

Simple AR models: (Regression with lagged variables.)

AR(1) model:

$$r_t = \phi_1 r_{t-1} + a_t,$$

where ϕ_1 is real number, which is referred to as parameters (to be estimated from the data in an application).

For example,

$$r_t = 0.2r_{t-1} + a_t.$$

$$(r_{t-1} = 0.2r_{t-2} + a_{t-1} \text{ and } r_{t+1} = 0.2r_t + a_{t+1})$$

$$1. \quad r_t = a_t + \phi_1 a_{t-1} + \cdots + \phi_1^m a_{t-m} + \phi_1^{m+1} r_{t-m-1}.$$

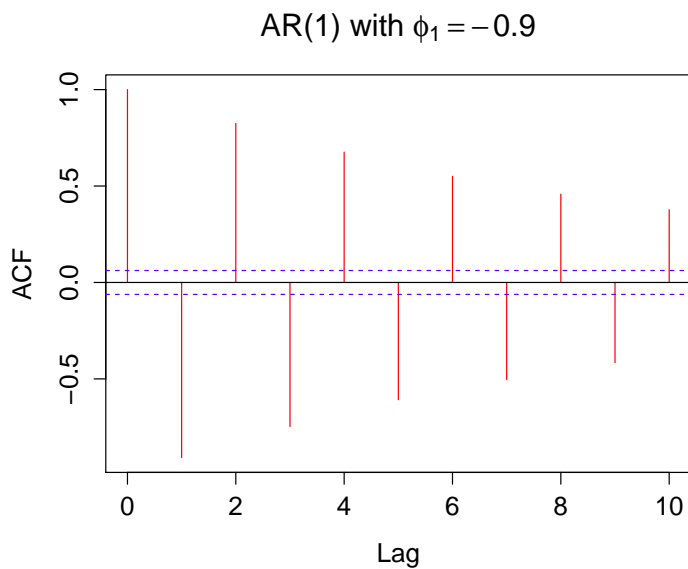
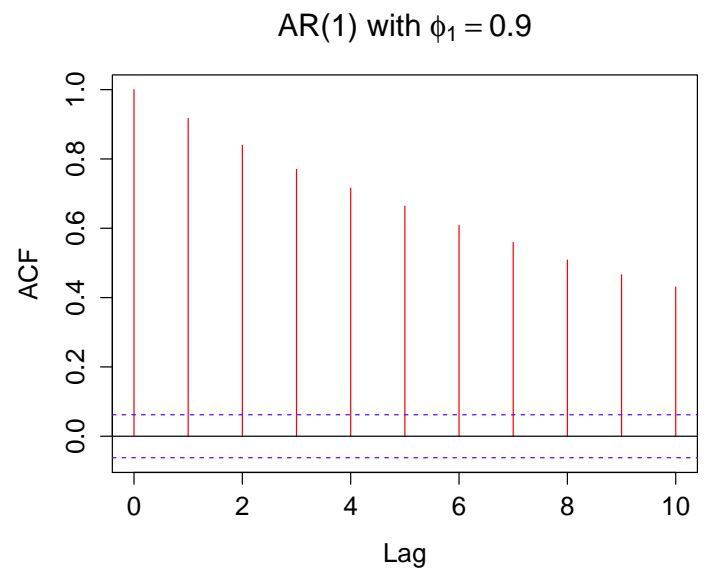
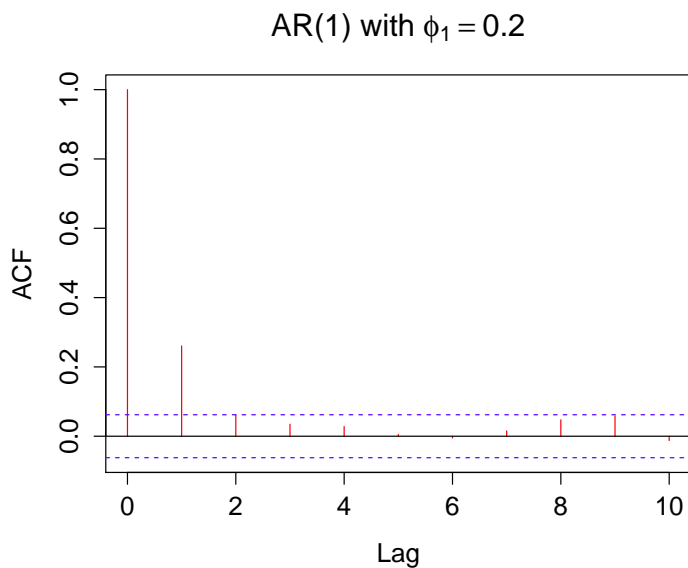
2. Stationarity: necessary and sufficient condition $|\phi_1| < 1$. Why?

$$r_t = a_t + \sum_{i=1}^{\infty} \phi_1^i a_{t-i}.$$

3. Mean: $E(r_t) = 0$.

4. Variance: $Var(r_t) = \frac{\sigma_a^2}{1-\phi^2}$.

5. Autocorrelations: $\rho_1 = \phi_1$, $\rho_2 = \phi_1^2$, etc. In general, $\rho_k = \phi_1^k$ and ACF ρ_k decays exponentially as k increases,



ACF of AR(1) model.

- When $|\phi_1| < 1$, the series $\{r_t\}$ exhibits mean-reverting behavior:

$$\text{Let } \Delta r_t = r_t - r_{t-1} = (\phi_1 - 1)r_{t-1} + a_t,$$

$$E(\Delta r_t | F_{t-1}) = (\phi_1 - 1)r_{t-1} < 0, \text{ if } r_{t-1} > 0;$$

$$E(\Delta r_t | F_{t-1}) = (\phi_1 - 1)r_{t-1} > 0, \text{ if } r_{t-1} < 0.$$

- When $\phi_1 = 1$, AR(1) model can be rewritten as the RW process:

$$r_t = a_t + a_{t-1} + \cdots + a_1 + r_0.$$

- When $|\phi_1| > 1$, e.g. $\phi = 3$,

$$r_t = a_t + 3a_{t-1} + 3^2a_{t-2} + \cdots + 3^{t-1}a_1 + 3^tr_0.$$

r_t is called the explosive process, in the sense that r_t diverges to ∞ .

AR(1) model with a drift:

$$r_t = \phi_0 + \phi_1 r_{t-1} + a_t,$$

where ϕ_0 and ϕ_1 are real numbers.

$$Er_t = \frac{\phi_0}{1 - \phi_1},$$

$Var(r_t)$ and ρ_k are not changed.

6. Forecast (minimum squared error):

(a) 1-step ahead forecast at time n , the forecast origin:

$$\hat{r}_n(1) = \phi_0 + \phi_1 \hat{r}_n.$$

(b) 1-step ahead forecast error:

$$e_n(1) = r_{n+1} - \hat{r}_n(1) = a_{n+1}.$$

Thus, a_{n+1} is the un-predictable part of r_{n+1} . It is the shock at time $n + 1$!

(c) Variance of 1-step ahead forecast error:

$$Var[e_n(1)] = Var(a_{n+1}) = \sigma_a^2.$$

(d) 2-step ahead forecast:

$$\hat{r}_n(2) = \phi_0 + \phi_1 \hat{r}_n(1).$$

(e) 2-step ahead forecast error:

$$e_n(2) = r_{n+2} - \hat{r}_n(2) = a_{n+2} + \phi_1 a_{n+1}.$$

(f) Variance of 2-step ahead forecast error:

$$\text{Var}[e_n(2)] = (1 + \phi_1^2) \sigma_a^2.$$

which is greater than or equal to $\text{Var}[e_n(1)]$, implying that uncertainty in forecasts increases as the number of steps increases.

(g) Behavior of l -step ahead forecasts:

Forecasting error:

$$e_n(l) = r_{n+l} - \hat{r}_n(l) = \sum_{j=0}^{l-1} \phi_1^j a_{n+l-j}.$$

Forecasting variance:

$$\text{Var}[e_n(l)] = \sigma_a^2 \sum_{j=0}^{l-1} \phi_1^{2j}.$$

Forecast interval (limit) (FI):

$$\left[\hat{r}_n(l) - N_{\frac{\alpha}{2}} \sigma_a \sqrt{\sum_{j=0}^{l-1} \phi_1^{2j}}, \hat{r}_n(l) + N_{\frac{\alpha}{2}} \sigma_a \sqrt{\sum_{j=0}^{l-1} \phi_1^{2j}} \right]$$

where $N_{\frac{\alpha}{2}}$ is the $\alpha/2$ -quantile of the standard normal distribution, i.e., $P(N > N_{\frac{\alpha}{2}}) = \alpha/2$.

When $\alpha = 0.05$, $N_{\frac{\alpha}{2}} = 1.96$.

7. A compact form: $(1 - \phi_1 B)r_t = \phi_0 + a_t$.

AR(2) model:

$$r_t = \phi_1 r_{t-1} + \phi_2 r_{t-2} + a_t,$$

$$\text{or } (1 - \phi_1 B - \phi_2 B^2)r_t = a_t.$$

1. Stationarity condition:

all the roots of $(1 - \phi_1 x - \phi_2 x^2) = 0$ lie outside the unit circle.

Decompose $1 - \phi_1 z - \phi_2 z^2 = (1 - \alpha_1 z)(1 - \alpha_2 z)$.

Then $|\alpha_1| < 1$ and $|\alpha_2| < 1$.

$$(1 - \alpha_1 B)(1 - \alpha_2 B)r_t = a_t.$$

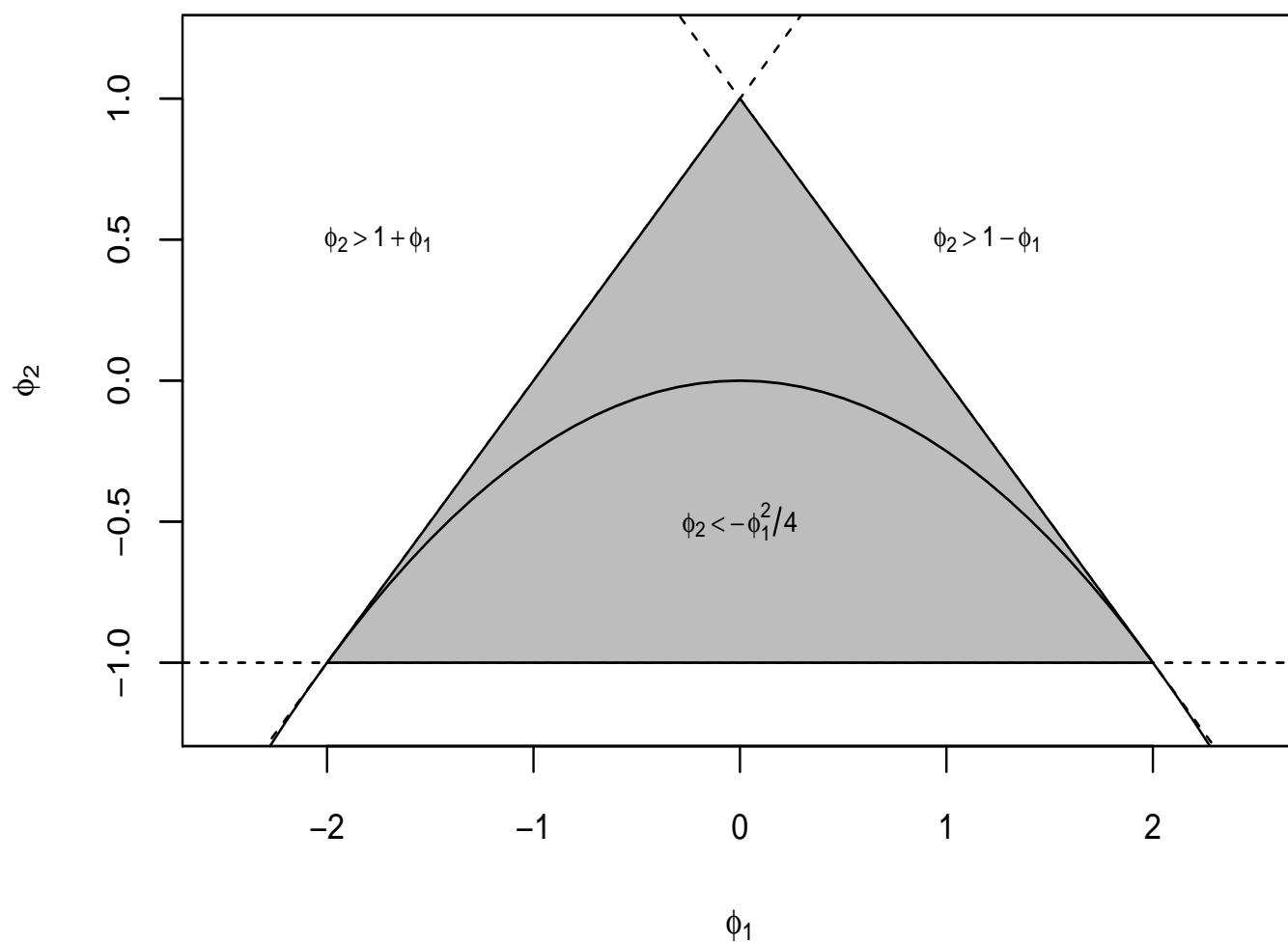
Let $u_t = (1 - \alpha_2 B)r_t$. Then

$$u_t = \alpha_1 u_{t-1} + a_t = a_t + \sum_{i=1}^{\infty} \alpha_1^i a_{t-i}.$$

$$r_t = \alpha_2 r_{t-1} + u_t = u_t + \sum_{j=1}^{\infty} \alpha_2^j u_{t-j}.$$

Stationarity condition is equivalent to

$$\begin{cases} \phi_2 + \phi_1 < 1, \\ \phi_2 - \phi_1 < 1, \\ -1 < \phi_2 < 1. \end{cases}$$



Stationary region for AR(2) model.

2. Mean: $E(r_t) = 0$

3. ACF: $\rho_0 = 1$, $\rho_1 = \frac{\phi_1}{1-\phi_2}$,

$$\rho_l = \phi_1 \rho_{l-1} + \phi_2 \rho_{l-2}, \quad l \geq 2.$$

The solutions of $\phi(z) = 1 - \phi_1 z - \phi_2 z^2 = 0$:

$$\omega_{1,2} = \frac{\phi_1 \pm \sqrt{\phi_1^2 + 4\phi_2}}{-2\phi_2}.$$

The ACF is a mixture of two exponential decays as the general solution of

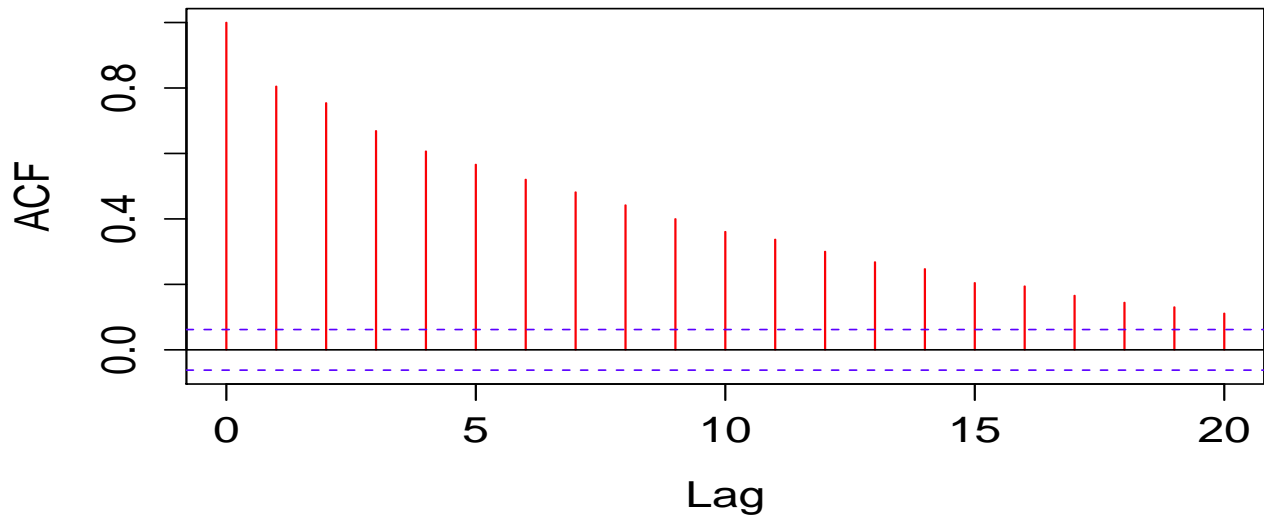
$$\rho_l = \frac{c_1}{\omega_1^l} + \frac{c_2}{\omega_2^l},$$

where c_1 and c_2 are unknown parameters depending on the initial conditions.

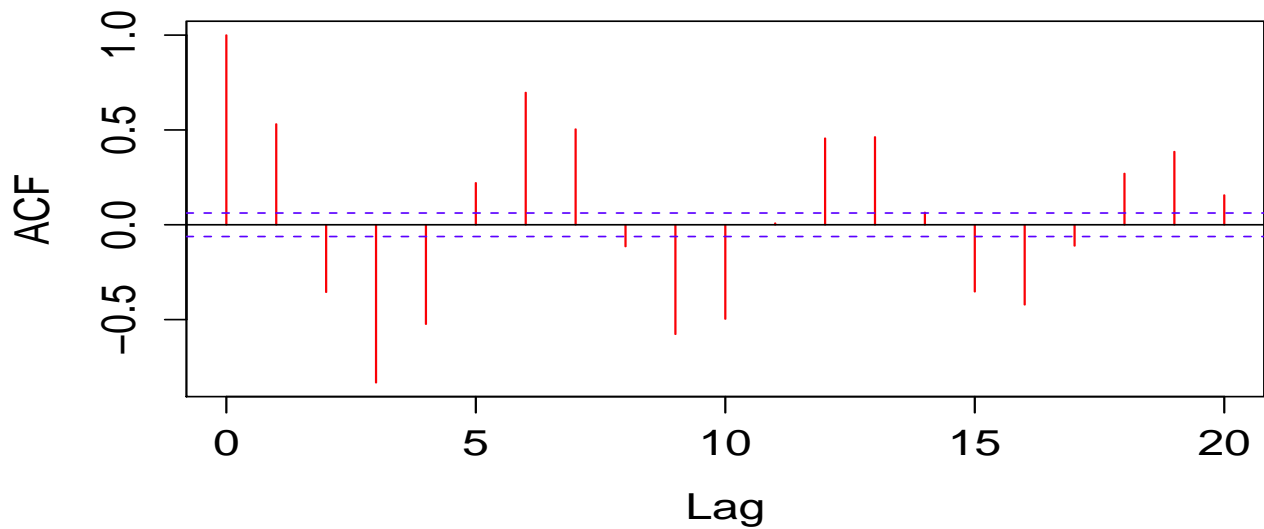
(a). If both roots ($\omega_{1,2}$) are real, then the ACF of AR(2) is a mixture of two exponential decay.

(b). If the roots are a complex conjugate pair, i.e. $\phi_1^2 + 4\phi_2 < 0$, then ACF would show a picture of damping sine and cosine waves.

AR(2) with $\phi_1 = 0.6$ and $\phi_2 = 0.3$



AR(2) with $\phi_1 = 1$ and $\phi_2 = -0.9$



The ACF of AR(2) model.

4. Stochastic business cycle: if $\phi_1^2 + 4\phi_2 < 0$, then r_t shows characteristics of business cycles with average length

$$k = \frac{2\pi}{\cos^{-1}[\phi_1/(2\sqrt{-\phi_2})]},$$

where the cosine inverse is stated in radian. If we denote the solutions of the polynomial as $a \pm bi$, where $i = \sqrt{-1}$, then we have $\phi_1 = 2a$ and $\phi_2 = -(a^2 + b^2)$ so that

$$k = \frac{2\pi}{\cos^{-1}(a/\sqrt{a^2 + b^2})}.$$

Example: For a fitted AR(2) model to one monthly data series denoted by

$$r_t - 1.3512r_{t-1} + 0.4612r_{t-2} = a_t.$$

Clearly, $\phi_1^2 - 4\phi_2 = 1.3512^2 - 4 \times 0.4612 = -0.0191 < 0$, which implies the existence of stochastic cycles. The average length of the stochastic cycles for this AR(2) model is approximately

$$k = \frac{2\pi}{\cos^{-1}[\phi_1/(2\sqrt{-\phi_2})]} = 61.71 \text{ months,}$$

which is about 5 years.

5. Forecasts: Similar to AR(1) models

AR(2) model with a drift:

$$r_t = \phi_0 + \phi_1 r_{t-1} + \phi_2 r_{t-2} + a_t,$$

$$E(r_t) = \frac{\phi_0}{1 - \phi_1 - \phi_2},$$

others are unchanged.

Building an AR model—a general form:

$$r_t = \phi_0 + \phi_1 r_{t-1} + \phi_2 r_{t-2} + \cdots + \phi_p r_{t-p} + a_t$$

Stationarity condition:

all the roots of $\phi_p(z) \equiv 1 - \phi_1 z - \phi_2 z^2 - \cdots - \phi_p z^p = 0$ lie outside the unit circle.

1. Order specification

Using Partial ACF (PACF): (naive, but effective)

Definition:

Let r_t be a stationary TS process. The conditional correlation

$$\begin{aligned} & \text{Corr}(r_t, r_{t+k} | r_{t+1}, \dots, r_{t+k-1}) \\ &= \frac{\text{Cov}[(r_t - \hat{r}_t)(r_{t+k} - \hat{r}_{t+k})]}{\sqrt{\text{Var}(r_t - \hat{r}_t) \text{Var}(r_{t+k} - \hat{r}_{t+k})}} \end{aligned}$$

is called the PACF of r_t and r_{t+k} , denoted by ϕ_{kk} , where \hat{Z}_t and \hat{Z}_{t+k} is the best linear estimate of Z_{t+k} given $\{Z_{t+1}, \dots, Z_{t+k-1}\}$ in mean square errors as $k \geq 2$. $\hat{Z}_t = \hat{Z}_{t+k} = EZ_t$ as $k = 1$.

Formula: $\phi_{11} = \rho_1,$

$$\phi_{kk} = \frac{\begin{vmatrix} 1 & \rho_1 & \rho_2 & \cdots & \rho_{k-2} & \rho_1 \\ \rho_1 & 1 & \rho_1 & \cdots & \rho_{k-3} & \rho_2 \\ & & \cdots & & & \\ & & \cdots & & & \\ \rho_{k-1} & \rho_{k-2} & \rho_{k-3} & \cdots & \rho_1 & \rho_k \end{vmatrix}}{\begin{vmatrix} 1 & \rho_1 & \rho_2 & \cdots & \rho_{k-2} & \rho_{k-1} \\ \rho_1 & 1 & \rho_1 & \cdots & \rho_{k-3} & \rho_{k-2} \\ & & \cdots & & & \\ & & \cdots & & & \\ \rho_{k-1} & \rho_{k-2} & \rho_{k-3} & \cdots & \rho_1 & 1 \end{vmatrix}}$$

Sample PACF: given observations r_1, \dots, r_n ,

$\rho_0, \dots, \rho_{n-1}$ are estimated by $\hat{\rho}_0, \dots, \hat{\rho}_{n-1}$.

ϕ_{kk} is estimated by

$$\hat{\phi}_{kk} = \frac{\begin{vmatrix} 1 & \hat{\rho}_1 & \hat{\rho}_2 & \cdots & \hat{\rho}_{k-2} & \hat{\rho}_1 \\ \hat{\rho}_1 & 1 & \hat{\rho}_1 & \cdots & \hat{\rho}_{k-3} & \hat{\rho}_2 \\ & & \cdots & & & \\ & & \cdots & & & \\ \hat{\rho}_{k-1} & \hat{\rho}_{k-2} & \hat{\rho}_{k-3} & \cdots & \hat{\rho}_1 & \hat{\rho}_k \end{vmatrix}}{\begin{vmatrix} 1 & \hat{\rho}_1 & \hat{\rho}_2 & \cdots & \hat{\rho}_{k-2} & \hat{\rho}_{k-1} \\ \hat{\rho}_1 & 1 & \hat{\rho}_1 & \cdots & \hat{\rho}_{k-3} & \hat{\rho}_{k-2} \\ & & \cdots & & & \\ & & \cdots & & & \\ \hat{\rho}_{k-1} & \hat{\rho}_{k-2} & \hat{\rho}_{k-3} & \cdots & \hat{\rho}_1 & 1 \end{vmatrix}}$$

- Key feature: PACF cuts off at lag p for an $AR(p)$ model, i.e. the lag- p sample PACF should not be zero.
- $\hat{\phi}_{k,k}$ should be close to zero for all $j > p$.
- Its distribution is

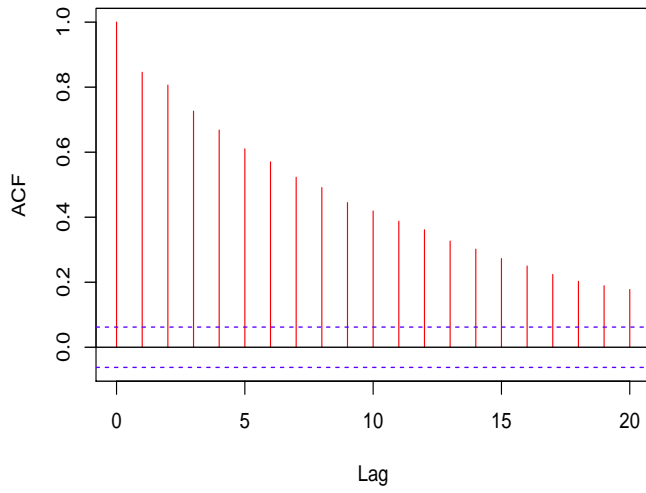
$$\hat{\phi}_{kk} \sim N(0, \frac{1}{n}).$$

Comparing $\hat{\phi}_{k,k}$ to $\pm 1.96/\sqrt{n}$ and decide if $\phi_{k,k} = 0$ is preferable.

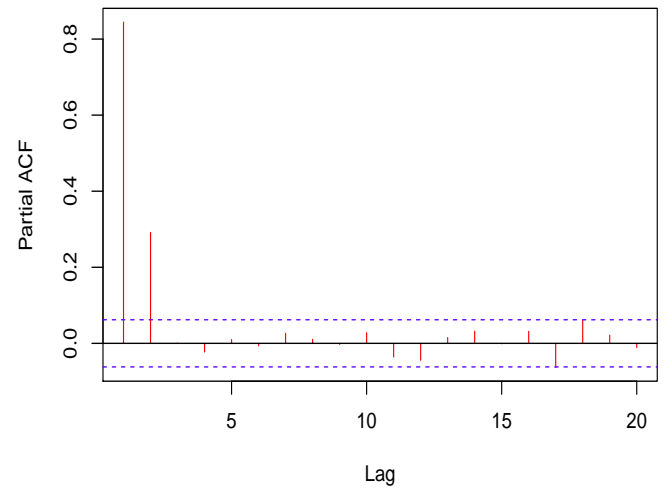
(See Text (p. 40) for details)

– Illustration

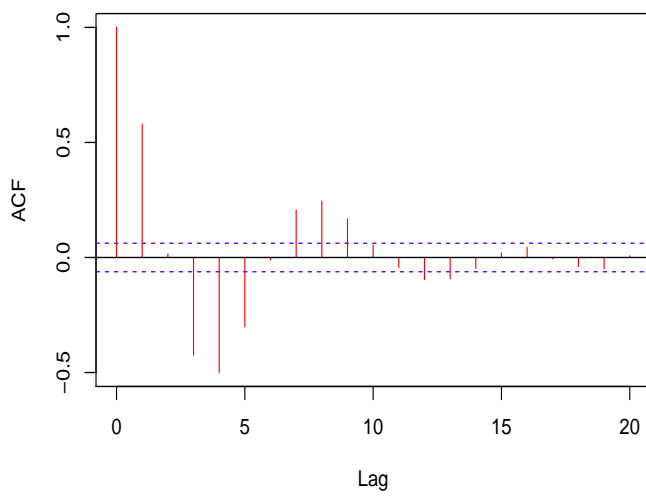
AR(2) with $\phi_1 = 0.6$ and $\phi_2 = 0.3$



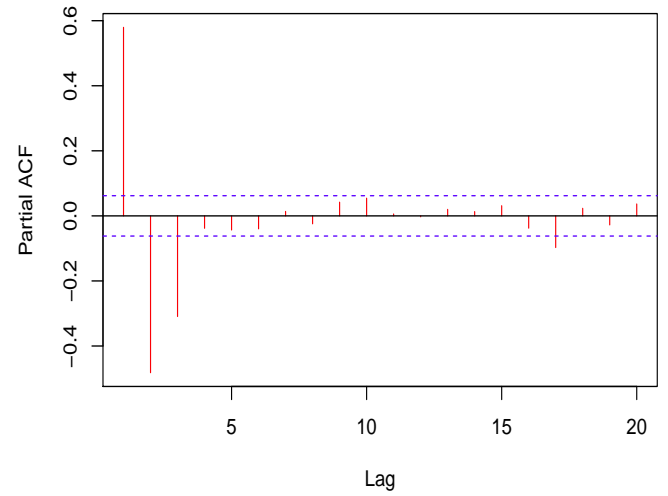
AR(2) with $\phi_1 = 0.6$ and $\phi_2 = 0.3$



AR(3) with $\phi = (0.7, -0.2, -0.3)$



AR(3) with $\phi = (0.7, -0.2, -0.3)$



PACF of AR(2) and AR(3) models.

2. Estimation: least squares estimator (LSE) or maximum likelihood estimator (MLE)

Assume $\{r_1, \dots, r_n\}$ is from AR(p) model with parameter $\tilde{\phi}_0 = (\phi_{00}, \phi_{01}, \dots, \phi_{0p})'$, i.e.

$$r_t = \phi_{00} + \phi_{01}r_{t-1} + \phi_{02}r_{t-2} + \dots + \phi_{0p}r_{t-p} + a_t$$

$\tilde{\phi}_0$ is called the true parameter.

Note that

$$r_t \neq \phi_0 + \phi_1 r_{t-1} + \phi_2 r_{t-2} + \dots + \phi_p r_{t-p} + a_t$$

if $\phi_i \neq \phi_{0i}$.

LSE:

$$\begin{aligned} a_t(\phi) &= r_t - \phi_0 - \phi_1 r_{t-1} - \phi_2 r_{t-2} - \dots - \phi_p r_{t-p} \\ &= r_t - \tilde{Z}'_{t-1} \phi, \end{aligned}$$

where $\tilde{Z}_{t-1} = (1, r_{t-1}, r_{t-2}, \dots, r_{t-p})'$, and $\phi = (\phi_0, \phi_1, \dots, \phi_p)'$ is called unknown parameter.

$$a_t(\phi) = a_t \text{ if and only if } \phi = \tilde{\phi}_0$$

$a_t(\phi)$ is called residuals.

$$S_*(\phi) = \sum_{t=1}^n a_t^2(\phi)$$

$$\begin{aligned} \frac{\partial S_*(\phi)}{\partial \phi} &= 2 \sum_{t=1}^n \frac{\partial a_t(\phi)}{\partial \phi} a_t(\phi) \\ &= -2 \sum_{t=1}^n \tilde{Z}_{t-1} r_t + 2 \sum_{t=1}^n \tilde{Z}_{t-1} \tilde{Z}'_{t-1} \phi \end{aligned}$$

Note that $\frac{\partial S_*(\phi)}{\partial \phi} \big|_{\phi=\hat{\phi}} = 0$, we have

$$\hat{\phi} = \left(\sum_{t=1}^n \tilde{Z}_{t-1} \tilde{Z}'_{t-1} \right)^{-1} \left(\sum_{t=1}^n \tilde{Z}_{t-1} r_t \right).$$

$$\frac{\partial^2 S_*(\phi)}{\partial \phi \partial \phi'} = 2 \sum_{t=1}^n \tilde{Z}_{t-1} \tilde{Z}'_{t-1} > 0.$$

$$a_t(\hat{\phi}) = r_t - \hat{\phi}_0 - \hat{\phi}_1 r_{t-1} - \hat{\phi}_2 r_{t-2} - \cdots - \hat{\phi}_p r_{t-p}.$$

is called the residual.

$$\hat{\sigma}_a^2 = \frac{1}{n} \sum_{t=1}^n a_t(\hat{\phi})^2.$$

Theory:

$$\sqrt{n}(\hat{\phi} - \phi_0) \sim N(0, \hat{\Omega}).$$

$$\sqrt{n}(\hat{\phi}_i - \phi_{0i}) \sim N(0, \hat{\sigma}_{ii}).$$

where

$$\hat{\Omega} = \left(\sum_{t=1}^n \tilde{Z}_{t-1} \tilde{Z}_{t-1}' \right)^{-1} \hat{\sigma}_a^2$$

and $\hat{\sigma}_{ii}$ is the (i, i) -element of $\hat{\Omega}$.

In R, s.e. is $\sqrt{\hat{\sigma}_{ii}/n}$.

$$H_0 : \phi_{0i} = 0 \text{ v.s. } H_a : \phi_{0i} \neq 0$$

If $|\hat{\phi}_i| > 1.96s.e. \approx 2s.e.$, then reject H_0 . ϕ_{0i} is considered to non-zero at significance level $\alpha = 0.05$.

3. Model checking:

Residual should be close to white noise if the model is adequate. Sample ACF of Residuals is

$$\hat{\rho}_k = \frac{\sum_{t=1}^{T-k} [a_t(\hat{\phi}) - \bar{a}][a_{t+k}(\hat{\phi}) - \bar{a}]}{\sum_{t=1}^T [a_t(\hat{\phi}) - \bar{a}]^2},$$

where $\bar{a} = \sum_{t=1}^n a_t(\hat{\phi})/n$. If model is correct, then the ACF of $\{\hat{\rho}_k\}$ is closed to that of $\{a_t\}$, i.e.,

$$\hat{\rho}_k \approx 0.$$

Joint test (Ljung-Box statistics):

$$Q(m) = T(T+2) \sum_{k=1}^m \frac{\hat{\rho}_k^2}{T-k} \sim \chi_{m-g}^2,$$

where g is the number of AR coefficients used in the model.

H_0 : model correct *vs* H_a : model incorrect.

Reject H_0 if

$$Q(m) > \chi_{m-g}^2(\alpha) \text{ or } p = P(\chi_{m-g}^2 > Q(m)) < \alpha.$$

4. Model selection:

(i). Akaike information criterion

$$AIC(l) = \ln(\hat{\sigma}_l^2) + \frac{2l}{T},$$

for an $AR(l)$ model, where $\hat{\sigma}_l^2$ is the MLE of residual variance.

Find the AR order p with

$$AIC(p) = \min_{l \in \{0, \dots, P\}} AIC(l),$$

where P is a given integer large enough.

(ii). BIC criterion:

$$BIC(l) = \ln(\hat{\sigma}_l^2) + \frac{l \ln(T)}{T}.$$

Find the AR order p with

$$BIC(p) = \min_{l \in \{0, \dots, P\}} BIC(l).$$

Needs a constant term? Check the sample mean.

Many software packages available, e.g. R, SCA, Splus, Eviews, SAS, SPSS, etc.