

Predictability Assessment in Stochastic Reaction Networks

Khachik Sargsyan, Bert Debusschere, Habib Najm

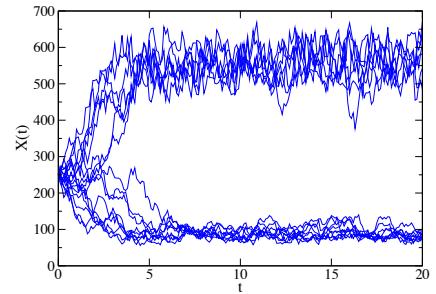
Sandia National Laboratories
Livermore, CA

Multiscale UQ Workshop
University of Southern California
March 7-8, 2011

Stochastic Reaction Networks

- Reaction networks involving small number of molecules necessitate the use of *stochastic* modeling instead of the *deterministic* one. E.g.

- Microbial processes
(bioenergy, bioremediation)
- Surface catalytic reactions
(fuel cells, batteries)
- Immune system signaling reactions

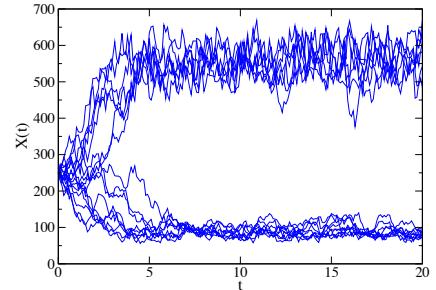


- SRNs are modeled as Jump Markov Processes

- Governed by Chemical Master Equation
 $\dot{P}(X(t) = n) = \sum_m A_{nm}P(X(t) = n)$
- Reduces to deterministic Rate Equations in the large volume limit
- Trajectories simulated by Gillespie's Stochastic Simulation Algorithm (SSA, Gillespie, 1977)

Stochastic Reaction Networks

- Reaction networks involving small number of molecules necessitate the use of *stochastic* modeling instead of the *deterministic* one. E.g.
 - Microbial processes
(bioenergy, bioremediation)
 - Surface catalytic reactions
(fuel cells, batteries)
 - Immune system signaling reactions
- SRNs are modeled as Jump Markov Processes
 - Governed by Chemical Master Equation
$$\dot{P}(X(t) = n) = \sum_m A_{nm}P(X(t) = n)$$
 - Reduces to deterministic Rate Equations in the large volume limit
 - Trajectories simulated by Gillespie's Stochastic Simulation Algorithm (SSA, Gillespie, 1977)



Objective: predictability in high-d

$$X(t, \theta, \lambda)$$

- Develop tools for *predictability*(λ) and *dynamical analysis*(t) of SRNs accounting for
 - Inherent stochasticity (θ)
 - Model/parameter uncertainty (λ)
 - Limited data

$$\mathcal{D} = \{(\lambda_i, X_i)\}_{i=1}^N$$

- Predictability assessment
 - Fix t , focus on λ dependence
 - Statistical properties $Y(\lambda) = \langle f(X(\theta, \lambda)) \rangle$ have sampling noise
 - How uncertainty in λ affects uncertainty in $Y(\lambda)$ given limited data
- High dimensionality of λ

Objective: predictability in high-d

$$X(t, \theta, \lambda)$$

- Develop tools for *predictability*(λ) and *dynamical analysis*(t) of SRNs accounting for
 - Inherent stochasticity (θ)
 - Model/parameter uncertainty (λ)
 - Limited data
- Predictability assessment
 - Fix t , focus on λ dependence
 - Statistical properties $Y(\lambda) = \langle f(X(\theta, \lambda)) \rangle$ have sampling noise
 - How uncertainty in λ affects uncertainty in $Y(\lambda)$ given limited data
- High dimensionality of λ

Objective: predictability in high-d

$$X(t, \boldsymbol{\theta}, \boldsymbol{\lambda})$$

- Develop tools for *predictability*($\boldsymbol{\lambda}$) and *dynamical analysis*(t) of SRNs accounting for
 - Inherent stochasticity ($\boldsymbol{\theta}$)
 - Model/parameter uncertainty ($\boldsymbol{\lambda}$)
 - Limited data
- Predictability assessment
 - Fix t , focus on $\boldsymbol{\lambda}$ dependence
 - Statistical properties $Y(\boldsymbol{\lambda}) = \langle f(X(\boldsymbol{\theta}, \boldsymbol{\lambda})) \rangle$ have sampling noise
 - How uncertainty in $\boldsymbol{\lambda}$ affects uncertainty in $Y(\boldsymbol{\lambda})$ given limited data
- High dimensionality of $\boldsymbol{\lambda}$

High-dimensional parametric uncertainty in stochastic systems

- Statistical property $Y(\boldsymbol{\lambda}) = \langle f(X(\boldsymbol{\theta}, \boldsymbol{\lambda})) \rangle$ of interest.
 - High-dimensional parametric uncertainty ($\boldsymbol{\lambda}$)
 - Sampling noise due to limited data $\{X_i\}$
- Expectation $\langle \cdot \rangle$ filters intrinsic noise.
 - Averaging over sample realizations of X
 - Still leftover noise, width $\sim 1/\sqrt{N}$
- Polynomial Chaos expansion to represent input-output relationship
 - Sensitivity analysis
 - Surrogate model for optimization or inverse problems
 - Identify key reaction mechanisms

Polynomial Chaos expansion represents a random variable as a polynomial of a standard random variable

- Truncated PCE: finite dimension n and order p

$$Y \simeq \sum_{k=0}^P c_k \Psi_k(\boldsymbol{\eta})$$

with the number of terms $P + 1 = \frac{(n+p)!}{n!p!}$.

- $\boldsymbol{\eta} = (\eta_1, \dots, \eta_n)$ standard i.i.d. r.v.
 Ψ_k standard orthogonal polynomials
 c_k spectral modes.
- Most common standard Polynomial-Variable pairs:
(continuous) Gauss-Hermite, Legendre-Uniform,
(discrete) Poisson-Charlier.

[Wiener, 1938; Ghanem & Spanos, 1991; Xiu & Karniadakis, 2002]

Polynomial Chaos expansion represents a random variable as a polynomial of a standard random variable

- Truncated PCE: finite dimension n and order p

$$Y \simeq \sum_{k=0}^P c_k \Psi_k(\boldsymbol{\eta})$$

with the number of terms $P + 1 = \frac{(n+p)!}{n!p!}$.

- $\boldsymbol{\eta} = (\eta_1, \dots, \eta_n)$ standard i.i.d. r.v.
- Ψ_k standard orthogonal polynomials
- c_k spectral modes.
- Most common standard Polynomial-Variable pairs:
(continuous) Gauss-Hermite, Legendre-Uniform,
(discrete) Poisson-Charlier.

[Wiener, 1938; Ghanem & Spanos, 1991; Xiu & Karniadakis, 2002]

Polynomial Chaos expansion represents a random variable as a polynomial of a standard random variable

- Truncated PCE: finite dimension n and order p

$$Y \simeq \sum_{k=0}^P c_k \Psi_k(\boldsymbol{\eta})$$

with the number of terms $P + 1 = \frac{(n+p)!}{n!p!}$.

- $\boldsymbol{\eta} = (\eta_1, \dots, \eta_n)$ standard i.i.d. r.v.
 Ψ_k standard orthogonal polynomials
 c_k spectral modes.
- Most common standard Polynomial-Variable pairs:
(continuous) Gauss-Hermite, Legendre-Uniform,
(discrete) Poisson-Charlier.

[Wiener, 1938; Ghanem & Spanos, 1991; Xiu & Karniadakis, 2002]

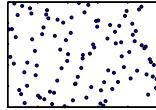
Alternative methods to obtain PC coefficients

$$Y \simeq \sum_{k=0}^P c_k \Psi_k(\boldsymbol{\eta}) \quad c_k = \frac{\langle Y(\boldsymbol{\eta}) \Psi_k(\boldsymbol{\eta}) \rangle}{\langle \Psi_k^2(\boldsymbol{\eta}) \rangle}$$

The integral $\langle Y(\boldsymbol{\eta}) \Psi_k(\boldsymbol{\eta}) \rangle = \int Y(\boldsymbol{\eta}) \Psi_k(\boldsymbol{\eta}) \pi(\boldsymbol{\eta}) d\boldsymbol{\eta}$ can be estimated by

- Monte-Carlo

$$\frac{1}{K} \sum_{j=1}^K Y(\boldsymbol{\eta}_j) \Psi_k(\boldsymbol{\eta}_j)$$



many samples from $\pi(\boldsymbol{\eta})$

- Quadrature

$$\sum_{j=1}^Q Y(\boldsymbol{\eta}_j) \Psi_k(\boldsymbol{\eta}_j) w_j$$

samples at quadrature

- *Bayesian inference*

$$P(c_k | Y(\boldsymbol{\eta}_j)) \propto P(Y(\boldsymbol{\eta}_j) | c_k) P(c_k)$$

any (number of) samples

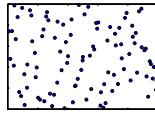
Alternative methods to obtain PC coefficients

$$Y \simeq \sum_{k=0}^P c_k \Psi_k(\boldsymbol{\eta}) \quad c_k = \frac{\langle Y(\boldsymbol{\eta}) \Psi_k(\boldsymbol{\eta}) \rangle}{\langle \Psi_k^2(\boldsymbol{\eta}) \rangle}$$

The integral $\langle Y(\boldsymbol{\eta}) \Psi_k(\boldsymbol{\eta}) \rangle = \int Y(\boldsymbol{\eta}) \Psi_k(\boldsymbol{\eta}) \pi(\boldsymbol{\eta}) d\boldsymbol{\eta}$ can be estimated by

- Monte-Carlo

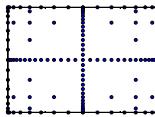
$$\frac{1}{K} \sum_{j=1}^K Y(\boldsymbol{\eta}_j) \Psi_k(\boldsymbol{\eta}_j)$$



many samples from $\pi(\boldsymbol{\eta})$

- Quadrature

$$\sum_{j=1}^Q Y(\boldsymbol{\eta}_j) \Psi_k(\boldsymbol{\eta}_j) w_j$$



samples at quadrature

- *Bayesian inference*

$$P(c_k | Y(\boldsymbol{\eta}_j)) \propto P(Y(\boldsymbol{\eta}_j) | c_k) P(c_k)$$

any (number of) samples

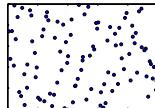
Alternative methods to obtain PC coefficients

$$Y \simeq \sum_{k=0}^P c_k \Psi_k(\boldsymbol{\eta}) \quad c_k = \frac{\langle Y(\boldsymbol{\eta}) \Psi_k(\boldsymbol{\eta}) \rangle}{\langle \Psi_k^2(\boldsymbol{\eta}) \rangle}$$

The integral $\langle Y(\boldsymbol{\eta}) \Psi_k(\boldsymbol{\eta}) \rangle = \int Y(\boldsymbol{\eta}) \Psi_k(\boldsymbol{\eta}) \pi(\boldsymbol{\eta}) d\boldsymbol{\eta}$ can be estimated by

- Monte-Carlo

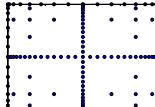
$$\frac{1}{K} \sum_{j=1}^K Y(\boldsymbol{\eta}_j) \Psi_k(\boldsymbol{\eta}_j)$$



many samples from $\pi(\boldsymbol{\eta})$

- Quadrature

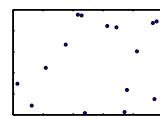
$$\sum_{j=1}^Q Y(\boldsymbol{\eta}_j) \Psi_k(\boldsymbol{\eta}_j) w_j$$



samples at quadrature

- *Bayesian inference*

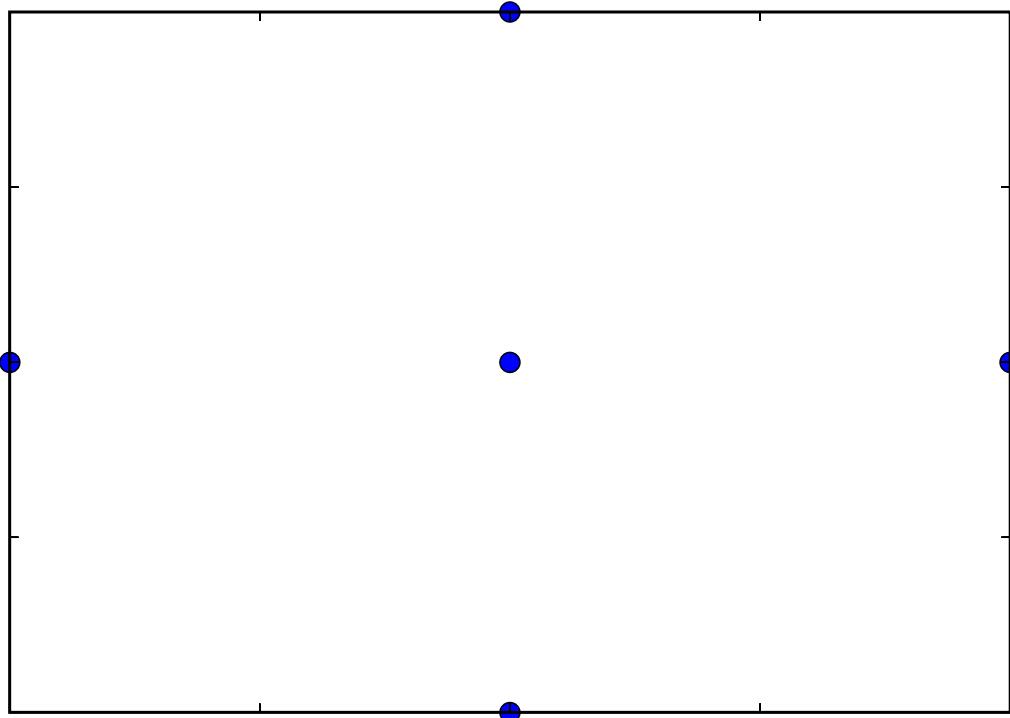
$$P(c_k | Y(\boldsymbol{\eta}_j)) \propto P(Y(\boldsymbol{\eta}_j) | c_k) P(c_k)$$



any (number of) samples

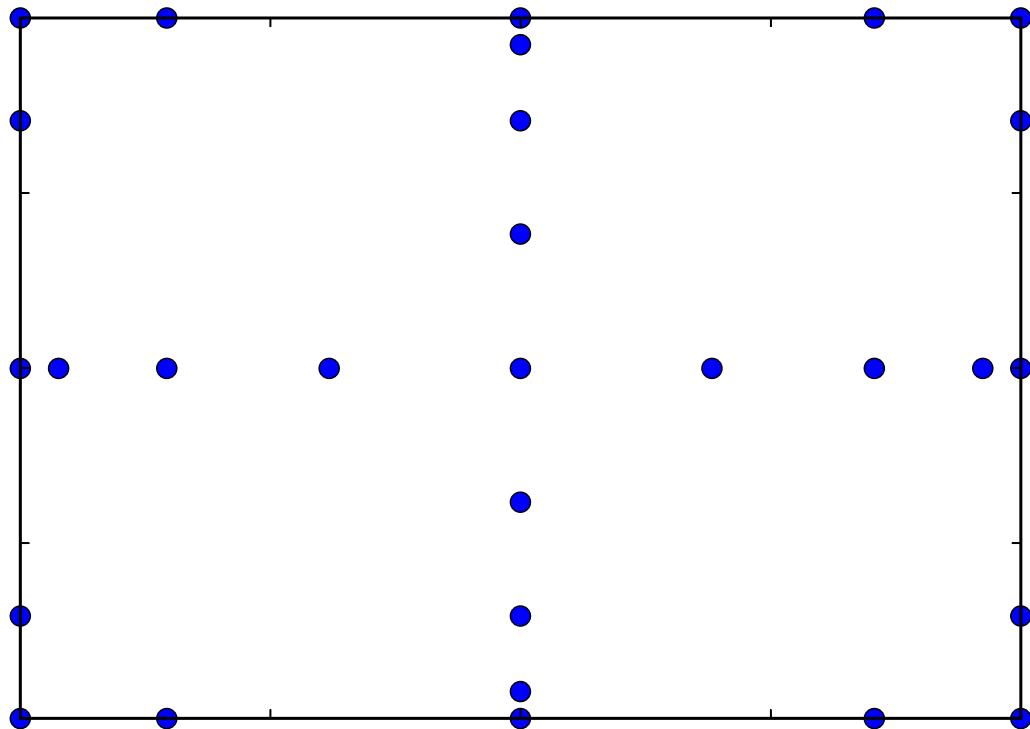
Sparse quadrature integration well-suited for high-dimensional *smooth* integrands

Clenshaw-Curtis sparse grid, Level = 1



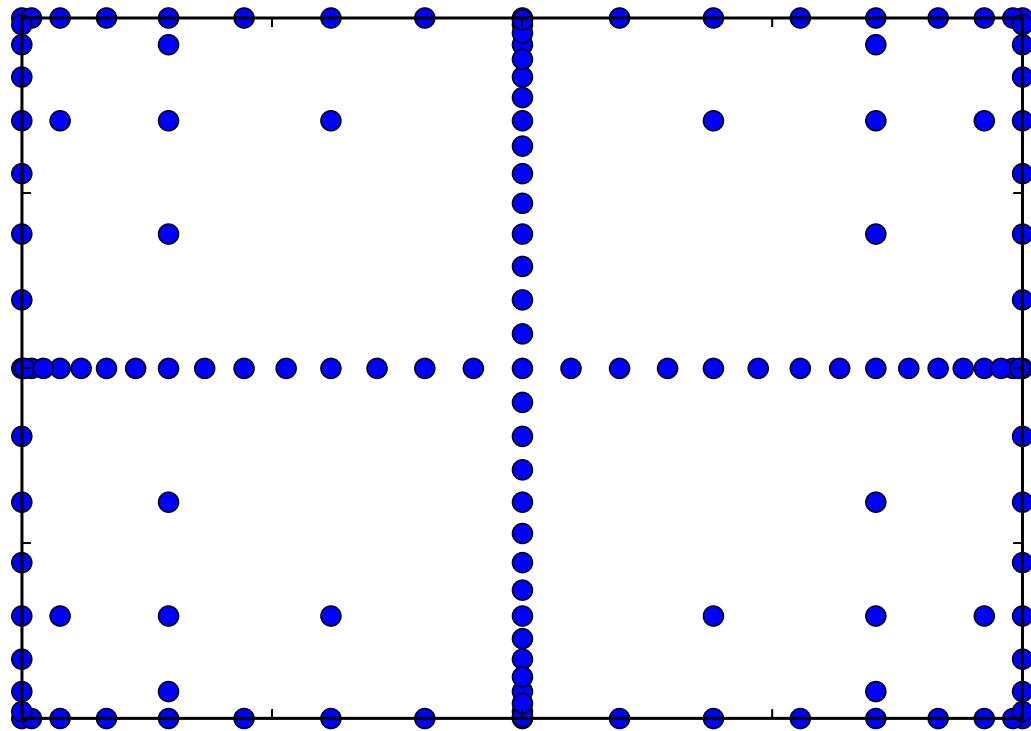
Sparse quadrature integration well-suited for high-dimensional *smooth* integrands

Clenshaw-Curtis sparse grid, Level = 3



Sparse quadrature integration well-suited for high-dimensional *smooth* integrands

Clenshaw-Curtis sparse grid, Level = 5



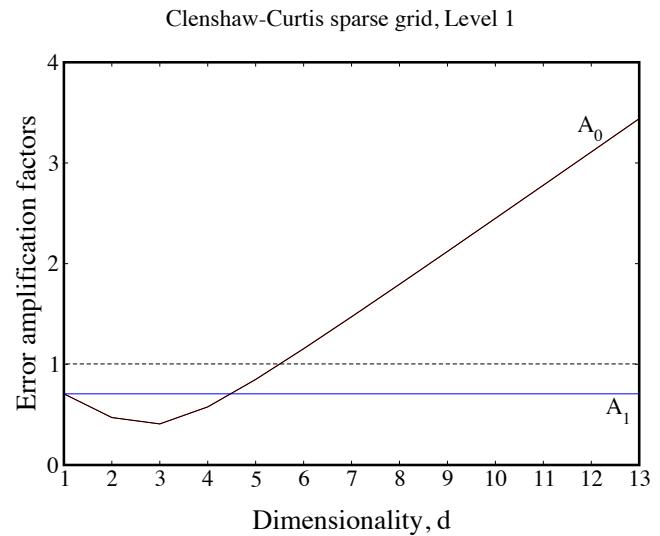
Sparse quadrature integration fails for noisy integrands

$$Y \simeq \sum_{k=0}^P c_k \Psi_k(\boldsymbol{\eta})$$

$$c_k = \frac{\langle Y(\boldsymbol{\eta}) \Psi_k(\boldsymbol{\eta}) \rangle}{\langle \Psi_k^2(\boldsymbol{\eta}) \rangle}$$

$$c_k \approx \frac{1}{\langle \Psi_k^2(\boldsymbol{\eta}) \rangle} \sum_{j=1}^Q Y(\boldsymbol{\eta}_j) \Psi_k(\boldsymbol{\eta}_j) w_j$$

Noise $Y \sim \sigma \implies$ Error $c_k \sim A_k \sigma$



- amplification factor A_k grows with dimensionality

- CC, level 1: $A_0 = \frac{1}{3} \sqrt{(d-3)^2 + \frac{d}{2}}, \quad A_1 = \frac{1}{\sqrt{2}}.$

- blame the negative weights.
- for full quadrature, $\frac{1}{n^{d/2}} \leq A_0 \leq 1$, no amplification!

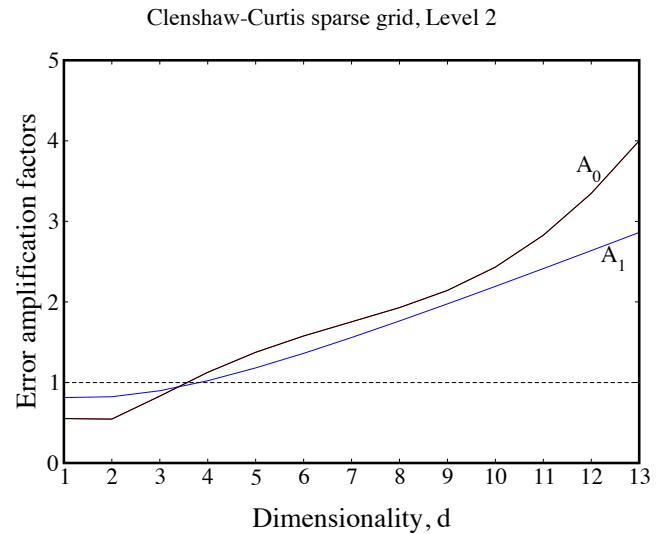
Sparse quadrature integration fails for noisy integrands

$$Y \simeq \sum_{k=0}^P c_k \Psi_k(\boldsymbol{\eta})$$

$$c_k = \frac{\langle Y(\boldsymbol{\eta}) \Psi_k(\boldsymbol{\eta}) \rangle}{\langle \Psi_k^2(\boldsymbol{\eta}) \rangle}$$

$$c_k \approx \frac{1}{\langle \Psi_k^2(\boldsymbol{\eta}) \rangle} \sum_{j=1}^Q Y(\boldsymbol{\eta}_j) \Psi_k(\boldsymbol{\eta}_j) w_j$$

Noise $_Y \sim \sigma \implies$ Error $_{c_k} \sim A_k \sigma$



- amplification factor A_k grows with dimensionality

- CC, level 1: $A_0 = \frac{1}{3} \sqrt{(d-3)^2 + \frac{d}{2}}, \quad A_1 = \frac{1}{\sqrt{2}}.$

- blame the negative weights.
- for full quadrature, $\frac{1}{n^{d/2}} \leq A_0 \leq 1$, no amplification!

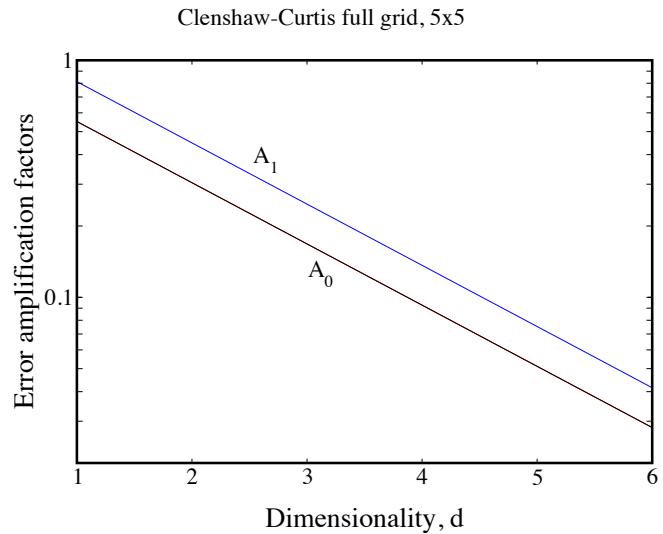
Sparse quadrature integration fails for noisy integrands

$$Y \simeq \sum_{k=0}^P c_k \Psi_k(\boldsymbol{\eta})$$

$$c_k = \frac{\langle Y(\boldsymbol{\eta}) \Psi_k(\boldsymbol{\eta}) \rangle}{\langle \Psi_k^2(\boldsymbol{\eta}) \rangle}$$

$$c_k \approx \frac{1}{\langle \Psi_k^2(\boldsymbol{\eta}) \rangle} \sum_{j=1}^Q Y(\boldsymbol{\eta}_j) \Psi_k(\boldsymbol{\eta}_j) w_j$$

Noise $_Y \sim \sigma \implies$ Error $_{c_k} \sim A_k \sigma$



- amplification factor A_k grows with dimensionality

- CC, level 1: $A_0 = \frac{1}{3} \sqrt{(d-3)^2 + \frac{d}{2}}, \quad A_1 = \frac{1}{\sqrt{2}}.$

- blame the negative weights.
- for full quadrature, $\frac{1}{n^{d/2}} \leq A_0 \leq 1$, no amplification!

Bayesian inference handles the intrinsic stochasticity well

$$Y = \langle X \rangle \simeq \sum_{k=0}^P c_k \Psi_k(\boldsymbol{\eta}) \quad \begin{array}{c} \text{Posterior} \\ \overbrace{P(\boldsymbol{c}|\mathcal{D})} \\ \times \end{array} \begin{array}{c} \text{Likelihood} \\ \overbrace{P(\mathcal{D}|\boldsymbol{c})} \\ \end{array} \begin{array}{c} \text{Prior} \\ \overbrace{P(\boldsymbol{c})} \end{array}$$

$$\text{Data} \quad \mathcal{D} \equiv \{(\tilde{\boldsymbol{\eta}}_1, X_1), \dots, (\tilde{\boldsymbol{\eta}}_N, X_N)\} = (\tilde{\boldsymbol{\eta}}, \mathbf{X})$$

$$L(\boldsymbol{c}) = P(\mathcal{D}|\boldsymbol{c}) = \left(\frac{1}{\sigma \sqrt{2\pi}} \right)^N \prod_{i=1}^N \exp \left(-\frac{(X_i - \sum_{k=0}^P c_k \Psi_k(\tilde{\boldsymbol{\eta}}_i))^2}{2\sigma^2} \right)$$

- Noise model is assumed gaussian with σ fixed or inferred
- Uniformly distributed priors
- Posterior exploration using Markov Chain Monte Carlo (MCMC)
- The whole posterior distribution is accessible,
i.e. uncertain response surface
- Input parameters can have arbitrary values

Bayesian inference handles the intrinsic stochasticity well

$$Y = \langle X \rangle \simeq \sum_{k=0}^P c_k \Psi_k(\boldsymbol{\eta}) \quad \begin{array}{c} \text{Posterior} \\ \widehat{P(\boldsymbol{c}|\mathcal{D})} \end{array} \propto \begin{array}{c} \text{Likelihood} \\ \widehat{P(\mathcal{D}|\boldsymbol{c})} \end{array} \begin{array}{c} \text{Prior} \\ \widehat{P(\boldsymbol{c})} \end{array}$$

$$\text{Data} \quad \mathcal{D} \equiv \{(\tilde{\boldsymbol{\eta}}_1, X_1), \dots, (\tilde{\boldsymbol{\eta}}_N, X_N)\} = (\tilde{\boldsymbol{\eta}}, \mathbf{X})$$

$$L(\boldsymbol{c}) = P(\mathcal{D}|\boldsymbol{c}) = \left(\frac{1}{\sigma \sqrt{2\pi}} \right)^N \prod_{i=1}^N \exp \left(-\frac{(X_i - \sum_{k=0}^P c_k \Psi_k(\tilde{\boldsymbol{\eta}}_i))^2}{2\sigma^2} \right)$$

- Posterior is a *multivariate normal*

$$\boldsymbol{c} \in \mathcal{N}(\underbrace{(\mathbf{P}^T \mathbf{P})^{-1} \mathbf{P}^T \mathbf{X}}_{\boldsymbol{\mu}}, \underbrace{\sigma^2 (\mathbf{P}^T \mathbf{P})^{-1}}_{\boldsymbol{\Sigma}}), \quad \text{where } \mathbf{P}_{ik} = \Psi_k(\tilde{\boldsymbol{\eta}}_i)$$

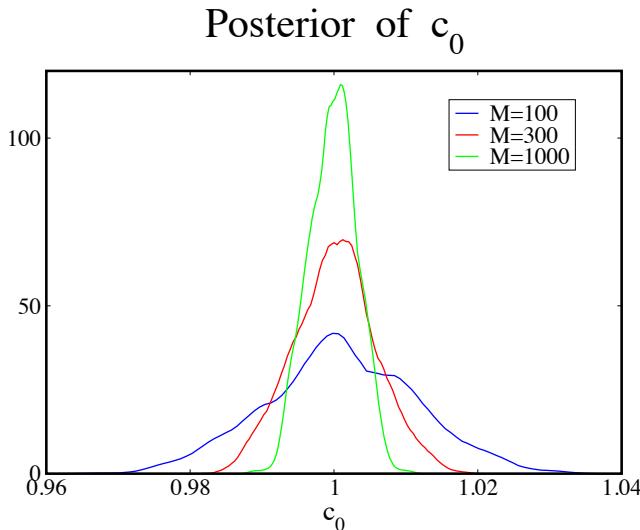
- The response surface is a *gaussian process*

$$\sum_{k=0}^P c_k \Psi_k(\boldsymbol{\eta}) = \boldsymbol{\Psi}(\boldsymbol{\eta})^T \boldsymbol{c} \in \mathcal{GP}(\boldsymbol{\Psi}(\boldsymbol{\eta})^T \boldsymbol{\mu}, \boldsymbol{\Psi}(\boldsymbol{\eta}) \boldsymbol{\Sigma} \boldsymbol{\Psi}(\boldsymbol{\eta}')^T)$$

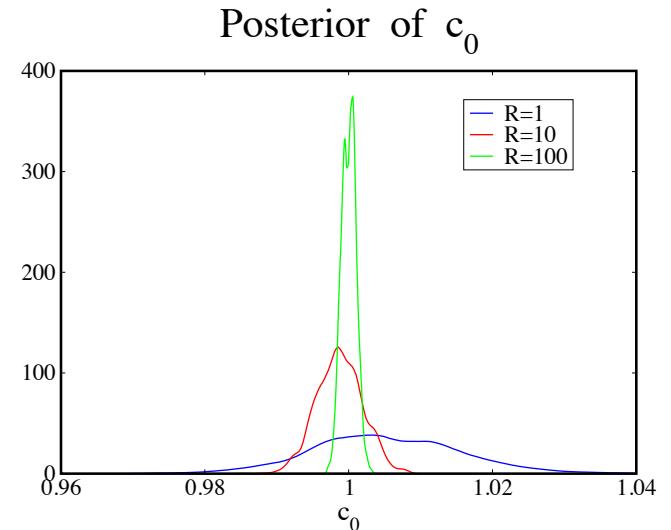
Posterior narrows around the true value as more samples are taken

- M parameter locations
- R replicas per parameter
- Second order Legendre polynomial expansion with unit coefficients.

No noise in function evaluations, $R = 1$



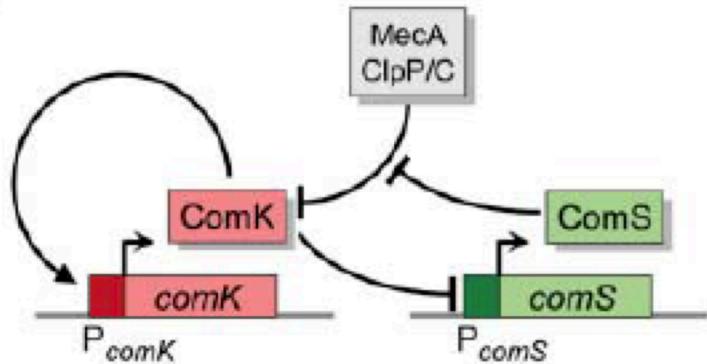
Noisy function evaluations, $M = 100$



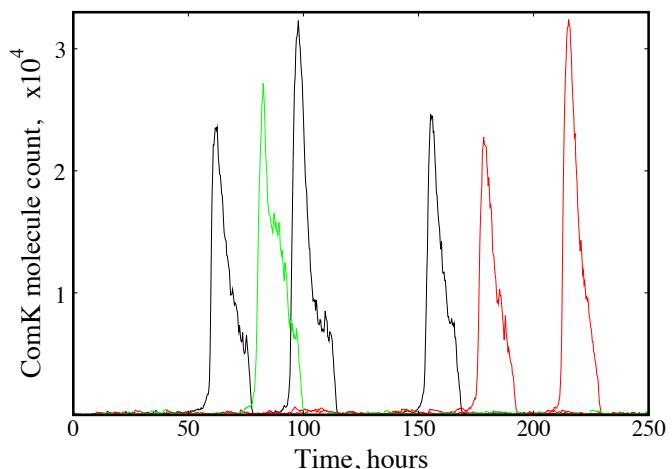
Bacillus subtilis is a soil bacterium

relevant to bioenergy and bioremediation

- 16 reactions, 11 species
- Competence in *B. Subtilis* allows uptake of external DNA
- Rapid rise in transcription factor comK molecules
- Vegetative → Competent state transition is driven by stochasticity
- Input parameters: rate constants of underlying reactions (high-d)
- Output observable: probability of competence $P_c = P(X_\infty > 5000)$



Süel et al., Science, 2007



Intrinsic stochasticity induces transition to competence

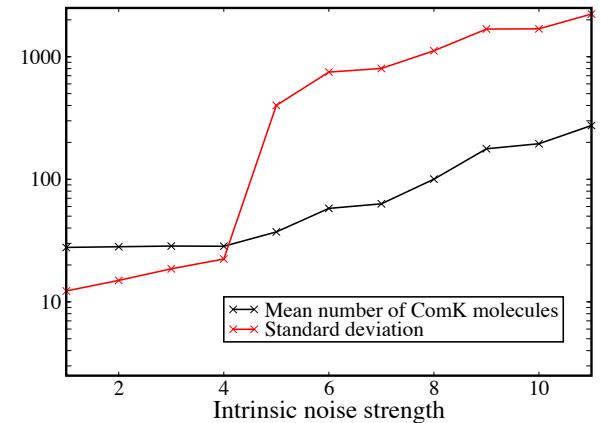
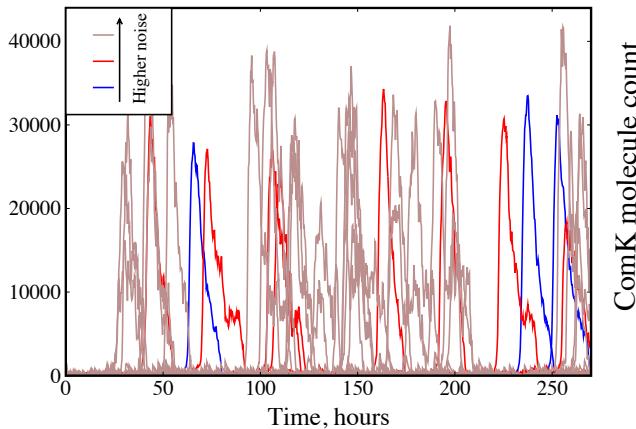
Chemical Master Equation (CME):

$$\frac{d}{dt}P(X(t) = n) = \sum_m A_{nm}P(X(t) = n)$$

Rate equation (ODE):

$$\frac{d}{dt}X(t) = \tilde{A}(X)$$

- No-noise or large volume limit (ODE) does not produce competence
- Many parameter combinations lead to the same ODE limit, but correspond to different effective volumes, i.e. intrinsic noise strength
- Increasing intrinsic noise leads to more frequent transitions



Intrinsic stochasticity induces transition to competence

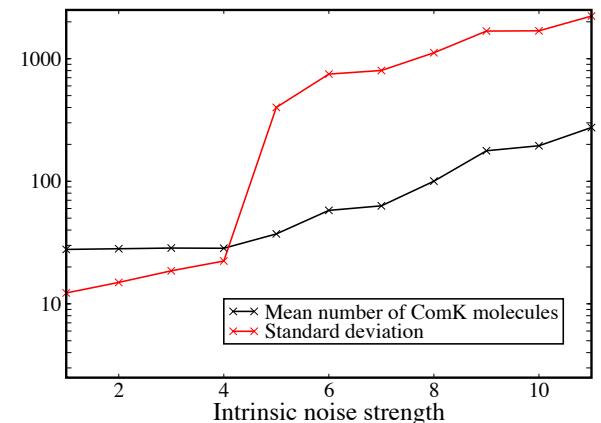
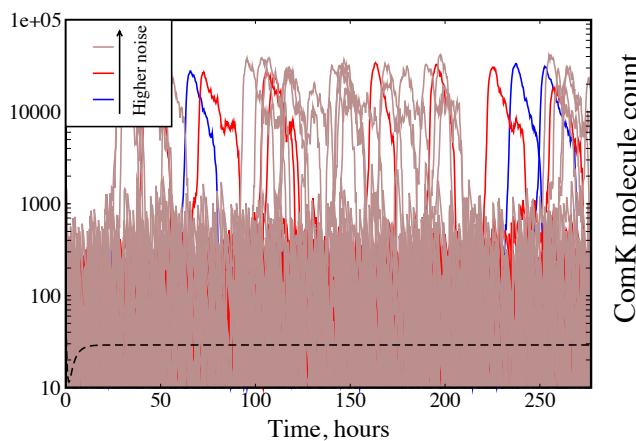
Chemical Master Equation (CME):

$$\frac{d}{dt}P(X(t) = n) = \sum_m A_{nm}P(X(t) = n)$$

Rate equation (ODE):

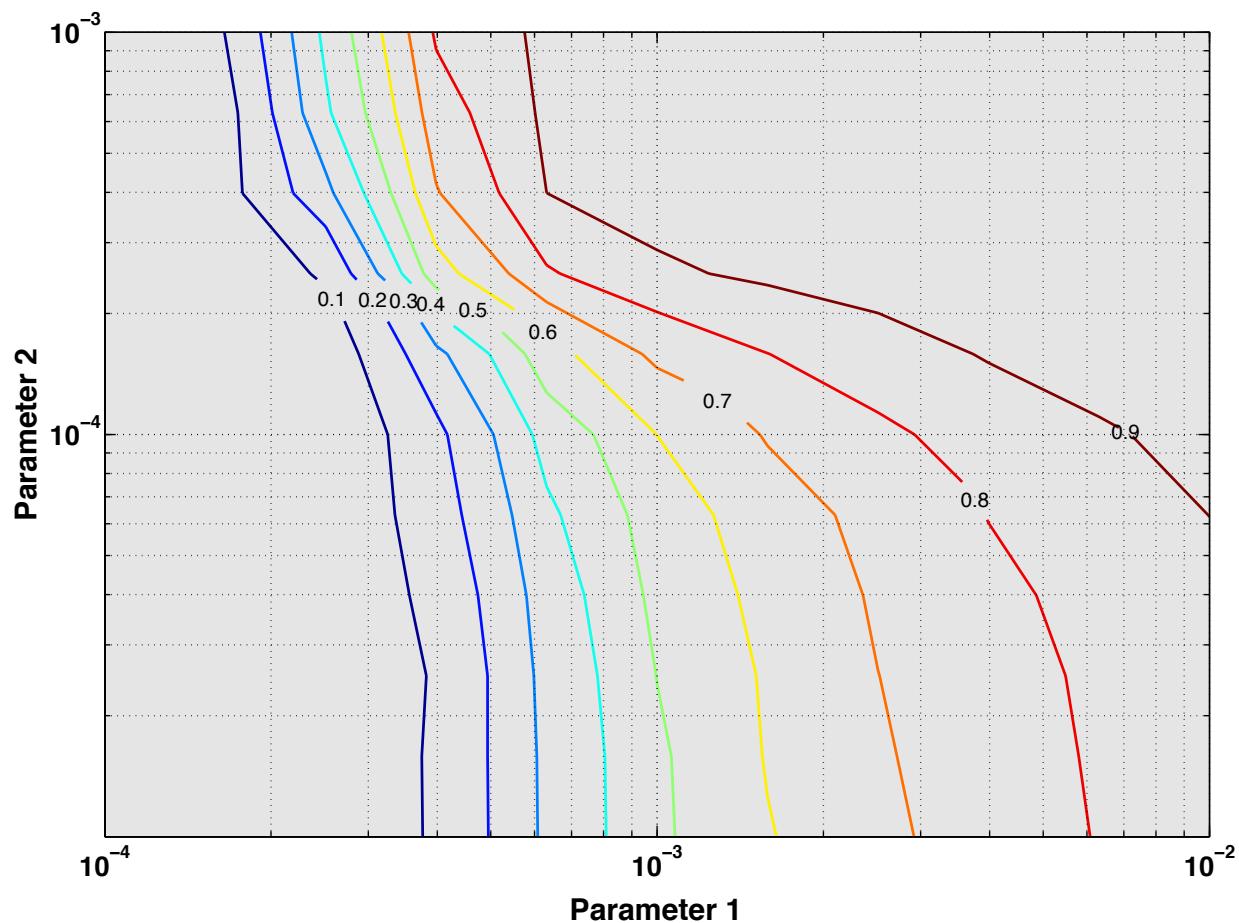
$$\frac{d}{dt}X(t) = \tilde{A}(X)$$

- No-noise or large volume limit (ODE) does not produce competence
- Many parameter combinations lead to the same ODE limit, but correspond to different effective volumes, i.e. intrinsic noise strength
- Increasing intrinsic noise leads to more frequent transitions



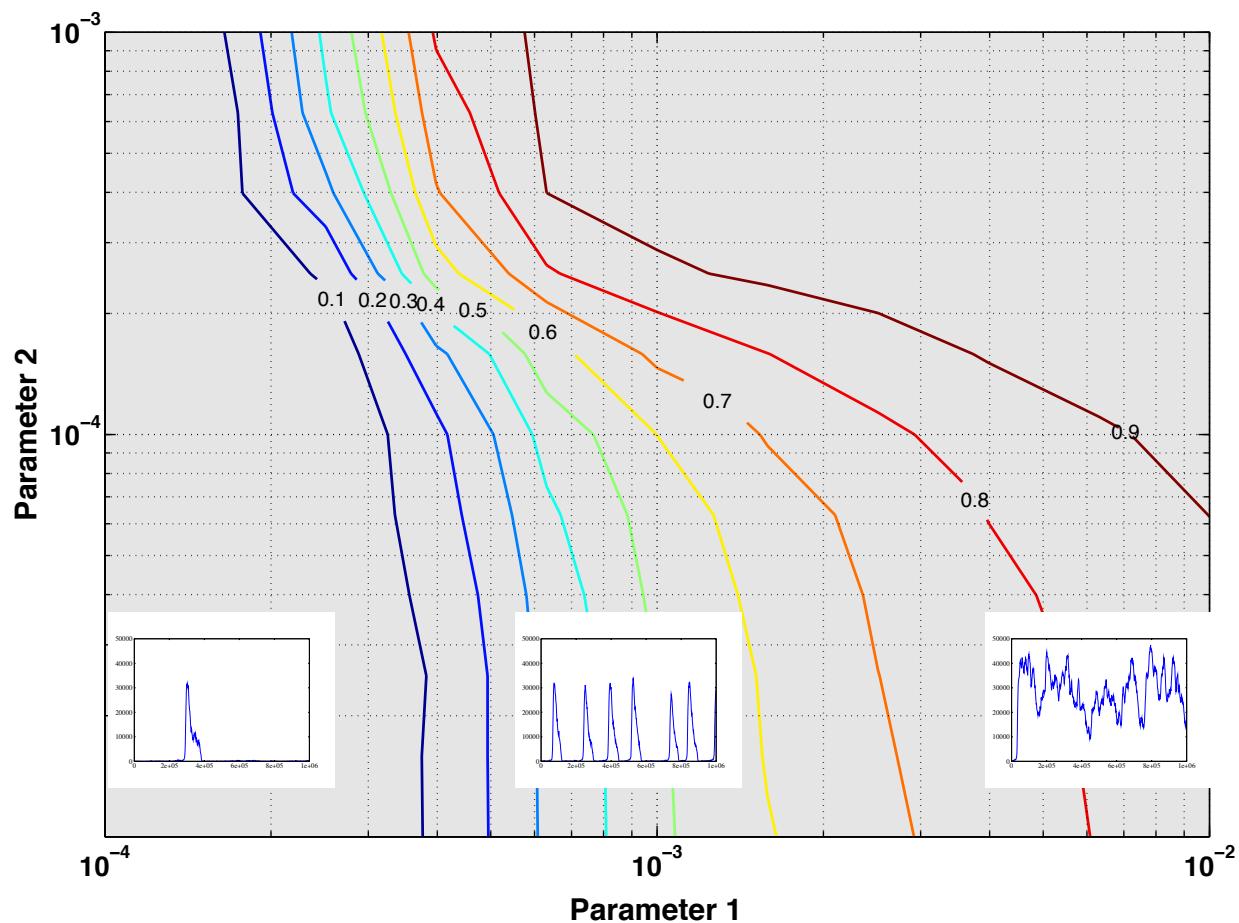
Various dynamical regimes revealed by exploring parameter space

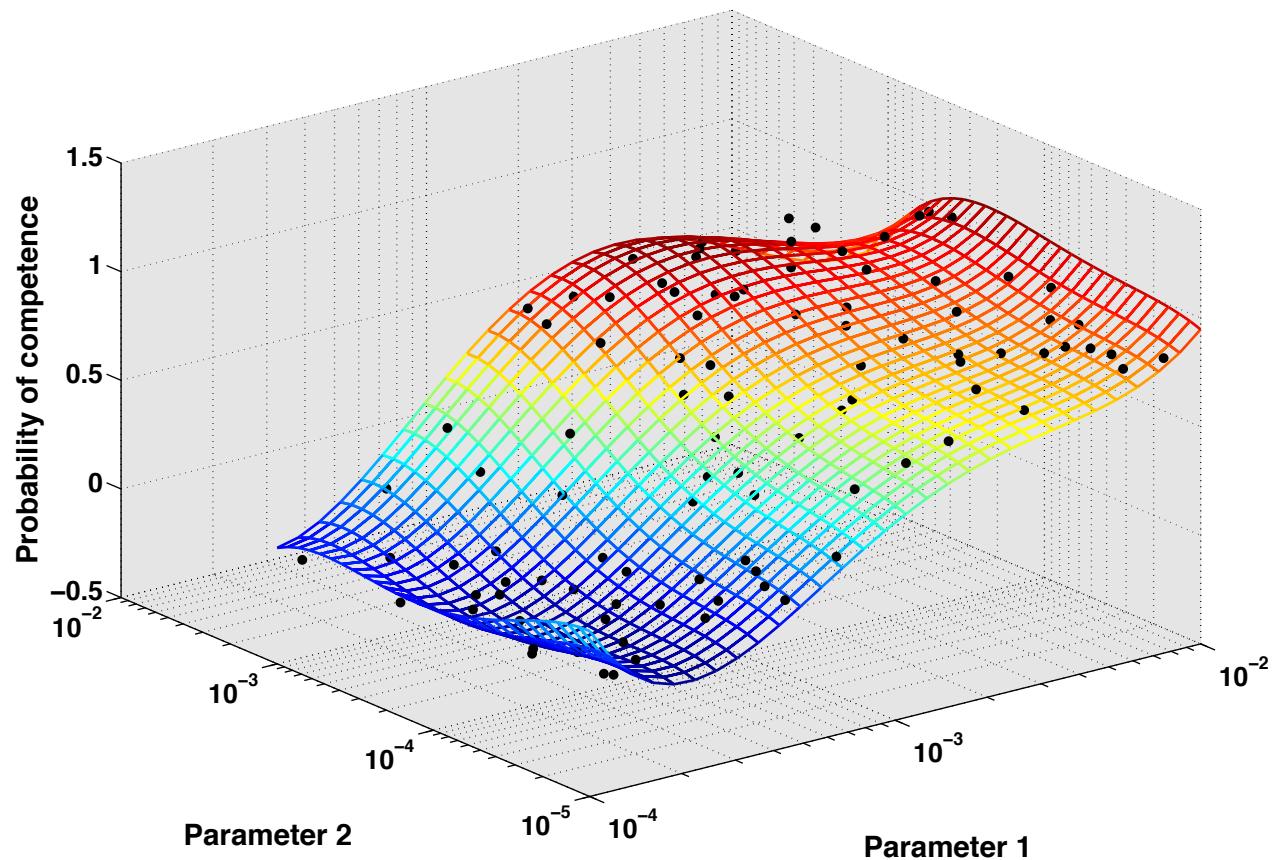
Contours of probability of competence P_c



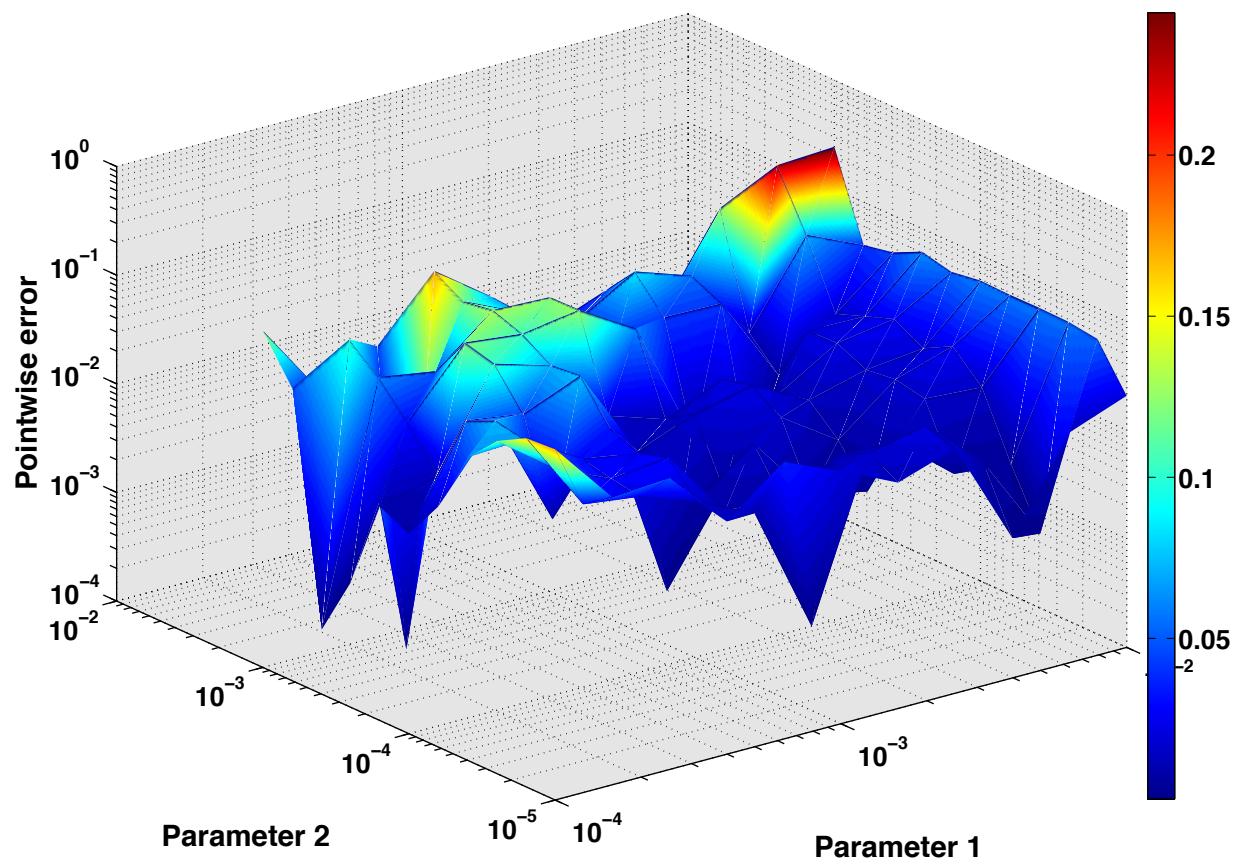
Various dynamical regimes revealed by exploring parameter space

Contours of probability of competence P_c



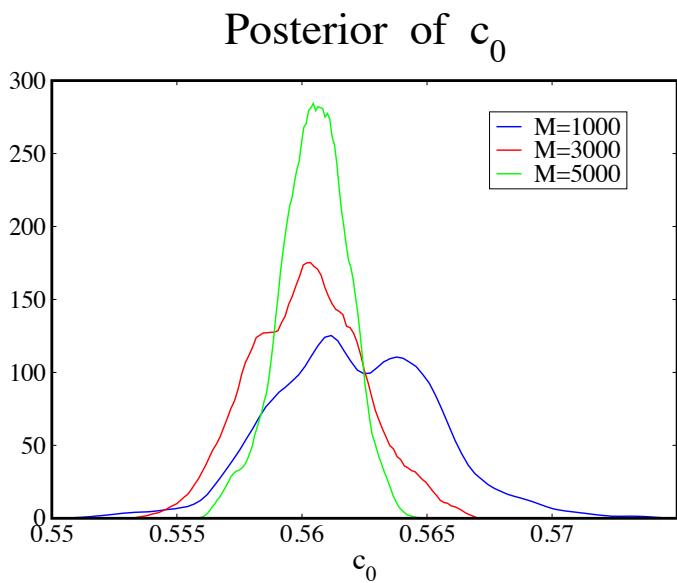


Pointwise error in MAP estimate of 4th order PC



Convergence both in posterior width and order

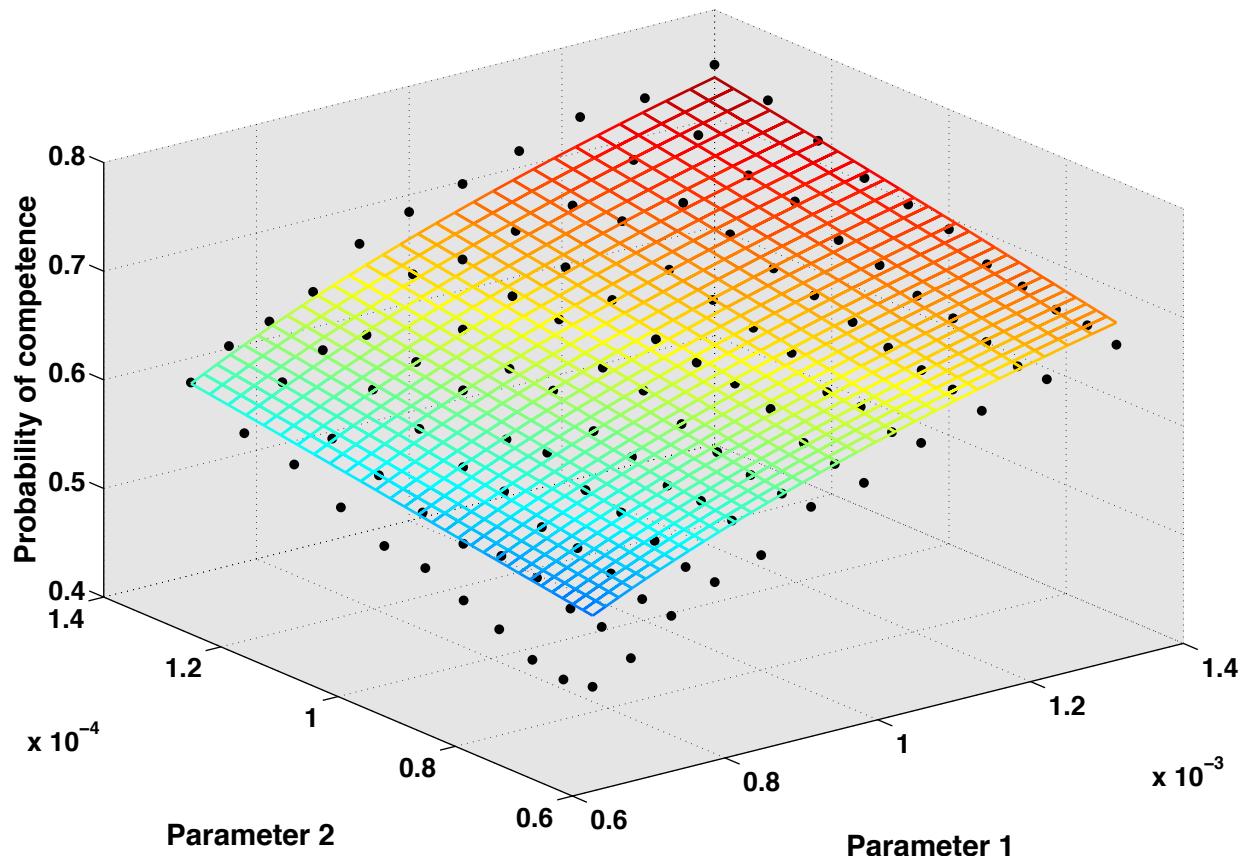
- With more input parameter samples, posterior narrows around the true value
- Convergence in PC order is established



Relative L_2 error



2-nd order response surface in 10-d case: 2-d slice



High Dimensional Model Representation (HDMR)

breaks the function into group-wise contributions of input variables

$$f(\boldsymbol{\lambda}) = f(\lambda_1, \dots, \lambda_d) = f_0 + \sum_i f_i(\lambda_i) + \sum_{i < j} f_{ij}(\lambda_i, \lambda_j) + \sum_{i < j < k} f_{ijk}(\lambda_i, \lambda_j, \lambda_k) + \dots$$

Component functions are found by

$$f_0 = \int_{R^d} f(\boldsymbol{\lambda}) d\boldsymbol{\lambda}, \quad f_i(\lambda_i) = \int_{R^{d-1}} f(\lambda_i, \boldsymbol{\lambda}_{\bar{i}}) d\boldsymbol{\lambda}_{\bar{i}} - f_0$$
$$f_{ij}(\lambda_i, \lambda_j) = \int_{R^{d-2}} f(\lambda_i, \lambda_j, \boldsymbol{\lambda}_{\bar{ij}}) d\boldsymbol{\lambda}_{\bar{ij}} - f_i(\lambda_i) - f_j(\lambda_j) - f_0$$

- Component function $f_{i_1 \dots i_s}(\lambda_{i_1}, \dots, \lambda_{i_s})$ is found by a $(d - s)$ -dimensional integral. Still too high-dimensional.
- Otherwise called ANOVA decomposition (analysis of variance)
- Exact in the limit, but not unique.

High Dimensional Model Representation (HDMR)

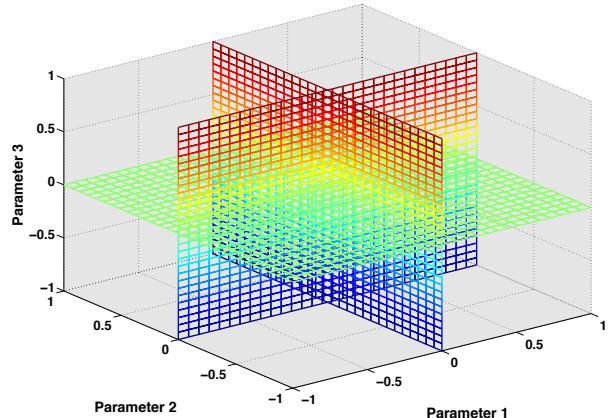
breaks the function into group-wise contributions of input variables

$$f(\boldsymbol{\lambda}) = f(\lambda_1, \dots, \lambda_d) = f_0 + \sum_i f_i(\lambda_i) + \sum_{i < j} f_{ij}(\lambda_i, \lambda_j) + \sum_{i < j < k} f_{ijk}(\lambda_i, \lambda_j, \lambda_k) + \dots$$

Component functions are found by

$$f_0 = \int_{R^d} f(\boldsymbol{\lambda}) d\boldsymbol{\lambda}, \quad f_i(\lambda_i) = \int_{R^{d-1}} f(\lambda_i, \boldsymbol{\lambda}_{\bar{i}}) d\boldsymbol{\lambda}_{\bar{i}} - f_0$$
$$f_{ij}(\lambda_i, \lambda_j) = \int_{R^{d-2}} f(\lambda_i, \lambda_j, \boldsymbol{\lambda}_{\bar{ij}}) d\boldsymbol{\lambda}_{\bar{ij}} - f_i(\lambda_i) - f_j(\lambda_j) - f_0$$

- Two major variants:
 - Cut-HDMR
 - RS-HDMR



High Dimensional Model Representation (HDMR)

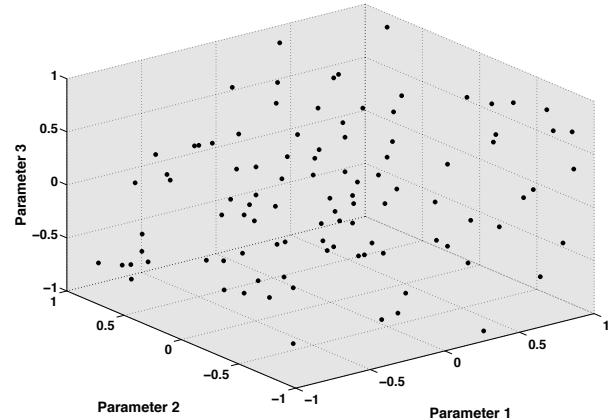
breaks the function into group-wise contributions of input variables

$$f(\boldsymbol{\lambda}) = f(\lambda_1, \dots, \lambda_d) = f_0 + \sum_i f_i(\lambda_i) + \sum_{i < j} f_{ij}(\lambda_i, \lambda_j) + \sum_{i < j < k} f_{ijk}(\lambda_i, \lambda_j, \lambda_k) + \dots$$

Component functions are found by

$$f_0 = \int_{R^d} f(\boldsymbol{\lambda}) d\boldsymbol{\lambda}, \quad f_i(\lambda_i) = \int_{R^{d-1}} f(\lambda_i, \boldsymbol{\lambda}_{\bar{i}}) d\boldsymbol{\lambda}_{\bar{i}} - f_0$$
$$f_{ij}(\lambda_i, \lambda_j) = \int_{R^{d-2}} f(\lambda_i, \lambda_j, \boldsymbol{\lambda}_{\bar{ij}}) d\boldsymbol{\lambda}_{\bar{ij}} - f_i(\lambda_i) - f_j(\lambda_j) - f_0$$

- Two major variants:
 - Cut-HDMR
 - RS-HDMR

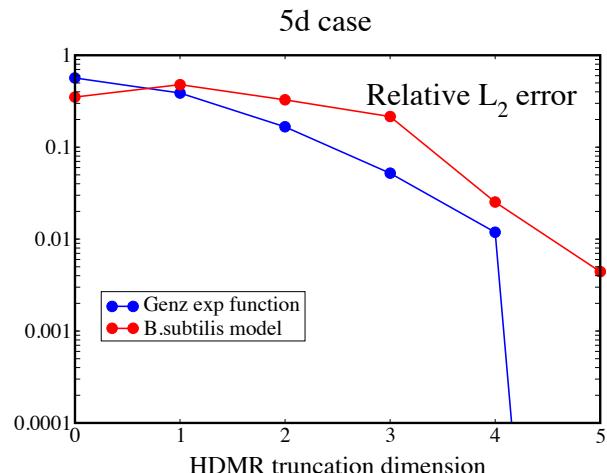
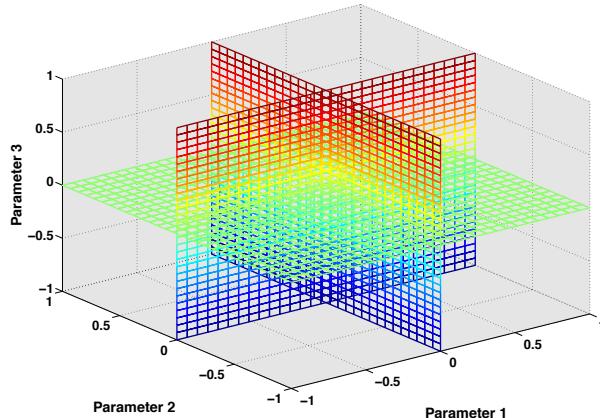


cut-HDMR disregards corners in the parameter space
and does not guarantee accuracy in general

- Component functions are

$$f_0 = f(\boldsymbol{\lambda}^a), \quad f_i(\lambda_i) = f(\lambda_i, \boldsymbol{\lambda}_{\bar{i}}^a) - f_0 \\ f_{ij}(\lambda_i, \lambda_j) = f(\lambda_i, \lambda_j, \boldsymbol{\lambda}_{\bar{i}\bar{j}}^a) - f_i(\lambda_i) - f_j(\lambda_j) - f_0$$

- Relies on values at lower-dimensional hyperplanes
- Depends on the anchor point $\boldsymbol{\lambda}^a$
- Does not account for ‘corners’

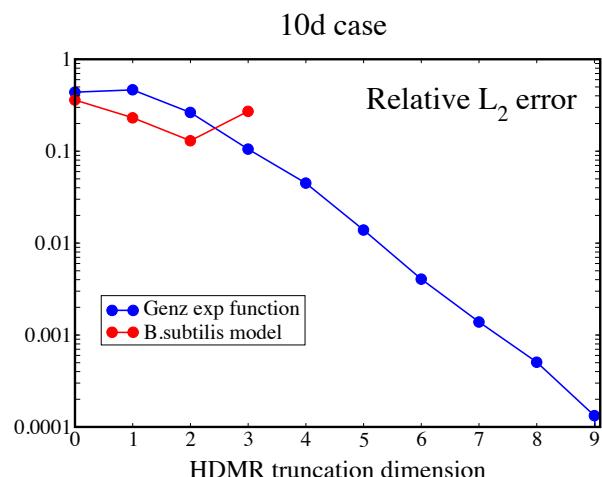
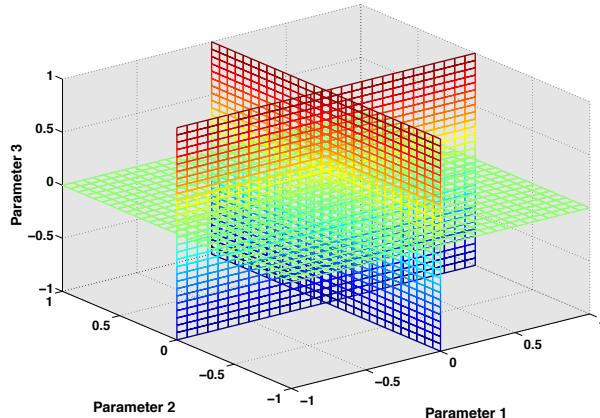


cut-HDMR disregards corners in the parameter space
and does not guarantee accuracy in general

- Component functions are

$$f_0 = f(\boldsymbol{\lambda}^a), \quad f_i(\lambda_i) = f(\lambda_i, \boldsymbol{\lambda}_{\bar{i}}^a) - f_0 \\ f_{ij}(\lambda_i, \lambda_j) = f(\lambda_i, \lambda_j, \boldsymbol{\lambda}_{\bar{i}\bar{j}}^a) - f_i(\lambda_i) - f_j(\lambda_j) - f_0$$

- Relies on values at lower-dimensional hyperplanes
- Depends on the anchor point $\boldsymbol{\lambda}^a$
- Does not account for ‘corners’



Random Sampling (RS) HDMR in principle equivalent to PC expansion with Monte-Carlo integration

$$f(\boldsymbol{\lambda}) = f(\lambda_1, \dots, \lambda_d) = f_0 + \sum_i f_i(\lambda_i) + \sum_{i < j} f_{ij}(\lambda_i, \lambda_j) + \sum_{i < j < k} f_{ijk}(\lambda_i, \lambda_j, \lambda_k) + \dots$$

- Component functions are found by

$$\begin{aligned} f_0 &= \int_{R^d} f(\boldsymbol{\lambda}) d\boldsymbol{\lambda}, & f_i(\lambda_i) &= \int_{R^{d-1}} f(\lambda_i, \boldsymbol{\lambda}_{\bar{i}}) d\boldsymbol{\lambda}_{\bar{i}} - f_0 \\ f_{ij}(\lambda_i, \lambda_j) &= \int_{R^{d-2}} f(\lambda_i, \lambda_j, \boldsymbol{\lambda}_{\bar{ij}}) d\boldsymbol{\lambda}_{\bar{ij}} - f_i(\lambda_i) - f_j(\lambda_j) - f_0 \end{aligned}$$

- MC integrals still too expensive (new random samples needed for each hyperplane)
- Represent component functions with a polynomial expansion and use same set of samples
- Equivalent to Monte-Carlo PC with reordered multiindices!
 - PC (total order): $f(\xi_1, \xi_2) = 1 + [\xi_1 + \xi_2] + [(\xi_1^2 - 1) + \xi_1 \xi_2 + (\xi_2^2 - 1)] + \dots$
 - RS-HDMR: $f(\xi_1, \xi_2) = 1 + [\xi_1 + (\xi_1^2 - 1)] + [\xi_2 + (\xi_2^2 - 1)] + \xi_1 \xi_2 \dots$
- In future: employ Bayesian inference on component functions.

Summary

- Polynomial Chaos expansions represent effects of uncertainties of input parameters to output statistical properties
 - Sensitivity analysis
 - Uncertainty quantification
 - Response surface construction
- Noise in function evaluations hampers quadrature methods
 - Sparse integration of noisy functions useless in high-d !
- HDMR constructions do not always guarantee accuracy with small computational effort
 - Generally still require high-d integrals
 - cut-HDMR overcomes this requirement but is not accurate enough
- Bayesian inference well-suited to handle noisy data

Acknowledgements

- Youssef Marzouk (MIT)
- Cosmin Safta (SNL)
- DOE Office of Science, Advanced Scientific Computing Research, Applied Mathematics.

Thank You!

Sandia National Laboratories is a multi-program laboratory operated by Sandia Corporation, a wholly owned subsidiary of Lockheed Martin Corporation, for the U.S. Department of Energy's National Nuclear Security Administration under contract DE-AC04-94AL85000.

Literature

- N. Wiener, "Homogeneous Chaos", *American Journal of Mathematics*, 60(4), pp. 897-936, 1938.
 - R. Ghanem and P. Spanos, "Stochastic Finite Elements: a Spectral Approach", Springer, 1991.
 - D. Xiu and G.E. Karniadakis, "The Wiener-Askey Polynomial Chaos for Stochastic Differential Equations", *SIAM J. Sci. Comput.*, 24(2), pp. 619-644, 2002.
 - O. Le Maître and O. Knio, "Spectral Methods for Uncertainty Quantification with Applications to Computational Fluid Dynamics", Springer, 2010.
-
- K. Sargsyan, B. Debusschere, H. Najm and O. Le Maître,
"Spectral representation and reduced order modeling of the dynamics of stochastic reaction networks via adaptive data partitioning".
SIAM Journal on Scientific Computing, 31:6, 2010.
 - K. Sargsyan, B. Debusschere, H. Najm and Y. Marzouk,
"Bayesian inference of spectral expansions for predictability assessment in stochastic reaction networks".
Journal of Computational and Theoretical Nanoscience, 6:10, 2009.