# Comparing One-stage cluster sampling to SRS

Week 6 (5.2)

Stat 260, St. Clair

# When is a one-stage cluster sample more precise than SRS?

When does

$$SE(\hat{t}_{\, cluster}) \stackrel{???}{<} SE(\hat{t}_{\, SRS})$$

answer: It depends on the measurement's Analysis of Variance (ANOVA)

#### SRS = 2.2 SRS = 2.2 St for cluster est is same street for same street Lohr Examples 5.6: design effect > svymean(~score, alg\_design, deff = TRUE) DEff mean score 62.5686 1.4916 2.245 > boxplot(score ~ class, data = algebra) 23 37 38 106 108 39 62 class

#### Population ANOVA

Let  $y_{ij}$  be your measurement of unit j in cluster i

ANOVA breaks the **total** sum of squares of *y* into **between cluster** and **within cluster** variation:

$$SST = SSB + SSW$$

For now, assume that cluster sizes are equal

$$M_i = M ext{ for all clusters } i = 1, \dots, N$$

# Population ANOVA

Source	df	Sum of Squares	Mean Square				
Between	N-1	$SSB = \sum_{i=1}^{N} M({ar{y}}_{i\mathcal{U}} - {ar{y}}_{\mathcal{U}})^2$	$MSB = rac{SSB}{N-1}$				
Within	N(M-1)	$SSW = \sum_{i=1}^N (M-1)S_i^2$	$MSW = rac{SSW}{N(M-1)}$				
total		$SSTot = \sum_{i=1}^{N} \sum_{j=1}^{M} (y_{ij} - ar{y}_{\mathcal{U}})^2$					
Between: Yn = overall near per SSU (pap.)  Yir = cluster i maan par SSU (pap.)							
Dithin: Si = pop. variance for cluster i							

#### Variance: SRS

**Equal cluster sizes:** We've sampled nM **observation units** (SSU) out of  $M_0=NM$  possible units.

For a SRS of  $\underline{nM}$  **observation units**, we can write the variance,  $SE^2$ , of  $\hat{t}_{SRS}$  as

$$Var(\hat{t}_{SRS}) = (NM)^2 \left(1 - rac{nM}{NM}
ight) rac{S^2}{nM}$$

where S is the SD of the measurements in the population.

#### Variance: One-stage cluster sample

**Equal cluster sizes:** Under this assumption the variance of  $\hat{t}_{unb}$  is equal to

$$Var(\hat{t}_{unb}) = N^{2} \left(1 - \frac{n}{N}\right) \frac{M \times MSB}{n}$$

$$Var(\hat{t}_{unb}) = N^{2} \left(1 - \frac{n}{N}\right) \frac{S_{\pm}^{2}}{N} \quad \text{variance of cluster total}$$

$$S_{\pm}^{2} = \frac{1}{N-1} \sum_{i=1}^{N} \left(\pm_{i} - \pm_{i}\right)^{2} = \frac{1}{N-1} \sum_{i=1}^{N} \left(H_{yin} - My_{in}\right)^{2}$$

$$t_{i} = \sum_{i=1}^{N} y_{ij} \times M = My_{in} \left[\frac{\sum_{i=1}^{N} y_{ij}}{N} \times NM - My_{in}\right]$$

$$S_{\pm}^{2} = \frac{M^{2}}{N-1} \sum_{i=1}^{N} \left(y_{in} - y_{in}\right)^{2} = \frac{M^{2}SB}{N-1} = M \times MSB$$

$$N \times MSB$$

#### Variance: SRS vs. Stratified sample

**Equal cluster sizes:** Under this assumption, the design effect for a one-stage cluster sample total estimate is

$$DEff(\hat{ar{y}}_{unb}) = DEff(\hat{t}_{unb}) = rac{Var(\hat{t}_{unb})}{Var(\hat{t}_{SRS})} = rac{MSB}{S^2}$$

#### Variance: SRS vs. Stratified sample

Cluster sampling is more precise than an equal sized SRS when

$$MSB < S^2$$

- $\Rightarrow$  between cluster variation is small
- $\Rightarrow$  measurements are heterogenous within clusters

Opposite result from strat. Sampling

-> strat. good when measurement within

strata are homogeneous.

#### Measuring homogeneity within clusters

• Intraclass correlation coefficient: for equal sized clusters

$$ICC = 1 - rac{M}{M-1} rac{SSW}{SSTot} \quad ext{where} \ - rac{1}{M-1} \leq ICC \leq 1$$

• Adjusted R-squared: can be used for unequal cluster sizes

$$R_a^2=1-rac{MSW}{S^2} \quad ext{where } 1-rac{NM-1}{N(M-1)} \leq R_a^2 \leq 1$$

- · For both:
  - values near 1 indicate <u>homogeneous</u> (similar) responses <u>within</u> clusters
     SSW ≈ MSW ≈ O
  - values near 0 indicate heterogeneous (dissimilar) responses within clusters

#### Design effect revisted

**Equal cluster sizes:** Under this assumption the vertex of  $\hat{t}_{unb}$  is equal to

$$egin{aligned} DEff(\hat{t}_{unb}) &= rac{MSB}{S^2} \ &= rac{MN-1}{M(N-1)}(1+(M-1)ICC) \ &= 1 + rac{N(M-1)}{N-1}R_a^2 \end{aligned}$$

### Design effect revisted

What is the design effect if

- N is big
- M = 11
- $R_a^2 = 0.5$

Deff(cluster) = 
$$1 + \frac{N(m-1)}{N-1}$$
  $P_a = \frac{1 + \frac{N}{N-1}(10)(\frac{1}{2})}{N - 1}$   
N big  $\approx 1 + 10(\frac{1}{2}) = 6 = \frac{Var(cluster)}{Var(SRS)}$   
 $Var(SRS) = \frac{Var(cluster)}{6} \Rightarrow n = 4 \text{ clusters Sampled}$   
 $n \times 6 = 4 \text{ clusters needed to sample to have Same St}$   
as an SRS of  $n \cdot 11$  observation unit  
 $equal SE$ :  $SRS n \cdot 11$  anite  $One - Stege n \cdot 6 \cdot 11$  anits

#### Big picture

- One-stage cluster sampling is "good" for precision if SSU within clusters have very **heterogeneous** responses
  - true whether or not cluster sizes are equal
- But often SSU within clusters have very **homogeneous** responses
  - clusters contain "similar" observation units
  - clusters defined for **cost-saving** reasons, not for precision

#### Post-Hoc comparison

Q: How do we compute the design effect with a **sample of data** from any one-stage cluster sample?

**1. (Any cluster sizes)** Use sampling weights to estimate  $Var(\hat{t}_{srs})$ 

• This is what the survey package when you use deff=TRUE

#### Post-Hoc comparison

Q: How do we compute the design effect with a **sample of data** from any one-stage cluster sample?

2. Equal cluster sizes: Estimate population sum of square values from sample mean square values msw and msb:

$$\widehat{SSW} = N(M-1)msw \quad \widehat{SSB} = (N-1)msb$$

The estimated design effect is

$$\widehat{DEff}(\hat{t}_{unb}) = \frac{\widehat{MSB}}{\hat{S}^2} = \frac{msb}{(\widehat{SSW} + \widehat{SSB})/(NM - 1)}$$

$$\widehat{S}^2 = \frac{\widehat{SSW} + \widehat{SSB}}{NM - 1}$$

$$NM - 1$$

# Estimating ICC and $R_a^2$

$$\widehat{SSW} = N(M-1)msw, \quad \widehat{SSB} = (N-1)msb, \quad \widehat{SST} = \widehat{SSB} + \widehat{SSW}$$

• Estimated *ICC* is

$$ICC = 1 - rac{M}{M-1}rac{\widehat{SSW}}{\widehat{SST}}$$

• Estimated  $R_a^2$  is

$${\hat R}_a^2=1-rac{msw}{{\widehat S}^2}$$

#### Example - GPA

$$N=100,\, n=5,\, M_i=4,\, M_0=400$$

	Suite 1	Suite 2	Suite 3	Suite 4	Suite 5
1	3.08	2.36	2.00	3.00	2.68
2	2.60	3.04	2.56	2.88	1.92
3	3.44	3.28	2.52	3.44	3.28
4	3.04	2.68	1.88	3.64	3.20
total	12.16	11.36	8.96	12.96	11.08

```
> dorm <- read.csv("http://math.carleton.edu/kstclair/data/Dorm_Cluster.
> dplyr::glimpse(dorm)
Rows: 20
Columns: 2
$ gpa <dbl> 3.08, 2.60, 3.44, 3.04, 2.36, 3.04, 3.28, 2.68, 2.00, 2
$ room <int> 1, 1, 1, 1, 2, 2, 2, 2, 3, 3, 3, 3, 4, 4, 4, 4, 5, 5, 5
```

#### Example - GPA

```
N=100, n=5, M_i=4, M_0=400
```

What is the design effect, ICC and  $R_a^2$  for estimating mean GPA?

```
> dorm_lm <- lm(gpa ~ factor(room), data = dorm)
> anova(dorm_lm)
Analysis of Variance Table

Response: gpa

Response: gpa

Df Sum Sq Mean Sq F value Pr(>F)

Custer factor(room) /4 2.7557 0.56392

With= Residuals 15 2.7756 0.18504

Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

msb = .564 msw = .185 pop. ss ssB = (N-1)msb = (100-1)(.569) = 55.83ssw = N(M-1)msw = (00(4-1)(.185) = 55.51

$$\hat{S}^{2} = \frac{S\hat{S}W + S\hat{S}B}{NM-1} = \frac{55.51 + 55.83}{100(4)-1} \approx ,776$$

$$DEPP = \frac{MSb}{\hat{S}^{2}} = \frac{.564}{.279} \approx 2.02$$

$$P_{a}^{2} = 1 - \frac{.185}{.279} \approx .34$$

$$ECC = 1 - \frac{M}{M-1} = \frac{.5Sw}{.279} = 1 - \frac{.4}{.55.51 + 55.83}$$

$$\approx .34$$
Survey package DEPP = 2.12

#### Lohr Examples 5.6: design effect of 2.245

What if cluster sizes are not equal?

```
> alg_lm <- lm(score ~ factor(class), data = algebra)</pre>
> anova(alg_lm)
Analysis of Variance Table
Response: score
              Df Sum Sq Mean Sq F value Pr(>F)
factor(class) 11 7086 644.14 2.1184 0.01915 *
Residuals 287 87270
                        304.08
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
> msb <- 644.14
> msw < -304.08
> msb/var(algebra$score) # rough DEff guess
[1] 2.03437
                   bies est. of 5.
```

#### Lohr Examples 5.6: design effect of 2.245

What if cluster sizes are not equal?