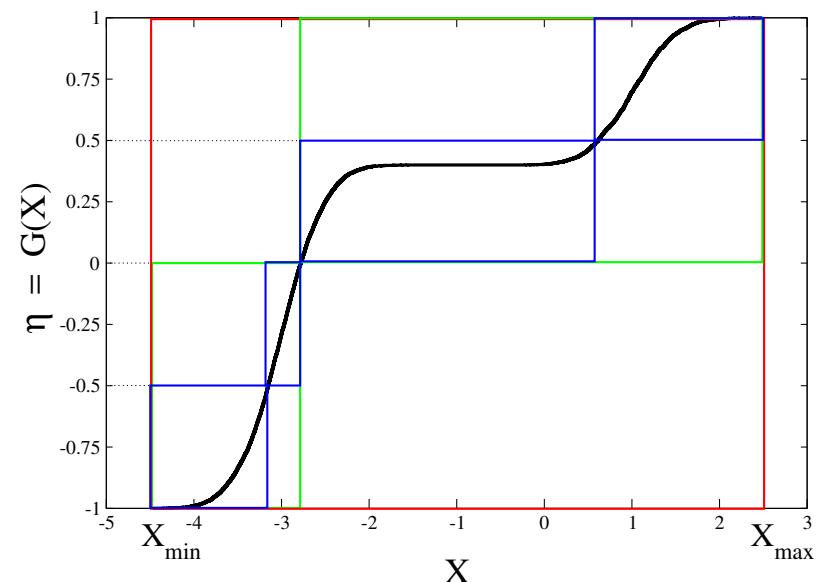
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Cumulative Distribution Function (CDF):  $F(x) = P(X < x)$ .

Rescaled CDF:  $\eta = G(X) \equiv 2F(X) - 1$  is Uniform[-1,1].



$X$ -domain decomposition



$\eta$ -domain decomposition

Le Maître et al., 2004 - Adaptive multi-wavelet basis.

Wan & Karniadakis, 2005 - Adaptive domain decomposition.



## Adaptive criterion is consistent with Kullback-Leibler ‘distance’

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Data:  $\mathcal{D} = \{X_i\}_{i=1}^N$

Model:  $X \simeq \sum_{k=0}^P c_k \Psi_k(\eta) = g_{\mathcal{D}}(\eta)$

MAP-PC samples:  $\{Y_i\}_{i=1}^N$ , where  $Y_i = g_{\mathcal{D}}(\eta_i)$

- Log-likelihood:

$$\log L = \log P(\text{Data}|\text{Model}) = \sum_{i=1}^N \log P_Y(X_i)$$

- Target log-likelihood (the *perfect match* log-likelihood, i.e. for  $\{Y_i\}_{i=1}^N = \mathcal{D}$ ):

$$\log L_T = \sum_{i=1}^N \log P_X(X_i)$$

- Kullback-Leibler divergence:

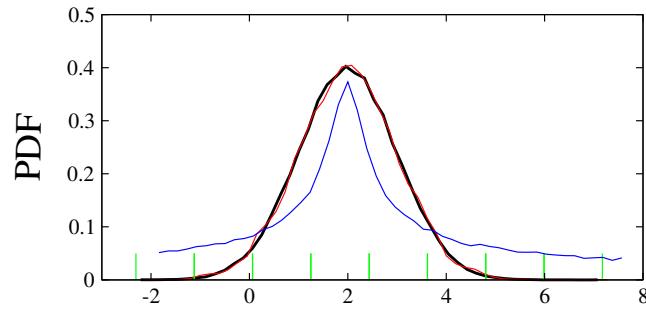
$$\rho(P_X, P_Y) = \int P(z) \log \frac{P_X(z)}{P_Y(z)} dz \simeq \frac{1}{N} \sum_{i=1}^N \log \frac{P_X(X_i)}{P_Y(X_i)}$$



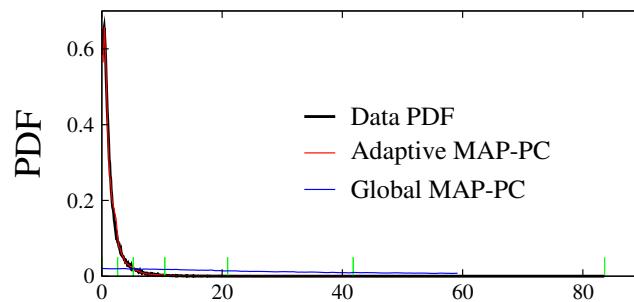
## Multi-domain PCE captures non-linearities and bimodalities well

### Legendre-Uniform multi-domain PC

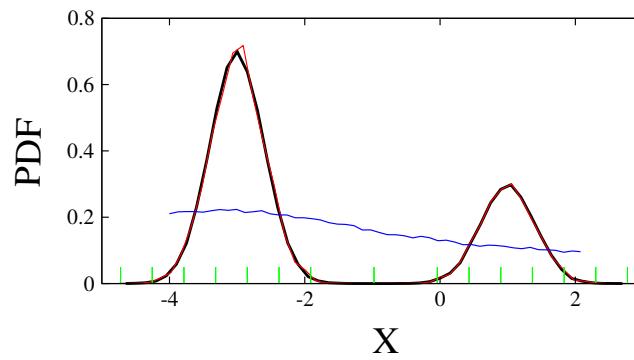
- Normal



- Lognormal

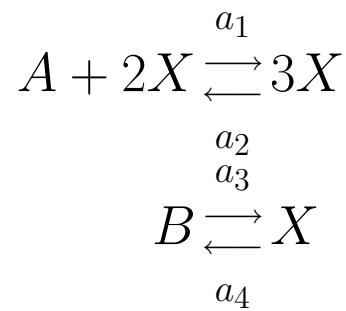


- Bimodal



## Schlögl model is a benchmark bistable process

- Reactions



- Propensities

$$a_1 = k_1 A X (X - 1)/2,$$

$$a_2 = k_2 X (X - 1)(X - 2)/6,$$

$$a_3 = k_3 B,$$

$$a_4 = k_4 X.$$

- Nominal parameters

$$k_1 A \quad 0.03$$

$$k_2 \quad 0.0001$$

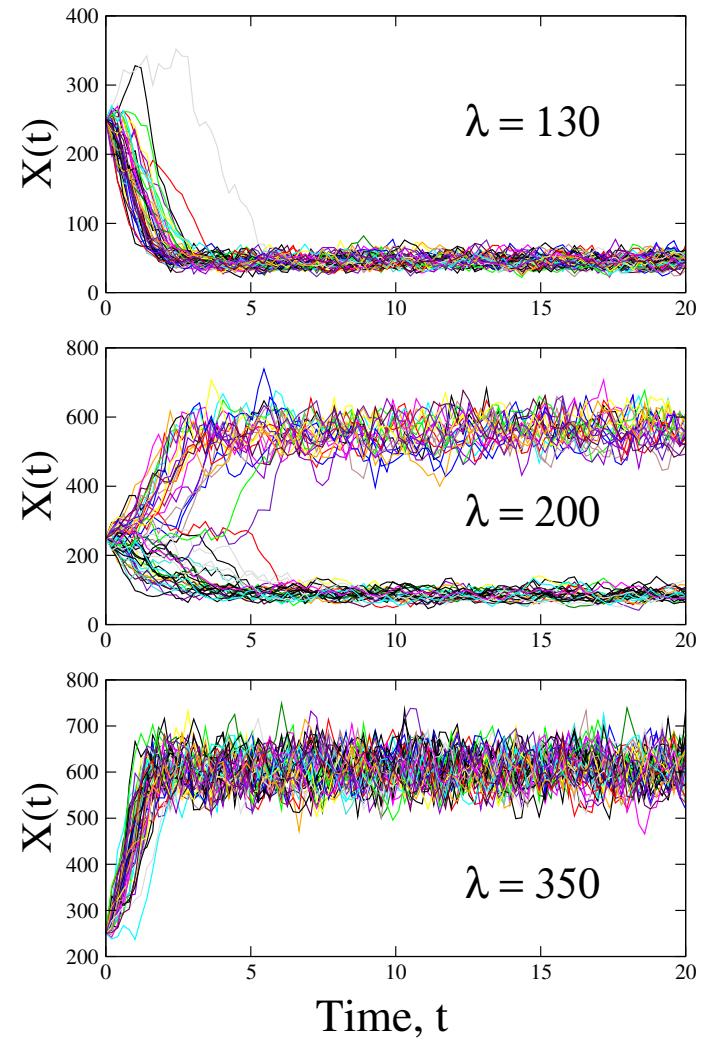
$$k_3 B = \lambda \quad 200$$

$$k_4 \quad 3.5$$

$$\frac{A}{10^5}$$

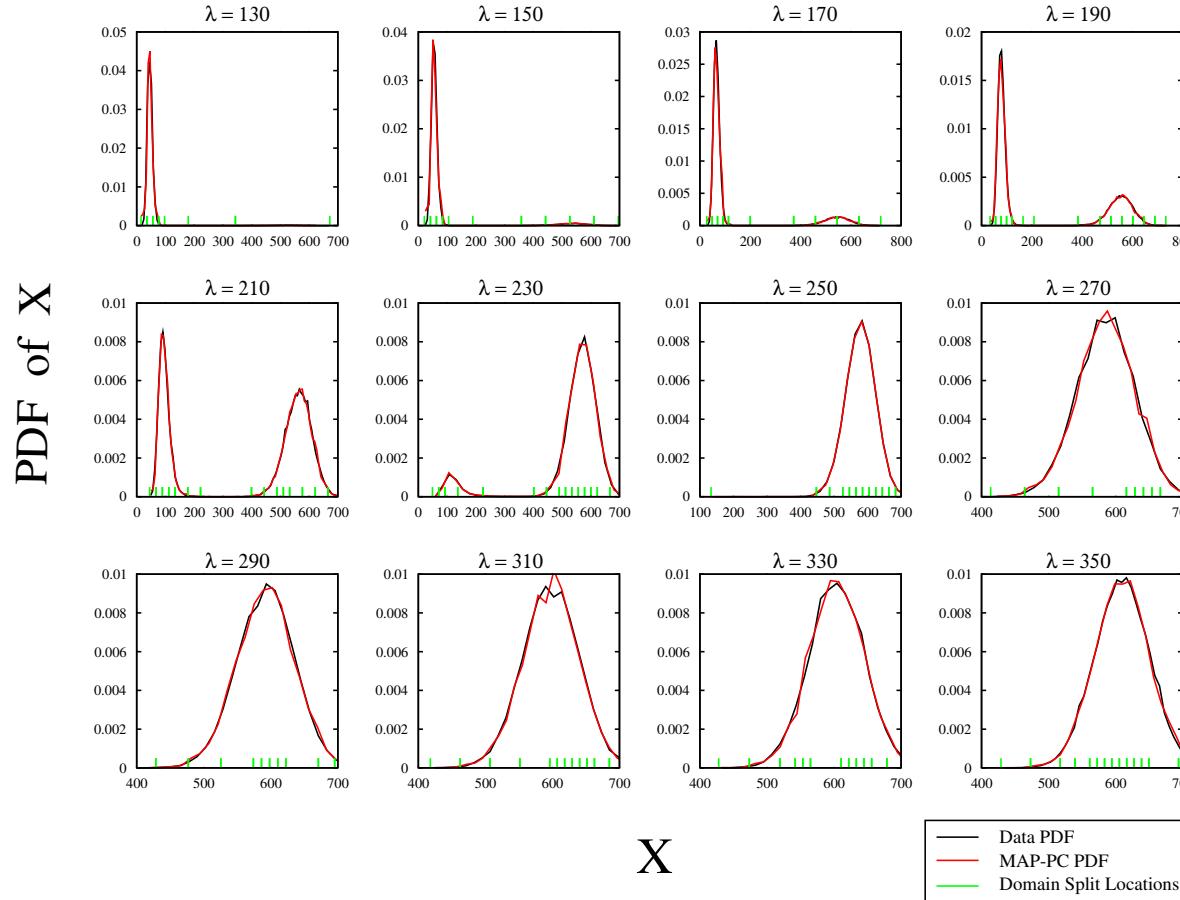
$$B \quad 2 \cdot 10^5$$

$$X(0) \quad 250$$



## PC Inference for fixed parameter values

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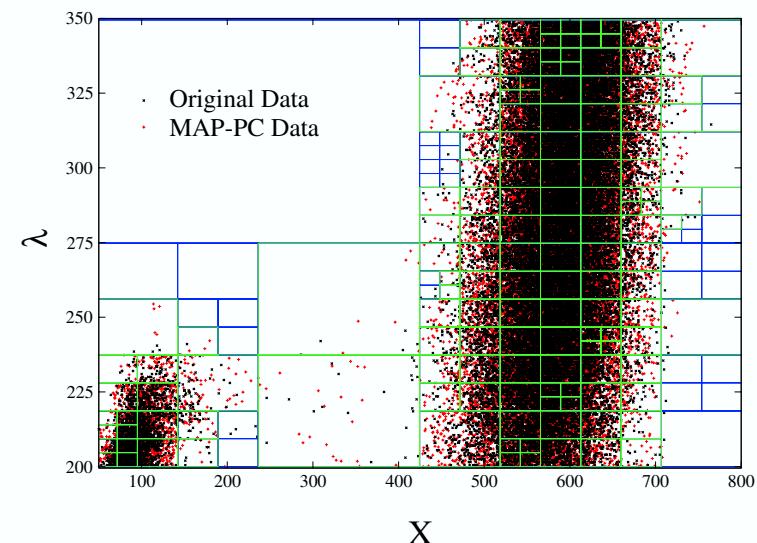
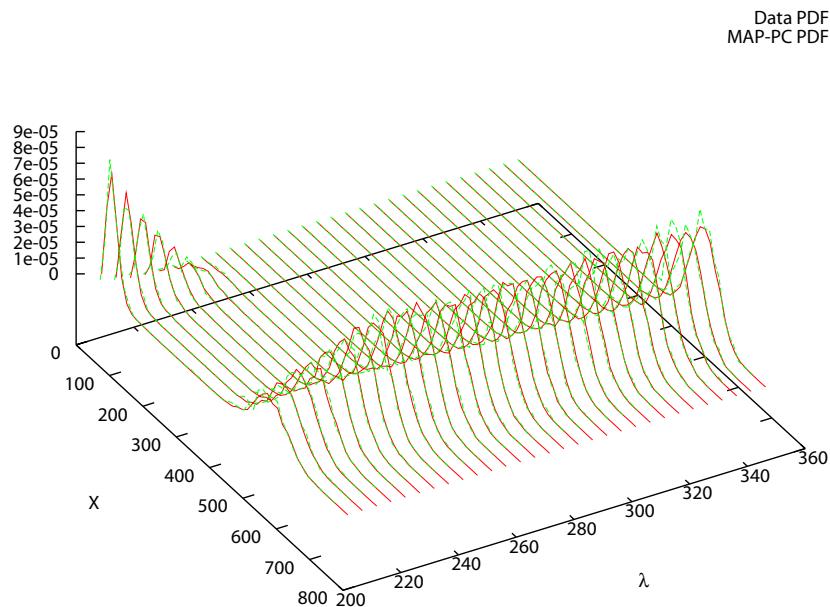


Qualitatively different behaviors across a range of  $\lambda$  values necessitates the parametric uncertainty introduction.



## Predictability: Parametric uncertainty propagation through PCE

- Postulate parametric uncertainty  $\lambda = \lambda_0 + \Delta\lambda\eta_1$
- Gather two-dimensional data  $\mathcal{D} = \{(X_i, \lambda_i)\}_{i=1}^N$
- Infer the model parameters  $c_k$ 's, where  $X = \sum_{k=0}^P c_k \Psi_k(\eta_1, \eta_2)$
- If the representation is not satisfactory (see the criterion), split the data domain and proceed recursively



## Dynamical Analysis: Big Picture

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Fix the parameter  $X(t, \boldsymbol{\theta}, \Lambda) \equiv X(t, \boldsymbol{\theta})$

$$\text{SSA} \longleftrightarrow \text{KL} \longleftrightarrow \text{PCE}$$

$$X(t, \boldsymbol{\theta}) \longleftrightarrow \xi_i(\boldsymbol{\theta}) (i = \overline{1, L}) \longleftrightarrow c_{ik} (i = \overline{1, L}, k = \overline{0, P})$$

$$X(t, \boldsymbol{\theta}) - \bar{X}(t) \simeq \sum_{i=1}^L \xi_i(\boldsymbol{\theta}) \sqrt{\lambda_i} f_i(t) \simeq \sum_{i=1}^L \left( \sum_{k=0}^P c_{ik} \Psi_k(\boldsymbol{\eta}) \right) \sqrt{\lambda_i} f_i(t)$$

**SSA**  $\longleftrightarrow$  **KL** : Karhunen-Loève (KL) decomposition of the stochastic process

**KL**  $\longleftrightarrow$  **PCE**: Polynomial Chaos expansion of each KL random variable



## Karhunen-Loève Decomposition: Intro

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- Separate the average:

$$X_0(t, \theta) = X(t, \theta) - \bar{X}(t)$$

- The covariance function is symmetric, bounded and positive definite. Hence, it can be expanded as a sum

$$C(t_1, t_2) = \langle X_0(t_1, \theta) X_0(t_2, \theta) \rangle = \sum_{n=1}^{\infty} \lambda_n f_n(t_1) f_n(t_2)$$

- Positive eigenvalues:

$$\int_0^T C(t_1, t_2) f_n(t_1) dt_1 = \lambda_n f_n(t_2).$$

- KL decomposition:

$$X(t, \theta) = \bar{X}(t) + \sum_{n=1}^{\infty} \xi_n(\theta) \sqrt{\lambda_n} f_n(t)$$



## Karhunen-Loève decomposition leads to reduced order modeling

- KL decomposition:

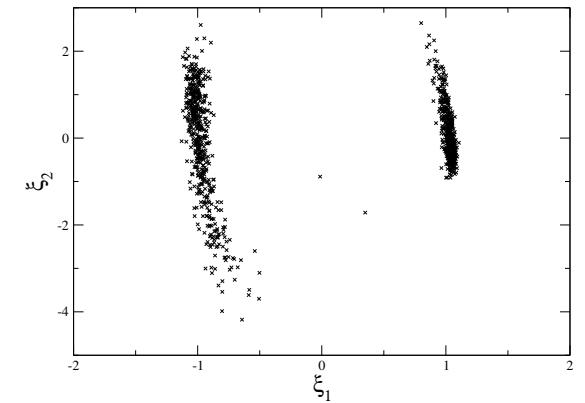
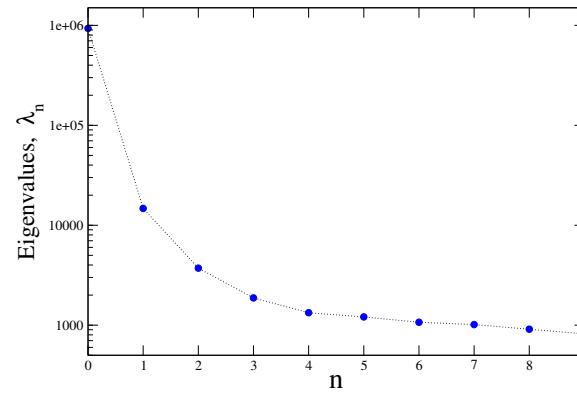
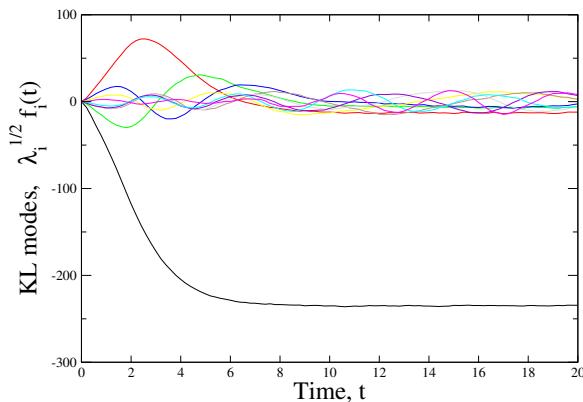
$$X(t, \theta) = \bar{X}(t) + \sum_{n=1}^{\infty} \xi_n(\theta) \sqrt{\lambda_n} f_n(t)$$

- Uncorrelated, zero-mean KL variables:

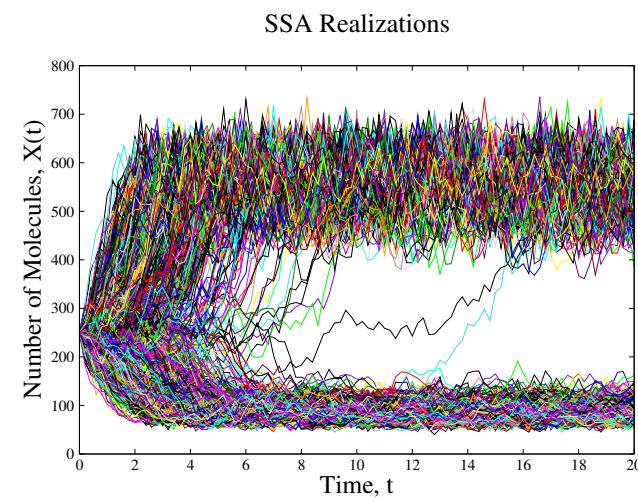
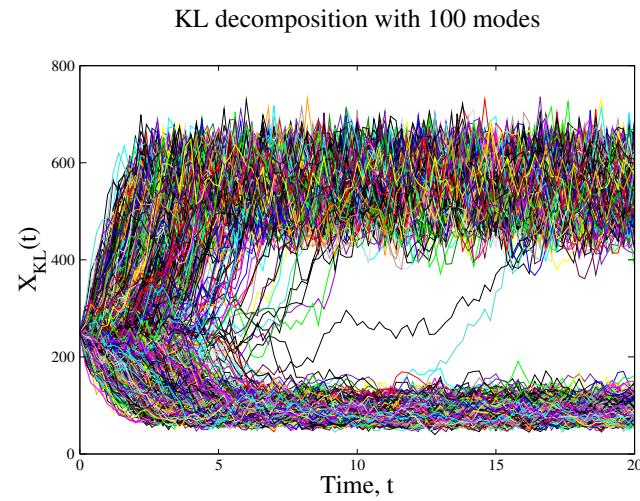
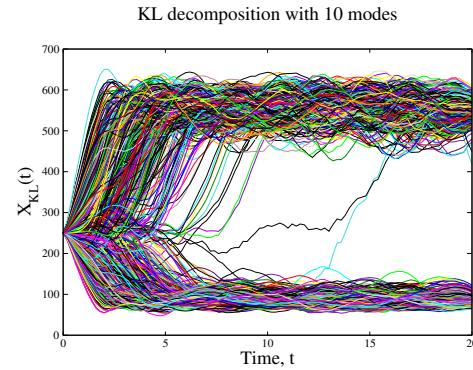
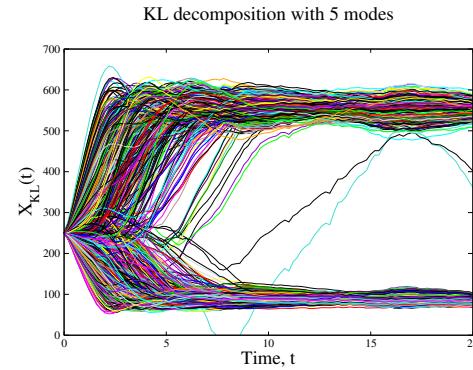
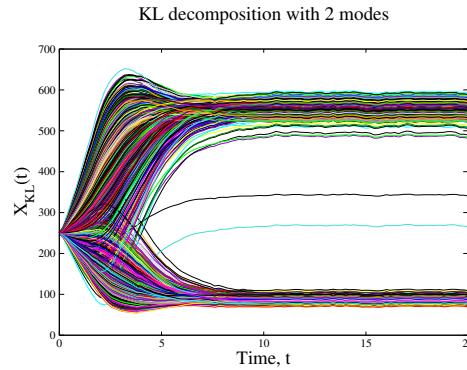
$$\langle \xi_n \rangle = 0, \quad \langle \xi_n \xi_m \rangle = \delta_{nm}$$

- SSA(continuum)  $\longleftrightarrow$  KL(discrete)

$$X(t) \longleftrightarrow \boldsymbol{\xi} = (\xi_1, \xi_2, \dots)$$

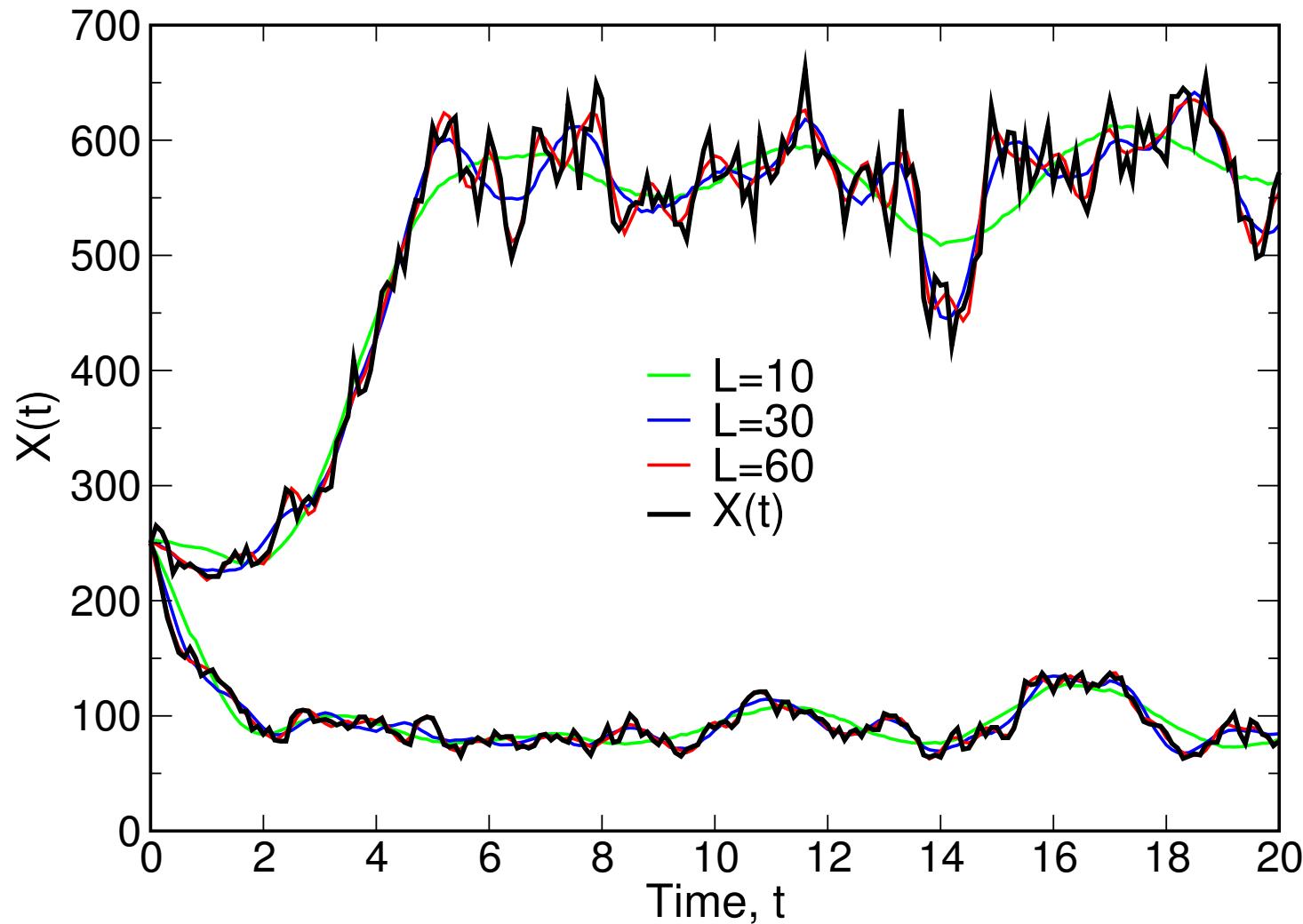


## Karhunen-Loève decomposition captures each realization



## Karhunen-Loève decomposition captures each realization

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## Challenges in PC expansions of KL random variables

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Need to PC-expand each of the KL random variables

$$\xi_i = \sum_{k=0}^P c_{ik} \Psi_k(\boldsymbol{\eta}), \text{ for } i = 1, \dots, L$$

- Quadrature-based non-intrusive spectral projection is not well-defined

$$c_{ik} = \frac{\langle \xi_i \Psi_k(\boldsymbol{\eta}) \rangle}{\langle \Psi_k^2(\boldsymbol{\eta}) \rangle}$$

- Employ (inverse) Rosenblatt transformation
- Multimodal variables not captured well
  - Use data clustering



## Rosenblatt Transformation

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- Rosenblatt transformation maps any (not necessarily independent) set of random variables  $(\xi_1, \dots, \xi_n)$  to uniform i.i.d.'s  $\{\eta_i\}_{i=1}^n$  (Rosenblatt, 1952).

$$\eta_1 = F_1(\xi_1)$$

$$\eta_2 = F_{2|1}(\xi_2|\xi_1)$$

$$\eta_3 = F_{3|2,1}(\xi_3|\xi_2, \xi_1)$$

⋮

$$\eta_n = F_{n|n-1, \dots, 1}(\xi_n|\xi_{n-1}, \dots, \xi_1)$$

- Inverse Rosenblatt transformation  $\boldsymbol{\xi} = R^{-1}(\boldsymbol{\eta})$  ensures a well-defined integration

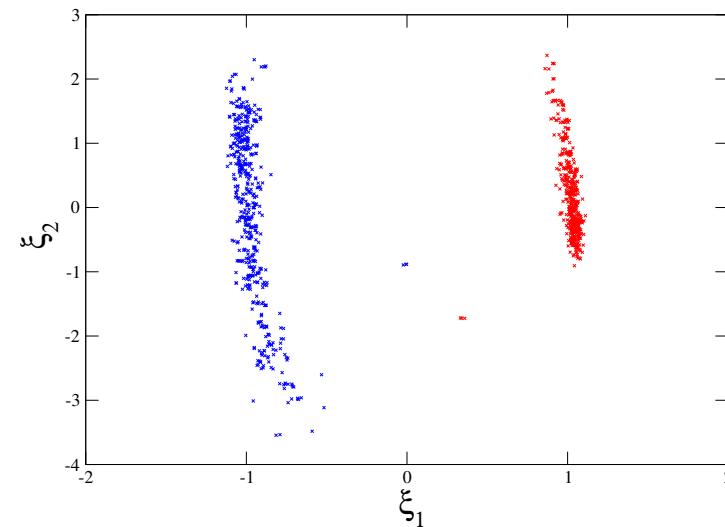
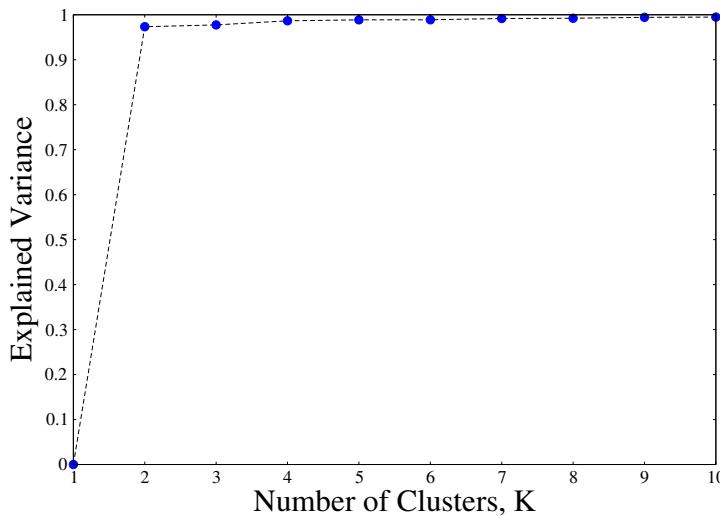
$$\langle \xi_i \Psi_k(\boldsymbol{\eta}) \rangle = \int (R^{-1}(\boldsymbol{\eta}))_i \Psi_k(\boldsymbol{\eta}) d\boldsymbol{\eta}$$



## Adaptive Data Clustering

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- Finite number of KL variables:  $\xi = (\xi_1, \xi_2, \dots, \xi_L)$
- Multidimensional data:  $\{\xi^{(i)}\}_{i=1}^N$
- K-Center clustering
- ‘Elbow’ criterion to pick initial number of clusters
- Adaptive clustering: split clusters if output is not good enough (check the Kullback-Leibler distance)



## PC Mixture Model

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- Divide data into  $K$  clusters with fractions  $p_j$ :

$$p_1 + p_2 + \cdots + p_K = 1$$

- Find PC expansion for  $\xi$  in each cluster:

$$\xi_{PC}^{(j)} = \sum_{k=0}^P \xi_k^{(j)} \Psi_k(\eta^{(j)})$$

- Superpose the results to obtain PC mixture model:

$$\xi = \xi_{PC}^{(j)} \text{ w. prob. } p_j$$

- Probability distribution function is a mixture of PC PDFs:

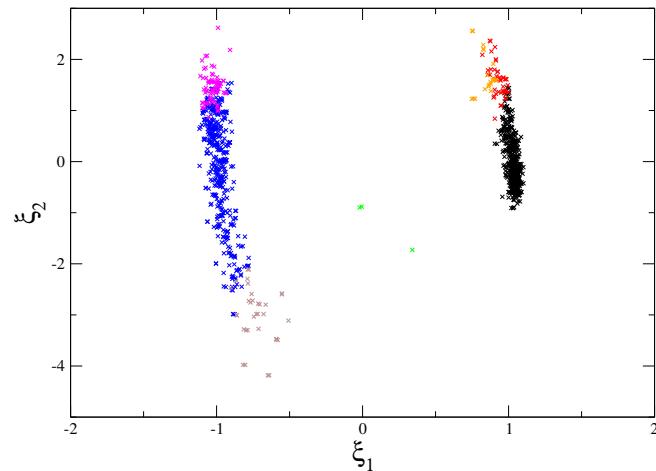
$$\text{Pdf}_\xi(x) = p_1 \text{Pdf}_{\xi_{PC}^{(1)}}(x) + \cdots + p_K \text{Pdf}_{\xi_{PC}^{(K)}}(x)$$



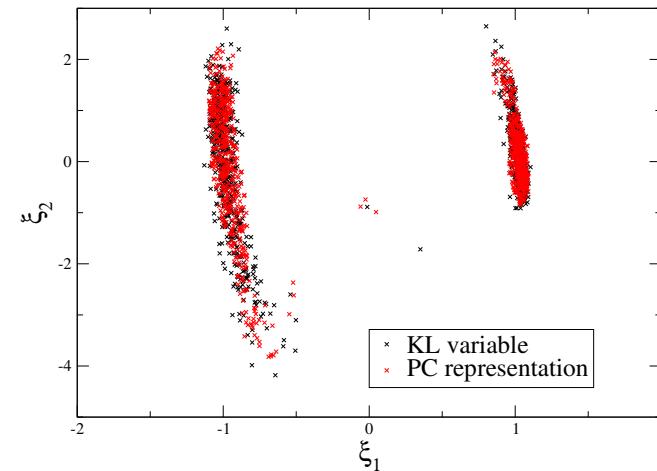
# Results for the Schlögl Model

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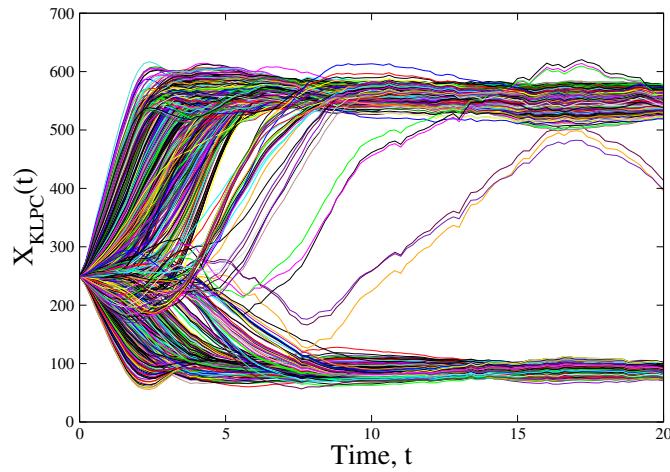
Adaptive Clustering



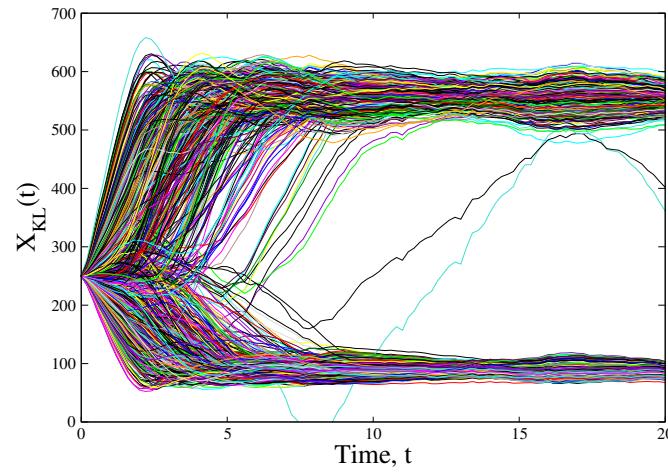
PC representation of first two KL variables



KL-PC representation, 5 KL modes, 3rd PC order



KL decomposition with 5 modes



## Conclusions and Ongoing Work

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- Adaptive domain decomposition with Bayesian inference allows representation of the state  $X(T, \theta, \lambda)$  across large range of parameter values
- PC mixture model with Karhunen-Loève decomposition represents the dynamics of the state  $X(t, \theta, \Lambda)$
  
- Dimensionality (complexity increase) studies
- Adaptive PC order
- Sparse quadrature integration or Latin Hypercube Sampling
- Combination of parameter uncertainties and time evolution



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## Multi-domain PC expansion

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- Partition:  $-1 = a_1 < b_1 = a_2 < b_2 = \dots = a_n < b_n = 1$

$$\mathcal{P} = \{[a_1, b_1), [a_2, b_2), \dots, [a_n, b_n]\}$$

- Linear map:  $f^I : I \equiv [a, b] \mapsto [-1, 1]$  from an interval  $[a, b]$  (subscripts dropped for simplicity) to  $[-1, 1]$ , by

$$f^I(\xi) = \tilde{\xi} = \frac{2}{b-a} \left( \xi - \frac{a+b}{2} \right)$$

- Multi-domain PC expansion

$$X \simeq g(\xi) = \sum_{I \in \mathcal{P}} \sum_{k=0}^P c_k^I \Psi_k^I(\xi),$$

where

$$\begin{aligned}\Psi_k^I(\xi) &\equiv 0, \quad \text{if } \xi \notin I \\ \Psi_k^I(\xi) &= \Psi_k(f^I(\xi)), \quad \text{if } \xi \in I\end{aligned}$$





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